

Are Measures of Early Education Programs Too Pessimistic? Evidence from a Large-Scale Field Experiment*

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Abstract

We estimate the direct and spillover effects of a large-scale early childhood intervention on the educational attainment of over 2,000 disadvantaged children. We show that failing to account for spillover effects results in a severe underestimation of the true policy impact. The intervention induced modest positive direct effects on both cognitive and non-cognitive test scores of children assigned to the treatment groups. We document large spillover effects on both treatment and control children who live near treated children. On average, spillover effects increase a child's non-cognitive (cognitive) scores by about 1.2 (0.6 to 0.7) standard deviations. The spillover effects are localized, decreasing with the spatial distance to treated neighbors. Our evidence suggests the spillover effect on *non-cognitive* scores are more likely to operate through the child's social network rather than the parents'. Our evidence also suggests influencing parental investment is an important channel through which *cognitive* spillover effects operate.

Keywords: early education, neighborhood, field experiment, spillover effects, non-cognitive skills.

JEL Classification: C93, I21, R1.

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“... I will emphasize again and again: that human capital accumulation is a social activity, involving groups of people in a way that has no counterpart in the accumulation of physical capital...” Lucas (1988)

1 Introduction

Evaluations of early childhood programs have played an important role in shaping policy debates on early education. For instance, the Head Start Impact Study (HSIS), a recent randomized control trial of Head Start, reported small effect sizes that fade out over a few years (Puma et al., 2010, 2012). These findings have heightened debate among academics over the cost effectiveness of Head Start (e.g. Barnett, 2011; Gibbs et al. 2011; Kline and Walters, 2016) and have been frequently cited by policymaker critics to argue Head Start is ineffective in achieving its mission and should be abandoned or seriously reformed.¹ Given the policy impact of the findings from early education interventions, accurate evaluation of these programs is crucial.

The standard approach in evaluating early education programs is to randomly assign children to treatment and control groups. From there, many analysts simply difference the mean outcomes and report the monetized treatment effect within a benefit-cost framework. This approach is based on the assumption that a child’s potential outcomes are independent of other participants’ treatment assignment; that is, no spillover effects occur. However, if the assignment of treatment to one subject alters the outcomes of control subjects, the difference between the control and treatment means would not reflect the true impact of the program. Specifically, if the intervention generates positive (negative) externalities benefiting (harming) the control children, ignoring spillover effects would result in an understatement (overstatement) of treatment effects and send an overly pessimistic (optimistic) signal of program quality. Accounting for spillover effects becomes even more important when a program is taken to scale. The level of social interactions—which are often a key channel through which spillover effects operate—increases at scale, resulting in an even larger bias in impact estimates if spillover effects are ignored (Deaton and Cartwright, 2018). In this paper, we provide the first empirical evidence on spillover effects from a large-scale early education intervention and show that disregarding these effects results in severe underestimation of the total impact.

Between 2010 and 2014, a series of early childhood programs were delivered to low-income families with young children in the Chicago Heights Early Childhood Center (CHECC). CHECC was located in Chicago Heights, IL, a neighborhood on Chicago’s South Side with characteristics similar to many other low-performing urban school districts. The goals of the intervention were to examine how investing in cognitive and non-cognitive skills of low-income children aged 3 to 4 affects their

¹ For example, a 2014 report from the House Budget Committee cites the findings of HSIS to conclude Head Start is “failing to prepare children for school” (see: http://budget.house.gov/uploadedfiles/war_on_poverty.pdf).

short- and long-term outcomes, and to evaluate the effectiveness of investing directly in the child’s education versus indirectly through the parents. To that end, families of over 2,000 disadvantaged children were randomized into (i) an incentivized parent-education program (Parent Academy), (ii) a high-quality preschool program (Pre-K), or (iii) a control group. The children’s cognitive and non-cognitive skills were assessed on a regular basis, starting before the randomization and continuing into the middle and end of the programs. Follow-up assessments were also conducted on a yearly basis.

Fryer, Levitt and List (2015) and Fryer et al. (2018) evaluated the impacts of various treatment arms of CHECC by comparing cognitive and non-cognitive developments between children who were randomized into the control and treatment groups. These studies reported modest yet positive treatment effects on cognitive and non-cognitive test scores. Both of these studies estimated treatment effects under the assumption that potential outcomes are independent of other children’s treatment assignment. Our goal is to relax this critical assumption and explore the spatial spillover effects on children’s cognitive and non-cognitive performance. In particular, we study how control and treated children indirectly benefit from this intervention through their “treated” neighbors. Furthermore, we shed light on the underlying mechanisms through which these spillover effects operate. Finally, we estimate the total impact of the intervention, accounting for spillovers, and illustrate how ignoring these indirect effects would bias our estimates of the program impacts.

Treatment effects from the programs offered at CHECC can spill over through two main channels. The first channel is the direct social interactions between children who were randomized during the intervention. Our analysis includes observations from early childhood (3 to 4 years of age), when peer influence at the neighborhood level starts, to middle childhood (8 to 9 years of age), when social interactions within neighborhoods increase dramatically as children enter school (Leventhal and Brooks-Gunn, 2000). Therefore, direct exposure to treated children who live in the same neighborhood is a likely mechanism that can generate spatial spillover effects.² The second channel is the interactions between parents. Observational studies have shown that neighborhoods can influence parental behavior and child-rearing practices (Leventhal and Brooks-Gunn, 2000), which play critical roles in early development (Cunha and Heckman, 2007; Waldfogel and Washbrook, 2011; Kautz et al. 2014; Fryer et al. 2015; Kalil, 2015). Because CHECC also offered education programs to parents, treatment effects can spill over through information and preference externalities, generated by social interactions with other parents who live close by.

We identify the spillover effects from CHECC, by exploiting some unique features of our data. First, conditional on the total number of neighbors who signed up to be a part of the intervention, the number of neighbors who were subsequently assigned to treatments is determined exogenously and through the randomization process. We leverage this experimentally induced variation in spatial

²See Epple and Romano (2011) and Sacerdote (2011) for recent reviews of the literature on peer effects.

exposure to treatments across children to estimate spillover effects. Second, our main identification strategy also exploits the panel nature of our data and the within-individual variations in exposure to treated neighbors, induced by delivery of programs over multiple years. Specifically, by including individual-specific fixed effects, our estimates control for any time-invariant individual, family, and neighborhood unobserved characteristics that might be correlated with spatial exposure to treatments. We also estimate the effects under a second model that relaxes the assumption of time-invariant omitted variables by controlling for the lagged dependent variables (LDV) and dispensing with the fixed effects. Whereas our main identification strategy uses *within-individual* variations to estimate the spillover effects, the LDV specification estimates the effects by exploiting both *within-individual* and *between-individual* variations in spatial exposure to treatments. Our findings, presented below, are robust to using this alternative specification.

We document large and significant spillover effects on both cognitive and non-cognitive skills and find the non-cognitive spillover effects are about two times larger than the cognitive spillover effects. Our estimates suggest that, on average, *each* additional treated neighbor residing within a three-kilometer radius of a child's home increases that child's cognitive score by 0.0033 to 0.0042 standard deviations (σ), whereas it increases her non-cognitive score by 0.0069σ to 0.0070σ . Given that an average child in our sample has 178 treated neighbors residing within a three-kilometer radius of her home—and assuming linearity—we can infer that, on average, a child gains between 0.6σ to 0.7σ in cognitive test scores and about 1.2σ in non-cognitive test scores in spillover effects from her treated neighbors. We find the spillover effects are localized and fall rapidly as the distance to a treated neighbor increases.

Fryer, Levitt, and List (2015) reported heterogeneity in treatment effects from the parental-education component of CHECC along the lines of gender and race. Through comparing outcomes between treatment and control children, Fryer, Levitt, and List (2015) find the Parent Academy increases test scores only for Hispanics and Whites and does not improve the outcomes of African American children. The authors do not find any significant differences by gender. These findings prompt us to explore whether such heterogeneities also exist in spillover effects. Our estimates suggest that non-cognitive spillover effects are significantly larger for African Americans than Hispanics. According to our fixed-effects estimates, an additional treated neighbor within a three-kilometer radius increases the non-cognitive test score of an African American child by 0.0100σ , whereas it increases the non-cognitive score of a Hispanic child by only 0.0045σ . We find no significant racial differences in cognitive spillover effects. Focusing on gender, our estimates suggest boys tend to benefit more than girls from cognitive and non-cognitive spillovers, although these gender differences are not significant at the conventional levels. This observation is consistent with previous empirical evidence on neighborhood effects, which tend to be larger for boys (Entwistle et al., 1994; Halpern-Felsher et al., 1997; Leventhal and Brooks-Gunn, 2000; Katz et al., 2001; Chetty and Hendren, 2018b).

To shed light on the mechanisms through which spillover effects operate, we start by comparing the effects from neighbors who were assigned to the parental-education programs with the effects from neighbors who were assigned to the preschool programs. Because, unlike in the Pre-K treatments, the focus of Parent Academies was on educating parents rather than children, if spillover effects are driven by interactions between parents, we might expect Parent Academy neighbors to generate larger effects than Pre-K neighbors.³ Alternatively, larger spillovers from Pre-K neighbors than from Parent Academy neighbors could imply the peer-influence channel plays an important role in generating the effects. Our estimates suggest non-cognitive spillovers are more likely to operate through preschool neighbors. According to our fixed-effects estimates, whereas an additional Parent Academy neighbor within three kilometers of a child's home induces a 0.0045σ increase in her non-cognitive score, an additional Pre-K neighbor living within the same distance increases her non-cognitive score by 0.0099σ . This finding suggests non-cognitive spillover effects are more likely to operate through children's rather than parents' social networks. We do not find any significant differences in *cognitive* spillover effects from Parent Academy and Pre-K neighbors.

Given our evidence suggesting peer influence at the neighborhood level is a key mechanism in generating non-cognitive spillover effects, we hypothesize that the racial differences in non-cognitive spillovers might be at least partially driven by differences in social interactions within neighborhoods. We explore this idea using data from the National Longitudinal Study of Adolescent Health Survey (Add Health). Our analysis confirms that African American adolescents are significantly more likely than Hispanics to (i) know most people in their neighborhoods, (ii) stop on the street and talk to someone from the neighborhood, and (iii) use recreation facilities in the neighborhood. Although these results cannot be interpreted as causal evidence, they are consistent with our previous finding that social interactions with peers within neighborhoods is a key channel in generating non-cognitive spillover effects.

Our estimates also suggest older children benefit more from non-cognitive spillovers. Indeed, we find the spillover effect on the non-cognitive test score of a 6-year-old child to be about 0.13σ to 0.14σ larger than the corresponding effect on a 5-year-old child. Because social interactions with peers in the neighborhood increase as a child gets older and enters middle childhood (Higgins and Parsons 1983; Cook and Cook, 2009), this observation is consistent with the hypothesis that *peer influence* is an important force in generating non-cognitive spillover effects.

Finally, our evidence suggests *cognitive* spillover effects are likely to operate—at least partially—through influencing the parents' decision to enroll their child in a (non-CHECC) preschool program. Using survey data, we show that families with more treated neighbors are significantly more likely to enroll their child in a preschool program (other than the ones offered at CHECC). Our evidence also

³This intuition does not rule out possible spillover effects from Pre-K neighbors that are generated through parental interactions. Afterall, parents of children who received the Pre-K treatments might also be impacted through the Pre-K programs. This intuition merely assumes Parent Academies affect parents *more* than Pre-K treatments do.

suggests children whose parents reported enrolling them in an alternative preschool program perform significantly better in cognitive assessments. Therefore, we conclude that influencing parental investment decisions—as measured by the choice to enroll one’s child in a preschool program—is a channel through which spillover effects on cognitive test scores operate.

We conclude our analysis by measuring the total impact of the intervention on children’s cognitive and non-cognitive performance, accounting for the spillover effects. Our estimates suggest that, on average, the intervention increased a treatment child’s cognitive (non-cognitive) test score by 0.82σ (1.32σ). Spillover effects make up a large portion of this total impact: whereas the average direct effect of the intervention on a *treatment* child’s cognitive (non-cognitive) score is 0.11σ (0.05σ), the corresponding indirect effect is 0.71σ (1.27σ). Control children also gain considerably as a result of the intervention: On average, the intervention increased a control child’s cognitive (non-cognitive) test score by 0.75σ (1.25σ). If we were to disregard the spillover effects on the control group and had simply based our estimates of the total impact on the outcome differences between the treatment and control children, we would have severely understated the total impact. Specifically, this approach would have misled us to conclude the intervention only improved the cognitive (non-cognitive) test scores of a treatment child by 0.06σ (0.07σ). Ignoring spillover effects would have also led us to wrongly conclude African American children did not benefit from the intervention. Accounting for spillover effects enables us to document a significant and large impact on non-cognitive performance that is significantly larger for African Americans than Hispanics.

We view our results speaking to three strands of literature. First, we contribute to the growing number of studies that measure *spillover effects* from programs and policy changes, designed to improve behaviors and outcomes in various domains such as the labor market (Ferracci, Jolivet, and van den Berg, 2010; Crepon et al., 2013; Lalivé, Landais, and Zweimüller, 2015; Gautier et al., 2017; Muralidharan, Niehaus, and Sukhtankar, 2018), health (Miguel and Kremer 2004; Janssen, 2011; Avitabile, 2012), compliance behavior (Rincke and Traxler, 2011; Boning et al., 2018), voting behavior (Sinclair, McConnell, and Green, 2012; Gine and Mansuri, 2018), retirement saving decisions (Duflo and Saez, 2003), and consumption (Angelucci and De Giorgi, 2009). We contribute to this literature by providing the first evidence on spillover effects from a large-scale early education intervention, shedding lights on the mechanisms, and estimating the total program impact when accounting for these spillover effects.

Second, we contribute to the literature that studies the role of *neighborhoods* in shaping children’s short- and long-term outcomes. The empirical evidence on how neighborhoods affect children comes from observational studies that document correlations between neighborhood characteristics and children’s outcomes, as well as studies that use experimental and quasi-experimental data to disentangle the causal effects of neighborhood from selection effects.⁴ We contribute to this

⁴See Leventhal and Brooks-Gunn (2000), Durlauf (2004), Minh et al. (2017) and Graham (2018) for reviews of

literature in two important ways. Our first contribution is to provide causal evidence on neighborhood effects by exploiting a unique form of exogeneity, which was induced by our intervention. The existing experimental and quasi-experimental evidence on how neighborhoods shape children’s outcomes identifies neighborhood effects using data from residential movers (e.g. Katz, Kling, and Liebman, 2001; Kling, Ludwig, and Katz, 2005; Damm and Dustmann, 2014; Chetty et al., 2016; Chyn, 2017; Chetty and Hendren, 2018a and 2018b). The identification of neighborhood effects in this literature relies on instruments such as randomly assigned housing vouchers, quasi-random assignment of immigrants to different neighborhoods, or public housing demolitions as sources of exogenous changes in neighborhood quality. We take a different approach in providing evidence on how neighborhoods influence children. Our identification strategy takes advantage of a large-scale intervention, delivered to a randomly chosen group of families, and exploits the experimentally induced variations—both within and between individuals—in spatial exposure to treated families to provide causal evidence on how neighbors influence children’s outcomes. Our second contribution to this literature is to shed light on the role of *neighbors* in generating neighborhood effects. Neighborhoods have multiple attributes, which can each influence a child’s outcomes, such as school quality, crime rate, neighbors, and so on. Unlike the previous estimates on neighborhood effects, which provide little empirical evidence on mechanisms, we are able to isolate and estimate the effect of *neighbors’ quality* as one of the many channels through which neighborhoods can influence children’s development. Specifically, our estimates suggest social interactions with other children in the neighborhood play an important role in the development of children’s non-cognitive skills.

Finally, our findings provide important insights for academics interested in modeling the formation of human capital. A growing body of literature develops dynamic models of skill formation to explore the role of various inputs in the production of cognitive and non-cognitive skills. Through structurally estimating such models, this literature has found inputs such as parental ability, home environment, and parental investments to be important determinants in the formation of future skills (e.g. Todd and Wolpin, 2007; Cunha and Heckman, 2008; Cunha, Heckman, and Schennach, 2010; Attanasio, Mehgir, and Nix, 2017; Attanasio et al. 2017). A recent study by Agostinelli (2018) develops a model of skill formation in which peer quality is an input into the production of skills and estimates this model using survey data from adolescents. Whereas Agostinelli exploits variations in cohort composition within schools to deal with the inherent endogeneity in the formation of peer groups in his data, our experimentally induced variation in the spatial exposure to treatments enables us to bypass such identification challenges. Agostinelli’s estimates suggest peer quality is an important input in the production of skills (as measured by grades and receptive vocabulary skills) for teenagers. Our results complement Agostinelli’s by providing empirical evidence for the role of peer influence at even younger ages. Our estimates suggest peer quality plays an especially important role in producing non-cognitive skills.

neighborhood effects on children.

The remainder of the paper is structured as follows. Section 2 summarizes key features of our intervention, randomization, and assessments. Section 3 describes our data and presents our estimation strategy. We present our main findings in section 4, where we report our fixed-effects estimates of spillover effects on cognitive and non-cognitive test scores, and explore heterogeneities by race and gender. Section 5 presents the lagged dependent variable (LDV) estimates of the spillover effects and discusses the robustness of our findings to using this alternative identification strategy. We discuss the mechanisms in section 6. In section 7, we estimate the total impacts of CHECC, break down these estimates into direct and indirect effects, and discuss how ignoring indirect effects would bias our estimates. We discuss policy implications and conclude in section 8.

2 Program Details

2.1 Overview of treatments

Between 2010 and 2014, a series of early childhood interventions were delivered to low-income families with young children in Chicago Heights Early Childhood Center (CHECC). The center was located in Chicago Heights, IL, which is a South Side, Chicago, neighborhood with characteristics similar to many other low-performing urban school districts. According to the 2010 Census, black and Hispanic minorities constituted about 80% of the population of Chicago Heights; its per-capita income was \$17,546 per year; and 90% of students attending the Chicago Heights School District were receiving free or reduced-price lunches.

The main goals of this large-scale intervention were (i) to examine how investing in cognitive and non-cognitive skills of low-income children 3 to 4 years of age affects their long-term outcomes, and (ii) to evaluate the effectiveness of investing directly in children’s education versus indirectly through their parents. To that end, families of over 2,000 children were randomized into either one of the four preschool programs (henceforth “Pre-K”) or one of the two parental-education programs (henceforth “Parent Academy”) or a control group.

The Parent Academy was designed to teach parents to help their child with cognitive skills, such as counting and spelling, as well as non-cognitive skills, such as working memory and self-control. The curriculum for Parent Academy was adapted from two effective preschool curricula: *Tools of the Mind*, which focuses on fostering non-cognitive skills, and *Literacy Express*, which focuses on improving cognitive skills.⁵ The curriculum was delivered to parents in 18, 90-minute sessions, which were held every two weeks over a nine-month period. Parent Academy families had the opportunity to earn up to \$7,000 per year and could participate until their child entered kindergarten. Earnings were based on parents’ attendance, their performance on homework, and their child’s performance

⁵See Fryer, Levitt, and List (2015) for more information on curriculum selection.

on the interim and end-of-year assessments. The two Parent Academy treatments differed only in how they administered incentives. Payments made to families in the “Cash” treatment were made via cash/direct deposits, whereas payments made to families in the “College” treatment were deposited into an account that could only be accessed once the child was enrolled in a full-time post-secondary institution. Fryer, Levitt, and List (2015) evaluated the Parent Academy treatments by comparing the outcomes of treatment and control children, under the assumption that treatments did not induce externalities to the control group, and found Parent Academies increase children’s non-cognitive scores by 0.203σ , but do not significantly impact cognitive scores. Moreover, the authors found that Parent Academies improve test scores for Hispanics and Caucasian children, but not for African Americans. Girls experienced greater benefits than boys, although the gender gap was not significant at the conventional levels.

Besides the Parent Academy, CHECC delivered four preschool programs in which children were treated *directly*. We refer to these programs as Pre-K treatments. These four treatments differed in their curricula, as well as the duration and intensity of delivery. “Tools,” “Literacy,” and “Preschool Plus” were nine-month full-day programs delivered during the school year, whereas “Kinderprep” was a two-month half-day program delivered during the summer before a child entered kindergarten.⁶ The curriculum for “Tools” was designed based on *Tools of the Mind*, which was focused on improving non-cognitive skills, whereas “Literacy” was based on *Literacy Express*, which focused on fostering cognitive skills. A new curriculum called “Cog-X” was developed for “Preschool Plus” and “Kinderprep”, which emphasized both cognitive and non-cognitive skills.⁷

2.2 Randomization

Between 2010 and 2013, 2,185 children from low-income families in South Side, Chicago were recruited and randomized into either one of the six treatments or the control group.⁸ The randomization took place once per year, at the beginning of each academic year.⁹ Some children were

⁶Preschool Plus and Kinderprep also offered a parental component that was much less extensive than the Parent Academies, both in terms of education time and incentives. Parent Academy parents could earn up to \$7,000 based on their attendance, their performance on homework, and their child’s performance on assessments, whereas Preschool Plus and Kinderprep parents could only earn up to \$900 and \$200, based merely on their attendance to parental workshops. Preschool Plus and Kinderprep treatments also offered fewer instruction time to parents. Whereas Parent Academy parents could spend 27 hours in parental workshop, Preschool Plus and Kinderprep parents were offered a maximum of 21 and 6 hours of parental education, respectively. The intensity of the preschool component of Preschool Plus was similar to that of “Tools” and “Literacy Express” in terms of instruction time.

⁷Fryer et al. (2018) evaluated Preschool Plus and Kinderprep under the assumption that the programs did not affect the outcomes of the control group. Through comparing the performance of treatment and control children, the authors found “Cog-X” treatments significantly improved cognitive scores (by about one quarter of a standard deviation), but failed to find any significant effects on non-cognitive scores. For more information on Pre-K programs, see Fryer et al. (2018).

⁸See Appendix B for maps of residential addresses.

⁹The exceptions were years three and four of the program during which randomization took place twice per year: In the first randomizations, children were randomized into either the nine-month preschool program, the

randomized during more than one year, mainly to encourage families who were initially placed in the control group to stay engaged with CHECC for assessments, by offering them a chance to participate in future years.¹⁰ The yearly randomization schedule created four cohorts of children we refer to by their year of randomization.¹¹ Table 1 summarizes the randomization schedule for each year of the program.

Table 1: Randomization by Year

	Control	Parent Academies	9-Month Pre-K	Kindergarten Preparation
Cohort-1 (2010)	242	153	172	0
Cohort-2 (2011)	443	216	166	0
Cohort-3 (2012)	422	0	196	107
Cohort-4 (2013)	376	0	104	99
Unique child	1270	317	539	206

Notes: The number of children randomized into each treatment group in each year of the program is reported. The bottom row presents the number of unique children in each group, over the course of four years.

2.3 Assessments

Our key outcome measures are children's performances in cognitive and non-cognitive assessments, which were used to evaluate the programs. These assessments consist of a *pre-assessment* administered to all incoming students prior to randomization, a *mid-assessment* between January and February, a *post-assessment*, which occurred in May, immediately after the school year ended, and a *summer assessment* at the end of the summer. Besides the assessments that took place during the program year, graduated children were also assessed annually in every April, starting the year after they finished the program. These assessments are referred to as *age-out assessments*. Appendix A presents the assessment schedule for all four cohorts.

Assessments included both cognitive and non-cognitive components and were administered by a

summer Kinderprep program, or the control group; and in the second, a smaller group of families were recruited and randomized into either the summer kindergarten preparation program or the control group. Table 1 combines the two randomizations.

¹⁰As a result, some children who were in the control group in earlier years were randomized into a treatment group in later randomizations. In a few cases, a child who was randomized into a treatment group in earlier years was assigned to a different (or the same) treatment in later randomizations. Overall, 1,675 children were randomized only once, 509 were randomized twice, and one child was randomized in three years.

¹¹Those children who were randomized in multiple years also appear in multiple cohorts.

team of trained assessors who were blind to the child’s treatment. The cognitive component used a series of nationally normed tests, measuring general intellectual ability and specific cognitive abilities such as reading, writing, and mathematics. The non-cognitive component included a combination of subtests measuring executive functions such as working memory, inhibitory control, and attention shifting, as well as a questionnaire completed by assessors, which measured self-regulation in emotional, attentional, and behavioral domains.

3 Data and the Econometric Model

3.1 Data

3.1.1 Construction of outcome variables

Our outcome measures are indexes generated from standardized test scores on cognitive and non-cognitive assessments.¹² The cognitive assessment included the Peabody Picture Vocabulary Test (PPVT), which assesses verbal ability and receptive vocabulary (Dunn et al., 1965), and four subtests of the Woodcock Johnson III Test of Achievement (WJ): (i) WJ-Letter and Word Identification (WJL), which measures the ability to identify letters and words; (ii) WJ-Spelling (WJS), which measures the ability to correctly write orally presented words; (iii) WJ-Applied Problems (WJA), which measures the ability to analyze and solve math problems; and (iv) WJ-Quantitative Concepts (WJQ), which assesses the knowledge of mathematical concepts, symbols, and vocabulary (Woodcock et al. 2001).

The non-cognitive component included the Blair and Willoughby Executive Function test (Willoughby et al. 2012), which is composed of three subtests assessing attention (Spatial Conflict), working memory (Operation Span), and attention shifting (Same Game) and the Preschool Self-Regulation Assessment (PSRA), which is designed to assess self-regulation in emotional, attentional, and behavioral domains.¹³

A cognitive index was made up of averaged percentile scores on each cognitive subtest, and a non-cognitive index was made up of average percent-correct scores on each non-cognitive subtest. The two indexes were then standardized by the type of assessment (pre-assessment, mid-assessment, etc.), including the entire study population (treated and control) who took that assessment, to obtain a zero mean and standard deviation of one.

¹²These indexes were constructed by Fryer et al. (2015, 2018) for the original evaluations of the programs.

¹³Because Blair and Willoughby tests are designed for preschool, a new test was added for assessments that were administered to older children (age-out assessments). For children in kindergarten or older, the Same Game test of Blair and Willoughby was replaced with a variant of Wisconsin Card Sort game, which measures attention shifting for children of that age.

To explore the spatial spillovers on both treatment and control children, we construct three samples: a pooled sample, including observations from both treatment and control groups; a control sample, including data from control children; and a treatment sample, including observations from treated children. Our treatment sample pools together observations from children who were randomized in any of the programs. We include observations from the baseline up to the fourth age-out assessment. Our final control, treatment, and pooled samples include 2,442, 3,074, and 5,208 observations, respectively.¹⁴ Appendix C presents the details regarding how we construct these three samples.

Table 2 provides summary statistics on the baseline demographic variables for our pooled sample. Note the majority (90%) of the children are either African American or Hispanic, and 53% live in families with an annual household income under \$35,000.

Table 2: Baseline Summary Statistics for the Pooled Sample

Variable	Share/Mean	Variable	Share/Mean
Gender		Mother's Education	
Male	.51	Less than high school	.08
Race		Some high school but no diploma	.12
Black	.40	High school diploma	.13
Hispanic	.50	Some college but no degree	.17
White	.09	College degree	.18
Other Race	.01	Other	.06
Missing Race	.01	Missing Mother's Education	.25
HH Income and Unemployment Benefits		Father's Education	
below 35K	.53	Less than high school	.09
36K-75K	.14	Some high school but no diploma	.1
75K+	.06	High school diploma	.13
Income Missing	.26	Some college but no degree	.12
Receives Unemployment Benefit	.09	College degree	.08
Missing unemployment benefit	.31	Other	.06
		Missing Father's Education	.44
Baseline Age (months)	45.32 (6.91)		

Notes: Summary statistics for baseline demographic variables are presented. For education levels, *Some high school but not diploma* includes parents with a GED or high school attendance without a diploma, *College degree* includes associate's, bachelor's and master's degrees, *Less than high school* includes an education level below 9th grade or no formal schooling, and *Other* includes vocational/technical or other unclassified programs. Standard deviations are reported in parentheses.

¹⁴Note that the number of observations in the pooled sample is smaller than the sum of the number of observations in our control and treatment samples. The reason is that in a few cases, when a child was first randomized into the control group and was placed into a treatment group in later randomizations, the pooled sample only includes observations that took place after the child was randomized into treatments. See Appendix C for more information.

3.1.2 Addresses and neighbor counts

To estimate the spatial spillover effects from the intervention, we calculate the number of treated neighbors of a child at a given time and use it as a measure of spatial *exposure* to treatments. To do so, we start by calculating commuting distances between the home locations of all pairs of children who were randomized during the intervention.¹⁵ Commuting distances are calculated by considering the street network structure and its restrictions (e.g., one-way roads, U-turns, etc.) and finding the closest driving distance between each pair. The average travel distance between a pair of children in our sample is 8.52 kilometers (std. dev.= 8.07), and 99.8% of the sample resides within 60 kilometers of each other. Figure 1 presents a histogram of travel distances between home locations of all children who were randomized during the intervention.

We define a pair as *neighbors* if the commuting distance between the two is less than “ r ” kilometers, and we call “ r ” the *neighborhood radius*. We conduct our analysis for various values of neighborhood radii. We then calculate the number of treated ($N_{i,t,r}^{treated}$) and control ($N_{i,t,r}^{control}$) neighbors of each child i at the time of her assessment t , and define the total number of CHECC neighbors of i as $N_{i,t,r}^{total} = N_{i,t,r}^{treated} + N_{i,t,r}^{control}$. Note that as more children are randomized into treatment and control groups over the four years of the intervention, the number of treated and control neighbors vary over time.¹⁶ Table 3 reports the summary statistics for $N_{i,t,r}^{treated}$ and $N_{i,t,r}^{control}$, and Figure 2 presents histograms of the exposure measure $N_{i,t,r}^{treated}$, for various values of neighborhood radii. Whereas for $r = 1$ kilometers, the variation in $N_{i,t,r}^{treated}$ is small, as the neighborhood radius increases to 3 kilometers and beyond, we gain considerable variations in the exposure measure.

Table 3: Neighbor Counts by Neighborhood Radius

	r = 1 km	r = 3 km	r = 5 km	r = 7 km
$N_r^{treated}$	27.89 (27.73)	178.13 (154.21)	325.63 (238.75)	422.81 (272.08)
$N_r^{control}$	29.49 (31.84)	183.47 (165.34)	333.73 (257.95)	437.79 (301.11)

Notes: This table presents the average number of treated and control neighbors of a child in our pooled sample, for various definitions of neighborhood radii. The numbers reflect all the observations in our pooled sample for which we observe both cognitive and non-cognitive scores. Standard deviations are reported in parentheses.

¹⁵Distances were calculated using the ArcGIS OD Cost Matrix Analysis tool.

¹⁶Because more children were receiving the treatments over the four-year span of the intervention, and a neighbor who was previously in the control group in an earlier randomization might be assigned into a treatment group in later years, $N_{i,t,r}^{control}$ can both increase or decrease over time. However, $N_{i,t,r}^{treated}$ can only increase over time, because no child who was already treated could be assigned into the control group in later years.

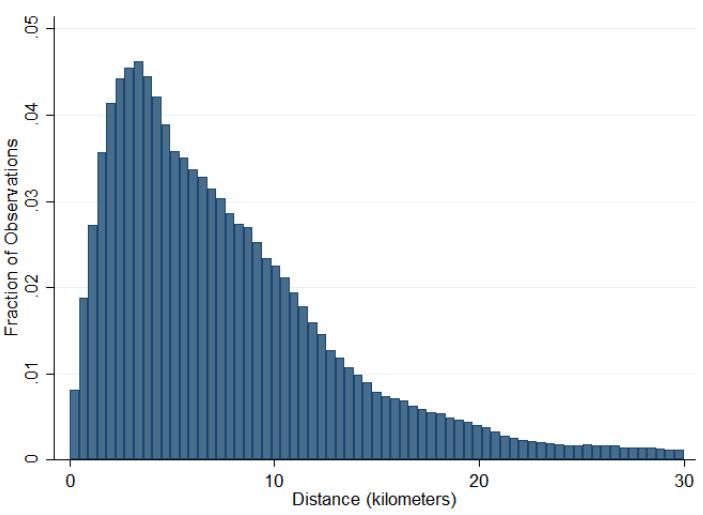


Figure 1: Histogram of distances between children in the study. The horizontal axis is cut at 30 kilometers.

3.2 Econometric model

We exploit three unique features of our data to estimate the spillover effects. First, conditional on the total number of a child's CHECC neighbors at a given point in time ($N_{i,t,r}^{total}$), the number of neighbors who are randomized into treatments ($N_{i,t,r}^{treated}$) is determined exogenously through the intervention. Second, the repeated assessment schedule generates a panel data, which enables us to track the performance of a child over time. Finally, multiple randomizations and the delivery of programs over the four years of the intervention create *within-individual* variations in our exposure measure $N_{i,t,r}^{treated}$.

Although the experimentally induced variation in our exposure measure serves as an important feature, which we exploit for identification, our estimation strategy does not exclusively rely on it. Given our limited sample size and the fact that the intervention was not designed to measure spatial spillovers, our exposure measure could be correlated with individual- or neighborhood-level unobservable characteristics. Therefore, we exploit the panel nature of our data to provide clean estimates of the spillover effects. The above three properties allow us to estimate the spillover effects using *within-individual* variations in our exposure measure ($N_{i,t,r}^{treated}$) through a fixed-effects specification. This technique uses the variations in spatial exposure over time and controls for any unobserved time-invariant individual-, family-, or neighborhood-level characteristics that might be correlated with $N_{i,t,r}^{treated}$.

We estimate spatial spillover effects from CHECC, using an individual fixed-effects specification of

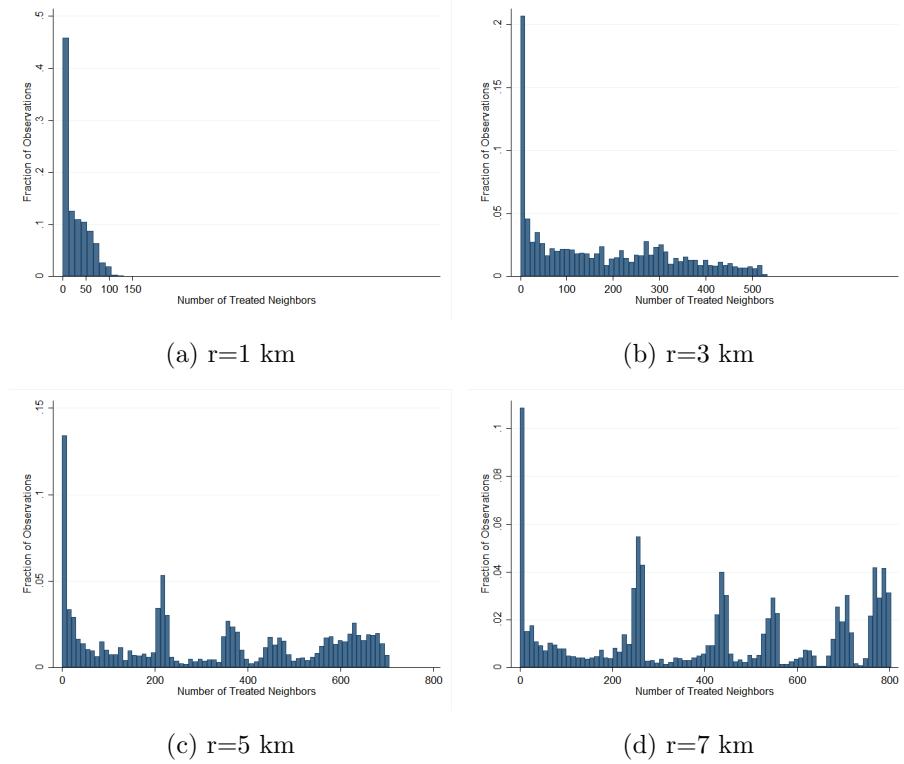


Figure 2: Histogram of $N_{i,t,r}^{treated}$ for $r = \{1, 3, 5, 7\}$ kilometers.

the form

$$Y_{i,t} = \beta_0 + \beta_1 N_{i,t,r}^{treated} + \beta_2 N_{i,t,r}^{total} + \gamma_i + \delta_t + \epsilon_{i,t,r}, \quad (1)$$

where $Y_{i,t}$ is the standardized cognitive or non-cognitive test score of a child i on test t , $N_{i,t,r}^{treated}$ represents the number of treated neighbors of i at time t as previously defined, and $N_{i,t,r}^{total}$ represents the total number of i 's neighbors who were randomized in the intervention by time t .¹⁷ γ_i and δ_t are individual and test (time) fixed effects. Under this specification, β_1 represents the average effect of moving one of the control neighbors of a child i to a treatment group, holding the total number of her neighbors constant. This measure (β_1) provides an intuitive estimate on the spillover effects from the intervention, because it enables a policymaker to weigh the benefits against the costs associated with treating an additional child in the neighborhood.

¹⁷ Miguel and Kremer (2004), Bobba and Gignoux (2016) and Gine and Mansouri (2018) use similar specifications to estimate spatial spillover effects. Like our specification, these studies use the number of treated individuals within a certain neighborhood radius as their measure of spatial exposure to treatments, and control for the total number of neighbors in their regression analyses. Unlike these three studies, which have relied exclusively on the experimentally induced variations in the distribution of treated neighbors across individuals, we are able to exploit the panel nature of our data and identify the spillover effects using within-individual variations in exposure to treatments and remove all unobserved time-invariant individual-level characteristics that might be correlated with the spatial exposure to treatments.

Section 5 presents an alternative model, which relaxes the assumption of time invariance for individual effects and shows the robustness of our findings to using this second specification.

4 Results

4.1 Main findings

We estimate spillover effects for the neighborhood radii of 3, 5, and 7 kilometers. As Figure 2 suggests, when neighborhood is defined too narrowly, the variation in $N_{i,t,r}^{treated}$ becomes too small, limiting our power to estimate the effects. Therefore, we define the neighborhood radii to be 3 kilometers and larger, which provides us with enough degree of variation in $N_{i,t,r}^{treated}$ to estimate the effects. Arguably, these choices of neighborhood radii are economically relevant. According to the National Household Travel Survey, the average commuting distance to school for a 6- to 12-years-old child is about 6 kilometers (3.6 miles).¹⁸ The average travel time from home to work for a Chicago Heights resident is estimated to be 26.1 minutes (US Census Bureau statistics), which translates to about 21 kilometers for a speed of 30 miles per hour.¹⁹ Because schools and workplaces provide natural interaction spaces for children and their parents, we can reasonably assume our choices of neighborhood radii are relevant distances within which social interactions can generate spillovers effects.

Table 4 presents estimated β_1 's from equation 1 for neighborhood radii of 3, 5, and 7 kilometers. Standard errors—clustered at the census-block-group level to allow for common error components within geographical units—are reported in parentheses below each point estimate. The left (right) panel presents the average spillover effect from treating an additional neighbor of a child on her standardized cognitive (non-cognitive) test score. Columns 1 and 4 report the pooled effects on both treatment and control children and reveal significant positive spillover effects on both cognitive and non-cognitive test scores. The effects on non-cognitive scores are over twice as large as the effects on cognitive scores: An additional treated neighbor within 3 kilometers of a child's home increases her cognitive score by 0.0033σ ($p < 0.01$), whereas it increases her non-cognitive score by 0.0069σ ($p < 0.01$). The differences in cognitive and non-cognitive spillovers are statistically significant.²⁰

Columns 2 and 3 break down the effects on cognitive scores by the treatment assignment of the

¹⁸<https://nhts.ornl.gov/briefs/Travel%20To%20School.pdf>

¹⁹ This estimate is consistent with a report by the National Household Travel Survey that suggests the average commuting distance from home to work in the US is about 19 kilometers (11.8 miles). For more information, refer to <https://www.bts.gov/sites/bts.dot.gov/files/docs/browse-statistical-products-and-data/national-transportation-statistics/220806/ntsentire2018q1.pdf> (page 73).

²⁰The p-values from the Wald test of the null hypothesis H0: $\beta_1^{cog} = \beta_1^{ncog}$ against H1: $\beta_1^{cog} \neq \beta_1^{ncog}$ are 0.001, 0.03, and 0.06 for neighborhood radii of 3, 5, and 7 kilometers, respectively.

receiving children. These estimates reveal that both treatment and control children benefit from living close to treated families. While the control group seems to benefit slightly more than the treatment group from cognitive spillover effects, the difference is not significant at the conventional levels.²¹ Columns 5 and 6 report the spillover in non-cognitive scores by the treatment assignment of the receiving child. These estimates illustrate that the treatment and control children both benefit from non-cognitive spillovers. The estimated spillover effects on non-cognitive scores on the control and treatment groups are very similar and are not significantly different across the two groups.²² These findings are robust to the choice of neighborhood radius, r .²³

In sum, we document significant positive spillover effects on both cognitive and non-cognitive test scores and find the effects are significantly larger for non-cognitive scores.²⁴

Table 4: Mean Effect Sizes on Cognitive and Non-cognitive Scores, Fixed-Effects Estimates

	Cognitive Scores			Non-cognitive Scores		
	Pooled (1)	Control (2)	Treatment (3)	Pooled (4)	Control (5)	Treatment (6)
$r = 3 \text{ km}$	0.0033*** (0.0010)	0.0038*** (0.001)	0.0016 (0.0010)	0.0069*** (0.0013)	0.0075*** (0.0015)	0.0064*** (0.0013)
$r = 5 \text{ km}$	0.0021*** (0.0006)	0.0023** (0.0008)	0.0010* (0.0006)	0.0043*** (0.0008)	0.0037*** (0.0011)	0.0034*** (0.0008)
$r = 7 \text{ km}$	0.0018*** (0.0005)	0.0021*** (0.0007)	0.0008* (0.0005)	0.0033*** (0.0007)	0.0025*** (0.0010)	0.0028*** (0.0007)
Obs.	5,208	2,442	3,074	5,208	2,442	3,074

Notes: Spillover effects from each additional treated neighbors ($\hat{\beta}_1$) estimated from equation 1 are presented. Columns 1-3 (4-6) represent the average spillover effects from an additional treated neighbor on a child's standardized cognitive (non-cognitive) score. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

4.2 Spatial fade-out

A closer look at the estimated β_1 's, reported in Table 4, suggests an important spatial pattern: The spillover effects from an additional treated neighbor becomes smaller as we broaden the neighborhood radius from 3 to 7 kilometers. To further explore this pattern and shed light on the relationship between spillover effects and distance, Figure 3 graphs the estimated β_1 's for a broader

²¹P-values from the Wald test of equal β_1^{cog} 's for the control and treatment group are 0.11, 0.16, and 0.13 for neighborhood radii of 3K, 5K, and 7K meters, respectively.

²²P-values from the Wald test of equal β_1^{ncog} for treatment and control group for neighborhood radii of 3K, 5K, and 7K meters are 0.80, 0.84, and 0.78.

²³Appendix D breaks down these effects by subtests and explores which components of the cognitive/non-cognitive index generate the effects.

²⁴Appendix I presents the estimated effects for a subsample of children who attend the majority of assessments.

range of r 's.²⁵ Note the effects on both cognitive and non-cognitive scores operate very locally. As we increase the neighborhood radius, the spillover effects from an additional treated neighbor monotonically decrease. Because a larger neighborhood radius corresponds to a longer average distance to neighbors, the negative relationship between $\hat{\beta}_1$ and r implies that as the distance between a child and her treated neighbor grows, the spillover effects on both cognitive and non-cognitive scores weakens. Specifically, the average spillover effect on a child's non-cognitive score, from a treated neighbor within a 3-kilometer radius is about twice as large as the effect from a treated neighbor who resides within a 7-kilometer radius (0.0069σ vs. 0.0033σ). Similarly, the average effect on the cognitive score from an additional treated neighbor who lives within 3 kilometers of a child is about twice as large as the effect from a treated neighbor who resides within a 7-kilometer radius (0.0033σ vs. 0.0018σ). In summary, we find spillover effects on both cognitive and non-cognitive scores are localized and decrease as the distance to a treated neighbor's home increases.

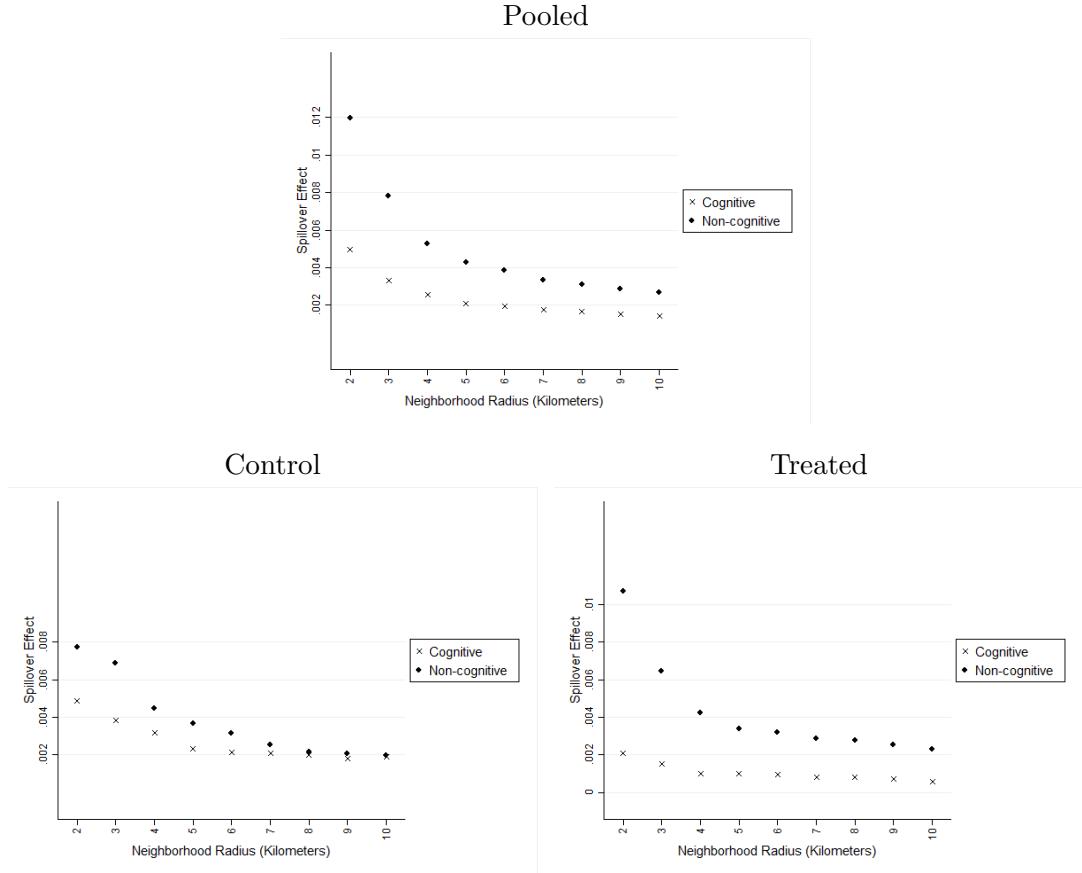


Figure 3: The spillover effect from having an additional treated neighbor on a child's standardized cognitive and non-cognitive scores, as functions of neighborhood radius.

²⁵The point estimates are reported in Appendix E.

4.3 Heterogeneous effects

Fryer, Levitt, and List (2015) evaluated the parent-education component of CHECC (Parent Academy). The authors reported positive treatment effects on cognitive and non-cognitive scores for Hispanics, but did not find any significant treatment effects on African American children. Parent Academy was also reported to have slightly larger effects on girls than boys, although the gender differences were not significant.

Motivated by the heterogeneity in treatment effects from the Parent Academy component of the intervention reported in Fryer, Levitt, and List (2015), we investigate whether children of different races (or gender) benefit differently from spillover effects. We do so by estimating equation 1, separately by race and gender.

Since African Americans and Hispanics make up over 90% of our sample, our analysis on heterogeneity along race focuses on these two groups. Panel (a) of Table 5 and Figure 4 present $\hat{\beta}_1$'s separately for African American and Hispanic receiving children. Comparing the effects across races, we find no significant differences in cognitive spillover effects between Hispanics and African Americans. In contrast to the effects on cognitive scores, spillovers on non-cognitive scores are significantly larger for African Americans than Hispanics.²⁶ On average, an additional treated neighbor increases the non-cognitive score of an African American child by about two to three times as much as a Hispanic child. For instance, an additional treated neighbor within a 3-kilometer radius increases the non-cognitive score of a Hispanic child by 0.0045σ , whereas it increases an African American child's non-cognitive score by 0.0100σ .

Panel (b) of Table 5 and Figure 4 present the estimated effects by the gender of a receiving child. Overall, boys seem to benefit more than girls from both cognitive and non-cognitive spillovers. However, these gender differences are not statistically significant.²⁷ In section 6, we shed light on the mechanism through which spillover effects operate, and discuss a potential source of heterogeneity in spillover effects along race and gender lines.

Note that our estimates cannot be directly compared with the heterogeneity in treatment effects reported by Fryer, Levitt, and List (2015) due to two main differences between our samples. First, unlike Fryer, Levitt, and List (2015) who report on heterogeneous effects from Parent Academy, our analysis uses data from all programs and considers heterogeneity in spillover effects on all children who were randomized during the intervention. Second, whereas Fryer, Levitt, and List (2015) base

²⁶P-values from Wald tests of the null of equal β_1 for Hispanics and African American children's cognitive (non-cognitive) test scores (in the pooled sample) are 0.15 (0.07), 0.10 (0.04), and 0.17 (0.07) for neighborhood radii of 3, 5, and 7 kilometers, respectively.

²⁷P-values from Wald tests of the null of equal β_1 for cognitive (non-cognitive) test scores for boys and girls (in the pooled sample) are 0.12 (0.47), 0.13 (0.50), and 0.10 (0.47) for neighborhood radii of 3, 5, and 7 kilometers, respectively.

Table 5: Mean Effect Sizes within Gender and Race Subgroups, Fixed-Effects Estimates

	(a) Race		(b) Gender	
	Cognitive Scores (1)	Non-cognitive Scores (2)	Cognitive Scores (3)	Non-cognitive Scores (4)
Black				
r = 3 km	0.0014 (0.0015)	0.0100*** (0.0024)	0.0048*** (0.0016)	0.0088*** (0.0019)
r = 5 km	0.0007 (0.0008)	0.0055*** (0.0015)	0.0029*** (0.0009)	0.0048*** (0.0012)
r = 7 km	0.0009 (0.0008)	0.0042*** (0.0012)	0.0024*** (0.0007)	0.0038*** (0.0010)
Obs.	2,087	2,087	2,583	2,583
Hispanic				
r = 3 km	0.0042*** (0.0013)	0.0045*** (0.0014)	0.0017 (0.0011)	0.0068*** (0.0019)
r = 5 km	0.0027*** (0.0009)	0.0019** (0.0008)	0.0013* (0.0007)	0.0037*** (0.0011)
r = 7 km	0.0023*** (0.0007)	0.0016** (0.0008)	0.0011 (0.0006)	0.0028*** (0.0009)
Obs.	2,580	2,580	2,625	2,625

Notes: The spillover effects from each additional treated neighbor on cognitive and non-cognitive test scores, estimated from the fixed-effects model (equation 1). Panel (a) presents the effects, separately for African American and Hispanic receiving children. Panel (b) reports the effects separately for boys and girls. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

their estimates on observations from the post program assessments (which took place immediately after the end of program year), our estimates use the pre-, mid-, and post-program assessment as well as up to four additional follow-up assessments, which were administered after a program year ended.

5 Robustness

Our identification strategy, presented in section 3 (individual fixed-effects model), was based on the assumption that individual effects are *time-invariant* omitted variables. In this section, we explore the robustness of our findings to a different estimation strategy, where we relax this assumption by directly controlling for the lagged dependent variables and dispensing with the fixed effects. Whereas the fixed-effects specification uses within-individual variations in spatial exposure to treatments, the lagged dependent variables (LDV) model exploits both within- and between-individual variations to estimate the spillover effects. Angrist and Pischke (2008) argue that fixed effects and LDV estimates have a bracketing property such that if the true model is LDV, fixed-effects estimates of a positive treatment effect provide upper bounds for the true effects, and if

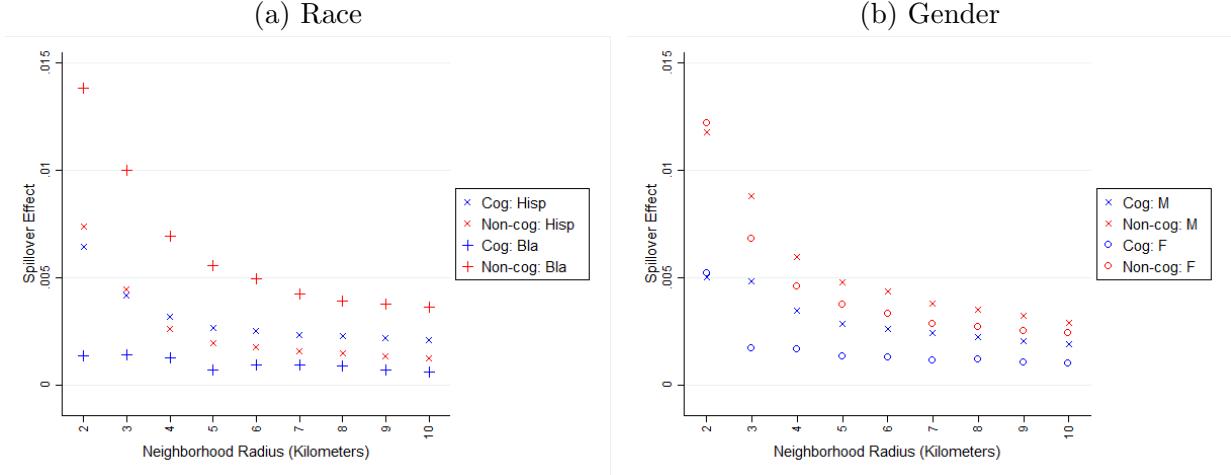


Figure 4: The spillover effect from an additional treated neighbor on a child’s standardized cognitive and non-cognitive scores, estimated from the fixed-effects model. Panel (a) presents the effects separately for African American and Hispanic children, and panel (b) presents the effects separately for males and females.

fixed effects is the true specification, the LDV estimates of a positive treatment effect provide lower bounds for the true effects. Therefore, the true effects are bounded to lie between the values estimated from the two models. We formulate this alternative specification as

$$Y_{i,t} = \beta_0 + \beta_1 N_{i,t,r}^{treated} + \beta_2 N_{i,t,r}^{total} + \eta Y_{i,t-1} + X_i \alpha + \sigma_b + \delta_t + \mu_c + \epsilon_{i,t,r}, \quad (2)$$

where $Y_{i,t}$, $N_{i,t,r}^{treated}$, and $N_{i,t,r}^{total}$ are defined as previously. We control for the lagged cognitive and non-cognitive test scores through $Y_{i,t-1}$, and include census-block group (σ_b) as well as time and cohort fixed effects (δ_t and μ_c). X_i represents a vector of time-invariant characteristics including gender, race, and age at the time of the baseline assessment. Under this specification, β_1 reflects the average spillover effect from an additional treated neighbor who lives within radius r on a child’s standardized cognitive or non-cognitive test score.

5.1 Main Effects

Table 6 presents β_1 ’s, estimated from equation 2, for three different neighborhood radii: 3, 5, and 7 kilometers. The standard errors, reported in parentheses, are clustered at the census-block-group level to allow for common error components within geographical units. Consistent with the results presented in section 4.1, we find significant spillover effects on both cognitive and non-cognitive test

scores, with larger effects on non-cognitive scores.²⁸ The comparison between the effects on treated and control children leads to a conclusion similar to the one we reported under the fixed-effects specification: We do not find any significant differences in spillover effects between children who were assigned to the treatment and control groups.²⁹ The estimates from the two models are also similar in magnitude: An additional treated neighbor within 3 kilometers of a child is estimated to increase that child's cognitive score by 0.0033σ under the fixed-effects specification and by 0.0042σ under the LDV model. Similarly, an additional treated neighbor residing within 3 kilometers is estimated to increase the child's non-cognitive score by 0.0069σ and 0.0070σ under the fixed-effects and LDV specifications, respectively.

Table 6: Mean Effect Sizes on Cognitive and Non-cognitive Scores, LDV Estimates

	Cognitive Scores			Non-cognitive Scores		
	Pooled (1)	Control (2)	Treatment (3)	Pooled (4)	Control (5)	Treatment (6)
r = 3 km	0.0042*** (0.0012)	0.0052** (0.0025)	0.0019 (0.0020)	0.0070*** (0.0015)	0.0059*** (0.0018)	0.0067** (0.0026)
r = 5 km	0.0033*** (0.0009)	0.0038** (0.0016)	0.0021* (0.0011)	0.0059*** (0.0014)	0.0037*** (0.0011)	0.0060*** (0.0023)
r = 7 km	0.0027** (0.0007)	0.0034*** (0.0013)	0.0015 (0.0010)	0.0054*** (0.0014)	0.0035*** (0.0013)	0.0053** (0.0022)
Obs.	3,403	1,495	2,093	3,403	1,495	2,093

Notes: Spillover effects from *each* additional treated neighbors ($\hat{\beta}_1$) estimated from equation 2 are presented. Columns 1-3 (4-6) represent the effects from an additional treated neighbor on a child's standardized cognitive (non-cognitive) score. Robust standard errors, clustered at the census-block-group level are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

5.2 Spatial fade-out

Figure 5 graphs the estimated effects from an additional treated neighbor, as a function of neighborhood radius. The point estimates are reported in Appendix E. Similar to Figure 3, the estimated effects from the LDV specification also exhibit a downward-sloping relationship with the neighborhood radius. The spatial fade-out is more salient for the effects on non-cognitive scores. Confirming our finding from estimating the fixed-effects model reported in section 4.2, the LDV estimates show that the spillover effect from a treated neighbor decreases as the spatial distance to that neighbor's home increases.

²⁸ P-values from the Wald test of the null hypothesis of $\beta_1^{cog} = \beta_1^{ncog}$ against the alternative of $\beta_1^{cog} \neq \beta_1^{ncog}$ are 0.045, 0.035, and 0.035 for neighborhood radii of 3, 5 and 7 kilometers, respectively.

²⁹ P-values from the Wald test of equal β_1^{cog} 's for the control and treatment group are 0.34, 0.39, and 0.22 for neighborhood radii of 3K, 5K, and 7K meters, respectively. The corresponding p-values for β_1^{ncog} 's are 0.82, 0.37, and 0.44.

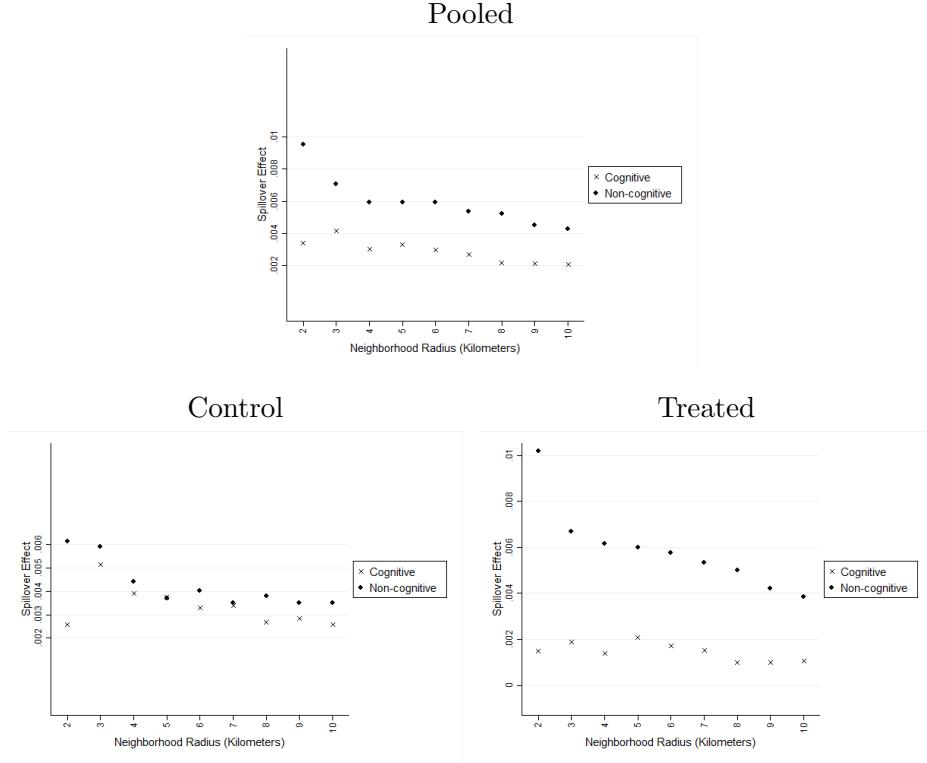


Figure 5: The spillover effects, estimated under the LDV model, from an additional treated neighbor on a child’s standardized cognitive and non-cognitive test scores, as functions of neighborhood radius.

5.3 Heterogeneous effects

Table 7 reports spillover effects by the race and gender of a receiving child, estimated from the LDV specification (equation 2). Focusing on race (panel-a), we find no significant differences in spillover on cognitive scores between African Americans and Hispanics.³⁰ Similar to the fixed-effects estimates, the LDV estimates of non-cognitive spillover effects reveal a large and significant racial gap, with African Americans benefiting about three to four times as much as Hispanics from their treated neighbors.³¹

Panel (b) of Table 7 reports the effects from the LDV specification, by the gender of a receiving child, and reveals that in general, boys tend to benefit slightly more than girls from spillover effects.

³⁰P-values from Wald tests of the null hypothesis of equal spillover effects on cognitive scores (β_1) for Hispanics and African Americans in the pooled sample are 0.59, 0.52, and 0.50 for neighborhood radii of 3, 5, and 7 kilometers, respectively.

³¹P-values from Wald tests of the null hypothesis of equal spillover effects on non-cognitive scores (β_1) for Hispanics and African Americans in the pooled sample are 0.001, 0.005, and 0.020 for neighborhood radii of 3, 5, and 7 kilometers, respectively.

Table 7: Mean Effect Sizes within Gender and Race Subgroups, LDV Estimates

	(a) Race		(b) Gender	
	Cognitive Scores (1)	Non-cognitive Scores (2)	Cognitive Scores (3)	Non-cognitive Scores (4)
Black				
r = 3 km	0.0056*** (0.0019)	0.0154*** (0.0034)	0.0054*** (0.0026)	0.0063*** (0.0023)
r = 5 km	0.0037*** (0.0013)	0.0108*** (0.0025)	0.0037*** (0.0011)	0.0059*** (0.0021)
r = 7 km	0.0032*** (0.0012)	0.0087*** (0.0022)	0.0033*** (0.0009)	0.0060*** (0.0021)
Obs.	1,312	1,312	1,693	1,693
Hispanic				
r = 3 km	0.0042** (0.0017)	0.0034** (0.0016)	0.0032** (0.0014)	0.0067*** (0.0017)
r = 5 km	0.0026** (0.0011)	0.0034** (0.0014)	0.0027*** (0.0010)	0.0053*** (0.0019)
r = 7 km	0.0022** (0.0009)	0.0031** (0.0016)	0.0017* (0.0009)	0.0044** (0.0018)
Obs.	1,760	1,760	1,710	1,710

Notes: Spillover effects from an additional treated neighbor on cognitive and non-cognitive test scores, estimated from LDV specification (equation 2). Panel (a) presents the effects, separately for African American and Hispanic receiving children. Panel (b) reports the effects separately for boys and girls. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

However, the gender differences are not significant at conventional levels.³²

6 Exploring the Mechanisms

6.1 Parent Academy versus Pre-K

Since the intervention offered education programs for both children and parents, one might expect the spillover effects to operate through the social network of children via direct interactions between them, or indirectly through their parents' social networks. To shed light on which channel generates stronger effects, we start by comparing spillovers from treated neighbors who were assigned to the parent-education programs (Parent Academy neighbors) with the effect from those who were assigned to the preschool programs (Pre-K neighbors). If spillovers mainly operate through parents' social network, we might expect larger effects from Parent Academy neighbors than from Pre-K neighbors, because Parent Academy treatments involved parents more directly and more intensely

³²P-values from Wald tests of the null hypothesis of equal spillover effects on cognitive scores (β_1) for boys and girls in the pooled sample are 0.29, 0.36, and 0.16 for neighborhood radii of 3, 5, and 7 kilometers, respectively. The corresponding p-values for spillover effects in non-cognitive scores are 0.88, 0.82, and 0.53.

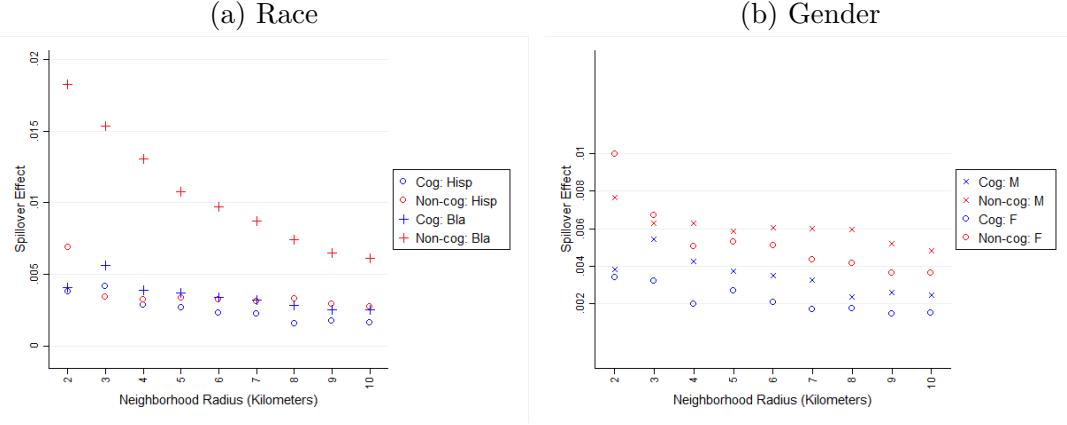


Figure 6: The spillover effect from an additional treated neighbor on a child’s standardized cognitive and non-cognitive scores, estimated from the LDV model. Panel (a) presents the effects separately for African American and Hispanics, and panel (b) presents the effects by gender.

than Pre-K treatments. On the other hand, larger spillover effects from Pre-K neighbors could imply the direct interactions between children play an important role in generating the effects.

To compare spillovers from Parent Academy and Pre-K neighbors, we estimate the fixed-effects and LDV specifications of the following forms:³³

$$Y_{i,t} = \beta_0 + \beta_p N_{i,t,r}^{Parent} + \beta_c N_{i,t,r}^{PreK} + \lambda N_{i,t,r}^{total} + \gamma_i + \delta_t + \epsilon_{i,t} \quad (3)$$

$$Y_{i,t} = \beta_0 + \beta_p N_{i,t,r}^{Parent} + \beta_c N_{i,t,r}^{PreK} + \lambda N_{i,t,r}^{total} + \eta Y_{i,t-1} + X_i \alpha + \sigma_b + \mu_c + \delta_t + \epsilon_{i,t}. \quad (4)$$

$N_{i,t,r}^{Parent}$ and $N_{i,t,r}^{PreK}$ represent the number of Parent Academy and Pre-K neighbors of a child i who reside within a distance r from i at the time of her assessment t , and $N_{i,t,r}^{total}$, $Y_{i,t}$, $Y_{i,t-1}$, X_i , γ_i , σ_b , μ_c and δ_t are defined as previously. To simplify the analysis and retain statistical power, we construct $N_{i,t,r}^{PreK}$ and $N_{i,t,r}^{Parent}$ by pooling neighbors who were assigned to any of the preschool programs as *Pre-K neighbors*, and pooling those who were assigned to any of the two parent-education programs as *Parent Academy neighbors*. Under the above specifications, β_p reflects the average spillover effect from an additional Parent Academy neighbor, holding $N_{i,t,r}^{PreK}$ and $N_{i,t,r}^{total}$ constant. In other words, β_p represents the average effect of substituting a control neighbor with a Parent Academy neighbor. Similarly, β_c represents the average spillover effect from an additional Pre-K neighbor on a child’s test scores.

Note that a child i may benefit from a Parent Academy neighbor k through two channels. The

³³In a few cases in which a treated neighbor k was first assigned to a Parent Academy (Pre-K) treatment and assigned to Pre-K (Parent Academy) in later years, k is counted as a Parent Academy (Pre-K) neighbor for the observations prior to the second randomization, and as a Pre-K (Parent academy) neighbor for the observations following the second randomization.

first channel is the parents' social networks: k 's parents may influence the behavior and decisions of i 's parents', which may in turn shape i 's development. Such effects can occur through information externalities (i.e., k 's parents share their acquired knowledge from Parent Academy with i 's parents) or preference externalities between parents. The second channel is peer influence: If Parent Academy improves k 's outcomes, child i might benefit from direct interactions with child k . The benefits from a Pre-K neighbor, however, are likely to spill over mainly through direct interactions between children (peer influence) because parents are not the main target of the Pre-K treatments.³⁴ Thus, although $\hat{\beta}_p$ might reflect spillovers through both the parents' and the child's social networks, $\hat{\beta}_c$ is more likely to reflect an effect that is mainly driven by direct interactions between children.

Table 8 reports estimated β_p and β_c for neighborhood radii of 3, 5, and 7 kilometers from the fixed-effects and LDV specifications. Focusing on *non-cognitive* scores, the estimates from both models suggest larger spillover effects from Pre-K neighbors than Parent Academy neighbors ($\hat{\beta}_p < \hat{\beta}_c$).³⁵ According to the fixed-effects estimates, an additional Pre-K neighbor within a 3-kilometer radius of a child increases her non-cognitive score by 0.0099σ , whereas an additional Parent Academy neighbor only induces a 0.0045σ increase in the non-cognitive test score. Similarly, the LDV estimates suggest an additional Pre-K neighbor within a 3-kilometer radius of a child increases her non-cognitive test score by 0.0108σ , whereas an extra Parent Academy neighbor within the same radius induces only a 0.0017σ increase in her non-cognitive test score. The larger spillovers from Pre-K neighbors than from Parent Academy neighbors suggests the non-cognitive effects are more likely to operate through direct social interactions between children rather than between parents.

Unlike the effects on non-cognitive scores, we find no significant differences in spillover effects on cognitive scores from Parent Academy and Pre-K neighbors.³⁶ Appendix G breaks down the estimated effects ($\hat{\beta}_p$ and $\hat{\beta}_c$), by the treatment assignment of the receiving children and confirms the above findings for both the treatment and control subgroups.

³⁴As previously described in section 2, two out of the four Pre-K treatments (Preschool Plus and Kinderprep) had parental-education components that were unincentivised, and were much less intensive than the one offered to Parent Academy families. Appendix F examines the spillover from Pre-K neighbors who were randomized to preschool programs with and without the parental component and shows that our conclusions are not sensitive to pooling these two treatments together.

³⁵P-values for the fixed-effects estimates from Wald tests of the null hypothesis $\beta_p = \beta_c$ against $\beta_p \neq \beta_c$ for r=3 km, r=5 km and r=7 km equal 0.004, 0.006, and 0.016, respectively. The corresponding p-values from the LDV estimates are 0.004, 0.010, and 0.03.

³⁶For the fixed-effects estimates, the p-values from Wald tests of the null hypothesis $\beta_p = \beta_c$ against $\beta_p \neq \beta_c$ for r=3 km, r=5 km and r=7 km are 0.82, 0.43, and 0.34, respectively. The corresponding p-values from the LDV estimates are 0.76, 0.14, and 0.14.

Table 8: Spillover Effects from Parent Academy and Pre-K Neighbors, Pooled Sample

	(a) Fixed Effects Estimates				(b) LDV estimates			
	Cognitive Scores		Non-cognitive Scores		Cognitive Scores		Non-cognitive Scores	
	β_p	β_c	β_p	β_c	β_p	β_c	β_p	β_c
$r = 3 \text{ km}$	0.0035***	0.0033***	0.0045***	0.0099***	0.0045***	0.0040***	0.0017	0.0108***
	(0.0012)	(0.0011)	(0.0008)	(0.0018)	(0.0016)	(0.0014)	(0.0017)	(0.0021)
$r = 5 \text{ km}$	0.0024***	0.0019***	0.0024***	0.0055***	0.0046***	0.0030***	0.0019	0.0068***
	(0.0008)	(0.0006)	(0.0006)	(0.0011)	(0.0013)	(0.0009)	(0.0017)	(0.0015)
$r = 7 \text{ km}$	0.0021***	0.0016***	0.0022***	0.0041***	0.0041***	0.0025***	0.0023	0.0058***
	(0.0005)	(0.0005)	(0.0005)	(0.0009)	(0.0013)	(0.0007)	(0.0020)	(0.0013)
Obs.	5,208		5,208		3,403		3,403	

Notes: Columns 1 and 3 (2 and 4) report the average effect of an additional Parent Academy (Pre-K) neighbor who resides within distance r of a child, on her standardized cognitive and non-cognitive scores. Panel (a) presents the estimates from the fixed-effects specifications (equation 3), and panel (b) reports LDV estimates (4). Robust standard errors, clustered at the census-block-group level, are in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$

6.2 Heterogeneity in neighborhood-level social interactions

Our estimates presented in sections 4 and 5 suggested that non-cognitive spillover effects are significantly larger for African Americans than Hispanics. We also found spillover effects to be larger for boys than girls, although the gender difference was not significant at the conventional levels. Given our evidence suggesting direct interactions with peers at the neighborhood level might be an important mechanism in generating non-cognitive spillover effects, one might hypothesize that the racial and gender differences in non-cognitive spillover effects originate from racial and gender differences in social interactions within neighborhoods. We explore this idea using data from the National Longitudinal Study of Adolescent Health survey (Add Health),³⁷ which includes measures of social interactions within neighborhoods. Add Health is a longitudinal study of a nationally representative sample of adolescents in grades 7 to 12 in the United States. Using the public-use data from Add Health, we explore racial and gender differences in variables that reflect children’s social interactions within neighborhoods. Table 9 presents our findings. Consistent with our hypothesis, we find African American children are significantly (at $p < 0.001$) more likely than Hispanics, and boys are significantly (at $p < 0.001$) more likely than girls, to (i) know most people in their neighborhood, (ii) stop on the street to talk to someone from the neighborhood, and (iii) use recreation facilities in the neighborhood. Although these data speak to a slightly older age group than the one in our study, we believe that these findings provide suggestive evidence that the higher level of

³⁷The AddHealth is a program project directed by Kathleen Mullan Harris and designed by J. Richard Udry, Peter S. Bearman, and Kathleen Mullan Harris at the University of North Carolina at Chapel Hill, and funded by grant P01-HD31921 from the Eunice Kennedy Shriver National Institute of Child Health and Human Development, with cooperative funding from 23 other federal agencies and foundations. Information on how to obtain the Add Health data files is available on the Add Health website (<http://www.cpc.unc.edu/addhealth>). No direct support was received from grant P01-HD31921 for this analysis.

neighborhood-level social interactions for African Americans (and for boys) is a possible cause of the larger non-cognitive spillover effects on this subgroup.

Table 9: Heterogeneity in neighborhood level social interactions (Add Health Data)

	African American	Hispanic	Male	Female
(1) Know most people in the neighborhood	0.78	0.69	0.79	0.71
(p-value)		(0.00)		(0.00)
(2) Stop on the street to talk to someone who lives in the neighborhood	0.84	0.78	0.86	0.79
(p-value)		(0.00)		(0.00)
(3) Use recreation/sports center in the neighborhood	0.25	0.18	0.28	0.17
(p-value)		(0.00)		(0.00)
Observations	1,121	537	787	871

Notes: The share of respondents who agreed with each statement. The data are restricted to African American and Hispanic respondents. Questions are from Wave II of the Add Health survey. Demographic information was obtained from Wave I of the survey. The precise wording of the survey questions are as follows: (1) You know most of the people in your neighborhood. (0 = false, 1 = true); (2) In the past month, you have stopped on the street to talk with someone who lives in your neighborhood. (0 = false, 1 = true); (3) Do you use a physical fitness or recreation center in your neighborhood? (0 = no, 1 = yes). P-values from a two-sample test of proportions under the null hypothesis of equal proportions against the alternative of unequal proportions between each pair of subgroups are reported in parentheses.

6.3 Spillovers and age

We might expect peer influence to play a more salient role in generating spillover effects for older than for younger children. Since at younger ages, children depend on their parents to put them in contact with peers, their friendship network tends to be mostly limited to children of their parents' acquaintances (Hartup, 1984). As children step into middle childhood and enter school, their interactions with peers increase dramatically and start growing independently from their caregivers' social network (Higgins and Parsons, 1983). Previous research has documented that children's social contact—both in terms of the number of peers as well as the time spent with peers—substantially increases between the ages of 6 and 12 (Cook and Cook, 2009). Following this line of reasoning, one might expect the peer influence-channel of the spillover effects to play a smaller role in earlier years when peer effects at the neighborhood level are more limited, and to become more pronounced as a child grows older. By exploring how the magnitude of cognitive and non-cognitive spillovers change as children age, we can gain deeper insight into the mechanisms generating the effects. We examine this idea by including an interaction term in equations 1 and 2:

$$Y_{i,t} = \beta_0 + \beta_1 N_{i,t,r}^{treated} + \beta_2 N_{i,t,r}^{treated} \times Age_{i,t} + \lambda N_{i,t,r}^{total} + \gamma_i + \delta_t + \epsilon_{i,t}. \quad (5)$$

$$Y_{i,t} = \beta_0 + \beta_1 N_{i,t,r}^{treated} + \beta_2 N_{i,t,r}^{treated} \times Age_{i,t} + \lambda N_{i,t,r}^{total} + \eta Y_{i,t-1} + X_i \alpha + \sigma_b + \mu_c + \delta_t + \epsilon_{i,t}. \quad (6)$$

where $Age_{i,t}$ represents the age of child i (in months) at the time of assessment t , and other arguments are defined as previously. Under the above specifications, β_2 reflects how spillover

effects change as a child grows. The fixed-effects and LDV estimates of β_2 are reported in Table 10. Note that under both specifications, estimated β_2 's for non-cognitive outcomes are positive and significant, suggesting older children benefit more than younger ones from non-cognitive spillover effects. According to our fixed-effects (LDV) estimates, on average, the spillover effect on the non-cognitive test score of a 6-year-old child from treated neighbors within a 3-kilometer radius is 0.13σ (0.14σ) larger than the corresponding effect on a 5-year-old child. Although the positive relationship between non-cognitive spillovers and age cannot be taken as causal evidence for mechanisms and should be interpreted with caution, it is consistent with the hypothesis that direct peer influence is a likely channel in generating non-cognitive spillover effects. The estimates of β_2 for cognitive spillovers are negative, which can be regarded as a decline in cognitive spillover effects with age. However, these point estimates are generally not significant at the conventional levels.

The results presented in sections 6.1, 6.2 and 6.3 provide suggestive evidence that peer influence at the neighborhood level serves as an important mechanism in generating *non-cognitive* spillover effects.

Table 10: Spillover Effects and Age

	Fixed Effects		LDV	
	Cognitive	Non-cognitive	Cognitive	Non-cognitive
	(1)	(2)	(3)	(4)
r = 3 km	-2.64e-06 (4.68e-06)	6.10e-05*** (1.50e-05)	-5.99e-07 (5.92e-06)	6.73e-05*** (1.68e-05)
r = 5 km	-6.34e-06 (4.26e-06)	5.41e-05*** (1.22e-05)	-4.03e-06 (4.75e-06)	5.22e-05*** (1.37e-05)
r = 7 km	-8.35e-06** (3.95e-06)	4.00e-05*** (1.08e-05)	-3.34e-06 (4.51e-06)	3.91e-05*** (1.26e-05)
Obs.	5,208	5,208	3,403	3,403

Notes: Estimated $\hat{\beta}_2$'s from equations 5 and 6 are presented. The point estimates represent how spillover effects from an additional treated neighbor change with a child's age (in months). Robust standard errors, clustered at the census-block-group level, are reported in parentheses; *** p<0.01, ** p<0.05, * p<0.1

6.4 Parental Investment

Spillover effects can also operate through influencing parental decisions, which affect children's development. Parents might learn from their neighbors about returns to investments and the most productive forms of investments in their children and adjust their choices accordingly. Our data include a self-reported measure of investment concerning parents' decision to enroll their child in preschool programs (other than CHECC). This variable was collected through a survey completed by parents, which was administered at the end of each program year (at the time of the children's

post-assessment).³⁸

We start by exploring whether this self-reported measure of parental investment has any predictive power regarding children's cognitive and non-cognitive performance, that is, whether enrolling one's child in a preschool program (other than CHECC) is productive in generating cognitive and non-cognitive skills. We do so by estimating an LDV model of the following form:³⁹

$$Y_{i,t} = \beta_0 + \kappa Z_{i,t} + \eta Y_{i,t-1} + X_i \alpha + \sigma_b + \mu_c + \delta_t + \epsilon_{i,t}, \quad (7)$$

where $Z_{i,t}$ is a binary variable indicating whether parents of a child i reported enrolling i in a preschool program during the school year prior to t .⁴⁰ All other terms are defined as previously. Under the above specifications, κ reflects whether a parent's decision to enroll her child in a preschool program is a relevant input in producing skills ($Y_{i,t}$). Table 11 presents κ 's estimated from equations 7. The estimates suggest a parent's decision to enroll her child in other (non-CHECC) preschool programs significantly increases the child's cognitive test score by 0.134σ , whereas the impact on non-cognitive test-scores is not significant at conventional levels. In the next step, we explore whether spillover effects occur on a parent's decision to enroll her child in other programs.

Table 11: Does a parent's decision to enroll a child in preschool programs affect the child's cognitive and non-cognitive scores?

	Cognitive	Non-cognitive
Other Programs	0.134*** (0.048)	0.127 (0.086)
Observations	655	655
R-squared	0.80	0.58

Notes: Estimated κ 's from equation 7 are presented. The point estimates reflect the effect of the parental decision to enroll children in preschool programs, and children's cognitive and non-cognitive test scores. Although we have about 1,000 observations on investment decisions, we observe lagged test scores for only about 60% of these observations. Robust standard errors, clustered at the census-block-group level, are reported in parentheses; *** p<0.01, ** p<0.05, * p<0.1

Our analysis of the spatial spillover effects on cognitive and non-cognitive test scores exploited the panel nature of our outcome variables, allowing us to estimate the causal spillover effects using individual fixed-effects or lagged dependent-variables specifications. Unfortunately, for a large majority (over 90%) of our sample, the data on parental decisions to enroll children in non-

³⁸For the last cohort of families who were randomized in the program, this information was also collected from parents at the time of pre-assessment (in addition to the post-assessment).

³⁹Since for the vast majority (over 90%) of our sample, we observe the investment measure only once, we cannot estimate their effects on test scores using within-individual variations.

⁴⁰Controlling for $N_{i,t,r}^{treated}$ and $N_{i,t,r}^{total}$ or the child's treatment assignment does not change our point estimates very much.

CHECC programs were collected only once (at the time of children's post-assessment). Therefore, we cannot use the previous identification strategies to estimate the causal spillover effects on parents' investment decision. Instead, we rely on the following OLS specification to explore this channel:

$$Z_{i,t} = \beta_0 + \beta_1 N_{i,t,r}^{treated} + \lambda N_{i,t,r}^{total} + X_i \alpha + \sigma_b + \mu_c + \delta_t + \epsilon_{i,t}, \quad (8)$$

where $Z_{i,t}$ represents parents' investment decision, and other variables are defined as previously. Under the above specification, β_1 represents the relationship between having an additional treated neighbor residing within distance r and a parent's decision to enroll her child in a non-CHECC preschool program. Table 12 presents estimated β_1 's from equation 8. Interestingly, our estimates suggest positive and significant spillover effects on this measure of parents' investment decision. Each additional treated neighbor residing within a 3-kilometer radius of a child's home increases the likelihood of the child's parents enrolling her in a preschool program by 0.55 percentage points. Given our previous finding that enrolling a child in a non-CHECC program significantly improves her cognitive performance, this result provides suggestive evidence that influencing the parental decision to enroll her child in a preschool program is an important channel through which cognitive spillover effects operate.

Table 12: Exposure to treatments and the parents' decision to enroll their child in preschool programs

Dependent Variable	Other Programs
$r = 3k$	0.0055** (0.0023)
$r = 5k$	0.0066* (0.0040)
$r = 7k$	0.010** (0.0041)
Observations	974

Notes: Estimated β_1 's from equation 8 are presented. Robust standard errors, clustered at the census-block-group level, are reported in parentheses; *** p<0.01, ** p<0.05, * p<0.1

The above finding should be interpreted with caution for two reasons. First, our measure of parental investment decision is self-reported, and therefore is not an objective measure of the actual choices. Second, given that the structure of our survey data prevents us from using individual fixed-effects or LDV methods in estimation, interpreting β_1 as the causal spillover effect on parents' decision would require an additional assumption that (conditional on $N_{i,t,r}^{total}$, X_i , σ_b , μ_c and δ_t) our exposure measure— $N_{i,t,r}^{treated}$ —is uncorrelated with unobserved individual-level characteristics. Recognizing these two caveats, we believe our results provide *suggestive* evidence that *cognitive* spillover effects might operate—at least partly—through influencing parents' decisions to enroll their child in a

preschool program.

7 Total Impact

After documenting positive spillover effects from the programs delivered at CHECC, we now turn to estimating the total impact of the intervention, accounting for these indirect effects. Beyond estimating the total impact, we are going to (i) disentangle the direct and the spillover effects from the intervention, and (ii) estimate the size of biases that would arise if we were to ignore the spillovers.

Before presenting our evaluation strategy, we should emphasize three key features of this exercise. First, whereas CHECC offered multiple education programs to both parents and children, the aim of this exercise is not to separately evaluate each program. Instead, we are going to provide an overall evaluation of the intervention as a whole by pooling together all treatments and accounting for spatial spillover effects. Second, given that we estimate spillover effects using panel data over multiple years, our total-impact evaluation would also be based on observations over multiple years, starting at the time of randomization and ending four years after the programs delivered to a child's cohort end. Therefore, our analysis would provide an average estimate of the total impact of CHECC over time, which would include the immediate as well as the longer-run effects.⁴¹ Finally, to fix ideas and simplify the presentation of our results, we set the neighborhood radii to 3 kilometers. Appendix H presents our estimates of the total impact for a broader range of neighborhood radii and discusses the robustness of our findings to varying r .

The total impact of the intervention (*Total*) on a child i who was randomized into one of the treatments (Parent Academies or Pre-K) can be expressed as the sum of the direct treatment effect (DE) and the spillover effects, which i receives from N other treated individuals in her vicinity (ST^N): $Total = Direct + ST^N$. If an evaluator ignores the spillover effects on control children (SC^N) and follows the standard approach, she would naively report the difference in the mean outcomes between the control and treated as the total impact. This standard approach would result in a biased estimate, which we call $Total^{Standard} = Total - SC^N = Direct + ST^N - SC^N$.

We evaluate the total impact of the intervention by estimating the following LDV model, using our pooled sample, which includes observations from both the treated and control children. We focus on the LDV specification for this estimation, because unlike the fixed-effects model, it allows us to

⁴¹For this reason, our estimates of the total impact cannot be directly compared to the ones reported in Fryer, Levit, and List (2015), which are based on test scores from the assessments administered immediately after the programs ended.

exploit between-individual variations to estimate the direct time-invariant treatment effect:

$$Y_{i,t} = \beta_0 + \theta T_i + \beta_1 N_{i,t,r}^{treated} + \lambda T_i \times N_{i,t,r}^{treated} + \beta_2 N_{i,t,r}^{total} + \eta Y_{i,t-1} + \alpha X_i + \delta_t + \mu_c + \sigma_b + \epsilon_{i,t}. \quad (9)$$

T_i is a treatment indicator, which equals 1 if i was assigned to a treatment group, and 0 otherwise, and $Y_{i,t}$, $Y_{i,t-1}$, $N_{i,t,r}^{treated}$, $N_{i,t,r}^{total}$, X_i , δ_t , μ_c , and σ_b are defined as previously. We include an interaction term ($T_i \times N_{i,t,r}^{treated}$) to allow for different spillover effects on treatment and control children. Under this specification, θ represents the average direct effect of the intervention on a treatment child (*Direct*), β_1 represents the average spillover effect from an additional treated neighbor on a control child (*SC*¹), and $\beta_1 + \lambda$ represents the average spillover effect from an additional treated neighbor on a treatment child (*ST*¹). Assuming linearity, the average total spillover effect from all treated neighbors on a control child (*SC* ^{N}) can be expressed as $\overline{N_r^{treated}} \times \hat{\beta}_1$, where $\overline{N_r^{treated}}$ represents the average number of treated neighbors who reside within distance r of a child.⁴² Likewise, the average total spillover effect from all treated neighbors on a child who was randomized into a treatment (*ST* ^{N}) can be expressed as $\overline{N_r^{treated}} \times (\hat{\beta}_1 + \hat{\lambda})$. Therefore, the total impact of the intervention on a treatment child (*Total*) is the sum of the direct and the spillover effects: $\hat{Total} = \hat{\theta} + \overline{N_r^{treated}} \times (\hat{\beta}_1 + \hat{\lambda})$.

Table 13 reports the estimated coefficients from equation 9 for $r = 3$ kilometers. The first two columns present the estimates for the pooled sample, and the last four report the estimates separately for African Americans and Hispanics. Focusing on the pooled sample, the average direct effects of being randomly assigned to a treatment group on a child's standardized cognitive and non-cognitive scores are 0.11 and 0.05, respectively. The average total spillover effects on a *control child*'s standardized cognitive and non-cognitive scores (*SC* ^{N}) are estimated to be 0.75 and 1.25. The corresponding spillover effects on a *treatment child* are 0.71σ and 1.27σ . The total impact of being assigned to a treatment (including both the direct and spillover effects) on a child's standardized cognitive and non-cognitive test scores is estimated to be 0.82 and 1.32, respectively. Note that the total spillover effects on both cognitive and non-cognitive scores of treatment children are larger than the direct treatment effects, suggesting a large portion of the total impact is due to the network effects that emerge from interactions with other treated individuals. This finding implies that if one were to treat a single child in isolation, the average cognitive and non-cognitive treatment effects would be about $(\frac{0.82}{0.11} \approx) 7$ and $(\frac{1.31}{0.05} \approx) 26$ times smaller than the estimated impacts in the presence of spillovers from other treated children.

The comparison across the last two rows of Table 13 shows that if we were to ignore the spillover effects altogether, we would have severely underestimated the total impact of the intervention. If we were to disregard the spillover effects, we would have wrongly concluded the intervention

⁴²Appendix J shows these findings do not change much if we relax the linearity assumption and allow for quadratic and cubic terms.

Table 13: Total Program Impact

	All		African American		Hispanic	
	Cognitive	Non-cognitive	Cognitive	Non-cognitive	Cognitive	Non-cognitive
	(1)	(2)	(3)	(4)	(5)	(6)
$\hat{\theta}$ (<i>Direct</i>)	0.1070*** (0.0393)	0.0456 (0.0642)	0.0145 (0.0667)	-0.0558 (0.111)	0.276*** (0.0671)	0.135 (0.0978)
$\hat{\beta}$	0.0042*** (0.0013)	0.0070*** (0.0015)	0.0056*** (0.0019)	0.0154*** (0.0034)	0.0044** (0.0018)	0.0034** (0.0016)
$\hat{\lambda}$	-0.0002* (0.0001)	0.0001 (0.0002)	0.0001 (0.0003)	0.0005 (0.0003)	-0.0007*** (0.0002)	-0.0000 (0.0003)
SC^N	0.75*** (0.22)	1.25*** (0.27)	0.72*** (0.24)	1.99*** (0.44)	1.27*** (0.48)	0.81** (0.38)
ST^N	0.71*** (0.22)	1.27*** (0.27)	0.74*** (0.25)	2.05*** (0.45)	1.05** (0.49)	0.80** (0.36)
<i>Total</i>	0.82*** (0.23)	1.32*** (0.27)	0.75*** (0.25)	1.99*** (0.44)	1.38*** (0.47)	0.93*** (0.37)
<i>Total^{Standard}</i>	0.06*** (0.02)	0.07** (0.03)	0.03 (0.05)	0.01 (0.07)	0.11*** (0.03)	0.12*** (0.04)
Observations	3,403	3,403	1,312	1,312	1,760	1,760
R-squared	0.659	0.384	0.684	0.441	0.613	0.362

Notes: Estimated coefficients from equation 9 are presented. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1. N includes all observations for which we have lagged cognitive and non-cognitive scores, and other regressors.
 $\overline{N^{treated}} = 178.13$; $\overline{N_{Black}^{treated}} = 128.83$; $\overline{N_{Hispanic}^{treated}} = 237.38$

only induced a 0.06σ and 0.07σ increase in cognitive and non-cognitive test scores, respectively, of a child who was randomly assigned to one of the treatment groups (*Total^{Standard}*).⁴³ These estimates on cognitive and non-cognitive scores are considerably smaller than our estimated effects, which account for spillovers to the control group (*Total*).

We report our estimates of impacts on African Americans and Hispanics, in columns 3-6 of Table 13.⁴⁴ The first observation is that the direct effects of being randomized into a treatment group (*Direct*) on both cognitive and non-cognitive test scores are larger on Hispanics than African Amer-

⁴³Note these estimates are considerably smaller than the estimated effects of the Parent Academy arm of the intervention previously reported in Fryer, Levitt, and List (2015). Fryer, Levitt, and List (2015)—who ignore spillover effects—report that Parent Academies increased cognitive and non-cognitive test scores by 0.119σ and 0.203σ , by the end of the program year. Our *Total^{Standard}* estimates of the whole intervention (which ignore spillover effects to the control group) on cognitive and non-cognitive scores are $1.99 (=0.119/0.06)$ and $2.9 (=0.203/0.07)$ times smaller than the reported effects by Fryer, Levitt, and List (2015). Two important factors cause the difference in these two sets of estimates: First, Fryer, Levitt, and List (2015) focus only on one treatment arm of the intervention (Parent Academies), whereas we estimate the impact of CHECC as a whole, and our sample includes observations from all children who were randomized over the four years of the intervention. Second, Fryer, Levitt and List (2015) focus on the short-term treatment effects and compare test scores from assessments that were administered immediately at the end of a program year. Our analysis, on the other hand, uses data from up to four years after a program ends, meaning we rely on both short- and longer-term outcomes to estimate the impact. In fact, the treatment effects wear off and the differences in test scores between the treatment and control groups become smaller over time, which suggests our estimates of *Total^{Standard}* should be smaller than the effects reported in Fryer, Levitt, and List (2015).

⁴⁴On average, an African American (Hispanic) child in our sample has 129 (237) treated neighbors within 3 kilometers of her home.

icans, although the difference is only significant for cognitive skills.⁴⁵ While the racial difference in cognitive spillovers are small and insignificant ($p > 0.40$), non-cognitive spillovers are significantly larger for African Americans ($p < 0.05$). Overall, the intervention increases the *cognitive* scores for African American and Hispanic children who were randomized into treatments by 0.75σ and 1.38σ , which are not significantly different from each other ($p = 0.45$). By contrast, African American children who were offered the chance to participate in one of the programs gain more than their Hispanic counterparts in *non-cognitive* skills as a result of the intervention. The average total program impact (*Total*) on the non-cognitive test score of an African American treatment child is 1.99σ , which is significantly larger than 0.93σ , the corresponding effect on a Hispanic treatment child ($p = 0.07$). Importantly, disregarding spillover effects and evaluating CHECC by naively differencing the outcomes between control and treatment children results in a serious misrepresentation of these findings. Our estimates, presented in the last row, suggest that ignoring spillovers would have led us to conclude that whereas Hispanic children benefited as a result of the intervention, African Americans gained nothing. Note that this conclusion was in fact the one Fryer, Levitt, and List (2015) reached in their evaluation of the parent-program components of the intervention. The above exercise illustrates how failing to account for spillover effects can lead to severe underestimation of a program impact, which could result in inaccurate policy conclusions.

8 Conclusions

The standard approach in evaluating early education programs is to randomly assign children to treatment and control groups and report the difference in outcomes as the program impact. This approach is based on the assumption that the program does not indirectly alter the outcomes of the control group; that is there are no spillover effects.

We provide the first empirical evidence on the spillover effects from a large-scale early childhood intervention. We document significant spatial spillover effects and find that living close to treated families improves the outcomes of control children. These effects are large: On average, the intervention increases the cognitive and non-cognitive test scores of a control child by 0.75σ and 1.25σ , respectively. We find these spillover effects are larger for African Americans than Hispanics. Our evidence suggests non-cognitive spillovers are more likely to operate through children's rather than parents' social networks. We also find suggestive evidence that cognitive spillover effects are generated through influencing the parents' decision to enroll their child in a preschool program.

Our findings have important policy implications for evaluating the impact of early education programs and for interpreting the estimated effects, when such programs are taken to scale. Our findings suggest the traditional measures of early education impacts, which ignore the externalities

⁴⁵The corresponding p-values for cognitive and non-cognitive scores are 0.006 and 0.17.

to the control group, are likely to be too pessimistic. Our findings suggest that disregarding the spillover effects from a large-scale early education intervention can result in a severe underestimation of the program impact. Accounting for spillover effects is even more important for understanding the impacts when a program is scaled up and the level of exposure to direct beneficiaries from the program increases.

Given the importance of non-cognitive skills in children's future labor market and educational outcomes (Heckman, Stixrud, and Urzua 2006; Borghans et al., 2008; Brunello and Schlotter, 2011), our findings provide practical insights into designing early interventions to better foster such skills. Specifically, our results suggest that interventions that promote social interactions both within participants and between participants and non-participants are likely to generate larger positive externalities on non-cognitive skills.

References

- [1] Abramitzky, Ran and Lavy, Victor and Pérez, Santiago (2018) “The Long-Term Spillover Effects of Changes in the Return to Schooling,” *National Bureau of Economic Research*.
- [2] Al-Ubaydli, O., List, J. A. and Suskind, D. (2017a), “What Can We Learn From Experiments? Understanding the Threats to the Scalability of Experimental Results,” Natural Field Experiments 00607, The Field Experiments Website.
- [3] Al-Ubaydli, O., List, J. A. LoRe, D. and Suskind, D. (2017b), “Scaling for Economists: Lessons from the Non-Adherence Problem in the Medical Literature,” *Journal of Economic Perspectives*, American Economic Association, 31(4), 125–144.
- [4] Almlund, M., Duckworth, A.L., J. Heckman and T. Kautz (2011), “Personality psychology and economics,” In: *Handbook of the Economics of Education, Vol. 4*, Amsterdam: Elsevier, pp. 1–181.
- [5] Angelucci, M., & De Giorgi, G. (2009). “Indirect effects of an aid program: how do cash transfers affect ineligibles’ consumption?” *American Economic Review*, 99(1), 486–508.
- [6] Angrist, J. D., and Pischke, J. S. (2008). “Mostly harmless econometrics: An empiricist’s companion.” Princeton university press.
- [7] Arzaghi, M. and J.V. Henderson (2008), “Networking off Madison avenue,” *Review of Economic Studies* 75(4), 1011–1038.
- [8] Åslund, O., Osth, J. and Y. Zenou (2010), “How crucial is distance to jobs for ethnic minorities? Old question – Improved answer,” *Journal of Economic Geography* 10, 389-422.
- [9] Avitabile, C. (2012). “Spillover effects in healthcare programs: evidence on social norms and information sharing” (No. IDB-WP-380). *IDB Working Paper Series*.
- [10] Aizer, A., & Currie, J. (2004). “Networks or neighborhoods? Correlations in the use of publicly-funded maternity care in California.” *Journal of public Economics*, 88(12), 2573–2585.
- [11] Barnett, W. S., (2011) “Effectiveness of Early Educational Intervention,” *Science*, 333, 975–978.
- [12] Bayer, P., Ross, S. L. and G. Topa (2008), “Place of work and place of residence: Informal hiring networks and labor market outcomes,” *Journal of Political Economy* 116(6), 1150–1196.
- [13] Bisztray, M., Koren, M., and A. Szeidl (2017), “Learning to import from your peers,” Unpublished manuscript, Central European University.

- [14] Bobba, M., & Gignoux, J. (2014). “Neighborhood effects and take-up of transfers in integrated social policies: Evidence from Progresa.”
- [15] Bobonis, G. J., & Finan, F. (2009). “Neighborhood peer effects in secondary school enrollment decisions.” *The Review of Economics and Statistics*, 91(4), 695–716.
- [16] Boning, W. C., Guyton, J., Hodge, R. H., Slemrod, J., & Troiano, U. (2018). “Heard it Through the Grapevine: Direct and Network Effects of a Tax Enforcement Field Experiment” (No. w24305). *National Bureau of Economic Research*.
- [17] Borghans, L., Duckworth, A.L., Heckman,J.J., and B. ter Weel (2008), “The economics and psychology of personality traits,” *Journal of Human Resources* 43, 972–1059.
- [18] Brunello, G., & Schlotter, M. (2011). “Non-cognitive skills and personality traits: Labour market relevance and their development in education & training systems” .
- [19] Carneiro, P. M., & Heckman, J. J. (2003). “Human capital policy.”
- [20] Chetty, R., & Hendren, N. (2018-a). The impacts of neighborhoods on intergenerational mobility I: Childhood exposure effects. *The Quarterly Journal of Economics*, 133(3), 1107-1162.
- [21] Chetty, R., & Hendren, N. (2018-b). The impacts of neighborhoods on intergenerational mobility II: County-level estimates. *The Quarterly Journal of Economics*, 133(3), 1163-1228.
- [22] Chetty, R., Hendren, N. and L.F. Katz (2016), “The effects of exposure to better neighborhoods on children: New evidence from the Moving to Opportunity experiment,” *American Economic Review* 106(4), 855–902.
- [23] Chyn, E. (2017). “Moved to Opportunity: The Long-Run Effect of Public Housing Demolition on Labor Market Outcomes of Children.”
- [24] Cook, J. L., and Cook, G. L. (2009). “Child development: Principles and perspectives.” Boston: Pearson A&B.
- [25] Crépon, B., Duflo, E., Gurgand, M., Rathelot, R.,& Zamora, P. (2013). “Do labor market policies have displacement effects? Evidence from a clustered randomized experiment.” *The Quarterly Journal of Economics*, 128(2), 531–580.
- [26] Cunha, F., & Heckman, J. (2007). “The technology of skill formation.” *American Economic Review*, 97(2), 31–47.
- [27] Cunha, F., Heckman,J.J., and S.M. Schennach (2010), “Estimating the technology of cognitive and noncognitive skill formation,” *Econometrica* 78, 883–931.

- [28] Currie, J., & Almond, D. (2011). “Human capital development before age five.” *In Handbook of labor economics*, 4, 1315–1486. Elsevier.
- [29] Damm, A.P. (2009), “Ethnic enclaves and immigrant labor market outcomes: Quasi-experimental evidence,” *Journal of Labor Economics* 27, 281–314.
- [30] Damm, A. P., & Dustmann, C. (2014). “Does growing up in a high crime neighborhood affect youth criminal behavior?.” *American Economic Review*, 104(6), 1806–32.
- [31] Dunn, L. M., Dunn, L. M., Bulheller, S., & Häcker, H. (1965). “Peabody picture vocabulary test.” *Circle Pines, MN: American Guidance Service*.
- [32] Duflo, E., & Saez, E. (2003). “The role of information and social interactions in retirement plan decisions: Evidence from a randomized experiment.” *The Quarterly journal of economics*, 118(3), 815–842.
- [33] Edin, P.-A., Fredriksson, P., Åslund, O. (2003), “Ethnic enclaves and the economic success of immigrants. Evidence from a natural experiment,” *Quarterly Journal of Economics* 118, 329–357.
- [34] Entwistle, D. R., Alexander, K. L., & Olson, L. S. (1994). “The gender gap in math: Its possible origins in neighborhood effects.” *American Sociological Review*, 822–838.
- [35] Epple, D., & Romano, R. E. (2011). “Peer effects in education: A survey of the theory and evidence.” In *Handbook of social economics*, 1, 1053–1163. North-Holland.
- [36] Ferracci, M., Jolivet, G., & van den Berg, G. J. (2010). “Treatment evaluation in the case of interactions within markets.”
- [37] Fryer, R.G. Jr. and L.F. Katz (2013), “Achieving escape velocity: Neighborhood and school interventions to reduce persistent inequality,” *American Economic Review: Papers & Proceedings* 103(3), 232–237.
- [38] Fryer Jr, R.G., Levitt, S.D. and List, J.A., (2015), “Parental incentives and early childhood achievement: a field experiment in Chicago heights,” National Bureau of Economic Research Working Paper No. 21477.
- [39] Fryer Jr, R.G., Levitt, S.D., List, J.A., and Samek, A. (2018), “Towards an Understanding of What Works in Preschool Education,” Unpublished manuscript, University of Chicago.
- [40] Gautier, P., Muller, P., Klaauw, B. V. D., Rosholm, M., & Svarer, M. (2012). “Estimating equilibrium effects of job search assistance.”

- [41] Gibbs, C., Ludwig, J., & Miller, D., (2011). “Does Head Start Do Any Lasting Good?” *NBER Working Paper* No. 17452.
- [42] Giné, X., & Mansuri, G. (2018). “Together we will: experimental evidence on female voting behavior in Pakistan.” *American Economic Journal: Applied Economics*, 10(1), 207–35.
- [43] Halpern-Felsher, B.L., Connell, J.P., Spencer, M.B., Aber, J.L., Duncan, G.J., Clifford, E., Crichlow, W.E., Usinger, P.A., Cole, S.P., Allen, L. and Seidman, E., (1997). “Neighborhood and family factors predicting educational risk and attainment in African American and White children and adolescents.” *Neighborhood poverty*, 1, 146–173.
- [44] Hartup, W. W. (1984). “The Peer Context In Middle Childhood.” In *Development During Middle Childhood: The Years From Six to Twelve*. National Research Council.
- [45] Heckman, J.J. and T. Kautz (2012), “Hard evidence on soft skills,” *Labour Economics* 19, 451–464.
- [46] Heckman, J.J., Malofeeva, L., Pinto, R. and P.A. Savelyev (2011), “Understanding the mechanisms through which an influential early childhood program boosted adult outcomes,” Unpublished manuscript, Department of Economics, University of Chicago.
- [47] Heckman, J.J., Moon, S.H., Pinto, R., Savelyev, P.A. and A. Yavitz (2010), “The rate of return to the High Scope Perry Preschool Program,” *Journal of Public Economics* 94, 114–128.
- [48] Heckman, J.J., Stixrud, J. and S. Urzua (2006), “The effects of cognitive and noncognitive abilities on labor-market outcomes and social behavior,” *Journal of Labor Economics* 24, 411–482.
- [49] Hellerstein, J.K., McInerney, M. and D. Neumark (2011), “Neighbors and coworkers: The importance of residential labor market networks,” *Journal of Labor Economics* 29, 659–695.
- [50] Higgins, E.T. and Parsons, J.E. (1983) “Social cognition and the social life of the child: Stages as subcultures” in E.T. Higgins, D.N. Ruble, W.W. Hartup (Eds.), *Social cognition and social development: A socio-cultural perspective*, Cambridge University Press, New York (1983), 15–62
- [51] Ioannides, Y.M. (2011), “Neighborhood effects and housing,” In: J. Benhabib, A. Bisin, and M.O. Jackson (Eds.), *Handbook of Social Economics*, Vol. 1B, Amsterdam: Elsevier Science, 1281–1340.
- [52] Ioannides, Y.M. and G. Topa (2010), “Neighborhood effects: Accomplishments and looking beyond them,” *Journal of Regional Science* 50, 343–362.

- [53] Janssens, W. (2011). “Externalities in program evaluation: The impact of a women’s empowerment program on immunization.” *Journal of the European Economic Association*, 9(6), 1082–1113.
- [54] Kalil, A. (2015). “Inequality Begins at Home: The Role of Parenting in the Diverging Destinies of Rich and Poor Children.” In Families in an Era of Increasing Inequality, *Springer International Publishing*. 63–82.
- [55] Katz, L.F., Kling, J.R. and J.B. Liebman (2001), “Moving to opportunity in Boston: Early results of a randomized mobility experiment,” *Quarterly Journal of Economics* 116, 607–654.
- [56] Kautz, T., Heckman, J. J., Diris, R., Ter Weel, B., & Borghans, L. (2014). “Fostering and measuring skills: Improving cognitive and non-cognitive skills to promote lifetime success.” *National Bureau of Economic Research*. (No. w20749).
- [57] Kim, J., Alderman, H., & Orazem, P. F. (1999). “Can private school subsidies increase enrollment for the poor? The Quetta Urban Fellowship Program.” *The World Bank Economic Review*, 13(3), 443-465.
- [58] Kline, P., and Walters, C. R. (2016). “Evaluating public programs with close substitutes: The case of Head Start.” *The Quarterly Journal of Economics*, 131(4), 1795–1848.
- [59] Kling, J.R., Ludwig, J. and L.F. Katz (2005), “Neighborhood effects on crime for female and male youth: Evidence from a randomized housing voucher experiment,” *Quarterly Journal of Economics* 120, 87–130.
- [60] Kling, J.R., Liebman, J.B. and L.F. Katz (2007), “Experimental analysis of neighborhood effects,” *Econometrica* 75, 83–119.
- [61] Kremer, Michael, Edward Miguel, and Rebecca Thornton (2009), “Incentives to Learn.” *Review of Economics and Statistics* 91 (1): 437- 456.
- [62] Lalive, R., & Cattaneo, M. A. (2009). “Social interactions and schooling decisions.” *The Review of Economics and Statistics*, 91(3), 457–477.
- [63] Lalive, R., Landais, C., & Zweimüller, J. (2015). “Market externalities of large unemployment insurance extension programs.” *American Economic Review*, 105(12), 3564–96.
- [64] Leventhal, T., & Brooks-Gunn, J. (2000). “The neighborhoods they live in: the effects of neighborhood residence on child and adolescent outcomes.” *Psychological bulletin*, 126(2), 309.
- [65] Miguel, E. and M. Kremer (2004), “Worms: identifying impacts on education and health in the presence of treatment externalities,” *Econometrica* 72, 159–217.

- [66] Minh, A., Muhajarine, N., Janus, M., Brownell, M., & Guhn, M. (2017). “A review of neighborhood effects and early child development: How, where, and for whom, do neighborhoods matter?” *Health & place*, 46, 155–174.
- [67] Muralidharan, K., Niehaus, P., & Sukhtankar, S. (2017). “General equilibrium effects of (improving) public employment programs: Experimental evidence from India” (No. w23838). *National Bureau of Economic Research*.
- [68] Oreopoulos, P. (2003), “The long-run consequences of living in a poor neighborhood,” *Quarterly Journal of Economics* 118, 1533–1575.
- [69] Puma, M., Bell, S., Cook, R., Heid, C., (2010) “Head Start Impact Study: Final Report” Washington, DC: U.S. Department of Health and Services, Administration for Children and Families,
- [70] Puma, M., Bell, S., Cook, R., Heid, C., Broene, P., Jenkins, F., Mashburn, A. & Downer, J. (2012). “Third Grade Follow-up to the Head Start Impact Study Final Report” OPRE Report # 2012-45, Washington, DC: Office of Planning, Research and Evaluation, Administration for Children and Families, U.S. Department of Health and Human Services.
- [71] Rincke, Johannes and Traxler, Christian, (2011), “Enforcement Spillovers,” *The Review of Economics and Statistics*, 93(4), 1224–1234.
- [72] Rosenthal, S. S. and W. C. Strange (2008), “The attenuation of human capital spillovers” *Journal of Urban Economics* 46(2), 373–389.
- [73] Sacerdote, B. (2011). “Peer effects in education: How might they work, how big are they and how much do we know thus far?” In *Handbook of the Economics of Education*. 3, 249–277. Elsevier.
- [74] Sanbonmatsu, L., Kling, J.R., Duncan, G.J. and J. Brooks-Gunn (2006), “Neighborhoods and academic achievement: Results from the Moving to Opportunity experiment,” *Journal of Human Resources* 41, 649–691.
- [75] Sinclair, B., McConnell, M., & Green, D. P. (2012). “Detecting spillover effects: Design and analysis of multilevel experiments.” *American Journal of Political Science*, 56(4), 1055–1069.
- [76] Waldfogel, J., & Washbrook, E. (2011). “Early years policy.” *Child Development Research*.
- [77] Willoughby, Michael T., R. J. Wirth, and Clancy B. Blair. (2012), “Executive function in early childhood: longitudinal measurement invariance and developmental change.” *Psychological assessment*, 24(2), 418.

- [78] Woodcock, Richard W., K. S. McGrew, and N. Mather. (2001), *Woodcock-Johnson tests of achievement.*, Itasca, IL: Riverside Publishing.

Spillover Effects in Education

Appendix

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A Assessment Schedule

Table A.1: Assessment Calendar

	Winter (Jan-Feb)	Early Spring (Apr)	Late Spring (May)	Late Summer (Jul-Sep)
2010				<i>Pre</i> ²⁰¹⁰
2011	<i>Mid</i> ²⁰¹⁰		<i>Post</i> ²⁰¹⁰ <i>Pre</i> ²⁰¹¹	<i>SL</i> ²⁰¹⁰
2012	<i>Mid</i> ²⁰¹¹	<i>AO1</i> ²⁰¹⁰	<i>Post</i> ²⁰¹¹ <i>Pre</i> ²⁰¹²	<i>SL</i> ²⁰¹¹
2013	<i>Mid</i> ²⁰¹²	<i>AO1</i> ²⁰¹¹ <i>AO2</i> ²⁰¹⁰	<i>Post</i> ²⁰¹² <i>Pre</i> ²⁰¹³	<i>SL</i> ²⁰¹²
2014	<i>Mid</i> ²⁰¹³	<i>AO1</i> ²⁰¹² <i>AO2</i> ²⁰¹¹ <i>AO3</i> ²⁰¹⁰	<i>Post</i> ²⁰¹³	<i>SL</i> ²⁰¹³
2015		<i>AO1</i> ²⁰¹³ <i>AO2</i> ²⁰¹² <i>AO3</i> ²⁰¹¹ <i>AO4</i> ²⁰¹⁰		
2016		<i>AO2</i> ²⁰¹³ <i>AO3</i> ²⁰¹² <i>AO4</i> ²⁰¹¹ <i>AO5</i> ²⁰¹⁰		

Notes: Superscripts are cohort identifiers (randomization years).

Pre= pre assessment; Mid= mid assessment; Post= post assessment; SL= summer loss assessment; AOx= age-out assessment x years after the treatment ended

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B Maps

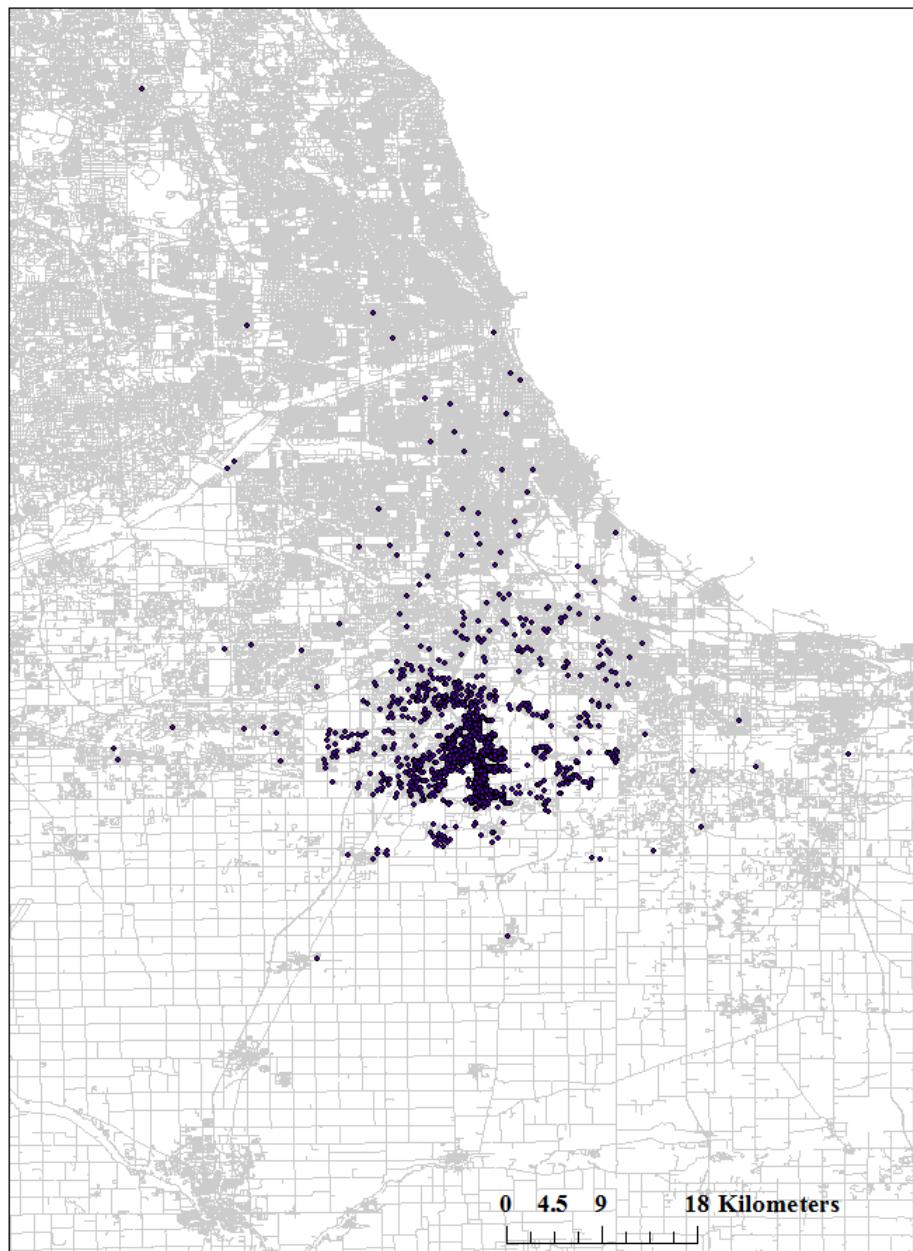


Figure B.1: Map of Home Locations

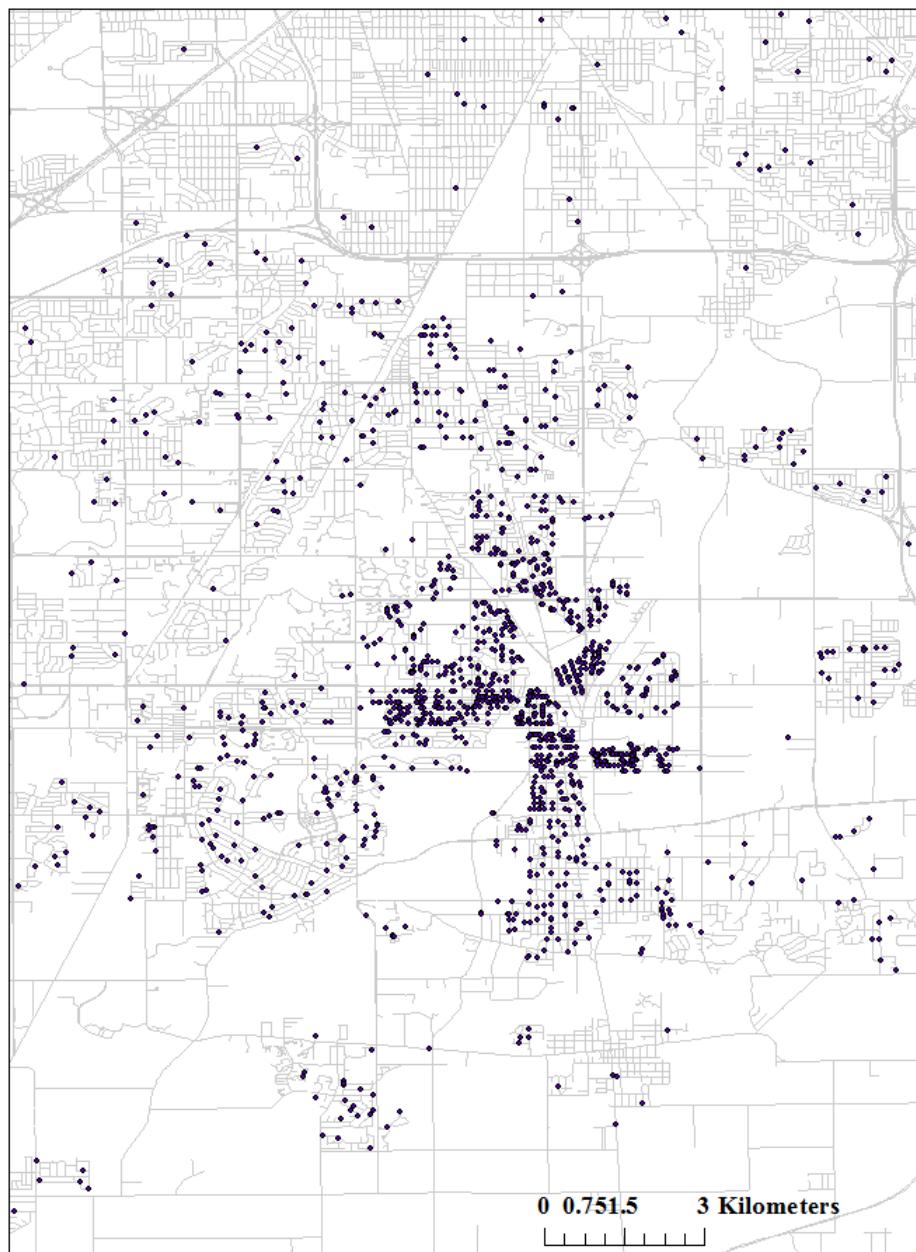


Figure B.2: Map of Home Locations, Zoomed in

C Details on Constructing the Samples

Here we present how we treat observations from children who were randomized in multiple years in constructing our panel data set for the control, treatment, and pooled samples.

We follow two rules in constructing the control sample: (i) For those control children who were randomized into treatments in later years, we only keep the observations that took place before their treatment started; (ii) for those who were randomized into the control group in more than one year, we only keep the observations corresponding to their first randomization.

In construction of our treatment sample, we follow two rules: (i) For those treatment children who were first randomized into the control and later into a treatment group, we only keep the observations after their treatment started; (ii) for those who were randomized twice and both times into a treatment group, we only keep the observations corresponding to the first randomization.

Finally, we use the following three rules in constructing our pooled sample: (i) For children who were randomized twice and both times into the control group, we only keep the observations that correspond to the first randomization; (ii) for children who were randomized twice and both times into a treatment group, we only keep the observations corresponding to their first randomization; and (iii) for those who were randomized twice, the first time into the control group and the second time into a treatment group, we only keep the observations corresponding to the second randomization.

Three factors can result in missing an observation for a child in our control sample: (*i*) The child was absent on the assessment day; (*ii*) the child was moved to a treatment group in a later randomization and thus her outcomes (for the times after she had entered the treatment groups) are not included in the sample; or (*iii*) the child belongs to later cohorts for which the assessment is taking place at a later date (April 2018 or after). Similarly, an observation from the treatment sample would be missing if (*i*) the child was absent on that assessment day; (*ii*) the child was previously in the control group and thus her outcomes (for the times before she entered the treatment group) are not included in the sample; or (*iii*) corresponding assessment is taking place at a later date (April 2018 or later).

D Sub-tests

The cognitive index was constructed by taking the average of a child's percentile scores on five subtests of the Peabody Picture Vocabulary Test (PPVT), WJ Letter-Word, WJ Spelling, WJ Applied Problems, and WJ Quantitative Concepts (the four subtests from the Woodcock Johnson III Test of Achievement). These scores were then standardized by assessment type (pre, mid, post,

etc.), including the entire study population. The non-cognitive index was constructed by taking the average of a child's percent-correct scores in the three subtests of the Blair and Willoughby (Spatial Conflict, Operation Span, and Same Game) and the Preschool Self-Regulation Assessment (PSRA) and standardizing with respect to the assessment type.⁴

To explore which components of the cognitive and non-cognitive measures are more important in generating the spillover effects, we estimate the following fixed-effects model:

$$Y_{i,t}^k = \beta_0 + \beta_1 N_{i,t,r}^{treated} + \beta_2 N_{i,t,r}^{total} + \gamma_i + \delta_t + \epsilon_{i,t}, \quad (\text{D.1})$$

where $Y_{i,t}^k$ is the standardized score of a child i at time t on subtest k , and $N_{i,t,r}^{treated}$ and $N_{i,t,r}^{total}$ represent the number of treated neighbors and total number of neighbors, as previously defined. We include the time and individual fixed effects, and cluster standard errors at the census-block-group level. Under this specification, β_1 represents the average spillover effect from an additional treated neighbor who resides within a radius r of a child.

Table D.1: Spillover Effects by Subtests—Fixed Effects

r (meters)	Cognitive Sub-scores					Non-cognitive Sub-scores				
	PPVT	WJL	WJA	WJS	WJQ	OSP	SPAT	PSRA	SAME	CARD
3000	.0025*	.0044***	-.0057***	.0044***	.0018	.0093***	.0142***	.0002	-.0052	.0082***
	(.0014)	(.001)	(.0011)	(.0011)	(.0019)	(.0026)	(.0031)	(.0021)	(.0128)	(.0023)
5000	.0014	.0026***	-.003***	.0031***	.0009	.0055***	.0089***	.0003	-.0105	.0051***
	(.0009)	(.0007)	(.0009)	(.0007)	(.001)	(.0014)	(.0019)	(.0013)	(.009)	(.0013)
7000	.0012*	.002***	-.0026***	.0025***	.0007	.0041***	.0071***	.0001	-.0077	.0045***
	(.0007)	(.0006)	(.0008)	(.0006)	(.0009)	(.0011)	(.0015)	(.001)	(.0078)	(.0011)

Notes: This table presents point estimates for the spillover effects from an additional treated neighbor on a child's standardized scores on each subtest, estimated from the fixed-effects model. Robust standard errors, clustered at the census-block-group level, are in parentheses;

*** p<0.01, ** p<0.05, * p<0.1

As point estimates for β_1 presented in Table D.1 suggest, the spillovers on cognitive skills are mainly driven by the effects on receptive vocabulary (PPVT), spelling ability (WJS), and the ability to identify letters and words (WJL). The spillovers on non-cognitive skills are driven by the effects on attention (Spatial Conflict), working memory (Operation Span), and attention shifting in children who are kindergarten age or older.

⁴For children in kindergarten or older, the Same Game test of Blair and Willoughby was replaced with a variant of Wisconsin Card Sort game.

E Spatial Fade-out

Table E.1: Mean Effect Sizes on Cognitive and Non-cognitive Scores, Fixed-Effects Estimates

	Cognitive Scores			Non-cognitive Scores		
	Pooled	Control	Treatment	Pooled	Control	Treatment
	(1)	(2)	(3)	(4)	(5)	(6)
r = 2 km	0.0050*** (0.0019)	0.0049** (0.0024)	0.0021 (0.0016)	0.0120*** (0.0025)	0.0077*** (0.0030)	0.0107*** (0.0023)
r = 3 km	0.0033*** (0.0010)	0.0038*** (0.0012)	0.0016 (0.0010)	0.0078*** (0.0013)	0.0069*** (0.0015)	0.0064*** (0.0013)
r = 4 km	0.0026*** (0.0007)	0.0032*** (0.0011)	0.0010* (0.0006)	0.0053*** (0.0009)	0.0045*** (0.0013)	0.0042*** (0.0010)
r = 5 km	0.0021*** (0.0006)	0.0023*** (0.0008)	0.0010* (0.0006)	0.0043*** (0.0008)	0.0037*** (0.0011)	0.0034*** (0.0008)
r = 6 km	0.0019*** (0.0006)	0.0022*** (0.0008)	0.0010* (0.0005)	0.0038*** (0.0007)	0.0031*** (0.0010)	0.0032*** (0.0008)
r = 7 km	0.0018*** (0.0005)	0.0021*** (0.0007)	0.0008* (0.0005)	0.0033*** (0.0007)	0.0025*** (0.0010)	0.0028*** (0.0007)
r = 8 km	0.0017*** (0.0005)	0.0020*** (0.0007)	0.0008* (0.0005)	0.0031*** (0.0006)	0.0021** (0.0009)	0.0028*** (0.0007)
r = 9 km	0.0015*** (0.0005)	0.0018*** (0.0007)	0.0007 (0.0005)	0.0029*** (0.00060)	0.0021** (0.0009)	0.0025*** (0.0007)
r = 10 km	0.0014*** (0.0005)	0.0019*** (0.0006)	0.0006 (0.0005)	0.0027*** (0.0006)	0.0020** (0.0009)	0.0023*** (0.0006)
	5,208	2,442	3,074	5,208	2,442	3,074

Notes: Columns 1-3 (4-6) represent the effect of an additional treated neighbor on a child's standardized cognitive (non-cognitive) score. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

Table E.2: Mean Effect Sizes on Cognitive and Non-cognitive Scores, LDV Estimates

r (meters)	Cognitive Scores			Non-cognitive Scores		
	Pooled (1)	Control (2)	Treatment (3)	Pooled (4)	Control (5)	Treatment (6)
r = 2 km	0.0034** (0.0017)	0.0026 (0.0039)	0.0015 (0.0028)	0.0095*** (0.0030)	0.0061** (0.0031)	0.0102** (0.0044)
r = 3 km	0.0042*** (0.0012)	0.0052** (0.0025)	0.0019 (0.0020)	0.0070*** (0.0015)	0.0059*** (0.0018)	0.0067** (0.0026)
r = 4 km	0.0030*** (0.0007)	0.0039** (0.0018)	0.0014 (0.0012)	0.0059*** (0.0013)	0.0037*** (0.0011)	0.0062*** (0.0023)
r = 5 km	0.0033*** (0.0009)	0.0038** (0.0016)	0.0021* (0.0011)	0.0059*** (0.0014)	0.0040*** (0.0013)	0.0060*** (0.0023)
r = 6 km	0.0030*** (0.0007)	0.0033** (0.0013)	0.0017 (0.0011)	0.0060*** (0.0016)	0.0035*** (0.0013)	0.0058** (0.0023)
r = 7 km	0.0027*** (0.0007)	0.0034*** (0.0013)	0.0015 (0.0010)	0.0054*** (0.0014)	0.0038*** (0.0012)	0.0053** (0.0022)
r = 8 km	0.0022*** (0.0006)	0.0027** (0.0011)	0.0010 (0.0010)	0.0052*** (0.0014)	0.0035*** (0.0012)	0.0050** (0.0020)
r = 9 km	0.0021*** (0.0006)	0.0028** (0.0011)	0.0010 (0.0008)	0.0045*** (0.0013)	0.0035*** (0.0011)	0.0042** (0.0019)
r = 10 km	0.0021*** (0.0006)	0.0026** (0.0010)	0.0011 (0.0008)	0.0043*** (0.0012)	0.0031*** (0.0011)	0.0039** (0.0018)
Obs.	3,403	1,495	2,093	3,403	1,495	2,093

Notes: Columns 1-3 (4-6) represent the effect from an additional treated neighbor on a child's standardized cognitive (non-cognitive) score. Robust standard errors, clustered at the census-block-group, are level in parentheses; *** p<0.01, ** p<0.05, * p<0.1

F Spillovers from Parent Academy, Preschool, and Cog-X Treatments

In section 6, we compared the spillover effects from the Parent Academy treatments, which exclusively offered education program for parents, to the four Pre-K treatments, which offered pre-school programs to children. Two of the four Pre-K treatment groups (Preschool-Plus and Kinderprep), also included parental components, which, compared to Parent Academies, were shorter and not incentivized. We refer to these two treatments as Cog-X treatments. In this section, instead of estimating the overall spillover effects from Pre-K treatments, we separately estimate the effects from Cog-X to the ones from the preschool treatments, which exclusively targeted children, using the following specification:

$$Y_{i,t} = \beta_0 + \beta_{parent} N_{i,t,r}^{Parent} + \beta_{parent_child} N_{i,t,r}^{Cogx} + \beta_{child} N_{i,t,r}^{Preschool} + \lambda N_{i,t,r}^{total} + \eta Y_{i,t-1} + X_i \alpha + \sigma_b + \mu_c + \delta_t + \epsilon_{i,t}, \quad (F.1)$$

where $N_{i,t,r}^{Parent}$, $N_{i,t,r}^{Cogx}$, and $N_{i,t,r}^{Preschool}$ represent the number of neighbors residing within distance r of a child i who were assigned to Parent Academy, Cog-X , and the two preschool treatments that

exclusively targeted children. All other arguments are defined as in section 6. Under the above specifications, β_{parent} , β_{parent_child} , and β_{child} represent the spillover effects from an additional treated neighbor who was assigned to Parent Academies, Cog-X, or the two preschool treatments with no parental components. Table F.1 presents the estimates of $\hat{\beta}_{parent}$, $\hat{\beta}_{parent_child}$ and $\hat{\beta}_{child}$.

Table F.1: Mean Effect Sizes on Cognitive and Non-cognitive Scores

	Cognitive Scores			Non-cognitive Scores		
	β_{parent} (1)	β_{parent_child} (2)	β_{child} (3)	β_{parent} (4)	β_{parent_child} (5)	β_{child} (6)
r = 3 km	0.0045** (0.0022)	0.0040*** (0.0013)	0.0040 (0.0026)	-0.0010 (0.0035)	0.0105*** (0.0023)	0.0132*** (0.0029)
r = 5 km	0.0074*** (0.0024)	0.0034*** (0.0009)	0.0005 (0.0024)	0.0007 (0.0045)	0.0067*** (0.0015)	0.0079* (0.0044)
r = 7 km	0.0074*** (0.0026)	0.0029*** (0.0007)	-0.0006 (0.0025)	0.0006 (0.0040)	0.0056*** (0.0013)	0.0074 (0.0045)
Obs.	3,403	3,403	3,403	3,403	3,403	3,403

Notes: Columns 1-3 (4-6) represent the effect of an additional treated neighbor of each type, on a child's standardized cognitive (non-cognitive) score. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

Our estimates suggest the programs that included parental-education components are likely to generate larger cognitive spillovers than those that exclusively targeted children.⁵ Focusing on non-cognitive spillovers, our estimates confirm our previous findings that programs that directly targeted children (Cog-X and preschool treatments) generate significantly larger non-cognitive spillovers than the ones that exclusively targeted parents ($p < 0.10$).

G Spillover effects from Pre-K versus Parent Academy Neighbors on Control and Treated Children

Table G.1 breaks down the estimated effects ($\hat{\beta}_p$ and $\hat{\beta}_c$), by the treatment assignment of a receiving child for neighborhood radii of 3, 5, and 7 kilometers. The estimates confirm that for both the treatment and control subgroups, the non-cognitive spillovers are larger from the Pre-K neighbors than the Parent Academy neighbors. Spillover effects on cognitive scores tend to be larger from Parent Academy than from Pre-K neighbors.

⁵The differences in spillover effects are insignificant for $r = 3$ km, but as we broaden neighborhood radii to 5 and 7 kilometers, the difference in spillovers on cognitive skills from Cog-X and Parent Academy to the ones from the preschool treatments become significant ($p < 0.10$).

Table G.1: Spillover Effects from Parent Academy and Pre-K Neighbors, on Treatment and Control

(a) Fixed Effects Estimates				(b) LDV estimates					
	Cognitive Scores		Non-cognitive Scores			Cognitive Scores		Non-cognitive Scores	
	β_p	β_c	β_p	β_c		β_p	β_c	β_p	β_c
	(1)	(2)	(3)	(4)		(1)	(2)	(3)	(4)
Effect on Control									
r = 3 km	0.0044**	0.0035***	0.0043***	0.0084***		0.0067*	0.0040*	0.0045	0.0070***
	(0.0017)	(0.0013)	(0.0016)	(0.0021)		(0.0038)	(0.0022)	(0.0027)	(0.0021)
r = 5 km	0.0028**	0.0020**	0.0022*	0.0047***		0.0045	0.0035**	0.0001	0.0044***
	(0.0011)	(0.0008)	(0.0012)	(0.0014)		(0.0027)	(0.0015)	(0.0027)	(0.0012)
r = 7 km	0.0024***	0.0018**	0.0018*	0.0032***		0.0051**	0.0033**	-0.0023	0.0040***
	(0.0009)	(0.0007)	(0.0010)	(0.0012)		(0.0025)	(0.0013)	(0.0032)	(0.0012)
Effect on Treated									
r = 3 km	0.0017	0.0014	0.0027**	0.0095***		0.0017	0.0020	-0.0009	0.0120***
	(0.0011)	(0.0013)	(0.0012)	(0.0020)		(0.0021)	(0.0023)	(0.0023)	(0.0033)
r = 5 km	0.0013**	0.0008	0.0014	0.0049***		0.0034**	0.0018	0.0001	0.0073***
	(0.0006)	(0.0008)	(0.0009)	(0.0012)		(0.0016)	(0.0012)	(0.0025)	(0.0021)
r = 7 km	0.0012**	0.0006	0.0015**	0.0037***		0.0030*	0.0013	0.0006	0.0059***
	(0.0005)	(0.0006)	(0.0008)	(0.0010)		(0.0017)	(0.0010)	(0.0027)	(0.0019)

Notes: Columns 1 and 3 (2 and 4) report the average spillover effect from an additional Parent Academy (Pre-K) neighbor, residing within a distance r of a child, on her standardized cognitive and non-cognitive scores (equations 3 and 4). The estimates are reported separately for the control and treatment groups. Panel (a) presents the fixed-effects estimates and panel (b) reports the LDV estimates. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

H Total Impact Evaluation and Neighborhood Radius

In section 7, we estimated the impact of the intervention under the neighborhood radius of 3 kilometers. Table H.1 presents the corresponding effects for two other neighborhood radii: $r=5$ and $r=7$ kilometers. Our estimates of the direct program effects on treatment children ($\hat{\theta}$) are not significantly different across various neighborhood radii.⁶

As we broaden the definition of the neighborhood, the estimated total spillover effect on both the control (SC^N) and treatment (ST^N) children increases. This finding is intuitive, because broadening the neighborhood would allow for neighbors who live farther away to also impact a child's outcomes. The increases in the total spillover effects result in larger estimates of the total impacts (*Total*) as we increase the neighborhood radius. The average estimated total impact of the intervention (*Total*) on a treatment child's cognitive scores is 0.82σ for $r = 3$ km and increases to 1.16σ and 1.24σ for $r = 5$ km and $r = 7$ km. Likewise, the average estimated total impact of the intervention (*Total*) on the non-cognitive score of a child increases from 1.32σ to 1.97σ and

⁶P-values from the Wald tests of $\hat{\theta}_{r1}^{Cog} = \hat{\theta}_{r2}^{Cog}$ against $\hat{\theta}_{r1}^{Cog} \neq \hat{\theta}_{r2}^{Cog}$ for r1=3K and r2=5K; r1=3K and r2=7K; and r1=5K and r2=7K are 0.28, 0.12, and 0.08. The corresponding p-values for non-cognitive scores ($\hat{\theta}^{N cog}$) are 0.75, 0.85, and 1.00.

Table H.1: Total Program Impact under r=3, 5, and 7 kilometers, LDV Estimates

$\frac{r}{N_d^{treated}}$	5 km		5 km		7 km	
	178.13		325.63		422.81	
$\hat{\theta}$ (<i>Direct</i>)	Cognitive 0.108*** (0.0393)	Non-cognitive 0.0456 (0.0642)	Cognitive 0.126*** (0.0458)	Non-cognitive 0.0540 (0.0731)	Cognitive 0.150*** (0.0528)	Non-cognitive 0.0541 (0.0837)
$\hat{\beta}$	0.0042*** (0.0017)	0.0070*** (0.0015)	0.0034*** (0.0009)	0.0058*** (0.0014)	0.0028*** (0.0007)	0.0053*** (0.0014)
λ	-0.0002* (0.0001)	0.0001 (0.0002)	-0.0002* (0.0001)	0.0001 (0.0002)	-0.0002* (0.0001)	0.0000 (0.0002)
SC^N	0.75*** (0.22)	1.25*** (0.27)	1.09*** (0.28)	1.90*** (0.46)	1.17*** (0.31)	2.25*** (0.61)
ST^N	0.71*** (0.22)	1.27*** (0.27)	1.03*** (0.27)	1.91*** (0.46)	1.09*** (0.30)	2.27*** (0.61)
<i>Total</i>	0.82*** (0.23)	1.32*** (0.27)	1.16*** (0.28)	1.97*** (0.45)	1.24*** (0.31)	2.32*** (0.60)
<i>Total^{Standard}</i>	0.06*** (0.02)	0.07** (0.03)	0.06*** (0.02)	0.07** (0.03)	0.07*** (0.02)	0.07** (0.03)
Observations	3,403	3,403	3,403	3,403	3,403	3,403
R-squared	0.659	0.384	0.660	0.387	0.660	0.387

Notes: Estimated coefficients from equation 9 are presented; N includes all the observations for which we have the lagged cognitive and non-cognitive scores, and other regressors. $\overline{N_{3km}^{treated}} = 178.13$; $\overline{N_{5km}^{treated}} = 325.63$; $\overline{N_{7km}^{treated}} = 422.81$; Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

2.32σ as we increase the radius from 3 to 5 and 7 kilometers. Finally, our estimates for program impacts if we were to ignore spillovers to control children (*Total^{Standard}*) are very similar and not significantly different across various radii.⁷

I Estimated Effects on a Restricted Sample

Sections 4 and 5 present the estimated spillover effects using all observations. One potential concern is the role of selection and how selection into taking assessments might be deriving our findings. We address this concern by estimating the main effects from equations 1 and 2 for a subset of our sample who take at least five out of the eight possible assessments. Our data include 1,792 observations from 313 children who attend a minimum of five assessments. Note these children represent under 20% of the total number of children in our pooled sample. Table I.1 presents the point estimates. Our estimates are similar in magnitude to the ones from the whole sample. The estimates on non-cognitive spillover effects are especially close to the ones presented in Table 4 and 6. The fixed-effects estimates of β_1 on cognitive spillover effects from the whole sample for neighborhood radii of 3, 5, and 7 kilometers are 0.0033σ , 0.0021σ , and 0.0018σ , respectively, whereas

⁷P-values from the Wald tests of $\hat{T}_{r1}^{Standard-Cog} = \hat{T}_{r2}^{Standard-Cog}$ against $\hat{T}_{r1}^{Standard-Cog} \neq \hat{T}_{r2}^{Standard-Cog}$ for $r1=3$ km and $r2=5$ km; $r1=3$ km and $r2=7$ km; and $r1=5$ km and $r2=7$ km are 0.71, 0.35, and 0.18. The corresponding p-values for non-cognitive scores are 0.99, 0.93, and 0.85.

the corresponding estimates from the restricted sample are 0.0040σ , 0.0021σ , and 0.0018σ . Likewise, the LDV estimates of β_1 on cognitive spillover effects from the whole sample for neighborhood radii of 3, 5, and 7 kilometers are 0.0042σ , 0.0033σ , and 0.0027σ . The corresponding estimates for the restricted subsample are 0.0029σ , 0.0017σ , and 0.0010σ . The fixed effects estimates on non-cognitive spillover effects from the whole sample for neighborhood radii of 3, 5 and 7 kilometers were 0.0069σ , 0.0043σ , and 0.0033σ , whereas the fixed-effects estimates from the restricted sample are 0.0070σ , 0.0036σ , and 0.0027σ . Similarly, the LDV estimates on non-cognitive spillover effects from the whole sample for neighborhood radii of 3, 5, and 7 kilometers are 0.0070σ , 0.0059σ , and 0.0054σ , whereas the corresponding fixed-effects estimates from the restricted sample are 0.0064σ , 0.0050σ , and 0.0045σ .

Table I.1: Estimates of spillover effects on the subsample with a minimum of five assessments

	Fixed Effects		LDV	
	Cognitive	Non-cognitive	Cognitive	Non-cognitive
r = 3 km	0.0040** (0.0011)	0.0070*** (0.0018)	0.0029** (0.0012)	0.0064*** (0.0019)
r = 5 km	0.0021*** (0.0007)	0.0036*** (0.0013)	0.0017* (0.0009)	0.0050*** (0.0015)
r = 7 km	0.0018*** (0.0006)	0.0027*** (0.0010)	0.0010 (0.0007)	0.0045*** (0.0013)
Obs.	1,792	1,792	1400	1400
Unique child	313	313	313	313

Notes: Estimated spillover effects from equations 1 and 2 for a subsample of observations from children who attended a minimum of five out of eight assessments. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

J Exploring Non-linearities in Measuring the Total Spillover Effects

We calculated our estimation of the total spillover effects under the assumption of linearity. In this section, we explore whether and how allowing for nonlinearities affects our estimates. We explore non-linearities by considering polynomial functional forms of up to degree 3.

$$Y_{i,t} = \alpha_0 + \theta T_i + \alpha_1 N_{i,t,r}^{treated} + \alpha_2 (N_{i,t,r}^{treated})^2 + \dots + \lambda_1 T_i \times N_{i,t,r}^{treated} + \lambda_2 T_i \times (N_{i,t,r}^{treated})^2 + \dots + \beta_2 N_{i,t,r}^{total} + \eta Y_{i,t-1} + X_i \alpha + \sigma_b + \delta_t + \mu_c + \epsilon_{i,t,r} \quad (J.1)$$

While under the linear specification, the marginal spillover effect from the j-th treated neighbor is given by α_1 for a control child and by $\alpha_1 + \lambda_1$ for a treated child, the corresponding effects under

polynomials of degrees 2 and 3 are given by $\alpha_1 + 2\alpha_2 j$ and $\alpha_1 + 2\alpha_2 j + 3\alpha_3 j^2$ for a control child, and $\alpha_1 + \lambda_1 + 2(\alpha_2 + \lambda_2)j$ and $\alpha_1 + 2(\alpha_2 + \lambda_2)j + 3(\alpha_3 + \lambda_3)j^2$ for a treated child. Therefore, the average spillover effects on a control child from all neighbors, using polynomials of degrees 1, 2, and 3, can be calculated as follows:⁸

$$\overline{SC_1^N} = \alpha_1 \overline{N^{tr}}$$

$$\overline{SC_2^N} = \alpha_1 \overline{N^{tr}} + 2\alpha_2 \sum_{j=1}^{\overline{N^{tr}}} j = \alpha_1 \overline{N^{tr}} + \alpha_2 \overline{N^{tr}} (\overline{N^{tr}} + 1)$$

$$\overline{SC_3^N} = \alpha_1 \overline{N^{tr}} + 2\alpha_2 \sum_{j=1}^{\overline{N^{tr}}} j + 3\alpha_3 \sum_{j=1}^{\overline{N^{tr}}} j^2 = \alpha_1 \overline{N^{tr}} + \alpha_2 \overline{N^{tr}} (\overline{N^{tr}} + 1) + 0.5\alpha_3 (\overline{N^{tr}} + 1)(2\overline{N^{tr}} + 1).$$

Table J.1: Estimates of spillover effects under non-linear specifications

	Cognitive			Non-cognitive		
	(1)	(2)	(3)	(4)	(5)	(6)
T_i	0.11*** (0.04)	0.08 (0.06)	0.05 (0.07)	0.05 (0.06)	0.12 (0.08)	0.08 (0.11)
$N_{i,t,r}^{treated}$	0.0042*** (0.0013)	0.0031 (0.0019)	0.0023 (0.0019)	0.0070*** (0.0015)	0.0080*** (0.0021)	0.0096*** (0.0028)
$(N_{i,t,r}^{treated})^2$		1.58e-06 (1.52e-06)	6.10e-06 (7.42e-06)		-1.42e-06 (1.39e-06)	-1.08e-05 (9.33e-06)
$(N_{i,t,r}^{treated})^3$			-5.44e-09 (9.01e-09)			1.04e-08 (1.07e-08)
$T_i \times N_{i,t,r}^{treated}$	-0.0002* (0.0001)	0.0001 (0.0006)	0.0007 (0.0014)	0.0001 (0.0002)	-0.0008 (0.0007)	0.0004 (0.0021)
$T_i \times (N_{i,t,r}^{treated})^2$		-7.08e-07 (1.12e-06)	-3.34e-06 (6.72e-06)		1.89e-06 (1.33e-06)	-4.41e-06 (1.04e-05)
$T_i \times (N_{i,t,r}^{treated})^3$			3.36e-09 (8.53e-09)			8.57e-09 (1.34e-08)
SC^N	0.75*** (0.22)	0.60* (0.31)	0.56** (0.29)	1.24*** (0.27)	1.37*** (0.35)	1.41*** (0.37)
ST^N	0.71*** (0.22)	0.55* (0.32)	0.60** (0.29)	1.27*** (0.27)	1.34*** (0.38)	1.39*** (0.38)
Total	0.82*** (0.23)	0.63* (0.34)	0.65** (0.28)	1.32*** (0.27)	1.45*** (0.38)	1.47*** (0.37)
Constant	0.469*** (0.113)	0.478*** (0.114)	0.491*** (0.123)	-2.431*** (0.180)	-2.500*** (0.186)	-2.449*** (0.188)
Observations	3,403	3,403	3,403	3,403	3,403	3,403
R-squared	0.659	0.659	0.660	0.384	0.384	0.385

Notes: Estimated coefficients from equation J.1 for neighborhood radius of $r = 3$ km. Columns 1 and 4 correspond to the linear specification; 2 and 5 correspond to polynomials of degree 2; columns 3 and 6 correspond to polynomials of degree 3. Robust standard errors, clustered at the census-block-group level, are in parentheses; *** p<0.01, ** p<0.05, * p<0.1

Table J.1 presents the estimated coefficients from equation J.1 for neighborhood radius of $r = 3$ km.

⁸Replacing α_i with $\alpha_i + \lambda_i$ ($\forall i \in \{1,2,3\}$) would give us the corresponding average spillover effects on a treatment child.

Note the coefficients of the quadratic and cubic terms are all insignificant, suggesting our linear specification is an appropriate representation. Our estimates of the average cognitive spillover effects from all treated neighbors become slightly smaller as we add quadratic and cubic terms, but the changes are small. The estimated non-cognitive spillover effects become larger as we move away from the linear specification. However, these increases are small. Overall, we find no strong evidence suggesting the spillover effect from an additional treated neighbor has a non-linear relationship with the number of treated neighbors.