# Providing Advice to Job Seekers at Low Cost: An Experimental Study on On-Line Advice.

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#### Abstract

We develop and evaluate experimentally a novel tool that redesigns the job search process by providing tailored advice at low cost. We invited job seekers to our computer facilities for 12 consecutive weekly sessions to search for real jobs on our web interface. For half, instead of relying on their own search criteria, we use readily available labor market data to display relevant alternative occupations and associated jobs. The data indicates that this broadens the set of jobs they consider and increases their job interviews especially for participants who otherwise search narrowly and have been unemployed for a few months.

Keywords: Online job search, occupational breadth, search design.

**JEL Codes**: D83, J62, C93

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#### 1 Introduction

Getting the unemployed back into work is an important policy agenda and a mandate for most employment agencies. In most countries, one important tool is to impose requirements on benefit recipients to accept jobs beyond their occupation of previous employment, at least after a few months. Yet there is little guidance how they should obtain such jobs and how one might advise them in the process. This reflects the large literature on active labor market policies which is predominantly silent about the effective provision of job search advice, where most studies do not distinguish between advice and enforcement. In their meta-study on active labor market policies Card et al. (2010) merge "job search assistance or sanctions for failing to search" into one category. Ashenfelter et al. (2005) assert a common problem that experimental designs "combine both work search verification and a system designed to teach workers how to search for jobs" so that it is unclear which element generates the documented success. Only few studies, reviewed in the next section, have focused exclusively on providing advice, mostly through labor-intensive counselling on multiple aspects of job search. Our study aims to contribute by providing and assessing low-cost, automated occupational advice to job seekers.

Even before evaluating the effects of advice on job search, a prime order question is what advice should be provided and how? In most countries, the provision of advice is usually done by trained advisors who meet job seekers on a regular basis, yet financial constraints often mean that such advice can only be limited in scope. Our first contribution is to propose an innovative low-cost way of providing tailored advice to job seekers online. It has long been argued that occupational information is something job seekers have to learn.<sup>3</sup> Recent evidence both for the US and the UK shows a pronounced occupational mismatch (Sahin et al. (2014), Patterson et al. (2016)): job seekers search in occupations with relatively few available jobs while at the same time other occupations with relatively more jobs are available but attract little interest. This "mismatch" has seen a further persistent increase since the great recession. Incomplete information could be a contributor if job seekers do not fully know which occupations currently have favorable conditions and whether their skills allow them to transit there. The tool we propose aims to address this by suggesting occupations (and shows the jobs that are currently available in them) using an algorithm based on representative labor market statistics. In a nutshell, it recommends additional occupations in which relevant other job seekers have successfully found jobs and where skill transferability is high, and visualises where market tightness is favorable.

Our second contribution is to evaluate how the advice provided through our tool affects job search behavior, i.e., to see if and how job seekers adjust their job search strategies in response to the suggestions they receive and whether this affects job interviews. To do this, we conduct a randomized study in a highly controlled and replicable environment. We recruited job seekers in Edinburgh from

<sup>&</sup>lt;sup>1</sup>See Venn (2012) for an overview of requirements across OECD countries.

<sup>&</sup>lt;sup>2</sup>See the clarification in Card et al. (2009), p. 6.

<sup>&</sup>lt;sup>3</sup> For example, Miller (1984), Neal (1999), Gibbons and Waldman (1999), Gibbons et al. (2005), Papageorgiou (2014) and Groes et al. (2015) highlight implications of occupational learning and provide evidence of occupational mobility consistent with a time-consuming process of gradual learning about the appropriate occupation.

local Job Centres and transformed the experimental laboratory into a job search facility resembling those in "Employability Hubs" which provide computer access to job seekers throughout the city. Participants were asked to search for jobs via our search platform from computers within our laboratory once a week for a duration of 12 weeks. The main advantage of this "field-in-the-lab" approach is that it allows us to obtain a complete picture of the job search process. Not only do we observe participants' activities on the job search platform, such as the criteria they use to search for jobs and which vacancies they consider; but we also collect information via weekly surveys on which jobs they apply for, whether they get interviews and job offers. Furthermore, we also collect information about other search activities that job seekers undertake outside the job search platform, which is important if one is worried that effects on any one search channel might simply induce shifts away from other search channels. This allows us to have measures of total search effort and total job interviews that include such effects. These are key advantages of this approach that complement the alternatives reviewed in the next section: Studies that rely on data from large on-line job search platforms typically do not have information on activities outside the job search platform nor on job search success, and currently lack a randomized design; studies that use administrative data usually only have information about final outcomes (i.e. job found) but know little about the job search process. However, because of the logistics required for our field-in-the-lab setup, our sample is limited to 300 participants. As a twelve week panel this is a large number for experimental work but limited relative to usual labor market studies, with associated limits in terms of power. Since it is the first study on the use of on-line advice, we found that the advantages warranted this approach.

Most of the literature in labor economics focuses on evaluating interventions that have been designed by policy makers or field practitioners. We add to this tradition here, not only by evaluating a novel labor market intervention, but also by leading the design of the intervention itself, using insights from labor economics to integrate existing labor market data right into a job search platform. To our knowledge our study is the first to use the expanding area of online search to provide advice by re-designing the jobs search process on the web, and allows for a detailed analysis of the effects on the job search "inputs" in terms of search and application behavior and the amount of interviews that participants receive.

Internet-based job search is by now one of the predominant ways of searching for jobs. Kuhn and Mansour (2014) document the wide use of the internet. In the UK where our study is based, roughly two thirds of both job seekers and employers now use the internet for search and recruiting (ONS (2013), Pollard et al. (2012)). We set up two search platforms for internet-based job search that access the database of live vacancies of Universal Jobmatch, the official job search platform provided by the UK Department of Work and Pensions, which features a vacancy count at over 80% of the official UK vacancies. One platform replicates "standard" designs where job seekers themselves decide which keywords and occupations to search for, similar to interfaces used on Universal Jobmatch and other commercial job search sites. The second "alternative" platform provides targeted occupational advice. It asks participants which (target) occupation they are looking for - which often coincides with the occupation of previous employment. Then a click of a button provides them with two lists containing the most related occupations. The first is based on common occupational transitions that people

who have worked in the target occupation make and the second contains occupations for which skill requirements are similar to that in the target occupation. Another click then triggers a consolidated query over all jobs that fall in any of these occupations within their geographic area. Participants can also take a look at maps to see in which occupations the ratio of unemployed workers to available jobs is more favorable - but data availability limits this to aggregated occupational groups. The maps provide direct information on the competition for jobs in an occupation, skill transferability provides information on the occupations in which the job seeker has realistic chances to fulfill the needs of a job opening, and information on successful transitions combines both because successful transitions require the availability of jobs in the new occupation and the skills to secure those jobs. The benefit of this intervention is that it provides job search advice in a highly controlled manner based on readily available statistical information, entails only advice and no element of coercion (participants were free to continue with the "standard" interface if they wanted to) and constitutes a low-cost intervention.

Job search occurs precisely because people lack relevant information that is costly and timeconsuming to acquire. The main benefit of the internet is precisely the ability to disseminate information at low cost, and our implementation makes wider occupational exploration easy. We investigate the following hypothesis about its effects: It should lead job seekers to consider a wider set of occupations beyond those they would consider anyhow, at least for those individuals that search only over a narrow set of occupations in the absence of our intervention. This should lead to more job interviews, especially for narrow searchers. For those who already explore many occupations without our intervention, predictions are less clear: if we propose a smaller set of occupations than they consider otherwise, they might stop exploring occupations that we do not feature as they appear less promising. How this affects job interviews depends on how they re-target their job search effort and could potentially reduce job interviews. Regarding the duration of unemployment, those with longer durations might be more open to new suggestions (e.g., if pressure on them to explore more occupations is higher as mentioned in the introductory paragraph) and our intervention has a larger chance to be valuable. While these predictions arise naturally, we provide an illustrative theory model that lays out these considerations in Section 6. We test these predictions relative to the obvious null hypothesis: there will be no effect if the information that we provide is already known to job seekers or if the real problem is incentives to search rather than information problems. Since our information is publicly available, it is conceivable that it is already known to individuals or their advisers at the job centre.

All participants searched with the standard interface for the first three weeks, which provides a baseline on how participants search in the absence of our intervention. After these initial three weeks, half of the participants continue with this interface throughout the study, while the other half was offered to try the alternative interface. We report the overall impact on the treatment group relative to the control group. We also compare treatment and control in particular subgroups of obvious interest: as indicated, our study has more scope to affect people who search narrowly prior to our intervention, and differential effects by duration of unemployment seem to be a particular policy concern. Overall, we find that our intervention exposes job seekers to jobs from a broader set of occupations, increasing our measure of breadth by 0.2 standard deviations which corresponds to the broadening that would occur naturally after an additional three months of unemployment. Job applications become broader,

and the total number of job interviews increases by 44%. These effects are driven predominantly by job seekers who initially search narrowly. They additionally apply closer to home, and experience a two-fold increase in total job interviews (compared to similarly narrow searchers in the control group). Among those, the effects are mostly driven by those with above-median unemployment duration (more than 80 days), for whom the effects on interviews are even larger. Since we collected information on job interviews obtained through other channels, we can assess possible spill-overs. We find positive effects for such other channels overall and within the aforementioned subgroups, which indicates that our information is helpful beyond the search on our particular platform. This re-enforcing effect is in contrast to crowding-out found in studies on monitoring and sanctions where improvements in monitored search activities led to offsetting reductions in other activities (Van den Berg and Van der Klaauw (2006)). In fact, the statistical significant impact on job interviews is driven by significantly larger reported interviews due to search outside the lab; the point estimates for increased interviews due to search in the lab are even larger but due to the lower base rate not significant (except for the group of initially-narrow longerterm-unemployed group where both are significant).

Across a number of robustness checks in terms of empirical specification we find similar overall patterns as in the baseline in terms of point estimates. Significance does depend on the specification and outcome variable. It is rather robust for increased occupational breadth of jobs that people are listing, and for the number of interviews for initially-narrow job seekers. As indicated, we do find heterogeneity in effects. For example, initially-broad job seekers significantly decrease their breadth of occupational search and we find no sign of increased interviews. In fact this group also uses the new interface less. This is in line with models such as Moscarini (2001) where individuals differ in their comparative advantage for searching in multiple occupations, which would provide different incentives to use the new interface. The heterogeneity in adoption and impact provides one reason why overall effects are weaker and lack robustness. While we do not find any significant negative effects of interviews for any subgroup some point estimates remain economically sizeable. This warrants further analysis and caution, and it might be promising to target advice to particular subgroups such as those who otherwise search narrowly and experienced somewhat longer unemployment. This is particularly interesting because targeting could be included directly into an online advice tool. Moreover, if the effects are positive either overall or for a targeted subgroup, the near zero marginal costs of our type of intervention should make it an attractive policy tool. Such a tool could be rolled out on large scale without much burden on the unemployment assistance system.

Yet, any of these conclusions needs to be viewed with caution. Apart from concerns about the power of our study, a true cost-benefit analysis would need further evaluation of effects on job finding probabilities as well as on whether additional jobs are of similar quality (e.g. pay similarly and can be retained for similar amounts of time). On that point, our approach shares similarities with the well-known audit studies (e.g. Bertrand and Mullainathan (2004)) on discrimination. The main outcome variable in these studies is usually the employer call-back rate rather than actual hiring decisions. As we elaborate in Section 5, it is evident that our study was not intended to pick up effects on job

<sup>&</sup>lt;sup>4</sup> Designing the alternative interface cost £20,000, and once this is programmed, rolling it out more broadly would have no further marginal cost of an existing platform such as Universal Johnatch.

finding because of its size compared to the very low baseline rate of job finding. We find indeed no indication of increased job finding - even in point estimates (though also no significant difference in point estimates in job finding compared to the large positive point estimates in job interviews). We acknowledge that this might not only be due to power issues, though. For example, the conversion rates of interviews into jobs in broader occupations could be lower.<sup>5</sup> A larger-scale assessment would be necessary here. Moreover, a broader roll-out in different geographic areas would also be needed to uncover any general equilibrium effects, which could reduce the effects if search by some job seekers negatively affects others, or could boost the effects if firms react to more efficient search with more job creation. Such general equilibrium effects may be important (as highlighted by Crépon et al. (2013) and Gautier et al. (2015)). We hope that future work with conventional large-scale search providers will marry the benefits of our approach with their large sample sizes.

The essence of our findings can be captured in a simple learning theory of job search that is presented in the pan-ultimate section. It also exposes why narrow individuals with slightly longer unemployment duration might be particularly helped by our intervention. In essence, after loosing their job individuals might initially search narrowly because jobs in their previous occupation appear particularly promising. If the perceived difference with other occupations is large, our endorsement of some alternative occupations does not make up for the gap. After a few months, unsuccessful individuals learn that their chances in their previous occupation are lower than expected, and the perceived difference with other occupations shrinks. Now alternative suggestions can render the endorsed occupations attractive enough to be considered. Our intervention then induces search over a larger set of occupations and increases the number of interviews. One can contrast this with the impact on individuals who already search broadly because they find many occupations roughly equally attractive. They can rationally infer that the occupations that we do not endorse are less suitable, and they stop applying there to conserve search effort. Their breadth declines, but effects on job interviews are theoretically ambiguous because search effort is better targeted, which might be the reason for the insignificant effects on job interviews for this group in our empirical analysis.

The subsequent section reviews related literature. Section 3 outlines how our study is set up. Section 4 sets the stage by providing basic descriptives about the job search process and the subject pool, covering also issues of representativeness, sample balance, and attrition. Section 5 assesses the impact of our intervention within our main empirical specification as well as in a number of robustness checks. Section 6 uses a simple model to illustrate the forces that might underlie our findings, and the final section concludes.

#### 2 Related Literature

As mentioned in the introductory paragraph, most studies on job search assistance evaluate a combination of advice and monitoring/sanctions. An example in the context of the UK, where our study is based, is the work by Blundell et al. (2004) that evaluates the Gateway phase of the New Deal for the

<sup>&</sup>lt;sup>5</sup>For example, Moscarini (2001) outlines a model where those who search narrow have particular advantages in those narrow sectors which would not extend equally to search over a broader set of occupations. This might be reflected only in lower interview rates, but could conceivably also affect the conversion rates.

Young Unemployed, which instituted bi-weekly meetings between long-term unemployed youth and a personal adviser to "encourage/enforce job search". The authors establish significant impact of the program through a number of non-experimental techniques, but cannot distinguish whether "assistance or the "stick" of the tougher monitoring of job search played the most important role" [p. 601]. More recently, Gallagher et al. (2015) of the UK government's Behavioral Insights Team undertook a randomized trial in Job Centres that re-focuses the initial meeting on search planning, introduced goal-setting but also monitoring, and included resilience building through creative writing. They find positive effects of their intervention, but cannot attribute it to the various elements.<sup>6</sup> Nevertheless, their study indicates that there might be room for effects of additional information provision as advice within the official UK system is limited since "many claimants' first contact with the job centre focuses mainly on claiming benefits, and not on finding work" (Gallagher et al. (2015)).

Despite the fact that a lack of information is arguably one of the key frictions in labor markets and an important reason for job search, we are only aware of a few studies that exclusively focus on the effectiveness of information interventions in the labor market. Prior to our study the focus has been on the provision of counseling services by traditional government agencies and by new entrants from the private sector. Behaghel et al. (2014) and Krug and Stephan (2013) provide evidence from France and Germany that public counseling services are effective and outperform private sector counseling services. The latter appear even less promising when general equilibrium effects are taken into account (Crépon et al. (2013)). Bennmarker et al. (2013) finds overall effectiveness of both private and public counseling services in Sweden. The upshot of these studies is their larger scale and the access to administrative data to assess their effects. The downside is the large costs that range from several hundred to a few thousand Euro per treated individual, the multi-dimensional nature of the advice and the resulting "black box" of how it is actually delivered and how it exactly affects job search. Our study can be viewed as complementary. It involves nearly zero marginal cost, the type of advice is clearly focused on occupational information, it is standardized, its internet-based nature makes it easy to replicate, and the detailed data on actual job search allow us to study the effects not only on outcomes but also on the search process.

Contemporaneously, Altmann et al. (2015) analyze the effects of a brochure that they sent to a large number of randomly selected job seekers in Germany. It contained information on i) labor market conditions, ii) duration dependence, iii) effects of unemployment on life satisfaction, and iv) importance of social ties. They find no significant effect overall, but for those at risk of long-term unemployment they find a positive effect between 8 months and a year after sending the brochure. In our intervention we also find the strongest effects for individuals with longer unemployment duration, but even overall effects are significant and occur much closer in time to the actual provision of information. Their study

<sup>&</sup>lt;sup>6</sup>This resembles findings by Launov and Waelde (2013) that attribute the success of German labor market reforms to service restructuring (again both advice and monitoring/sanctions) with non-experimental methods.

<sup>&</sup>lt;sup>7</sup>There are some indirect attempts to distinguish between advice and monitoring/sanction. Ashenfelter et al. (2005) cite experimental studies from the US by Meyer (1995) which have been successful but entailed monitoring/sanctions as well as advice, and they then provide evidence from other interventions that monitoring/sanctions are ineffective in isolation. This leads them indirectly to conclude that the effectiveness of the first set of interventions must be due to the advice. Yet subsequent research on the effects of sanctions found conflicting evidence: e.g., Micklewright and Nagy (2010) and Van den Berg and Van der Klaauw (2006) also find only limited effects of increased monitoring, while other studies such as Van der Klaauw and Van Ours (2013), Lalive et al. (2005) and Svarer (2011) find strong effects.

has low costs of provision, is easily replicable, treated a large sample, and has administrative data to assess success. On the other hand, it is not clear which of the varied elements in the brochure drives the results, there are no intermediate measures on how it affects the job search process, and the advice is generic to all job seekers rather than tailored to the occupations they are looking for.

Our study is also complementary to a few recent studies which analyze data from commercial online job boards. Kudlyak et al. (2014) analyze U.S. data from Snagajob.com and find that job search is stratified by educational attainment but that job seekers lower their aspirations over time. Faberman and Kudlyak (2014) analyze the same data source to see if the declining hazard rate of finding a job is driven by declining search effort. They find little evidence for this. The data lacks some basic information such as employment/unemployment status and reason for leaving the site, but they document some patterns related to our study: Occupational job search is highly concentrated and absent of any exogenous intervention it broadens significantly but only slowly over time, with 60% of applications going to the modal occupation in week 2 and still 55% going to the modal occupation after six months.<sup>8</sup>

Marinescu and Rathelot (2014) investigate the role of differences in market tightness as a driver of aggregate unemployment. They measure the geographic breadth of search by using U.S search data from Careerbuilder.com and concur with earlier work that differences in market tightness are not a large source of unemployment. In their dataset search is rather concentrated, with the majority of applications aimed at jobs within 25km distance and 82% of applications staying in the same city (Core-Based Statistical Area), even if some 10% go to distances beyond 100km. Using the same data source, Marinescu (2014) investigates equilibrium effects of unemployment insurance by exploiting state-level variation of unemployment benefits. The level of benefits affects the number of applications, but effects on the number of vacancies and overall unemployment are limited. Marinescu and Wolthoff (2014) document that job titles are an important explanatory variable for attracting applications in Careerbuilder.com, that they are informative above and beyond wage and occupational information, and that controlling for job titles is important to understand the remaining role of wages in the job matching process. As mentioned, these studies have large sample size and ample information of how people search on the particular site, but none involves a randomized design nor do they have information on other job search channels. Also, their focus is not on advice.

Our weekly survey of job search activity outside the lab over a panel of twelve weeks is related to the seminal panel study by Card and Mueller (2016) that conducted weekly interviews regarding reservation wages with a panel of job seekers in the US over the course of half a year. Our study has a slightly different focus, and uses the survey as a complement to the direct measures of job search activity within our job search platform and within a controlled randomized trial.

Our recommendation to target occupational information to job seekers that otherwise search narrowly is in the spirit of recent discussion of profiling in active labor market policy. Profiling singles out

<sup>&</sup>lt;sup>8</sup>The modal occupation is the occupation to which the individual sends the largest share of her applications.

<sup>&</sup>lt;sup>9</sup>These numbers are based on Figure 5 in the 2013 working paper. Neither paper provides numbers on the breath of occupational search. The "distaste" for geographical distance backed out in this work for the US is lower than that backed out by Manning and Petrongolo (2011) from more conventional labor market data in the UK, suggesting that labor markets in the UK are even more local.

subsets of individuals for treatment according to a probabilistic assessment of the benefits (see, e.g., Berger et al. (2000) for a comprehensive discussion). Interestingly, in our environment the profiling could be integrated directly into a standard job search engine in which individuals first search "normally" and subsequently, depending on the breadth of their search, occupational information could be offered or not.

To our knowledge, our study is the first that undertakes job-search platform design and evaluates it. The randomized setup allows for clear inference. While the rise in internet-based search will render such studies more relevant, the only other study of search platform design that we are aware of is Dinerstein et al. (2014), who study a change at the online consumer platform Ebay which changed the presentation of its search results to order it more by price relative to other characteristics. This lead to a decline in prices, which is assessed in a consumer search framework. While similar in broad spirit of search design, the study obviously differs substantially in focus.

## 3 The Set-Up of the Study

Two main contributions underlie our study: first, we design of a novel online tool that provides labor market information that is readily available to researchers but usually not to job seekers. The aim is to make this available in an easily accessible cost-effective form and to enable a direct link to the potential jobs. Second, we evaluate the new tool experimentally in a randomized controlled experiment for which we invited job seekers in the area of Edinburgh to our computer facilities for a period of 12 weeks during two waves, one in the fall of 2013 and one in the spring of 2014. We used a "standard" interface for comparison, which relies on a keyword search as in most existing job search platforms. All participants started with the standard search platform. Half of the sample was exposed to the new tool after 3 weeks. We now describe the experimental design in more detail. Descriptives on the job search process and on the sample are provided in the next section, followed by the empirical evaluation.

#### 3.1 Description of the Advice Interface

We designed an on-line job search interface in collaboration with professional programmers from the IT Applications Team at the University of Edinburgh. The main feature of the interface is to provide a tailored list of suggestions of possible alternative occupations that may be relevant to job seekers, based on a preferred occupation that job seekers pre-specify (but can change at any time). As mentioned in the introduction we provide advice on occupations for multiple reasons: First, recent influential work has argued that the great recession has led job seekers to increasingly concentrate too much search effort on occupations with too few vacancies (Sahin et al. (2014)). These findings for the US have been replicated for the UK (Patterson et al. (2016)), and one explanation might be a lack of information about labor market conditions or about skill transferability. Moreover, it has long been argued that learning about occupations might play a substantial role in job search, suggesting a role for information provision.<sup>10</sup> Finally, occupations are an observational unit with sufficient employment so that we can exploit existing representative surveys in order to provide advice.

<sup>&</sup>lt;sup>10</sup>See, for example, the citations in Footnote 3.

We use two methodologies to compile a list of alternative occupations to the preferred occupation specified by the job seeker. The first methodology builds on the idea that successful labor market transitions experienced by people with a similar profile contain useful information about occupations that may be suitable alternatives to the preferred occupation: the fact that others found jobs there indicates that skills might be transferable and job available. It is based on the standard idea in the experimentation literature that others have already borne the cost of experimentation and found suitable outcomes, and this knowledge would be useful to reduce the experimentation costs of a given job seeker.

To do this, we use information on labor market transitions observed in the British Household Panel Survey and the national statistical database of Denmark (because of larger sample size). <sup>11</sup> Both databases follow workers over time and record in what occupation they are employed. We then match the indicated preferred occupation to the most common occupations to which people employed in the preferred occupation transition to. For each occupation, we created a list of three to five common transitions. The list contained all occupations that occur in the top-10 common transitions in both datasets (if there were more than five of these, we selected the five highest occurring occupations). In case this resulted in less than 3 occupations, we added the highest ranked transitions from each dataset until the list contained at least three occupations.

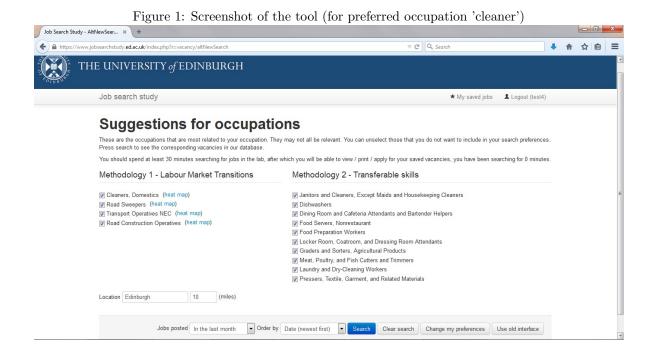
This methodology has the advantage of being highly flexible and transportable. Many countries now have databases that could be used to match this algorithm. That is, the tool we propose can easily be replicated and implemented in many different settings.

The second methodology uses information on transferable skills across occupations from the US based website O\*net, which is an online "career exploration" tool sponsored by the US department of Labor, Employment & Training Administration. For each occupation, they suggest up to 10 related occupations that require similar skills. We retrieved the related occupations and presented the ones related to the preferred occupation as specified by the participant. This provides information on skill transferability only, not on job availability.

The tool is directly embedded in the job search interface. That means that once participants have specified their preferred occupation, they could then click "Save and Start Searching" and were taken to a new screen where a list of suggested occupations was displayed. The occupations were listed in two columns: The left column suggests occupations based on the first methodology (based on labor market transitions). The right column suggests occupations based on the second methodology (O\*net related occupations). Figure 1 shows a screenshot of the tool, with suggestions based on the preferred occupation 'cleaner'. Participants were fully informed of the process by which these suggestions came about, and could select or unselect the occupations they wanted to include or exclude in their search. By default all were selected. If they then click the "search" button, the program searches through the same underlying vacancy data as in the control group but selects all vacancies that fit any of the selected occupations in their desired geographic area.<sup>12</sup>

<sup>&</sup>lt;sup>11</sup>The name of the database is IDA - Integrated Database for Labour Market Research administered by Statistics Denmark. We are grateful to Fayne Goes for providing us with the information.

<sup>&</sup>lt;sup>12</sup>Occupations in O\*net have a different coding and description and have a much finer categorization than the threedigit occupational code available in the British Household Panel Survey (BHPS) and in Universal Johnatch vacancy

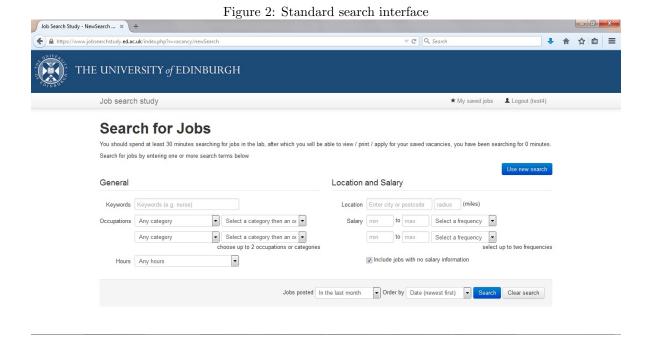


In addition to these suggestions, the interface also provides visual information on the tightness of the labor market for broad occupational categories in regions in Scotland. The goal here is to provide information about how competitive the labor market is for a given set of occupations - which is closest to the idea of search mismatch in Sahin et al. (2014) and provides information on the competition for jobs but not on skill transferability. We constructed "heat maps" that use recent labor market statistics for Scotland and indicate visually (with a colored scheme) where jobs may be easier to get (because there are many jobs relative to the number of interested job seekers). These maps were created for each broad occupational category (two-digit SOC codes).<sup>13</sup> Participants could access the heat maps by clicking on the button "heat map" which was available for each of the suggested occupations based on labor market transitions. So they could check them for each broad category before actually performing a search, not for each particular vacancy.

In principle this tool can be used with any database of vacancies that includes occupational codes; for our experimental approach we combine it with one of the largest database in the UK.

data. We therefore asked participants twice for their preferred occupation, once in O\*net form and BHPS form. The query on the underlying database relies on keyword search, taking the selected occupations as keywords, to circumvent problems of differential coding.

<sup>&</sup>lt;sup>13</sup>These heat maps are based on statistics provided by the Office for National Statistics, (NOMIS, claimant count, by occupations and county, see https://www.nomisweb.co.uk/). We created the heat maps at the two-digit level because data was only available on this level. Clearly, this implies that the same map is offered for many different 4-digit occupations, and job seekers might see the same map several times, which limits the value of this approach relative to the earlier ones. Obviously a commercial job search site could give much richer information on the number of vacancies posted in a geographic area and the number of people looking for particular occupations in particular areas. An example of a heat map is presented in the Online Appendix 8.2.6.



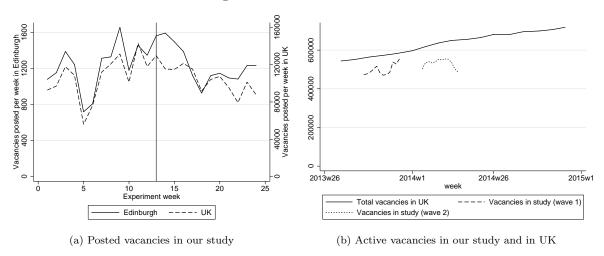
#### 3.2 Control Treatment: Standard Search Interface

We designed a standard job search engine that replicates the search options available at the most popular search engines in the UK (such as Monster.com and Universal Jobmatch), again in collaboration with the IT Applications Team at the University of Edinburgh. As in the treatment group this allowed us to record precise information about how people search for jobs (what criteria they use, how many searches they perform, what vacancies they click on and what vacancies they save), as well as collecting weekly information (via the weekly survey) about outcomes of applications and search activities outside the laboratory.

Figure 2 shows a screenshot of the main page of the standard search interface. Participants can search using various criteria (keywords, occupations, location, salary, preferred hours), but do not have to specify all of these. Once they have defined their search criteria, they can press the search button at the bottom of the screen and a list of vacancies fitting their criteria will appear. The information appearing on the listing is the posting date, the title of the job, the company name, the salary (if specified) and the location. They can then click on each individual vacancy to reveal more information. Next, they can either choose to "save the job" (if interested in applying) or "not save the job" (if not interested). If they choose not to save the job, they are asked to indicate why they are not interested in the job from a list of possible answers.

As in most job search engines, they can modify their search criteria at any point and launch a new search. Participants had access to their profile and saved vacancies at any point in time outside the laboratory, using their login details. They could also use the search engine outside the laboratory. We recorded all search activity on our platform including those that take place outside the lab. The latter is, however, only a very small share compared to the search activities performed in the lab.

Figure 3: Number of vacancies



The key feature of this interface is that job seekers themselves have to come up with the relevant search criteria. This is shared by commercial sites like Universal Jobmatch or Monster.com at the time of our study, which also provide no further guidance to job seekers on things such as related occupations.

#### 3.3 Vacancies

In order to provide a realistic job search environment, both the new tool and the standard search interface access a local copy of the database of real job vacancies of the government website Universal Jobmatch. This is a very large job search website in the UK in terms of the number of vacancies. This is a crucial aspect in the setup of the study, because results can only be trusted to resemble natural job search if participants use the lab sessions for their actual job search. The large set of available vacancies combined with our carefully designed job search engine assures that the setting was as realistic as possible. Panel (a) of Figure 3 shows the number of posted vacancies available through our search engine in Edinburgh and in the UK for each week of the study (the vertical line indicates the start of wave 2). Each week there are between 800 and 1600 new vacancies posted in Edinburgh. Furthermore, there is a strong correlation between vacancy posting in Edinburgh and the UK. In panel (b) the total number of active vacancies in the UK is shown over the second half of 2013 and 2014. As a comparison the total number of active vacancies in the database used in the study in both waves is shown. It suggests that the database contains over 80% of all UK vacancies, which is a very extensive coverage compared to other online platforms. It is well-known that not all vacancies

<sup>&</sup>lt;sup>14</sup>Panel (b) is based on data from our study and data from the Vacancy Survey of the Office of National Statistics (ONS), dataset "Claimant Count and Vacancies - Vacancies", url: www.ons.gov.uk/ons/rel/lms/labour-market-statistics/march-2015/table-vacs01.xls

<sup>&</sup>lt;sup>15</sup>For comparison, the largest US jobsearch platform has 35% of the official vacancies; see Marinescu (2014), Marinescu and Wolthoff (2014) and Marinescu and Rathelot (2014). The size difference might be due to the fact that the UK platform is run by the UK government.

on online job search platforms are genuine, so the actual number might be somewhat lower. We introduced ourselves a small number of additional posts (below 2% of the database) for a separate research question (addressed in a separate paper). 17

#### 3.4 Job Seekers

To study the effect of information provision through the new interface, we recruited job seekers in the area of Edinburgh in two waves: wave 1 was conducted in September 2013 and wave 2 in January 2014. Labor market conditions in Edinburgh are broadly consistent with national ones: the unemployment rate in the UK overall and in Edinburgh in particular between 2011 and 2014 is shown in part a) of Figure 4 where the vertical lines indicate the start of each wave. These statistics are based on the Labour Force Survey and not the entire population. Therefore we present the number of job search assistance (JSA) claimants in the Edinburgh and the UK in panel (b), which is an administrative figure and should be strongly correlated with unemployment. The number of JSA claimants is decreasing monotonically between 2012 and 2015, and the Edinburgh and UK figures follow a very similar path.

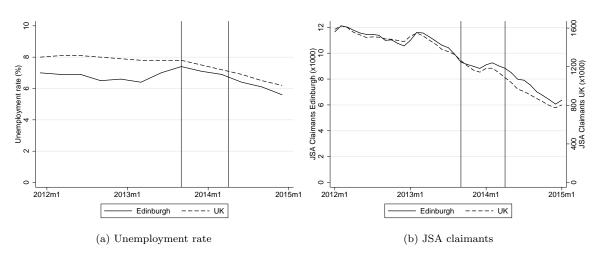


Figure 4: Aggregate labor market statistics

The eligibility criteria for participating to the study were: being unemployed, searching for a job

<sup>&</sup>lt;sup>16</sup> For Universal Jobmatch evidence has been reported on fake vacancies covering 2% of the stock posted by a single account (Channel 4 (2014)) and speculations of higher total numbers of fake jobs circulate (Computer Business Review (2014)). Fishing for CV's and potential scams are common on many sites, including Carreerbuilder.com (The New York Times (2009a)) and Craigslist, whose chief executive, Jim Buckmaster, is reported to say that "it is virtually impossible to keep every scam from traversing an Internet site that 50 million people are using each month" (The New York Times (2009b)).

<sup>&</sup>lt;sup>17</sup>Participants were fully informed about this. They were told that "we introduced a number of vacancies (about 2% of the database) for research purposes to learn whether they would find these vacancies attractive and would consider applying to them if they were available". They were asked for consent to this small percentage of research vacancies and were informed about the true nature of such vacancies if they expressed interest in the vacancy before any actual application costs were incurred, so any impact was minimized. This small number is unlikely to affect job search, and there is no indication of differential effects by treatment group: In an exit survey the vast majority of participants (86%) said that this did not affect their search behavior, and this percentage is not statistically different in the treatment and control group (p-value 0.99). This is likely due to the very low numbers of fake vacancies and to the fact that fake advertisements are common in any case to online job search sites (see footnote 16) and that this is mentioned to job seekers in many search guidelines (see e.g. Joyce (2015)).

for less than 12 weeks (a criterion that we did not enforce), and being above 18 years old. We imposed no further restrictions in terms of nationality, gender, age or ethnicity. We aimed to recruit 150 participants per wave. Compared to the stock of JSA claimants that constitutes about 2%. <sup>19</sup>

As a background on the institutional setting of our study, individuals on job seeker allowance (JSA) receive between £52.25 and £72 per week depending on age. Eligibility depends on sufficient contributions during previous employment or when income is sufficiently low.<sup>20</sup> This is linked to the requirement to be available and actively looking for work. In practice, this implies committing to agreements made with a work coach at the job centre, such as meeting the coach at regular (usually bi-weekly) intervals, applying to suggested vacancies, or participating in suggested training. They are not entitled to reject job offers because they dislike the occupation or the commute, except that the work coach can grant a period of up to three months to focus on offers in the occupation of previous employment, and required commuting times are capped at 1.5 hours per leg. The work coach can impose sanctions on benefit payments in case of non-compliance to any of the criteria.

We obtained the collaboration of several local public unemployment agencies (called Jobcentre Plus) to recruit job seekers on their premises during a two-week window prior to each wave. This window is suitable since most individuals on job seeker allowance meet their advisers bi-weekly, which gives us a chance to encounter most of them. The counselors were informed of our study and were asked to advertise the study. We also placed posters and advertisements at various public places in Edinburgh (including libraries and cafes) and posted a classified ad on a popular on-line platform (not limited to job advertisements) called Gumtree. Table 1 presents the sign up and show up rates. Of all participants, 86% were recruited in the Jobcentres. Most of the other participants were recruited through our ad on Gumtree. We approached all visitors at the Jobcentres during two weeks. Out of those we could talk to and who did not indicate ineligibility, 43% percent signed up. Out of everyone that signed up, 45% showed up in the first week and participated in the study, which is a substantial share for a study with voluntary participation. These figures display no statistically significant difference between the two waves of the study.

We also conducted an online study, outside the laboratory, in which job seekers were asked to

<sup>&</sup>lt;sup>18</sup>We do drop the observations on one participant from our sample because this participant had been unemployed for over 30 years and was therefore an extraordinary outlier in our sample. We only include participants who search at least once, which excludes two participants who showed up once without searching and never returned. Including them the analysis has no effects on the qualitative findings.

<sup>&</sup>lt;sup>19</sup>The number of JSA claimants in Edinburgh during our study is approximately 9,000, the monthly flow of new JSA claimants in Edinburgh during the study is around 1,800.

<sup>&</sup>lt;sup>20</sup>Benefits of £56.25 per week apply to those aged up to age 24, and £72 per week for those aged 25 and older. Contribution-based JSA is given to Individuals if they have contributed sufficiently through previous employment, and benefits last for a maximum of 6 months. Afterwards - or in the absence of sufficient contributions - income-based JSA applies, with identical weekly benefits but with extra requirements. The amount is reduced if they have other sources of income, if they have savings or if their partner has income. Once receiving JSA, the recipient is not eligible for income assistance, however they may receive other benefits such as housing benefits.

<sup>&</sup>lt;sup>21</sup> The sign up rate at Jobcentres for the lab study in wave 2 is based on only one day of recruitment for the following reason: We asked our assistants to write down the number of people they talked to and the number that signed up. Unfortunately these have not been separated for the online study and the lab study. In the first wave there were different assistants for the two studies, such that we can compute the sign up shares separately. In the second wave we asked assistants to spend parts of their time per day exclusively on the lab study and parts exclusively on the online study, so we only have sign-ups for the total number. One day was an exception, as recruitment was done only for the lab study on this day, such that we can report a separate percentage based on this day. We do not have a separate number for sign-up for the online study.

complete a weekly survey about their job search. These participants did not attend any sessions, but simply completed the survey for 12 consecutive weeks. This provides us with descriptive statistics about job search behavior of job seekers in Edinburgh and it allows us to compare the online participants with the lab participants. These participants received a £20 clothing voucher for each 4 weeks in which they completed the survey. The online participants were recruited in a similar manner as the lab participants, which means most of them signed up at the Jobcentres.<sup>22</sup> The sign up rate at the Jobcentres was slightly higher for the online survey (58%), however out of those that signed up, only 21% completed the first survey. This was partly caused by the fact that about one-fourth of the email addresses that were provided was not active.

In Section 4.1 we discuss in more detail the representativeness of the sample, by comparing the online and the lab participants with population statistics.

#### 3.5 Experimental Procedure

Job seekers were invited to search for jobs once a week for a period of 12 weeks (or until they found a job) in the computer facilities of the School of Economics at the University of Edinburgh. We conducted sessions at six different time slots, on Mondays or Tuesdays at 10 am, 1 pm or 3:30 pm. Participants chose a slot at the time of recruitment and were asked to keep the same time slot for the twelve consecutive weeks.

Participants were asked to search for jobs using our job search engine for a minimum of 30 minutes.<sup>23</sup> After this period they could continue to search or use the computers for other purposes such as writing emails, updating their CV, or applying for jobs. They could stay in our facility for up to two hours. We emphasized that no additional job search support or coaching would be offered.

All participants received a compensation of £11 per session attended (corresponding to compensation for meal and travel expenses as advized by Jobcentre Plus) and we provided an additional £50 clothing voucher for job market attire for participating in 4 sessions in a row. Our study did not affect the entitlements or obligations that participants face at the local Jobcentre.<sup>24</sup>

Participants were asked to register in a dedicated office at the beginning of each session. At the first session, they received a unique username and password and were told to sit at one of the computer desks in the computer laboratory. The computer laboratory was the experimental laboratory located at the School of Economics at the University of Edinburgh with panels separating desks to minimize interactions between job seekers. They received a document describing the study as well as a consent

 $<sup>^{22}</sup>$ Participants were informed of only one of the two studies, either the on-site study or the on-line study. The did not self-select into one or the other.

<sup>&</sup>lt;sup>23</sup>The 30 minute minimum was chosen as a trade-off between on the one hand minimizing the effect of participation on the natural amount of job search, while on the other hand ensuring that we obtained enough information. Given that participants spent around 12 hours a week on job search, a minimum of half an hour per week is unlikely to be a binding constraint on weekly job search, while it was a sufficient duration for us to collect data. Furthermore, similar to our lab participants, the participants in the online survey (who did not come to the lab and had no restrictions on how much to search) also indicate that they search 12 hours per week on average. Among this group, only in 5% of the cases the reported weekly search time is smaller than 30 minutes. In the study, the median time spent in the laboratory was 46 minutes. We made sure that participants understood that this is not an expectation of their weekly search time, and that they should feel free to search more and on different channels.

<sup>&</sup>lt;sup>24</sup> All forms of compensation effectively consisted of subsidies, i.e. they had no effect on the allowances the job seekers were entitled to. The nature and level of the compensation were discussed with the local job centres to be in accordance with the UK regulations for job seeker allowances.

Table 1: Recruitment and show-up of participants

	Full sample	Wave 1	Wave 2
Recruitment channel participants:			
Job centres	86%	83%	89%
Gumtree or other	14%	17%	11%
Sign up rate jobcentre for lab study $^a$	43%	39%	$47\%^c$
Show up rate lab study	45%	43%	46%
Sign up rate jobcentre for online study <sup>a</sup>		60%	
Show up rate online study <sup><math>b</math></sup>	21%	21%	21%

<sup>&</sup>lt;sup>a</sup> Of those people that were willing to talk to us about the study, this is the share that signed up for the study. <sup>b</sup> About a fourth of those that signed up for the online study had a non-existing email address, which partly explains the low show up rate. <sup>c</sup> Based on only one day of recruitment - see Footnote 21 for explanation.

form that we collected before the start of the initial session (the form can be found in the Online Appendix 8.2.1). We handed out instructions on how to use the interface, which we also read aloud (the instructions can be found in the Online Appendix 8.2.2). We had assistance in the laboratory to answer clarifying questions. We clarified that we were unable to provide any specific help for their job search, and explicitly asked them to search as they normally would.

Once they logged in, they were automatically directed to our own website. They were first asked to fill in a survey. The initial survey asked about basic demographics, employment and unemployment histories as well as beliefs and perceptions about employment prospects, and measured risk and time preferences. From week 2 onwards, they only had to complete a short weekly survey asking about job search activities and outcomes. For vacancies saved in their search in our facility we asked about the status (applied, interviewed, job offered). We asked similar questions about their search through other channels than our study. The weekly survey also asked participants to indicate the extent to which they had personal, financial or health concerns (on a scale from 1 to 10). The complete survey questionnaires can be found in the Online Appendices 8.2.4 and 8.2.5.

After completing the survey, the participants were re-directed towards our search engine and could start searching. A timer located on top of the screen indicated how much time they had been searching. Once the 30 minutes were over, they could end the session. They would then see a list of all the vacancies they had saved and were offered the option of printing these saved vacancies. This list of printed vacancies could be used as evidence of required job search activity at the Jobcentre. It was, however, up to the job seekers to decide whether they wanted to provide that evidence or not. We also received no additional information about the search activities or search outcomes from the Jobcentres. We only received information from the job seekers themselves. This absence of linkage was important to ensure that job seekers did not feel that their search activity in our laboratory was monitored by the employment agency. They could then leave the facilities and receive their weekly compensation.<sup>25</sup> Those who stayed could either keep searching with our job search engine or use the computer for other

<sup>&</sup>lt;sup>25</sup>Participants were of course allowed to leave at any point in time but they were only eligible to receive the weekly compensation if they had spent 30 minutes searching for jobs using our search engine.

Table 2: Rar	ndomization s	cheme
	Wave 1	Wave 2
Monday 10 am	Control	Treatment
Monday 1 pm	Treatment	Control
Monday 3:30 pm	Control	Treatment
Tuesday 10 am	Treatment	Control
Tuesday 1 pm	Control	Treatment
Tuesday 3:30 pm	Treatment	Control

purposes (such as updating their CV, applying on-line or using other job search engines). We did not keep track of these other activities. Once participants left the facility, they could still access our website from home, for example in order to apply for the jobs they had found.

#### 3.6 Randomization

All participants used the standard interface in the first 3 weeks of the study. Half of the participants was offered the "alternative" interface, which incorporates our new tool (as shown in Figure 1), from week 4 onwards. Participants were randomized into control (no change in interface) and treatment group (alternative interface) based on their allocated time slot. We randomized the first time slot into treatment and control, and assigned each following time slot in an alternating pattern, to avoid any correlation between treatment status and a particular time slot. Each time slot that was allocated to control (treatment) in the first wave was assigned to treatment (control) in the second wave. Table 2 presents the assignment of sessions to control and treatment groups. Note that the change of interface was not previously announced, apart from a general introductory statement to all participants that included the possibility to alter the search engine over time.

Participants received a written and verbal instruction of the alternative interface (see Online Appendix 8.2.3), including how the recommendations were constructed, in the fourth week of the study before starting their search. For them, the new interface became the default option when logging on. It should be noted, though, that it was made clear to participants that using the new interface was not mandatory. Rather, they could switch back to the previous interface by clicking a button on the screen indicating "use old interface". If they switched back to the old interface, they could carry on searching as in the previous weeks. They could switch back and forth between interfaces. This ensures that we did not restrict choice, but rather expanded their means of searching for a job.

#### 3.7 Measures of Job Search

The main goal of the study is to evaluate how tailored advice affects job search strategies. Our data allow us to examine each step of the job search process related to the search on our platform: the listing of vacancies to which job seekers are exposed, the vacancies they apply to and the interviews they receive. In the weekly survey that participants complete before starting to search, we ask about applications and interviews through channels other than our study. The intervention may affect these outcomes as well, since the information provided in the alternative interface could influence people's job search strategies outside the lab. Therefore we also document the weekly applications and interviews

	Table 3: Outcome variables	
	Search activity in the lab	Search activity outside the lab
Listed vacancies		
Occupational Breadth	$\checkmark$	
Geographical Breadth	$\sqrt{}$	
Number	$\checkmark$	
Applications		
Occupational Breadth	$\sqrt{}$	
Geographical Breadth		
Number		$\checkmark$
Interviews		
Number	$\sqrt{}$	$\checkmark$
Core and non-core occupation	ns $\sqrt{}$	

through other channels. Of course, ultimately one would also like to evaluate the effects on job finding and the characteristics of the job found (occupation, wage, duration, etc.), which would be important to evaluate the efficiency implications of such an intervention. This is however not the prime goal of this study and given the small sample of participants, we should be cautious when interpreting results on job finding as we discuss in a separate part in Section 5.5.

We summarize in Table 3 the outcome variables of interest. All measures are defined on the set of vacancies retrieved in a given week, independent of whether they arose due to many independent search queries or few comprehensive queries. The main outcome variables relate to (1) listed vacancies, (2) applications and (3) interviews. The exact definition of each of these is presented next.

The most immediate measure of search relates to listed vacancies, i.e., the listing of vacancies that appears on the participants' screen as a return to their search queries in a given week. To be precise, when a participant hits the search button on either the standard or the new interface, all vacancies that fall under the search criteria are retrieved. Up to 25 of these vacancies are shown immediately on the computer screen, ordered by default according to the most recent date of posting (but alternative orderings can be chosen such as location or salary). The displayed vacancies are recorded as "listed" in this week. If the initial query returned more vacancies and the participant wants to see them, he has to actively move to the next screen where again up to 25 additional vacancies are shown. These again are recorded as "listed" for this week. This means that vacancies are only recorded as listed if the applicant had them on the screen, and vacancies that are e.g. older and were not consulted by the participant are excluded. If the applicant hits the search button again for a new query, again those vacancies that appear on his screen are added to the "listed" vacancies for that week. That means that all our analyses are at the weekly level and, thus, we group all listings in a week together. <sup>26</sup> We note that listings are not mechanical even in the treatment group but, rather, remain an outcome of their choice: on the new interface users still decide how many pages of results to move through, which

<sup>&</sup>lt;sup>26</sup>The alternative interface tends to necessitate less search queries than the standard interface to generate the same number of vacancies because on the alternative interface one query is intended to also return vacancies for other related occupations. For that reason the weekly analysis seems more interesting compared to results at the level of an individual query. This also means that in a given week each vacancy is counted at most once, even if it is returned as a result to multiple queries.

geographical radius to explore, how many recommended alternative occupations to keep, and how many preferred occupations and associated alternatives to explore in a given week - not to mention that participants can revert back to standard keyword search to explore some options more deeply (we document the use of each interface later on).

The second measure of search behavior relates to applications, which we consider a more direct measure of interest as compared to viewed vacancies (vacancies that the job seeker clicks on in order to view all job details) and saved vacancies to which the job seeker might want to apply later.<sup>27</sup> For applications we have information about applications based on search activity conducted inside the laboratory as well as outside the laboratory which we collected through the weekly surveys. For the applications based on search in the laboratory, we asked participants to indicate for each vacancy saved previously whether they actually applied to it or not.<sup>28</sup> We can therefore precisely map applications to the timing of the search activity. This is important as there may be a delay between the search and the actual application; so applications that are made in week 4 and after could relate to search activity that took place before the actual intervention. For the applications conducted based on search outside the laboratory, we do not have such precise information. We asked how many applications job seekers made in the previous week but we do not know the timing of the search activity these relate to. For consistency, we assume that the lag between applications and search activity is the same inside and outside the laboratory (which is one week) and assign applications to search activity one week earlier. As a result, we cannot use information on search activity in the last week of the experiment, as we do not observe applications related to this week.

For listed vacancies and applications we look at the number as well as measures of breadth (occupational and geographical). For occupational breadth we focus on the UK Standard Occupational Classification code (SOC code) of a particular vacancy, which consists of four digits.<sup>29</sup> The structure of the SOC codes implies that the more digits two vacancy codes share, the more similar they are. Our measure of diversity within a set of vacancies is based on this principle, defining for each pair within a set the distance in terms of the codes. The distance is zero if the codes are the same, it is 1 if they only share the first 3 digits, 2 if they only share the first 2 digits, 3 if they share only the first digit and 4 if they share no digits. This distance, averaged over all possible pairs within a set, is the measure that we use in the empirical analysis, but discuss robustness to alternative measures in Section 5.6. Note that this distance is increasing in breadth (diversity) of a set of vacancies. We compute this measure for the set of listed and applied vacancies in each week for each participant. For geographical breadth we use a simple measure. Since a large share of searches restricts the location to Edinburgh, we use the weekly share of a participant's searches that goes beyond Edinburgh as the measure of geographical breadth.<sup>30</sup>

 $<sup>\</sup>overline{^{27}}$  Not surprisingly, results for viewed and saved vacancies are reminiscent of those for listed and applied vacancies and are omitted for brevity.

<sup>&</sup>lt;sup>28</sup>If they have not applied, they are asked whether they intend to apply and only if they answered affirmatively they were asked again next week whether they did apply or not. A similar procedure is followed for interviews.

<sup>&</sup>lt;sup>29</sup>The first digit of the code defines the "major group", the second digit defines the "sub-major group", the third digit defines the "minor group" and the fourth digit defines the "unit group" which provides a very specific definition of the occupation. Some examples are "Social science researchers" (2322), "Housekeepers and related occupations" (6231) and "Call centre agents/operators" (7211).

<sup>&</sup>lt;sup>30</sup>Note that the direct surroundings of Edinburgh contain only smaller towns. The nearest large city is Glasgow, which

Our third outcome measure is interviews - which is the measure most closely related to job prospects. As was done for applications, we assign interviews to the week in which the search activity was performed, and assign interviews through channels other than the lab to search activity two weeks earlier. As a result we do not use information on search activity in weeks 11 and 12 of the experiment, because for job search done in these weeks we do not observe interviews. We have information on the number interviews, but the number is too small on average to compute informative breadth measures. As an alternative, we asked individuals at the beginning of the study about three "core" occupations in which they are looking for jobs, and we can estimate separate treatment effects for interviews in core and non-core occupations.

#### 3.8 Professionalism of Search Interfaces

In order for the study to provide a valid environment to study search behavior, it is important that participants themselves take it seriously and do not view our service as inferior to search environments in the overall marketplace. In an exit survey we asked participants to evaluate the interface and found that participants evaluated it very positively. The responses to the question "How would you rate the search interface compared to other interfaces?" were: Poor (7%) Below average (7%) Average (14%) Good (46%) Very Good (26%). These responses were very similar in treatment and control groups.

# 4 Descriptive Statistics on Job Seeker Characteristics and Job Search Behavior

This section provides descriptive statistics about the characteristics of the sample of job seekers in our study and provides an overview about how they search for jobs. We also use this to indicate how our experimental sample compares to the (limited) information we have on the overall set of JSA claimants in Edinburgh and to those participating in the online survey, and to demonstrate balance between treatment and control group. For the latter, we can not only compare basic characteristics, but also their job search behavior in the first three weeks where individuals in both treatment and control group use the same standard interface and share the same instructions. The control group faces no intervention throughout the study, and we document how they change their job search over time. And for the treatment group we document to which extent they adopt the new interface. Finally, we present data on attrition.

# 4.1 Job Seeker Characteristics and Job Search History: summary, representativeness and balance

Demographic variables, based on the first week baseline survey, show that 43% of the lab participants are female, the average age is 36 and 43% have some university degree. 80% classify themselves as 'white' and 27% have children. This is summarized in Table 4. We can compare this to aggregate statistics about the population of job seekers available from The Office of National Statistics (NOMIS)

takes about 1-1.5 hours of commuting time.

where we truncate unemployment duration to obtain a sample with similar median.<sup>31</sup> Unfortunately this provides only few variables presented in the last column of the table. It indicates that we oversample women and non-whites, while the average age is very similar. Another comparison group are the participants in our online survey which arguably face a lower hurdle to participation in the study. Results are presented in the intermediate columns, and in column 9 the p-value of a two-sided t-test for equal means relative to the lab participants is shown. The online survey participants differ somewhat in composition: they are more likely to be female, they are slightly younger and they have less children.

Table 4: Characteristics of lab participants and online survey participants (based on the first week initial survey)

	La	ab part	icipant	S	(	Online	survey		T-test <sup>a</sup>	$\operatorname{Pop.}^{b}$
	mean	$\operatorname{sd}$	$\min$	max	mean	$\operatorname{sd}$	$\min$	max	pval	
Demographics:										
gender $(\%)$	43	50	0	1	52	50	0	1	.09	33
age	36	12	18	64	34	12	18	64	.08	35
high educ (%)	43	50	0	1	43	50	0	1	1.00	
white $(\%)$	80	40	0	1	77	42	0	1	.43	89
number of children	.53	1	0	5	.28	.57	0	2	.02	
couple (%)	23	42	0	1	23	42	0	1	.96	
any children $(\%)$	27	45	0	1	23	42	0	1	.41	
Job search history:										
vacancies applied for	64	140	0	1000	75	187	0	1354	.53	
interviews attended	.48	0.84	0	6	2.7	4	0	20	.00	
jobs offered	.42	1.1	0	8	.51	1.6	0	10	.52	
at least one offer $(\%)$	20	40	0	1	24	34	0	1	.36	
days unempl. (mean)	260	620	1	5141	167	302	8	2929	.15	111
days unempl. (median)	80				118					81
less than 183 days (%)	76	43	0	1	75	44	0	1	.76	
less than 366 days (%)	85	35	0	1	91	28	0	1	.13	
job seekers allowance (£)	52	75	0	1005	58	42	0	280	.49	
housing benefits $(\pounds)$	64	129	0	660	48	95	0	400	.36	
other benefits $(\pounds)$	14	65	0	700	12	56	0	395	.81	
01	205				100					
Observations	295				103					

<sup>&</sup>lt;sup>a</sup> P-value of a t-test for equal means of the lab and online participants. <sup>b</sup> Average characteristics of the population of job seeker allowance claimants in Edinburgh over the 6 months of study. The numbers are based on NOMIS statistics, conditional on unemployment duration up to one year. <sup>c</sup> High educated is defined as a university degree.

The lower part of Table 4 shows variables related to job search history, also based on the first week baseline survey. The lab participants have on average applied to 64 jobs during the unemployment spell preceding the participation in our study. These led to 0.48 interviews and 0.42 job offers.<sup>32</sup> Only 20% received at least one offer. Mean unemployment duration at the start of the study is 260 days, while the median is 80 days. About three-fourth of the participants had been unemployed for less

<sup>&</sup>lt;sup>31</sup>Source: Office for National Statistics: NOMIS Official Labour Market Statistics. Dataset: Claimant Count conditional on unemployment duration
12 months, average over the duration of the study. Restricting attention to less than
12 months ensures similar median unemployment duration between the NOMIS query and our dataset.

 $<sup>^{32}</sup>$ We censor the response to the survey question on the number of previous job offers at 10.

than half a year. Participants typically receive job seekers allowance and housing allowance, while the amount of other benefits received is quite low. The online survey participants are not significantly different on most dimensions, except that they attended more job interviews.

To check the balance between treatment and control group we also report demographics and job search history separately by group in Table 5. Only one out of 19 variables - the number of children - displays significant differences between the groups. This indicates balance of the sample. Balance is further corroborated by the fact that also none of the 14 measures of search behavior during the first three weeks of the study shown in the lowest panel in Table 5 displays any significant differences. We discuss these further in the next subsection. A more formal assessment of balance through a Holm-Bonferroni test across either all 19 baseline variables or across all 33 variables including initial job search does not reject equality between the groups even at the 10% level. We will now turn to document basic patterns of job search amongst the unemployed in our sample.

#### 4.2 Descriptives of Job Search Behavior During the Study

In terms of job search behavior in our study over the first three weeks, we find that the control group lists on average 493 vacancies, of which 25 are viewed, and 10 are saved (see third panel in Table 5). Out of these, participants report to have applied to 3 and eventually get an interview in 0.1 cases. Furthermore, they report about 9 weekly applications through channels outside our study, leading to 0.5 interviews on average. For the sets of listed vacancies and applications we compute a measure of occupational breadth (as described in subsection 3.7), of which the average values are also shown. Participants in the control group report 11 hours of weekly job search in addition to our study. In the weekly survey, participants were also asked to rate to what extent particular problems were a concern to them. On average, health problems are not mentioned as a major concern, while financial problems and strong competition in the labor market seem to be important. Finally, about 30% met with a case worker at the Jobcentre in a particular week. The values for job search behavior during the first three week for the treatment group are very similar.

Comparing job search behavior and outcomes after week three between treatment and control group is at the heart of the empirical assessment of the next section. Here we simply report some additional observations to provide some background.

First, about a third of job seekers search for jobs in the exact same occupation of their previous employment. We compare the occupations that they list in their employment history (obtained in the initial survey) with the three "preferred occupations" that they list when asked in which occupations they would prefer to find a job.<sup>33</sup> To be precise, we compute the share of their previous occupations that are listed as preferred (future) occupations. This provides a measure of how close job search is to one's work history. We find that for 35%, all of their previous occupations are now listed as preferred occupations.<sup>34</sup> For 27%, some of their previous occupations are listed as preferred occupations, and

<sup>&</sup>lt;sup>33</sup>Note that these 3 preferred occupations are good proxies for actual search. We show this by comparing them to the first occupation that is specified in the alternative interface (unfortunately we can only do so for the treatment group). For 51% of the job seekers this first occupation is one of the three preferred occupations (at the 4-digit level). At the two-digit level 69% selects one of their 3 preferred occupations in the job search interface.

<sup>&</sup>lt;sup>34</sup>Note that about half of all participants only indicate one previous occupations.

Table 5: Characteristics of the treatment and control group

	(	Contro	l group		T	reatme	nt grou	.p	T-test
	mean	$\operatorname{sd}$	min	max	mean	$\operatorname{sd}$	$\min$	max	p-value
Demographics:									_
female (%)	42	0.5	0	1	43	0.5	0	1	0.83
age	36	11	18	62	36	12	18	64	0.85
high educ $a$ (%)	44	0.5	0	1	41	0.49	0	1	0.63
survey qualification level	4.2	1.9	1	8	4.4	1.9	2	8	0.36
white $(\%)$	80	0.4	0	1	80	0.4	0	1	0.97
number of children	0.66	1.1	0	5	0.38	0.81	0	5	0.01
couple (%)	25	0.43	0	1	21	0.41	0	1	0.41
any children (%)	31	0.46	0	1	24	0.43	0	1	0.17
Job search history:									
expect job within 12 weeks $(\%)$	0.59	0.49	0	1	0.58	0.5	0	1	0.93
vacancies applied for	75	156	0	1000	53	120	0	1000	0.18
interviews attended	0.43	0.71	0	5	0.54	0.95	0	6	0.28
jobs offered	0.37	0.97	0	5	0.48	1.2	0	8	0.43
at least one offer $(\%)$	20	0.4	0	1	20	0.4	0	1	0.91
days unemployed (mean)	290	674	1	5028	228	558	1	5141	0.39
days unemployed (median)	81		1	5028	77		1	5141	
less than 183 days	0.75	0.43	0	1	0.78	0.42	0	1	0.60
less than 366 days	0.84	0.37	0	1	0.87	0.34	0	1	0.54
job seekers allowance $(\mathfrak{L})$	49	41	0	225	56	100	0	1005	0.46
housing benefits $(\pounds)$	65	124	0	600	62	135	0	660	0.90
other benefits $(\pounds)$	9.7	39	0	280	18	84	0	700	0.41
Weekly search activities in weeks 1-3:									
listed	493	399	4.3	3049	493	374	1	1966	1.00
viewed	25	14	3	86	26	18	0	119	0.57
saved	10	10	0	65	11	12	0	79	0.54
applied	3.3	5.8	0	45	2.5	4.3	0	33	0.14
interview	0.098	0.34	0	3.3	0.083	0.24	0	1.5	0.66
applications other	9.3	11	0	68	7.4	8.3	0	37	0.13
interviews other	0.54	0.71	0	4	0.47	0.77	0	5	0.48
broadness listed <sup><math>b</math></sup>	3.2	0.61	0	3.7	3.3	0.56	1	3.7	0.50
broadness applied <sup><math>b</math></sup>	3	0.95	0	4	3.2	0.9	0	4	0.34
hours spend <sup><math>c</math></sup>	11	8.3	0.5	43	12	10	1	43	0.15
concern health (scale 1-10)	1.5	2.6	0	10	1.7	2.7	0	10	0.48
concern financial (scale 1-10)	7.2	2.7	0	10	7	3.1	0	10	0.47
concern competition (scale 1-10)	7.4	2.3	0	10	7.2	2.2	0.5	10	0.43
met caseworker (%)	0.32	0.37	0	1	0.28	0.39	0	1	0.48
Observations	152				143				

Demographics and job search history values are based on responses in the baseline survey from the first week of the study. Search activities are mean values of search activities over the first 3 weeks of the study. <sup>a</sup> High educated is defined as a university degree. <sup>b</sup> Occupational broadness, as defined in section 3.7. <sup>c</sup> The number of hours spend on job search per week, as filled out in the weekly survey, averaged over week 2 and 3.

for 38%, none of their previous occupations are indicated to be preferred occupations. On average, each participant lists 46% of their previous occupations as preferred occupations. These numbers are computed using 4-digit occupation codes. If we use 3-digit codes, 51% of previous occupations are listed as preferred occupations, while this figure is 61% for 2-digit codes and 69% for 1-digit codes.

Second, most applications go to recently posted vacancies. The median age of a vacancy at the time of an application is 12 days. Of all applications to jobs from our search interface, 85% goes to a vacancy that is at most three weeks old at the time the application is reported. Only 7% goes to a vacancy that has been posted more than four weeks earlier. Since applications are reported once per week retrospectively, the age of vacancies is even slightly overestimated. In Figure 14 in the online appendix we show the full distribution of vacancy age at the time of the application.

Third, the breadth and the number of vacancies that job seekers list increase over time, while the numbers of applications and interviews decrease over time. There is no significant trend for breadth of applications though this is imprecisely measured, nor on weekly hours spent on job search or on the mean wage of jobs to which applications are sent. These results follow from regressing the outcome on a linear (weekly) time trend using only the control group and including individual fixed effects. The focus on the control group is to avoid any confounding with the treatment. The results are presented in Table 6. In column (1) we find no significant trend in the number of (self-reported) hours spent on job search per week. In column (2) we find that breadth of listed vacancies increases significantly each week with 0.015 which is about 2.2% of a standard deviation. Also the total number of listed vacancies in a week increases significantly (with 9 vacancies per week, see column (3)). The effect on breadth of applications (column (4)) is insignificant but so imprecisely measured that it is not statistically significantly different from the effect on breadth of listings. As we mentioned in the literature review, in a much larger dataset from a selected US job board Faberman and Kudlyak (2014) find a significant albeit slow increase in occupational breadth of applications as measured as the fraction not sent to the modal occupation. These trends contrast the intensity measures: the weekly number of applications through the lab (column (5)) and job interviews through the lab (column (6)) decrease significantly. Column (7) shows the estimate from a regression on the average wage of the applications. We find no significant time trend, but it should be noted that a large share of vacancies does not report a wage and thus this result could suffer from selection bias. One may worry that the results in Table 6 are affected by dynamic selection as some participants leave the study over time. In Table 16 in the Online Appendix we show the results of columns (1)-(5) for the subsample of participants that are still present in the final weeks of the study (i.e., attended at least one session in week 10, 11 or 12), and results are very robust. To sum up, breadth of listings increases but interviews decrease over time. This pattern is likely to be driven by duration per se, and obviously our experiment that attempts to increase breadth through information provision might generate very different relationships.

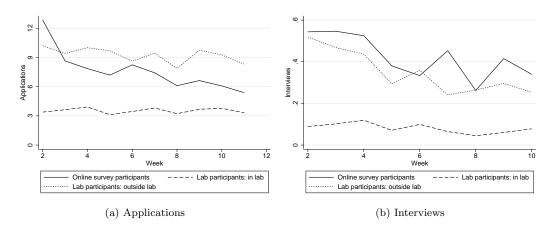
Forth, we investigate whether the requirement to search on our platform has an effect on job search per se by comparing the patterns we just described for the control group to those for online participants. Both groups face no intervention but one has to come physically to our lab to search on our standard interface. The online survey includes a question asking for the weekly number of applications sent and the weekly number of job interviews. As explained in Section 3.7 we associate applications and

Table 6: Job search activity over time (only control group)

			•	•	(1		
	(1)	(2)	(3)	(4)	(5)	(9)	
	Hours search	Breadth of	Number of	Breadth of	Number of	Number of	
	per week	listed vac.	listed vac.	applications	applications	interviews	•••
Time trend	0.040	0.015***	8.91**	-0.0052	-0.15**	-0.0072*	
	(0.063)	(0.0048)	(4.10)	(0.014)	(0.059)	(0.0043)	(60.2)
Individual FE	yes	yes	yes	yes	yes	yes	
Mean of dep. var.	12.2	3.29	536.1	3.07	3.38	0.082	19711.7
Weeks	1-12	1-12	1-12	1-11	1-11	1-10	1-11
Z	1040	1193	1196	504	1125	1049	654

All regressions contain only control group individuals. "Time trend" is a linear weekly trend. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01

Figure 5: Jobsearch behavior online and lab participants



interviews that result from search in the lab to the week in which the search activity was performed, and for reports on applications and interviews we adjust by the same average delay (one week for applications and two weeks for interviews). With this in mind, the average number of applications are shown in panel (a) of Figure 5 and the average number of interviews in panel (b) of Figure 5. For lab participants we observe both the number of applications from job search in the lab, and the number of applications reported through other job search activities. The number of applications outside the lab is quite similar to the number reported by the online participants, while the sum of the two types of applications for lab participants is higher than for the online participants. This difference could be the result of additional search induced through our intervention, even though we cannot rule out that it is the result of selection of more motivated participants into the lab study. In panel (b) we find that the sum of interviews in- and outside the lab is very similar to the number reported by the online participants.<sup>35</sup>

#### 4.3 Attrition

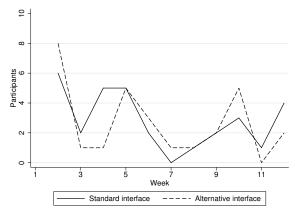
The study ran for 12 weeks, but job seekers could obviously leave the study earlier either because they found a job or for other reasons. Whenever participants dropped out, we followed up on the reasons for dropping out. In case they found a job, we asked for details, and in many cases we were able to obtain detailed information about the new job. Since job finding is a desirable outcome related to the nature of our study, we present attrition excluding job finding in of Figure 6. An exit from the study is defined to occur in the week after the last session in which the individual attended a lab session. In most weeks, we lose between 2 and 4 participants, and these numbers are very similar in control and treatment groups. On average, we have 8.3 observations per participant.<sup>36</sup>

We now investigate whether the composition of the control and treatment group changes over time

<sup>&</sup>lt;sup>35</sup>In Figure 13 in the online appendix we show the weekly sum of the two sources of applications and interviews for lab participants and include confidence intervals. The number of applications differs significantly for online and lab participants in most weeks while the number of interviews is never significantly different.

<sup>&</sup>lt;sup>36</sup> In the online appendix we show the distribution of the number of attended weeks per participant, split by preintervention (weeks 1-3) and post-intervention (weeks 4-12). See figures 10, 11 and 12.

Figure 6: Attrition of participants in the standard and alternative interface groups (excluding job finding)



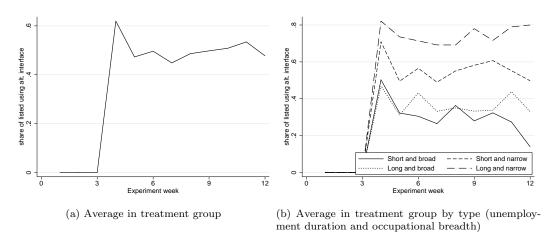
due to attrition, by looking at observable characteristics of those that remain in the study. We compute mean values of the same set of variables as in Table 5, for individuals remaining in the study in week 1, 4 and 12. For each of these groups of survivors, we test whether the treatment and control group are significantly different. Since we present 32 variables for three groups of survivors, this implies 96 tests. The resulting p-values are presented in Table 32 in the Online Appendix. Only 6 of the p-values are smaller than 0.10, so there is no indication that attrition leads to systematic differences in the composition of the treatment and control group. Also a Holm-Bonferroni test for joint significance does not reject the null hypothesis of identical values.

The apparent lack of selection is on the one hand helpful to study how the intervention may have affected search outcomes, on the other hand it hints that we are unlikely to capture differences in job finding rates, which are low overall. We will come back to the analysis of drop out and job finding in more detail in Subsection 5.5.

#### 4.4 Use of alternative interface

An obvious question regarding our treatment intervention is whether participants actually use the alternative interface. They are free to revert back to the standard interface, and in this sense our intervention can be considered an intention-to-treat. We are hesitant to adopt this interpretation since all participants in the treatment group used the alternative interface at least once and were therefore exposed to recommendations and suggestions based on their declared "desired" occupation. It could be that they used this information when they revert back to searching with the standard interface. With this in mind, we report information on actual usage. Panel (a) of Figure 7 plots the fraction of users of the alternative interface over the 12 weeks. On average we find that around 50% of the listed vacancies of the treated participants come from searches using the alternative interface over the 8 weeks and this fraction remains quite stable throughout. This does not mean that only 50% of the treatment group is treated, though. As long as participants use the interface at least once, they will have been exposed to a set of suggestions they may incorporate in their future search, whether

Figure 7: Share of listed vacancies that results from using the alternative interface



they continue searching with the new tool or not. $^{37}$  We discuss panel b) that considers subgroups of participants later on.

## 5 Analysis and Results

As outlined in the introduction, the hypothesis behind the intervention is that providing information about other occupations will allow individuals to explore vacancies from a larger set of occupations. This should hold in particular for individuals that otherwise explore few occupations. Exploring more occupations could go along with more search, or with the same search effort concentrated on more occupations but in a closer geographic region. The hypothesis is that this leads to more job interviews. For job seekers who already explore many occupations our intervention could backfire if they reduce their occupational breadth and their search effort. For job seekers who are longer unemployed appear more inclined to consider a larger set of occupations (recall column two in Table 6) and since their institutional incentives to do so are larger (recall Section 3.4), a differential effect by unemployment duration can be expected. An illustrative model behind these intuitive predictions is provided in Section 6. The null hypothesis against which we test is that information provision has no effect, which is conceivable if job seekers or their advisers at the job centre are already aware of the (publicly available) information that we provide. The following lays out the empirical strategy to investigate this.

<sup>&</sup>lt;sup>37</sup>The variation in usage results from both between and within users. The participants in the treatment group use the alternative interface for at least one search in 75 % of the weeks on average, and in 35 % of the weeks the alternative interface was used solely. These findings are shown in Figure 16 (panel (a)) in the Online Appendix. When aggregating all listed vacancies of treatment group users over week 4-12, we find that 22% have all vacancies returned by the alternative interface, while 76% have vacancies returned from both interfaces (see panel (b)). In panel (c) we investigate whether this pattern changes over time. We show that the shares of users in the treatment group that (i) uses only the alternative interface, (ii) uses only the standard interface and (iii) uses both interfaces are all very constant across the nine experiment weeks.

#### 5.1 Econometric Specification

Our data is a panel and our unit of observation is at the week/individual level. That is, we compute a summary statistic for each individual of her search behavior (vacancies listed, applications, interviews) in a given week; see Section 3.7 for a description of the outcome measures of interest. Since it is a randomized controlled experiment in which we observe individuals for three weeks before the treatment starts, the natural econometric specification is a model of difference-in-differences. To take account of the panel structure we include individual random effects. By design, there should be no correlation between individual characteristics (observable and unobservable) and treatment assignment, at least initially. To test whether the Random Effects specification is appropriate for the entire duration of the study, we have estimated a fixed effects model and performed a Hausman test for each of the main specifications. In none of the cases we could reject that the random effects model is consistent, such that we decide in favor of the random effects model for increased precision.<sup>38</sup> We discuss robustness at the end of this section (Subsection 5.6) where we show that point estimates are similar when using individual fixed effects yet precision is lower. As has been emphasized by Bertrand et al. (2004), serial correlation is an issue in difference-in-differences models. We follow their suggestion and average the weekly observations into two observations per individual, one before (weeks 1-3) and one after the intervention (weeks 4-12), but again report robustness to alternative specifications at the end of this section.

We compare a variable measuring an outcome (Y) in the control and treatment group before and after the week of intervention, controlling for time period fixed effects  $(\alpha_t)$ , before or after the intervention), time–slot × wave fixed effects  $(\delta_g)$  and a set of baseline individual characteristics  $(X_i)$ to increase the precision of the estimates. The treatment effect is captured by a dummy variable  $(T_{it})$ , equal to 1 for the treatment group in the period after the intervention. The specification is:

$$Y_{it} = \alpha_t + \delta_q + \gamma T_{it} + X_i \beta + \eta_i + \epsilon_{it} \tag{1}$$

where *i* relates to the individual, *t* to the time period and  $\eta_i + \epsilon_{it}$  is an error term consisting of an individual specific component  $(\eta_i)$  and a white noise error term  $(\epsilon_{it})$ . Individual characteristics  $X_i$  include gender, age and age squared, unemployment duration and unemployment duration squared<sup>39</sup> and dummies indicating financial concerns, being married or cohabiting, having children, being highly educated and being white. Standard errors are clustered at the individual level in the regressions, to account for any remaining correlation of an individual's observations.

As mentioned earlier, one important challenge with such approach has to do with attrition. If there is differential attrition between treatment and control groups, it could be that both groups differ in unobservables following the treatment. We proceed in two ways to address this potential concern. First, in Section 4.3 we document attrition across treatment and control groups and find no evidence of asymmetric attrition in terms of observable characteristics. Second, our panel structure allows us to control for time-invariant heterogeneity and use within-individual variation. When we estimate

<sup>&</sup>lt;sup>38</sup>We performed a Hausman test, testing for a difference between the treatment coefficient estimates in a random effect and a fixed effects model. Results can be found in the online appendix in Table 18.

<sup>&</sup>lt;sup>39</sup>Unemployment duration is defined as the reported duration at the start of the study.

a random and fixed effects model, as mentioned above the Hausman test fails to reject the latter. Even though the treatment itself is assigned at the group-level and it is unlikely to be correlated with unobserved individual characteristics, differential attrition could create correlation between the treatment and unobservable individual characteristics. This would then lead to rejection of the random-effects model. The fact that we can never reject this model is thus another indication against differential attrition between treatment and control groups.

Another important aspect relevant for the econometric specification is the potential heterogeneity of effects across individuals. Given the nature of the intervention, it is likely that the treatment affects different individuals differentially. In order for our intervention to affect job search and job prospects, it has to open new search opportunities to participants and participants have to be willing to pursue those opportunities. Participants may differ in terms of their search strategies. We expect our intervention to broaden the search for those participants who otherwise search narrowly, which we will measure by their search in the weeks prior to the intervention. For those who are already searching broadly in the absence of our intervention it is not clear whether we increase the breadth of their search. We therefore estimate heterogeneous treatment effects by initial breadth (splitting the sample at the median level of breadth over the first three weeks).<sup>40</sup>

Second, the willingness to pursue new options depends on the incentives for job search, which change with unemployment duration for a variety of reasons. Longer-term unemployed might be those for whom the search for their preferred jobs turned out to be unsuccessful and who need to pursue new avenues, while they are also exposed to institutional incentives to broaden their search (the Jobcentres require job seekers to become broader after three months). Note again that we are always comparing otherwise identical individuals in the treatment and control groups, so the incentives to broaden their search by themselves would not be different, but the information we provide to achieve this differs. We therefore also interact the treatment effect with a dummy for above median unemployment duration. In the subsequent section we provide a simple theoretical model formalizing the channels that may explain differential effects.<sup>41</sup>

Apart from these dimensions for which we have clear reasons for separate investigation we do not explore other dimensions of heterogeneity for which we have less clear reasons for investigation to avoid data mining. Nevertheless it might be interesting to know whether breadth of search is correlated with other factors that might drive the observations we report. We investigate this by

$$Y_{it} = \theta D_i + \alpha_{1t} D_i + \alpha_{2t} (1 - D_i) + \delta_g + \gamma_1 T_{it} D_i + \gamma_2 T_{it} (1 - D_i) + X_i \beta + \eta_i + \epsilon_{it}$$
(2)

Thus, the specification contains an additional baseline difference between the groups  $(\theta)$ , differential time period effects for the the groups  $(\alpha_{1t}$  and  $\alpha_{2t})$  and differential treatment effects between the groups  $(\gamma_1$  and  $\gamma_2)$ . Note that since we average observations into two period (before and after the intervention),  $\alpha_{1t}$  and  $\alpha_{2t}$  simply contain a time effect for the second period. Note also that, just as in the baseline model, the specification contains time-slot X wave dummies  $(\delta_g)$  and since treatment is assigned at the time-slot-level, these control for any baseline differences between the control and treatment group.

<sup>&</sup>lt;sup>40</sup>To check the robustness of our classification of job seekers as narrow or broad searchers, we used three different ways of doing this classification (based on listed vacancies in week 1, week 2 and week 3) and checked whether the classifications are consistent. We find that the classifications of week 1 and 2 agree on 69 % of the job seekers, those of week 1 and 3 agree on 67 % of the job seekers and those of week 2 and 3 agree on 86% of the job seekers.

 $<sup>^{41}</sup>$ When estimating heterogeneous effects we adapt our specification to include all necessary additional terms. Define  $D_i$  to be an indicator equal to one for individuals belonging to group 1 (for example narrow searchers) and equal to zero for individuals belonging to group 2 (for example broad searchers). We estimate:

Table 7: Effect of intervention on listed vacancies

	Bread	lth of	Number of
	listi	ings	listings
	(1)	(2)	$\overline{}(3)$
	Occupational	Geographical	Lab
Treatment	0.13**	-0.01	-34.99
	(0.06)	(0.02)	(52.09)
Treatment X occupationally broad	-0.07**	0.01	-23.71
	(0.04)	(0.03)	(90.08)
X occupationally narrow	0.34*** (0.10)	-0.03 $(0.03)$	-41.84 (64.01)
Model	Linear	Linear	Linear
Observation weeks	1-12	1-12	1-12
N	540	541	541

Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01

regressing it on a number of individual characteristics. Results are presented in columns (1) and (2) in Table 17 in the online appendix. We find that breadth of search is not easily predicted based on individual characteristics. Almost all variables are not statistically different from zero, and the  $R^2$  of the regression is low (0.18). The same holds for unemployment duration (columns (3) and (4)).

For the sake of brevity, in the main body we only present the results on the treatment effect ( $\gamma$ ) as well as the interaction effects between the treatment and the subgroups of interest. In Table 23 in the Online Appendix we report full results including all other covariates for the main regressions. We report results without adjusting for the actual use of the interface in the control group, but discuss an alternative empirical specification in which treatment assignment is used as an instrument for usage to capture intention-to-treat under robustness at the end of this section - with all the caveats mentioned in Section 4.2.

#### 5.2 Effects on Listed vacancies

We first look at the effects on listed vacancies - both in terms of number and breadth. We have two variables measuring how broad participants search, one in terms of occupation (as described in Section 3.7), the other in terms of geography (fraction of vacancies outside Edinburgh metropolitan area). We also measure the number of vacancies that were listed.

We estimate a linear model with individual random effects (equation (1)). The results are presented

in Table 7. The first row presents a significant positive overall effect on breadth of search in terms of occupation. The breadth measure increases with 0.13, which amounts to approximately one-fifth of a standard deviation. Another way to assess the magnitude of this effect is to compare it to the natural increase in breadth of listings over time (as discussed in the previous section), which implies that the treatment effect is equivalent to the broadening that on average happens over 9 weeks. We find no significant evidence of an overall effect on geographical breadth or on the number of listed vacancies.

In rows two and three in Table 7 we split the sample according to how occupationally broad job seekers searched in the first three weeks. We find clear heterogeneous effects: those who looked at a more narrow set of occupations in the first three weeks become broader, while those who were broad become more narrow as a result of the intervention. Note that these effects are not driven by 'regression to the mean' since we compare narrow/broad searchers in our treatment group to similarly narrow/broad searchers in our control group. 42 We again find no significant effects on the geographic distance of job search nor on the number of job applications. 43,44 Nevertheless, the point estimates would be consistent with the view that the narrow subgroup extends their geographic breadth by searching geographically closer to home, possibly because they are shown more vacancies in their local vicinity. This could explain how they can search occupationally broader without looking at more listings, though there might also be a substitution of broader listings away from the previous narrow listings. The total effects on job prospects remains in either case an empirical matter that we take up in subsequent sections.

The different effects on occupational breadth can be reconciled in a setting where broad searchers find many occupations plausible and use the additional information to narrow down the suitable set, while narrow searchers find few occupations suitable and use the additional information to broaden this set. This mechanism is more formally described in Section 6.

Finally, we split the effect further depending on how long job seekers have been searching for a job and present the results in Table 8. We interact the intervention effect with two groups: short term unemployed (with unemployment duration of less than the median of 80 days) and long term unemployed (with unemployment duration above the median). The effect is estimated for four groups: interactions of occupational breadth and unemployment duration. We find that results do not change much, though standard errors are larger. We still find that occupationally narrow searchers become broader while those that were already broad become more narrow, irrespective of unemployment duration. Shorter unemployed job seekers seem to consult less listings in the treatment group, significantly so for broader ones. If the new information allowed them to focus their search better this might not necessarily harm their job prospects, as outlined in our theoretical model, but nevertheless this remains

<sup>&</sup>lt;sup>42</sup> In Figure 15 in the online appendix we show the mean breadth of the different groups before and after the intervention to clarify further that these results are not caused by regression to the mean.

<sup>&</sup>lt;sup>43</sup>In the Online Appendix we also report estimates where we split the sample according to breadth along the geographical dimension at the median (see Table 20). The results are similar (those who were searching broadly become more narrow and vice versa, and there is some trade-off with occupational breadth). This could still be driven by initial occupational breadth, since this is negatively correlated with initial geographical breadth (coefficient -0.36) and is not controlled for. Indeed, when we split both by occupational and geographical breadth the effects are driven by the occupational dimension, which we will henceforth focus on.

<sup>&</sup>lt;sup>44</sup>The difference in the number of observations between the columns in Table 7 and similar tables that follow is due to the fact that we can only compute the occupational (geographical) breadth measure if the number of listed is two (one) or larger, which excludes different numbers of observations depending on the variable of interest.

Table 8: Effect of intervention on listed vacancies - interactions

		dth of	Number of
		ings	listings
	(1)	(2)	(3)
	Occupational	Geographical	Lab
Treatment			
X long unempl. and occ. broad	-0.10**	0.06	189.12
	(0.05)	(0.04)	(135.01)
X short unempl. and occ. broad	-0.05 $(0.05)$	-0.04 $(0.05)$	-252.80** (120.19)
X long unempl. and occ. narrow	0.36** (0.15)	-0.04 $(0.05)$	23.35 $(62.51)$
X short unempl. and occ. narrow	0.32** (0.13)	-0.01 (0.05)	-112.82 (116.52)
Model Observation weeks	Linear 1-12	Linear 1-12	Linear 1-12
N	540	541	541

Each column represents one regression. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

a concern that we return to when we consider the effect on job interviews.

#### 5.3 Effects on Applications

The second measure of search behavior relates to applications. We have information about applications based on search activity conducted inside the laboratory as well as outside the laboratory which we collected through the weekly surveys. The distribution of applications contains a large share of zeros (in almost 50% of the weekly observations there are zero applications through the lab). Therefore we estimate a negative binomial model, with individual random effects.<sup>45</sup> For these models we report  $[\exp(\text{coefficient}) - 1]$ , which is the percentage effect.

The results are presented in Table 9. We find no overall treatment effect on applications, except for a decrease in their geographical breadth (approximately one-fifth of a standard deviation). When we split the sample according to initial occupational breadth, we find a similar pattern as for listings. Those who searched more narrowly in terms of occupation become occupationally broader, while those that searched broadly become more narrow. The estimates are significantly different from zero at the 10 % level. We find no effects on the number of applications for either group (columns (3) - (5)),

<sup>&</sup>lt;sup>45</sup>Due to overdispersion in the distribution of applications, we prefer a negative binomial model over a Poisson model. However, negative binomial regressions are sometimes less robust and in addition no consensus exists on how to include fixed effects (Allison and Waterman (2002)). Furthermore, we can not cluster standard errors with the random effects negative binomial regressions. Therefore we also report results from Poisson regressions in the Online Appendix (Table 19). The findings are similar.

Table 9: Effect of intervention on applications

		dth of ations		Number of applications	
	(1) Occupational	(2) Geographical	(3) Lab	(4) Outside lab	(5) Total
Treatment	0.03 (0.20)	-0.06* (0.03)	0.09 (0.16)	-0.03 (0.09)	0.01 (0.09)
Treatment					
X occupationally broad	-0.43*	-0.02	-0.08	-0.06	-0.05
	(0.22)	(0.05)	(0.19)	(0.13)	(0.12)
X occupationally narrow	0.49* (0.29)	-0.09** (0.04)	0.27 $(0.27)$	-0.02 (0.13)	0.08 $(0.13)$
Model	Linear	Linear	Neg. Bin.	Neg. Bin.	Neg. Bin.
Observation weeks $N$	1-11 305	1-11 363	1-11 541	1-11 490	1-11 487

Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (3)-(5) are Negative Binomial regression models where we report [exp(coefficient) - 1], which is the percentage effect. Standard errors in parentheses (clustered by individual in column (1) and (2)). \* p < 0.10, \*\*\* p < 0.05, \*\*\* p < 0.01

though point estimates might indicate a pattern where the initially-narrow group expands broadness through more applications and vice versa for the initially-broad subgroup. There is also a negative effect on geographical breadth for the occupationally narrow job seekers (column (2)).<sup>46</sup>

Again, we split these effects by the duration of unemployment and report results in Table 10. In column (1), we find that occupational breadth goes down significantly for long term unemployed broad searchers. It increases most for long term unemployed narrow searchers, yet this is insignificant due to large standard errors. This increase is accompanied by a significant decrease in geographical distance. There is a significant reductions in the occupational breadth for longerterm unemployed broad participants. Estimates on the number of applications are insignificant, though point estimates are economically large. As noted earlier, even decreases in occupational breadth can be beneficial if job search becomes better targeted.

#### 5.4 Effects on Interviews

We now turn to interviews, the variable that is most closely related to job prospects. Since the number of interviews per week is always very small, we cannot compute breadth measures. So we only look at a measure of the number of interviews obtained as a result of search conducted inside the laboratory

<sup>&</sup>lt;sup>46</sup>When splitting the sample according to how narrow people searched in terms of geography, we find no evidence of heterogeneous effects. Results are presented in the Online Appendix in Table 21.

Table 10: Effect of intervention on applications - interactions

		th of ations		Number of applications	
	(1)	(2)	(3)	(4)	(5)
	Occupational	Geographical	Lab	Outside lab	Total
Treatment					
X long unempl. and occ. broad	-0.67***	-0.07	-0.24	-0.22	-0.20
	(0.25)	(0.06)	(0.21)	(0.14)	(0.13)
X short unempl. and occ. broad	-0.18	0.02	0.17	0.17	0.17
	(0.33)	(0.07)	(0.36)	(0.22)	(0.21)
X long unempl. and occ. narrow	0.51	-0.10**	0.42	-0.11	0.00
	(0.34)	(0.05)	(0.40)	(0.16)	(0.17)
X short unempl. and occ. narrow	0.40	-0.08	0.25	0.14	0.22
	(0.41)	(0.06)	(0.37)	(0.20)	(0.21)
Model	Linear	Linear	Neg.	Neg.	Neg.
			Bin.	Bin.	Bin.
Observation weeks	1-11	1-11	1-11	1-11	1-11
N	305	363	541	490	487

Each column represents one regression. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (3)-(5) are Negative Binomial regression models where we report [exp(coefficient) – 1], which is the percentage effect. Standard errors in parentheses (clustered by individual in column (1) and (2)). \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

and outside the laboratory.<sup>47</sup> Because of the large share of zeros, we estimate a Poisson model with individual random effects. Again we report [ $\exp(\text{coefficient}) - 1$ ], which is the percentage effect.

Results are presented in Table 11. There is a positive effect of the treatment of 44% on the total number of interviews, which is significant at the 10% level. We also find positive effects on interviews on the two separate dimensions of search in the lab and search outside the lab, but even though the point estimate for the effect within the lab is highest only the increase in out-of-lab interviews is statistically significant. This can be explained by the difference in base rate which is lower in the lab making statistical inference more difficult: In the pre-treatment period the number of interviews through the lab was 0.09, while the number of interviews through other channels was 0.53.

When splitting the sample according to breadth of search, we find that the effect is entirely driven by those who searched narrowly in terms of occupation. For this group the number of interviews increases for search activity conducted both in the lab and outside (though again, only the increase of the out-of-lab interviews is statistically significant). This seems to indicate that the additional

 $<sup>^{47}</sup>$  For interviews reported outside the lab we censor observations at 3 interviews per week, because of some outliers. Results are similar when no such restriction is imposed. As a check of consistency, we also check whether interviews are ever reported without preceding applications. We find that in 98.2 % of the weeks in which an interview is reported, a positive number of applications was reported in at least one of the two preceding weeks.

Table 11: Effect of intervention on interviews

	Number of			
		interviews	;	
	(1)	(2)	(3)	
	Lab	Survey	Total	
Treatment	0.61	0.40*	0.44*	
	(0.79)	(0.27)	(0.28)	
Treatment				
X occupationally broad	-0.37	-0.00	-0.07	
	(0.43)	(0.28)	(0.24)	
X occupationally narrow	1.13	0.86**	1.03***	
- •	(1.26)	(0.47)	(0.55)	
Model	Poisson	Poisson	Poisson	
Observation weeks	1-10	1-10	1-10	
N	540	466	464	

Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (1)-(3) are Poisson regression models where we report [exp(coefficient) - 1], which is the percentage effect. Standard errors clustered by individual in parentheses. \*  $p < 0.10, \, *** \, p < 0.05, \, **** \, p < 0.01$ 

information is not only helpful for search on our platform, but also guides behavior outside.<sup>48</sup> The point estimates for the occupationally broad group are all insignificant and in absolute value much smaller, but point estimates are negative.

When we further split the sample according to length of unemployment duration, we find that the positive treatment effects on the narrow searchers is mainly driven by the long term unemployed narrow searchers. This group gets a significant increase in the number of interviews both as a result of search activity done inside the lab and outside the lab.<sup>49</sup> These findings highlight that our intervention is particularly beneficial to people who otherwise search narrowly and who have been unemployed for some months. It might be encouraging that there are no significant negative effects on the groups that became occupationally narrower, but some of the negative point estimates should require further investigation.

The set of weekly interviews is too small to compute breadth measures. We did, however, ask

<sup>&</sup>lt;sup>48</sup>We find some evidence of heterogeneity in treatment effects when we split the sample according to initial geographical breadth, with a large positive significant treatment effect for those who searched broadly geographically. Results are presented in the Online Appendix in Table 22.

<sup>&</sup>lt;sup>49</sup>The extremely large value of the increase in lab interviews for the long term narrow searchers is partly due an individual outlier that reported an average of 3.5 interviews per week in the treatment period. If we exclude this individual, the coefficient is still large, positive and statistically significant (6.75\*\*\*).

Table 12: Effect of intervention on interviews - interactions

	Number of			
		interviews		
	(1)	(2)	(3)	
	Lab	Survey	Total	
Treatment				
X long unempl. and occ. broad	-0.27	-0.23	-0.21	
	(0.72)	(0.25)	(0.26)	
X short unempl. and occ. broad	-0.37	0.17	0.01	
	(0.52)	(0.47)	(0.37)	
X long unempl. and occ. narrow	13.12***	2.44***	3.39***	
	(9.25)	(1.19)	(1.42)	
X short unempl. and occ. narrow	-0.26	0.31	0.30	
	(0.51)	(0.40)	(0.44)	
Model	Poisson	Poisson	Poisson	
Observation weeks	1-10	1-10	1-10	
N	540	466	464	

Each column represents one regression. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (1)-(3) are Poisson model regressions where we report [exp(coefficient) – 1], which is the percentage effect. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\*\* p < 0.05, \*\*\* p < 0.01

individuals at the beginning of the study to indicate three core occupations in which they search for jobs, and we observe whether an interview was for a job in someone's core occupation or for a job in a different occupation. We had seen earlier that the alternative interface was successful in increasing the occupational breadth of listed vacancies and applications, and separate treatment effects on interviews in core vs non-core occupations allow some assessment of whether this lead to more "breadth" in job interviews. Results are presented in Table 13. We indeed find that the increase in the number of interviews relative to the control group comes from an increase in non-core occupations that were not their main search target at the beginning of our study, though due to low precision the effect is not statistically significant. As the number of interviews becomes small when splitting between core and non-core, we cannot split the sample further by subgroups.

One may worry that the increase in interviews in non-core occupations is associated with different quality of the interviews. For example, the suggestions could lead to interviews for jobs with different wages. We have investigated this by comparing the average wage of listed vacancies, applications and interviews and find that the alternative interface does not significantly change the wage of any of these. $^{50}$ 

 $<sup>^{50}</sup>$ We computed for every individual in every week the average wage of listed vacancies, applications or interviews and performed regressions similar to our main specifications.

Table 13: Effect of intervention on interviews: core and non-core occupations

	Number of i	nterviews (in the lab)
	(1)	(2)
	Core	Non-core
Treatment	-0.14	0.75
	(0.72)	(0.85)
Model	Poisson	Poisson
Observation weeks	1-10	1-10
N	540	540

Each column represents one regression. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (1)-(2) are Poisson model regressions where we report [exp(coefficient) – 1], which is the percentage effect. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Our findings suggest that the alternative interface may be more beneficial to those that search narrowly and have been relatively long unemployed. This finding is supported by statistics on usage of the interface over time. Panel (b) of Figure 7 shows the evolution of the fraction of treated participants using the interface, splitting the sample by occupational breadth and unemployment duration. We find that long term narrow searchers are indeed using the interface more than the other groups (with around 75% of them using the interface in contrast to around 45% for the other groups), and this difference is statistically significant. The fractions remain quite stable over the 8 weeks. This finding supports the intuition that some groups of job seekers benefit more from the intervention and are therefore more willing to use the alternative interface. This group, the long-term unemployed narrow searchers is exactly the group for which we find the most pronounced positive effects. The idea that these groups are more willing to use the alternative interface is supported by responses from the baseline survey in the first week. The participants were asked to specify how long they expected it would take to find a job. Within the group of short-term unemployed the median response is "less than 3 months" which might indicate a rather clear idea of how to obtain a job, while for the long-term unemployed group the median response is "less than 6 months" which might indicate a less clear view and more scope to provide successful alternatives.<sup>51</sup>

#### 5.5 Effects on Job Finding

We now briefly turn the analysis of job finding. As mentioned earlier, the study was not designed to evaluate effects on job finding and, given the size of the sample, we should be cautious in interpreting any results we have. Also, one should keep in mind that attrition from one week to the next for

<sup>&</sup>lt;sup>51</sup>We only asked this question once, in the first week. Asking it on a weekly basis might have affected people's behavior by structurally emphasizing that they had not yet succeeded in finding employment.

Table 14: Summary statistics on job finding and drop out for weeks 3 and 12

	In Study - No Job	Found a Job	Out of Study	Job fin	ding week <sup>+</sup>
				mean	(std)
Week 3					
Standard interface	130	12	9	2.2	(0.6)
Alternative interface	128	10	6	2.1	(0.7)
Week $12^{++}$					
Standard interface	72	36	19	7.6	(2.2)
Alternative interface	79	27	18	8.1	(2.6)

 $<sup>^{+}</sup>$  Job finding week conditional on finding a job by the respective week.  $^{++}$  Outcome by week 12 for individuals that were still present in week 4.

unexplained reasons is low but of the same order of magnitude as the confirmed job finding rate.<sup>52</sup>

We classify job seekers in three categories depending on the information recorded in week 3 (before the intervention) and week 12 (last week of the intervention): Job seekers are either (1) present in the study and having no job ("no job"), (2) not present in the study and unclear outcome ("out of study"), (3) not present in the study and having found a job ("job").

Table 14 presents the distribution of job seekers across categories, as well as the average length (in weeks) job finders had to wait to find a job. Note that we record the week they accepted a job offer, not the week the job actually started. For week 12, we report the conditional distribution for those who were still in the study in week 4 and have therefore been exposed to the new interface if they were in the treatment group. There is indication that the job finding rate is slightly higher in the standard interface than in the alternative interface already in week 3, however this appears more pronounced in week 12, but since we have around 15% of our sample who dropped out and we do not know if they found a job or not, it is difficult to draw conclusions based on these numbers.

These numbers are nevertheless useful to get a sense of the sample size one would need in order to capture significant effects on job finding. We perform a simple sample size calculation to illustrate how the required sample size for finding an effect on job finding exceeds the sample size required for finding an effect on the number of interviews. To detect a 44% increase in interviews due to the intervention (see Table 11), a sample size of 70 observations per treatment is required (so 140 in total). For job finding, detecting a similar sized effect requires around 3794 observations per treatment, due to a much lower base rate.<sup>53</sup> Even if one takes the (at most) 8 observations per individual in our study into account, it is clear that we lack power to identify any realistic effect on job finding.

Bearing this in mind, we estimate a simple duration model where the duration is the number of

 $<sup>^{52}</sup>$ We tried to follow-up by calling them at least 3 times, though for a non-trivial share of the attrition we still do not observe perfectly whether the person found a job or just quit the study.

 $<sup>^{53}</sup>$ The precise computation is as follows. We observe in the first three weeks that, on average, participants have a total of 0.61 interviews per week through the lab and other channels (see Table 5). To detect a 44% increase in interviews due to the intervention (see Table 11), such that the interview rate becomes 0.89, a sample size of 70 observations per treatment is required (so 140 in total). This number is based on an one-sided test with type-I error probability  $\alpha = 0.10$  and power  $1 - \beta = 0.80$ . The standard deviation is assumed to be 0.75 in both groups, based on the numbers reported in Table 5. For job finding, we observe 19 people finding a job in the first 3 weeks, which implies a weekly job finding rate of approximately 0.02. If we make the (strong) assumption that the additional interviews are equally likely to result in a job as the initial interviews, we would expect a 44% increase in job finding. Note that this is a conservative choice as this would be a very large effect. Still, to be able to pick up the increase in job finding from 0.02 to 0.0288 requires a sample size of 3794 people per treatment (similar test as for interviews).

Table 15: Treatment effects on job finding rate

	(1)	(2)
Theatment	-0.14	-0.18
Treatment	-	
	(0.25)	(0.31)
Treatment x Occupationally narrow		0.09
		(0.56)
N	253	253

Proportional Cox Hazard model, with time-slot fixed effects, and individual characteristics. We exclude observations censored at 3 weeks or less. Reported values are coefficients. \* p < 0.10.

weeks we observe an individual until she/he finds a job. Since we know when each individual became unemployed, we can calculate the total unemployment duration and use this as a dependent variable. This variable is censored for individuals who drop out of the study or who fail to find a job before the end of the study. We estimate a proportional Cox hazard model with the treatment dummy as independent variable, controlling for additional individual characteristics and group session dummies.

We report estimates for the entire sample and for the sub-samples conditioning on initial search type (narrow vs broad search). The results are presented in Table 15. We fail to find significant differences in the hazard rates across treatments. That is, we have no evidence that the job seekers exposed to the alternative interface were more or less likely to find a job (conditionally on still being present in week 4). Despite the negative point estimate for the treatment group, even increases in the hazard of the treatment group of the magnitude of the increase in interviews overall (29%) or for narrow individuals (52%) are well within the confidence interval of these estimates. That is not to say that lack of power is the only plausible reason for finding no effect. As mentioned in the introduction, interviews in broader occupations might not convert to jobs at the same rate. We return to advocating larger studies in the conclusion.

#### 5.6 Robustness: Alternative Specifications

In our analysis we made some choices regarding the specification (of the empirical model and of the definition of some variables). Below we discuss alternative choices and investigate whether our results are robust to these specifications. We consider (1) individual fixed effects instead of random effects,

- (2) weekly observations instead of aggregated data, (3) linear models instead of count data models,
- (4) excluding the last one or two observations per individual, (5) an alternative breadth measure and
- (6) IV regressions with the use of the alternative interface as the treatment intensity.<sup>54</sup>

<sup>54</sup> We also thank an anonymous referee for requesting additional analysis of heterogeneous effects by educational level. We don't find pronounced differences in the effectiveness of the intervention by education level. As mentioned in Section

As we discuss in section 5.1, we used individual random effects models in all empirical analysis up to this point to increase precision. A Hausman test does not reject validity of the random effects model. In Table 24 of the Online Appendix we show our baseline regressions using individual fixed effects instead of random effects. We include regressions with outcome variables: breadth of listed vacancies, breadth of applications, number of applications (in the lab, outside the lab, and total) and the number of interviews (in the lab, outside the lab, and total). For each outcome we show the overall effect and the effect by initial occupational breadth. We find very similar overall pattern but reduced precision and significance. Occupational breadth of listed vacancies increases significantly for narrow searchers, and decreases (slightly) for broad searchers. For breadth of applications the estimates suggest a similar pattern, but none of these are statistically significant. Similar to our baseline estimates we find no effect on the number of applications. For interviews we find large positive coefficients for narrow searchers, but due slightly reduced precision these are not statistically significant.

While we have (at most) 12 weekly observations per individual, we use data in all estimations that has been aggregated into two observations per individual (before and after the intervention). We do so to minimize problems related to serial correlation (as suggested by Bertrand et al. (2004)). We can estimate the same regressions including all observations. The specification is identical except that we now include 12 time fixed effects instead of 2. We present the key results in the online appendix in Table 25. We find that patterns are very similar: breadth of listed vacancies increases (strongly for narrow searchers), the same happens with breadth of applications (but not significantly) with no significant effect on the number of applications. The point estimate for the number of interviews remains economically large but slightly lower, and retains significance only for narrow searchers. For them it is significant both inside and outside the lab.

As a third robustness check we consider the model specification for the number of applications and interviews. Since these are count variables and contain many zeros, we use Poisson regressions or negative binomial regressions in the main analysis. One might wonder whether the use of these models drives our results. In Table 26 in the online appendix we present linear regressions for the main specifications in which we used non-linear models. These are the number of applications (in the lab, outside the lab, and total) and the number of interviews (in the lab, outside the lab, and total). We find similar patterns when using simple linear regression: there is no clear impact on applications, but the point estimate for interviews is economically large, and significant for narrow searchers.

The fourth robustness check considers the way we obtain our data on applications and interviews in the lab. As discussed, participants can save a vacancy if they are interested, and will be asked whether they applied in subsequent weeks. Once they have applied, they will be asked whether they received an interview. Most applications are sent in the first week after saving the vacancy (86%), while most interviews are obtained in the first two weeks (83%). As a result, we must observe an individual one week after saving a vacancy to obtain information about applying and two weeks after saving a vacancy to obtain information about an interview. This is the reason that we exclude week 12 in regressions for applications and weeks 11 and 12 in regressions for interviews. Alternatively,

<sup>5.1,</sup> we prefer to focus the analysis in the paper only on two obvious dimensions of heterogeneity though, to prevent data mining.

we can exclude for each individual his/her last one or two observations. The results from the main specification when using this approach are shown in the online appendix in Table 27. The results are very similar, with again no impact on applications overall though broad individuals apply significantly less broad and narrow individuals apply broader (but insignificantly). Also the significantly positive effects on interviews overall and for narrow individuals in particular remain.

Fifth, we consider our method for defining occupational breadth of job search. In our approach the distance between two occupations is based on the number of common digits of the two occupational codes (see section 3.7 for the detailed description). Alternatively, one can focus on a particular digit of the occupational code and call occupations identical if they share the same code up to that digit. Broadness can then be defined by the well-known Gini-Simpson index. The several different measures of breadth are highly correlated; for example, for listed vacancies our measure has a correlation above 0.95 with 4 different Gini-Simpson measures (see Tables 30 and 31 in the Online Appendix). Not surprisingly we find very similar results when we adopt this alternative measure. A more elaborate alternative is to use empirically observed transitions between occupations in labor market surveys to measure the "closeness" of the two occupations.<sup>55</sup> We apply this approach to measure the breadth of listed vacancies and the breadth of applications.<sup>56</sup> In addition, we use the breadth of listed vacancies in the first three weeks (as measured by this method) to define the groups "narrow" and "broad" searchers (as we do in all main analysis). The results of the main regressions are presented in Table 28. We find that the effect on breadth of listed is similar: breadth increases significantly for the full sample and the effect is larger for narrow searchers (though not significant due to slightly lower precision). Note that the new breadth measure has a different scale and to interpret the magnitude of the effect we include the standard deviation of the dependent variable in the table. We find that the effect of the intervention is about 1/6 of a standard deviation, which is very similar to the effect of 1/5 of a standard deviation in our baseline. For applications we find that results are very similar to our baseline results. The coefficients suggest an increase in breadth for narrow searchers and a decrease for broad searchers, though neither is statistically significant. We find no effect on the number of applications, and a significant increase in interviews (both in the lab and outside the lab) for narrow searchers.

Finally we consider an interpretation that all our results are intention-to-treat effects. Since using the alternative interface was voluntary for all individuals in the treatment group, some changed back to the normal interface quickly while other used it continuously for 9 weeks (we show the extent to which users use the alternative interface in Figure 7). One might argue that not all job seekers in the treatment group were treated (with the same intensity). We are hesitant to emphasize this interpretation too much, because suggestions about alternative occupations can affect job seekers even after a user switches back to the standard interface. They might simply search for the suggested occupations on the standard interface. The suggestions might even affect job search through other

 $<sup>\</sup>overline{^{55}\text{We thank an anonymous referee for this suggestion}}$ .

 $<sup>^{56}</sup>$ We use occupational transitions in the BHPS (that we also apply to generate suggestions in the search interface). The advantage of this approach is that theoretically this creates a continuous measure of closeness between two occupations and that this measure is based on real-world transitions. In praxis there is a downside due to sample size: the transitions identify a limited number of occupations to which transitions are somewhat common (often no more than 5) and assign a zero to the rest. The reason is the limited size of the BHPS relative to the large number of possible transitions (353 occupations lead to  $353^2 = 124,609$  possible transitions).

channels. However, for the sake of comparison, we can consider treatment assignment as an instrument for actual usage when estimating our empirical models. We define actual usage as the share of listings that is performed using the alternative interface in a particular week. This share is around 50% for the treatment group and differs substantially depending on the different groups (as is shown in figure 7). The results of estimating the effect of alternative interface usage, using treatment assignment as an instrument, are presented in Table 29. As expected, the estimates are larger in magnitude. We find that breadth of listed vacancies increases (with a coefficient of 0.24\*\* compared to 0.13\*\* in the baseline results). Additionally, we also find that breadth of applications increases significantly for narrow searchers. The number of applications is unaffected, and interviews increase significantly for narrow searchers.

## 6 An Illustrative Model

In the empirical section we saw that our information intervention increases occupational breadth: listings are broader and more job interviews are obtained, possibly driven by jobs outside the core occupations. Job interviews are increased particularly for long-term but narrow searchers, and there is an indication that they they apply more. Searchers who already search broadly without our intervention decrease their breadth. While it is obvious why narrow individuals are affected differently from broad ones, it might be less obvious why it is the longer-term unemployed that seem to react stronger to our information intervention. Here we briefly sketch a very stylized occupational job search model that is capable of organizing our thoughts about the driving forces. It is based on the idea that workers learn about the occupations in which they search for jobs, in the spirit of e.g. Neal (1999), with the difference that workers start with heterogeneous beliefs about different occupations and that we study information provision. The goal is not to provide the richest framework, but to provide a simple setup that captures the previous intuitive arguments in a coherent framework. Among other simplifications, we only model "breadth" in a crude way (neither distinguishing listings vs applications as these are qualitatively similar, nor incorporating geography).

A job seeker can search for jobs in different occupations, indexed  $i \in \{1, ..., I\}$ . For each occupation she decides on the level of search effort  $e_i$ . Returns to searching in occupation i are given by an increasing but concave function  $f(e_i)$ .<sup>58</sup> The returns to getting a job are given by wage w and are the same across occupation, and b denotes unemployment benefits. The cost of search is given by an increasing and convex function  $c(\sum e_i)$ .<sup>59</sup> A limiting case is a fixed total search effort  $\bar{e}$ , such that

<sup>&</sup>lt;sup>57</sup>Note that we use linear models for all instrumental variable specifications.

<sup>&</sup>lt;sup>58</sup>The decreasing returns capture that the number of job opportunities within an occupation may be limited. We are focusing on the individual worker's search here, and do not additionally model the aggregate matching function that might depend on the total number of vacancies and the number of other job seekers who explore the same occupation. All of this is suppressed as the individual takes it as given. For simplicity we also abstract from heterogeneity in occupations which might make the return to search occupation-specific. As mentioned, we also abstract from geography, though effects on breadth consistent with our empirical findings could easily be obtained by assuming that more search effort in a given occupation means applying to jobs that are further away geographically and that the benefit of a job equals the wage minus geographical distance.

 $<sup>^{59}</sup>$ In models with only one occupation it is immaterial whether c is convex or f concave or both. With multiple occupations, we chose a setup where the costs are based on total effort, which links the various occupations, while the return to search is occupation specific. In this setting, if returns were linear all search would be concentrated in only one market. If costs were linear, then changes in one market would not affect how much individuals search in other markets.

costs are zero up to that point and infinite thereafter.

The individual is not sure of her job prospects within the various occupations. Her job prospects are either good (in which case we denote her H - high - type) or bad (in which case we denote her L- low - type). If her job prospects are good she obtains a job in occupation i with arrival probability  $a_H f(e_i)$ , otherwise she obtains a job with probability  $a_L f(e_i)$ , where  $a_H > a_L = 0$ , where the equality is assumed only for simplicity. The uncertainty can be about whether the skills of the job seeker (still) meet the requirements of the occupation. The individual does not know whether she is a high or low type in occupation i, but assigns probability  $p_i$  to being a high type. So the individual's type is a vector  $(p_1, p_2, ..., p_I)$  of probabilities for each occupation, and is all that is relevant for the decision of the individual in this environment with binary outcomes in each occupation. Still, when we introduce the information content of the alternative interface later on, it will be convenient to make the additional assumption that the individual is unsure of the exact value of the probability in each of the occupations, and only knows its distribution  $Q_i$  with support  $[\underline{q}_i, \overline{q}_i]$  among people that are like her. Then  $p_i$  can be interpreted as the average belief according to  $Q_i$ . For technical convenience assume that types are not too good, i.e.,  $\bar{q}_i \leq 1/2$ , so that the average belief is also bounded by this number. This ensures that an occupation with higher belief also has higher variance and both increase the incentives to search in this occupation in such a simple bandit problem, which makes search incentives monotone in  $p_i$ .

Given this average prior and her effort, her expected chances of getting a job offer in occupation i are

$$h(p_i, e_i) = f(e_i)(p_i a_H + (1 - p_i) a_L).$$

Given a vector of beliefs  $p = (p_1, ..., p_I)$  and a vector of search effort in the various occupations  $e = (e_1, ..., e_I)$ , the overall expected probability of being hired in some occupation is

$$H(p,e) = 1 - \prod_{i} (1 - h(p_i, e_i))$$

where the product gives the probability of not getting a job offer in any occupation.

Assume the unemployed job seeker lives for T periods, discounts the future with factor  $\delta$ , if she finds a job this is permanent and pays wage w per period, and if she remains unemployed she obtains benefits b for that period. Obviously searching in an occupation changes the beliefs about it. An individual who has a prior  $p_i^t$  at the beginning of period t and spends effort  $e_i^t$  during the period but does not get a job will update her beliefs about the chance of being a high type in occupation i by Bayes rule. Let  $B(p_i^t, e_i^t)$  denote this new belief. For interior beliefs we have<sup>60</sup>

$$p_i^{t+1} = B(p_i^t, e_i^t) = \begin{cases} = p_i^t & \text{if } e_i^t = 0\\ < p_i^t & \text{if } e_i^t > 0, \end{cases}$$
 (3)

since there is no learning without effort, and the individual becomes more pessimistic if she does put effort but does not get a job. Let  $B(p, e) = (B(p_1, e_1), ..., B(p_I, e_I))$  denote the vector of updates.

So both play a separate role here.

<sup>&</sup>lt;sup>60</sup>The exact formula in this case is  $B(p_i^t, e_i^t) = p_i^t [1 - f(e_i^t)a_H]/[1 - p_i^t f(e_i^t)a_H - (1 - p_i^t)f(e_i^t)a_L]$ . Note also that beliefs do go up if the person finds a job, but under the assumption that the job is permanent this does no longer matter.

The state variable for an individual is the time period t because of her finite life-time, and her belief vector at the beginning of this period  $p = p^t$ . Given this, she chooses her search effort vector  $e = e^t$  to maximize her return. She obtains for sure her outside option of doing nothing in the current period: her current unemployment benefit payment and the discounted value of future search. Additionally, if she finds a job, she gets the lifetime value of wages  $(W_t)$  to the extent that they exceed her outside option. Finally, she has to pay the search effort costs. So the return to search is given by

$$R_t(p) = \max_{e} \left( b + \delta R_{t+1}(B(p, e)) + H(p, e) \left( W_t - (b + \delta R_{t+1}(B(p, e))) \right) - c(\sum_{i} e_i) \right)$$
(4)

The model implies that an individual may search in multiple occupations due to decreasing returns in each one. The distribution of her effort across occupations depends on the set of priors  $p_i$ ,  $i \in 1,...,I$ . For our purposes a two-period model suffices (for which  $R_3 = 0$ ,  $W_2 = w$  and  $W_1 = w(1 + \delta)$ ).<sup>61</sup> The first period captures the newly unemployed, and the second period the longer-term unemployed.

The unanticipated introduction of the alternative interface provides an additional source of information on occupations. It displays a list of occupations suitable for someone like her. In general, this implies that for these occupations the individual may update her beliefs positively, while for those not on the list she may update her beliefs downwards. To formalize this mechanism, assume that an occupation is only featured on the list if the objective probability  $q_i$  of having good job prospects exceeds a threshold  $\hat{q}$ . In the first period of unemployment this means that for any occupation on the list the individual updates her belief upward to the average of  $q_i$  conditional on being larger than  $\hat{q}$  (i.e.,  $p_i^1 = \int_{\hat{q}}^{\bar{q}_i} q_i dQ_i / \int dQ_i$ ). For occupations that are not on the list her beliefs decline to the average of  $q_i$  conditional of  $q_i$  being below  $\hat{q}$  (i.e.,  $p_i^1 = \int_{\underline{q}_i}^{\hat{q}} q_i dQ_i / \int dQ_i$ ). Obviously these updates also apply if the alternative interface is introduced at a later period of unemployment as long as the individual has not yet actively searched in this occupation.<sup>62</sup> The alternative interface induces an update in belief  $p_i^t$  when it is introduced, but given this update problem (4) continues to characterize optimal behavior.

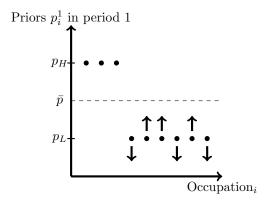
In order to gain some insights in how this affects the occupational breadth of search, consider for illustration two types of occupations. Occupations  $i \in 1, ..., I_1$  are the "core" ones where the job seeker is more confident and holds first period prior  $Q_i = Q_H$  leading to average belief  $p_i = p_h$ , while she is less confident about the remaining "non-core" occupations to which she assigns prior  $Q_j = Q_L$  with average  $p_j = p_L$  such that  $p_L \leq p_H$ . Assume further that core occupations enter the list in the alternative interface for sure (i.e.,  $\underline{q}_H > \hat{q}$ ), which means that the alternative interface provides no information content for them. For non-core occupations we assume that there is information content (i.e.,  $\hat{q} \in (\underline{q}_L, \overline{q}_L)$ ) so that the alternative interface changes the prior positively if this occupation is featured on the alternative interface and negatively if it is not. For ease of notation, denote by  $e_H$  the search effort in the first period in core occupations, and by  $e_L$  the same for non-core occupations.

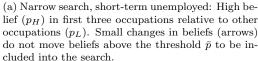
The following results are immediately implied by problem (4): given the search period, the number

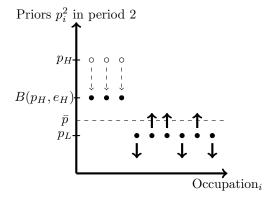
<sup>&</sup>lt;sup>61</sup>Infinitely lived agents would correspond to a specification with  $W_t = w/(1-\delta)$  and  $R_t(p) = R(p)$ .

<sup>62</sup> If the individual has already exerted search effort the updating is more complicated but obviously being on the list continues to be a positive signal. Consider a period t with prior  $p_i^t$ . The information that occupation i is on the list in the alternative interface can be viewed as changing the very first prior  $p_i^t$ , and this translates into the updated prior in period t by successively applying the updating formula (3), using the efforts that have been exerted in the interim.

Figure 8: Model Illustration: narrow search







(b) Narrow search, longer-term unemployed: Update in first three occupations leads to lower belief in these (dashed arrows for  $p_H$  in first three occupations). This brings threshold  $\bar{p}$  closer to the beliefs in other occupations  $(p_L)$ , so that some additional information moves some occupations above the threshold and broadens search (arrows for occupations 4-10).

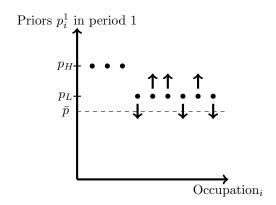
of core occupations and the current belief about them, there exists a level  $\bar{p}$  such that the individual puts zero search effort on the non-primary occupations iff  $p_i^t \leq \bar{p}$  for each non-core occupation i. Intuitively, when the average belief about being a high type in the non-core occupations is sufficiently close to zero, then it is more useful to search in the core occupations and search effort in non-core occupations is zero. The level of  $\bar{p}$  is increasing in the average belief about the core occupations (if core occupations are more attractive search is expanded there, which drives up the marginal cost of any further search in non-core occupations) and in the number of core occupations (again core occupations as a whole attract more search effort).

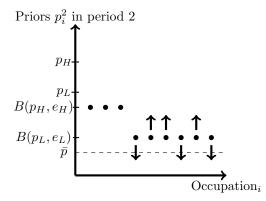
We depict our notion of an individual who is recently unemployed and narrow in Figure 8 (a). The person is narrow because her beliefs in her core occupations  $(p_H)$  are high enough that she does not want to search in the secondary occupations  $(\bar{p} > p_L)$ . This individual concentrates so much effort onto the primary occupations that marginal effort costs are large, and therefore she does not want to explore the less likely occupations. In fact, the distance in employment prospects is so large that small changes in the prior  $p_L$  induced by the alternative interface - indicated by the thick arrows in the figure - do not move them above the threshold  $\bar{p}$ .<sup>63</sup> So there would be no difference in search behavior with or without the alternative interface.

In panel (b) we depict our notion of the same individual after a period of unemployment. Her prior at the beginning of the second period is derived by updating from the previous one. After unsuccessful search in the core occupations it has fallen there, as indicated by the lower priors for

 $<sup>^{63}</sup>$ Whether the information of the alternative interface leads to small changes in the prior or large ones depends on the informativeness of the alternative interface. We consider here the case where the informativeness is low enough (e.g.,  $\overline{q}_L - \underline{q}_L < \epsilon$  for sufficiently small  $\epsilon$  so that the support of initial beliefs is not very dispersed, which bounds the possible change in priors due to additional information). We do not explore informativeness that yields to large changes in the prior here, as it would have the counterfactual implication that recently unemployed individuals already become broad due to the alternative interface.

Figure 9: Model Illustration: broad search





- (a) Broad search, short-term unemployed: Beliefs are rather similar and all beliefs are above the search cutoff  $\bar{p}$ . Small changes in beliefs (arrows) can move some occupations below this cutoff, making the person narrower.
- (b) Broad search, longer-term unemployed: Similar to part (a) but at a lower level of beliefs.

the first three occupations. Since she did not search in non-core occupations, her prior about them remains unchanged. So the beliefs are now closer together, and since they are the only source of heterogeneity the utility of applying to either of them are also closer. (If one were to additionally model penalties for failing to search broader over time, this would reinforce the effect since it would also reduce the perceived distance in utility between these occupations.) In a model with multiple rounds beliefs about the core occupations would eventually fall so low that individuals would start searching more broadly even without access to the alternative interface (as we see in our data for the control group, see section 4). Panel (b) depicts a shorter time frame where beliefs did not fall that much and  $p_L$  remains below the new  $\bar{p}$  so that the individual remains narrow. But since the distance is closer those with access to the alternative interface obtain information that moves some of their beliefs about non-core occupations above the threshold  $\bar{p}$ , which makes it attractive to search there and they become broader than their peers without such information. These increased opportunities materialize in a higher shadow-cost of remaining at the current level of search effort. Therefore, search effort weakly increases relative to the control group without alternative interface, and strictly so if the cost function is smooth. In turn it must lead to better job prospects (as the higher search effort needs to be compensated to make it individually optimal). So this rationalizes why longer-unemployed individuals in the treatment group become broader and their number of interviews increases, relative to the control group. It also implies a weak increase in search effort relative to the control group. At low unemployment durations to the contrary there is little effect.

Figures 9 (a) and (b) depict individuals who are already broad in the absence of an information intervention, since the threshold  $\bar{p} < p_L$ . This could be because an individual has rather equal priors already early in unemployment, as shown in panel (a). Alternatively it could be a person whose beliefs fell over the course of the unemployment spell to a more even level, as shown in (b) (possibly from an initially uneven profile such as in Figure 8 (a)). In both cases, the person already searches in

all occupations, but additional negative information (i.e., occupations that are not included in the list that is recommended in the alternative interface) might move the prior of those occupations so low that the person stops searching there and becomes narrow. Effects on search effort (in case it is flexible) and job prospects are ambiguous: search effort can now be concentrated more effectively on promising occupations which raises effort and job prospects; alternatively the negative information on some occupations can translate simply into reduced search effort which is privately beneficial but reduces job prospects. Depending on parameters, either can dominate.<sup>64</sup> This can rationalize why otherwise broad searchers become narrower in our treatment group, without significant effects on job prospects.

Thus, the model is able to replicate differential effects by breadth and unemployment duration. In this model, as in all models of classical decision theory, more information can only improve the expected utility for the individual. When total search effort is fixed (at  $\bar{e}$ ) then both the increase in breadth for narrow searchers as well as the decrease in breadth for broad searchers have to raise job prospects, as this is the only remaining interest of job seekers. But even if some groups like those that already search broadly would cut back on search effort in a way that reduces job prospect, this has to be overcompensated for them by savings on their search effort - i.e., it is privately efficient. Obviously socially, when taking into account unemployment benefit payments, this could lead to costs if some of the broad searchers have parameters that lead them to cut back on search effort in non-core occupations in a way such that their job prospects decline. We find fairly limited evidence on reduced effort and no significant negative effects on job interviews for any subgroup, but since we find negative point estimates of non-trivial magnitude more studies will be necessary to confirm both the empirical findings and our rationalization here.

## 7 Conclusion

We provided an information intervention in the labor market by redesigning the search interface for unemployed job seekers. Compared to a "standard" interface where job seekers themselves have to specify the occupations or keywords they want to look for, the "alternative" interface provides suggestions for occupations based on where other people find jobs and which occupations require similar skills. It provides this information in an easily accessible way by showing two lists and links to maps with market tightnesses, and provides all associated vacancies at the click of a button. While the initial costs of setting up such advice might be non-trivial, the intervention shares the concept of a "nudge" in the sense that the marginal cost of providing the intervention to more individuals is essentially costless

<sup>&</sup>lt;sup>64</sup>Which effect dominates depends importantly on the curvature of the total cost function  $c(\sum e_i)$  and of the returns to occupational effort  $f(e_i)$ . Consider the extreme case of extremely convex effort costs: costs are zero up to some threshold  $\sum e_i = \bar{e}$  and infinite thereafter. In this case clearly workers expend exactly  $\bar{e}$  units of total effort. Better information does not alter this but targets this effort better, so job prospects increase. Consider alternatively an economy with strictly curved f(.) and linear c(.) and let  $e_L$  denote the search effort in a non-core occupation for a broad individual. Now replace the returns function  $f(e_i)$  with a function  $\tilde{f}(e_i)$  that is identical for  $e_i < e_L$  but beyond that level marginal returns are zero  $(\tilde{f}(e_i)=f(e_i^*)$  for  $e_i > e_L)$ . Clearly the new function is more concave than the old one. Under this extreme return function, the effects of additional information are clear. For occupations with negative news the individual cuts his effort, without expanding it in occupations with positive news as the additional benefits are zero. Clearly search effort and job prospects fall as a consequence of additional information. While these extreme cases help to build intuition, less extreme cases display similar effects and we do not have a full characterization of when job prospects will increase.

and individuals are free to opt out and continue with the standard interface. There is currently strong interest in interventions of this kind.<sup>65</sup> While our intervention has a clear information component that falls within classical economic theory, a major aim of the intervention was to keep things simple for participants so little cognitive effort is required to learn on the alternative interface, which might be considered a nudge element.

We find that the alternative interface significantly increases the overall occupational breadth of job search in terms of listed vacancies. In particular, it makes initially narrow searchers consider a broader set of options, but decreases occupational breadth for initially broad searchers, even though overall the former effect dominates. Overall we find a positive effect on job interviews especially for those which otherwise search narrowly and have an above-median unemployment duration. The effect of unemployment duration is illustrated in our model where those who just got unemployed concentrate their efforts on those occupations where they have most hopes in and are not interested in investing time into new suggestions. If this does not lead to success, their confidence in these occupations declines and they become more open to new ideas.

Some words of caution in line with those in the introduction are warranted. While we find no statistically significant negative effects on job interviews for any subgroup, we cannot rule out that some of them get hurt through less interviews. Moreover, the size of the current study precludes any precise assessment of the effects on job finding, and currently we find no evidence of improvements on this dimension. We have limited information on the types of job found, which jeopardizes our ability to provide a convincing analysis on the duration and quality of jew jobs. At this stage, we can therefore not conclude that the increase in interviews is beneficial. Finally, additional larger-scale roll-out of such assistance would be required to document the full employment effects. The current study does not allow the assessment of equilibrium effects that would arise if everyone obtained information.

With these caveats in mind, our findings suggest that targeted job search assistance to those who otherwise search narrowly and with somewhat longer unemployment duration could be effective, in a cost-efficient way. The programming for the study cost £20,000 (\$30,000). If a large-scale website such as Universal Jobmatch would roll out such a scheme for millions of job seekers, it is obvious that the cost per participant is at the order of a few pence. So any meaningful positive employment effects would swamp the costs. As a first study on job search design on the web, it offers a new route how to improve market outcomes in decentralized environments and hopefully opens the door to more investigations in this area.

<sup>&</sup>lt;sup>65</sup>We thank the Behavioral Insights team of the UK cabinet office, the Department of Work and Pensions, and the researchers at the Welsh government for their interest in our work.

<sup>&</sup>lt;sup>66</sup>The study also devoted substantial resources (£80,000/\$120,000) to attracting participants, compensating participants, and for research assistants to carry out these activities (see also Footnote 4). An existing website would not need to incur such costs as they already have job seekers who search on their site. These numbers do not include the salaries of the authors. Even if the latter were included, the cost per participant at a large website would still only be some pence.

# References

- Allison, P. D. and Waterman, R. P. (2002). Fixed effects negative binomial regression models. *Sociological Methodology*, 32(1):247–265.
- Altmann, S., Falk, A., Jäger, S., and Zimmermann, F. (2015). Learning about job search: A field experiment with job seekers in Germany. IZA Working Paper No 9040.
- Ashenfelter, O., Ashmore, D., and Deschênes, O. (2005). Do unemployment insurance recipients actively seek work? Evidence from randomized trials in four U.S. states. *Journal of Econometrics*, 125(1-2):53 75.
- Behaghel, L., Crépon, B., and Gurgand, M. (2014). Private and public provision of counseling to job-seekers: Evidence from a large controlled experiment. *American Economic Journal: Applied Economics*, 6(4):142–174.
- Bennmarker, H., Grönqvist, E., and Öckert, B. (2013). Effects of contracting out employment services: Evidence from a randomized experiment. *Journal of Public Economics*, 98:68 84.
- Berger, M. C., Black, D., and Smith, J. (2000). Evaluating profiling as a means of allocating government services. In Lechner, M. and Pfeiffer, F., editors, *Econometric Evaluation of Labour Market Policies*, pages 59–84. Physica, Heidelberg.
- Bertrand, M., Duflo, E., and Mullainathan, S. (2004). How much should we trust differences-in-differences estimates? *The Quarterly Journal of Economics*, 119(1):249–275.
- Bertrand, M. and Mullainathan, S. (2004). Are Emily and Greg more employable than Lakisha and Jamal? A field experiment on labor market discrimination. *American Economic Review*, 4(94):991–1013.
- Blundell, R., Dias, M. C., Meghir, C., and van Reenen, J. (2004). Evaluating the employment impact of a mandatory job search program. *Journal of the European Economic Association*, 2(4):569–606.
- Card, D., Kluve, J., and Weber, A. (2009). Active labor market policy evaluations: A meta-analysis. IZA Discussion Paper Series, No. 4002. http://ftp.iza.org/dp4002.pdf.
- Card, D., Kluve, J., and Weber, A. (2010). Active labor market policy evaluations: A meta-analysis. The Economic Journal, 120:452–477.
- Card, D. and Mueller, A. (2016). A contribution to the empirics of reservation wages. *American Economic Review*, 1(8):142–179.
- Channel 4 (2014). Why is government website carrying fake jobs? http://www.channel4.com/news/why-is-government-website-carrying-fake-jobs. Posted 07-02-2014. Last accessed 28-09-2015.

- Computer Business Review (2014). Universal jobmatch here to stay despite fake job adverts. http://www.cbronline.com/news/cloud/aas/universal-jobmatch-here-to-stay-but-future-of-provider-monster-is-unclear-4204007. Written by Joe Curtis. Posted 26-03-2014; Last accessed 28-09-2015.
- Crépon, B., Duflo, E., Gurgand, M., Rathelot, R., and Zamora, P. (2013). Do labor market policies have displacement effects: Evidence from a clustered randomized experiment. *Quarterly Journal of Economics*, 128(2):531–580.
- Dinerstein, M., Einav, L., Levin, J., and Sundaresan, N. (2014). Consumer price search and platform design in internet commerce. National Bureau of Economic Research Working Paper 20415.
- Faberman, J. R. and Kudlyak, M. (2014). The intensity of job search and search duration. Working Paper 14-12 Federal Reserve Bank of Richmond.
- Gallagher, R., Gyani, A., Kirkman, E., Nguyen, S., Reinhard, J., and Sanders, M. (2015). Behavioural insights and the labour market: Evidence from a randomised controlled pilot study and a large stepped-wedge controlled trial. Mimeo.
- Gautier, P. A., Muller, P., van der Klaauw, B., Rosholm, M., and Svarer, M. (2015). Estimating equilibrium effects of job search assistance. *CESifo Working Paper Series*, No. 5476.
- Gibbons, R., Katz, L. F., Lemieux, T., and Parent, D. (2005). Comparative advantage, learning, and sectoral wage determination. *Journal of Labor Economics*, 23(4):681–724.
- Gibbons, R. and Waldman, M. (1999). A theory of wage and promotion dynamics inside firms. Quarterly Journal of Economics, 114(4):1321–1358.
- Groes, F., Kircher, P., and Manovskii, I. (2015). The u-shapes of occupational mobility. *Review of Economic Studies*, 82(2):659–692.
- Joyce, S. P. (2015). How to avoid 5 major types of online job scams. http://www.job-hunt.org/onlinejobsearchguide/job-search-scams.shtml. Last accessed 28-09-2015.
- Krug, G. and Stephan, G. (2013). Is contracting-out intensified placement services more effective than provision by the PES? Evidence from a randomized field experiment. IZA Discussion Paper No. 7403.
- Kudlyak, M., Lkhagvasuren, D., and Sysuyev, R. (2014). Systematic job search: New evidence from individual job application data. Mimeo.
- Kuhn, P. and Mansour, H. (2014). Is internet job search still ineffective? *Economic Journal*, 124(581):1213–1233.
- Lalive, R., van Ours, J. C., and Zweimüller, J. (2005). The effect of benefit sanctions on the duration of unemployment. *Journal of the European Economic Association*, 3(6):1386–1417.

- Launov, A. and Waelde, K. (2013). Thumbscrews for agencies or for individuals? How to reduce unemployment. RePEc Discussion Paper 1307.
- Manning, A. and Petrongolo, B. (2011). How local are labor markets? Evidence from a spatial job search model. CEPR Discussion Paper No. 8686.
- Marinescu, I. (2014). The general equilibrium impacts of unemployment insurance: Evidence from a large online job board. Mimeo.
- Marinescu, I. and Rathelot, R. (2014). Mismatch unemployment and the geography of job search. Mimeo.
- Marinescu, I. and Wolthoff, R. (2014). Opening the black box of the matching function: The power of words. Mimeo.
- Meyer, B. D. (1995). Lessons from U.S. unemployment insurance experiments. *Journal of Economic Literature*, 33(1):91–131.
- Micklewright, J. and Nagy, G. (2010). The effect of monitoring unemployment insurance recipients on unemployment duration: Evidence from a field experiment. *Labour Economics*, 17(1):180–187.
- Miller, R. A. (1984). Job matching and occupational choice. *Journal of Political Economy*, 92(6):1086–1120.
- Moscarini, G. (2001). Excess worker reallocation. Review of Economic Studies, 3(68):593-612.
- Neal, D. (1999). The complexity of job mobility among young men. *Journal of Labor Economics*, 17(2):237–261.
- ONS (2013). Internet access households and individuals, 2013. UK Office for National Statistics Statistical Bulletin.
- Papageorgiou, T. (2014). Learning your comparative advantages. Review of Economic Studies, 8(3):1263–1295.
- Patterson, C., Sahin, A., Topa, G., and Violante, G. (2016). Working hard in the wrong place: A mismatch-based explanation to the u.k. productivity puzzle. *European Economic Review*, (84):42–56.
- Pollard, E., Behling, F., Hillage, J., and Speckesser, S. (2012). Jobcentre plus employer satisfaction and experience survey 2012. UK Department of Work and Pensions Research Report No 806.
- Sahin, A., Song, J., Topa, G., and Violante, G. (2014). Mismatch unemployment. American Economic Review, (104):3529–3564.
- Svarer, M. (2011). The effect of sanctions on exit from unemployment: Evidence from Denmark. *Economica*, 78(312):751–778.

- The New York Times (2009a). Company rarely placed clients in jobs, former employees say. http://www.nytimes.com/2009/08/17/us/17careerbar.html. Written by Michael Luo. Posted 16-08-2009; Last accessed 28-09-2015.
- The New York Times (2009b). Online scammers prey on the jobless. http://www.nytimes.com/2009/08/06/technology/personaltech/06basics.html. Written by Riva Richmond. Posted 05-08-2009; Last accessed 28-09-2015.
- Van den Berg, G. J. and Van der Klaauw, B. (2006). Counseling and monitoring of unemployed workers: Theory and evidence from a controlled social experiment. *International Economic Review*, 47(3):895–936.
- Van der Klaauw, B. and Van Ours, J. C. (2013). Carrot and stick: How re-employment bonuses and benefit sanctions affect exit rates from welfare. *Journal of Applied Econometrics*, 28(2):275–296.
- Venn, D. (2012). Eligibility criteria for unemployment benefits: Quantitative indicators for OECD and EU countries. OECD Social, Employment and Migration Working Papers, No. 131, OECD Publishing.

# 8 Appendix - For Online Publication

## 8.1 Extended results

Figure 10: Histogram of the total attendance in weeks per individual

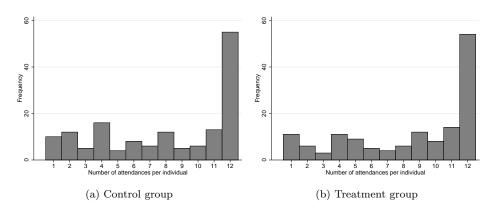


Figure 11: Histogram of the attendance in weeks 1-3 per individual

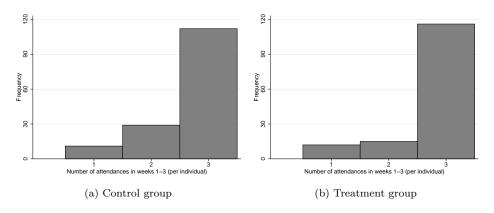


Figure 12: Histogram of the attendance in weeks 4-12 per individual

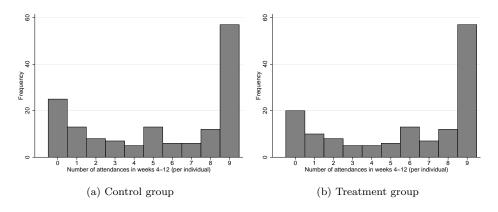


Figure 13: Applications and interviews of lab participants and online survey participants with 95% confidence interval

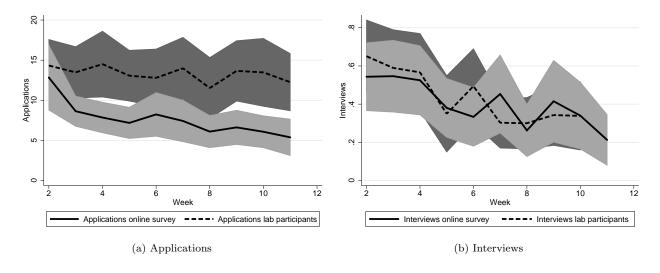
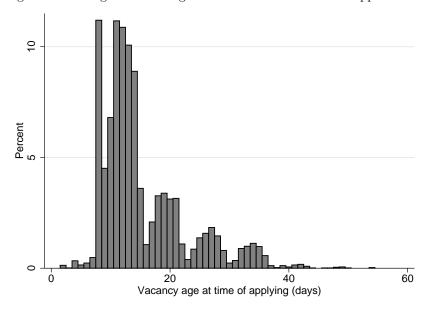


Figure 14: Histogram of the age of vacancies at the time of applications



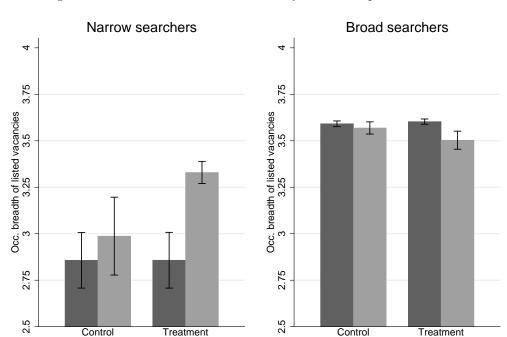


Figure 15: Mean values breadth of listed by initial occupational breadth

Figure 15 shows that both in the treatment and in the control group there is regression to the mean at least in point estimates: Narrow searchers before the intervention become broader after the intervention both in the control and in the treatment group, and broad searchers before the intervention become narrower after the intervention both in the control and treatment group. But the magnitude is larger for the treatment group.

Before

After

After

Before

7D 11 10 T 1	1 4		/ 1 , 1	•	1.1 1.10
Table Ib. Job	search activity	z over time	ioniv control	group survivors	s until week 10)
10010 10. 000	Doditon accivity	, over unite	(OIII) COIICIOI	Stoup but ition	differ week io

	(1)	(2)	(3)	(4)	(5)
	Hours search per week	Breadth of listed vac.	Number of listed vac.	Breadth of applications	Number of applications
Time trend	0.057	0.014***	7.86*	-0.0046	-0.12*
	(0.066)	(0.0050)	(4.28)	(0.015)	(0.064)
Individual FE	yes	yes	yes	yes	yes
Mean of dep. var.	12.1	3.29	542.4	3.07	3.86
Weeks	1-12	1-12	1-12	1-11	1-11
N	833	918	920	418	849

All regressions contain only control group individuals. "Time trend" is a linear weekly trend. Standard errors clustered by individual in parentheses. Sample contains only control group individuals that attended at least one session in week 10, 11 or 12. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01

Table 17: Relation between breadth/unemployment duration and individual characteristics

	(1)	(2)	(3)	(4)
	Breadth	Breadth	Unemployment	Unemployment
	dummy	continuous	duration dummy	duration continuous
Age	-0.04**	0.01	0.01	3.10
	(0.02)	(0.02)	(0.02)	(3.21)
$Age^2$	0.03	-0.04	-0.02	-4.35
	(0.02)	(0.03)	(0.02)	(4.12)
Gender	0.07	0.08	0.04	0.91
	(0.06)	(0.07)	(0.06)	(10.78)
Weeks unemployed	0.00	-0.00		
	(0.00)	(0.00)		
Weeks unemployed <sup>2</sup>	-0.00	0.00		
	(0.00)	(0.00)		
Financial problems	0.04	0.04	0.10	-14.67
	(0.06)	(0.07)	(0.06)	(10.60)
Married/cohabiting	-0.01	-0.05	0.06	-16.71
	(0.07)	(0.09)	(0.07)	(12.61)
Children	-0.08	-0.05	-0.13*	22.56*
	(0.07)	(0.09)	(0.08)	(13.45)
High educated	-0.08	-0.05	-0.01	23.14**
	(0.06)	(0.08)	(0.06)	(11.17)
White	0.02	0.22**	0.16**	4.09
	(0.07)	(0.09)	(0.08)	(13.41)
Constant	1.44***	3.36***	0.15	-19.97
	(0.32)	(0.41)	(0.34)	(59.68)
Observations	295	295	295	295
$R^2$	0.178	0.213	0.044	0.044

Standard errors in parentheses. The dependent variable is a dummy for searching broad in weeks 1-3 in column (1), a continuous breadth measure in column (2), a dummy for having unemployment duration above the median in column (3) and the continuous unemployment duration (in weeks) in column (4). \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Table 18: Random effects vs Fixed Effects: Hausman tests

	Listed		Appl	ications			Interviews	
	(1)	(2)	(3)	(4)	(5)		(7)	(8)
	$\mathbf{Breadth}$	Breadth	In lab	Outside lab	Total		Outside lab	Total
Treatment (fe model)	0.13**	0.062	0.066	-0.031	0.011		0.25	0.29
	(0.062)	(0.21)	(0.16)	(0.095)	(0.094)	(0.95)	(0.38)	(0.35)
Treatment (re model)	0.15***	0.12	0.0081	-0.077	-0.033		0.18	0.20
	(0.00)	(0.15)	(0.14)	(0.09)	(0.09)		(0.22)	(0.21)
P-val Hausman test <sup>a</sup>	0.58	0.69	0.28	0.18	0.19		0.83	0.74
Model	Linear	Linear	Neg. Bin	Neg. Bin	Neg. Bin		Poisson	Poisson
Included weeks	1-12	1-11	1-11	1-11	1-11		1-10	1-10
N	540	305	410	428	424		306	314

concern applications and columns (6)-(8) concern interviews. We report  $[\exp(\operatorname{coefficient}) - 1]$  in columns (3)-(8), which is the percentage effect. Estimates from the random effects models differ from other tables because no other variables are included here (individual characteristics and time-slot fixed effects). Standard errors in parentheses. A time period dummy is included in all regressions (but not reported). <sup>a</sup> P-value of a chi-squared test of equal estimates for the treatment effect. Column (1) concerns listed vacancies, columns (2)-(5)

Table 19: Effect of intervention on the number of applications - alternative specifications  ${\bf r}$ 

	Num	ber of Applica	tions
	(1)	(2)	(3)
	In lab	Outside lab	Both
Treatment	-0.01 (0.17)	-0.07 (0.11)	-0.02 (0.11)
Treatment			
X occupationally broad	-0.08	-0.05	-0.02
A occupationally broad	(0.23)	(0.18)	(0.19)
X occupationally narrow	0.05	-0.09	-0.04
	(0.25)	(0.12)	(0.13)
Model	Poisson	Poisson	Poisson
	RE	RE	RE
Observation weeks	1-11	1-11	1-11
Observations	541	490	487

Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. We report [exp(coefficient) - 1], which is the percentage effect. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01

Table 20: Effect of intervention on listed vacancies - extensions (split by geographical breadth)

		dth of	Number of
		ings	listings
	(1)	(2)	(3)
	Occupational	Geographical	Lab
Treatment	0.13*** (0.06)	-0.01 (0.02)	-34.99 (52.09)
	(* )	( )	()
Treatment			
X geographically broad	0.22**	-0.03	30.68
	(0.09)	(0.04)	(66.28)
X geographically narrow	0.02	0.03	-111.03
	(0.06)	(0.03)	(81.42)
Treatment X occ. broad and geo. broad	-0.07	0.00	126.87
	(0.05)	(0.06)	(168.51)
X occ. broad and geo. narrow	-0.09*	0.04	-123.80
	(0.05)	(0.04)	(98.25)
X occ. narrow and geo. broad	0.40***	-0.05	-11.08
Ţ.	(0.13)	(0.05)	(56.80)
X occ. narrow and geo. narrow	0.21*	0.01	-83.32
	(0.11)	(0.03)	(141.01)
Model	Linear	Linear	Linear
Observation weeks	1-12	1-12	1-12
N	540	541	541

Each column represents three separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Table 21: Effect of intervention on applications - extensions (split by geographical breadth)

		dth of cations		Number of applications	
	(1) Occupational	(2) Geographical	(3) Lab	(4) Outside lab	(5) Total
Treatment	0.03 (0.20)	-0.06* (0.03)	0.09 (0.16)	-0.03 (0.09)	0.01 (0.09)
Treatment					
X geographically broad	-0.03 $(0.26)$	-0.10** (0.04)	$0.06 \\ (0.21)$	-0.07 $(0.12)$	$0.00 \\ (0.12)$
X geographically narrow	$0.10 \\ (0.25)$	-0.00 (0.04)	0.12 $(0.24)$	0.01 $(0.14)$	$0.03 \\ (0.13)$
Treatment					
X occ. broad and geo. broad	-0.65** $(0.30)$	-0.10 (0.08)	$0.08 \\ (0.33)$	-0.17 $(0.16)$	-0.08 $(0.17)$
X occ. broad and geo. narrow	-0.17 (0.28)	$0.03 \\ (0.06)$	-0.16 (0.23)	$0.05 \\ (0.18)$	-0.03 (0.16)
X occ. narrow and geo. broad	0.41 $(0.36)$	-0.11** (0.05)	$0.05 \\ (0.27)$	0.01 $(0.16)$	$0.06 \\ (0.16)$
X occ. narrow and geo. narrow	0.65* (0.36)	-0.04 (0.04)	0.71 $(0.58)$	-0.05 $(0.20)$	$0.12 \\ (0.22)$
Model	Linear	Linear	Neg. Bin.	Neg. Bin.	Neg. Bin.
Observation weeks $N$	1-11 305	1-11 363	1-11 541	1-11 490	1-11 487

Each column represents three separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (3)-(5) are negative binomial model regressions where we report [exp(coefficient) – 1], which is the percentage effect. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01

Table 22: Effect of intervention on interviews - extensions (split by geographical breadth)

		Number of interviews	
	(1) Lab	(2) Survey	(3) Total
Treatment	0.61 (0.79)	0.40* (0.27)	0.44* (0.28)
Treatment			
X geographically broad	1.90** (1.47)	0.65** $(0.40)$	0.85*** $(0.40)$
X geographically narrow	0.19 (0.81)	0.14 $(0.33)$	$0.12 \\ (0.36)$
Treatment			
X occ. broad and geo. broad	0.99 $(2.00)$	$0.41 \\ (0.56)$	$0.42 \\ (0.50)$
X occ. broad and geo. narrow	-0.75* (0.20)	-0.27 $(0.24)$	-0.37 $(0.20)$
X occ. narrow and geo. broad	1.99* (1.67)	0.83** (0.53)	1.14*** (0.58)
X occ. narrow and geo. narrow	$0.65 \\ (1.36)$	$0.85 \\ (0.85)$	0.87 $(0.93)$
	Poisson 1-10 540	Poisson 1-10 466	Poisson 1-10 464

Each column represents three separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (1)-(3) are Poisson regression models where we report [exp(coefficient) – 1], which is the percentage effect. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Table 23: Effect of intervention - all coefficients

	(1)	(0)	(9)
	(1) Number of	(2) Total number of	(3) Total number of
	listed	applications	interviews
Treatment	-34.99	-0.02	0.44*
Heatment	(52.09)	(0.11)	(0.28)
	(52.09)	(0.11)	(0.28)
Age	4.71	0.04	-0.01
	(14.04)	(0.04)	(0.04)
$Age^2$	-12.52	-0.07	-0.01
0.	(18.49)	(0.05)	(0.06)
Gender	72.41	-0.16	0.29
Gender			
	(47.31)	(0.11)	(0.22)
Weeks unemployed	-0.71	0.00	-0.01*
	(0.66)	(0.00)	(0.00)
Weeks unemployed <sup>2</sup>	0.01	0.00	0.00
r	(0.09)	(0.00)	(0.00)
Financial problem	101.74*	0.12	0.26
F	(52.81)	(0.14)	(0.19)
Q 1	70.41	0.00	0.00
Couple	-73.41	-0.20	0.38
	(48.40)	(0.11)	(0.31)
Children	-84.87	0.12	0.05
	(56.63)	(0.17)	(0.19)
High educated	-24.84	-0.11	0.23
8	(59.23)	(0.12)	(0.23)
XX71.:4 -	FA C4		0.09
White	54.64	-0.20	-0.03
	(69.10)	(0.14)	(0.18)
Constant	584.41**	8.15***	-0.13
	(276.25)	(7.20)	(0.72)
Model	Linear	Poisson	Poisson
Observation weeks	1-12	1-11	1-10
N	541	487	464

Each column represents one regression. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup) and individual random effects. Columns (2) and (3) are Poisson regression models where we report  $[\exp(\text{coefficient})-1]$ , which is the percentage effect. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\*\* p < 0.05, \*\*\*\* p < 0.01

Table 24: Robustness: intervention effect using individual fixed effects

	Listed		Appli	Applications			Interviews	
	$(1) \\ \mathbf{D}_{\text{magg}} \mathcal{A}_{\text{th}}$	$(2) \\ D_{-0.00} \frac{1}{441}$	(3) $1 - 1$	(4)	(5)	(9)	(7)	(8)
	breadtn	breadth	In Iab	Outside lab	Botn	III Iab	Outside lab	Botn
Treatment	0.13**	0.062	990.0	-0.031	0.011	0.57	0.25	0.29
	(0.062)	(0.21)	(0.16)	(0.095)	(0.094)	(0.73)	(0.25)	(0.26)
X occupationally broad	**090.0-	-0.098	-0.070	-0.019	-0.026	-0.38	0.25	0.050
•	(0.029)	(0.24)	(0.20)	(0.15)	(0.14)	(0.41)	(0.43)	(0.32)
X occupationally narrow	0.32***	0.23	0.19	-0.056	0.029	1.08	0.22	0.47
	(0.11)	(0.35)	(0.26)	(0.12)	(0.12)	(1.21)	(0.30)	(0.41)
Model	Linear	Linear	Neg. bin.	Neg. bin.	Neg. bin.	Poisson	Poisson	Poisson
	(Ind. FE)	(Ind. FE)	(Ind. FE)	(Ind. FE)	(Ind. FE)	(Ind. FE)	(Ind. FE)	(Ind. FE)
Observation weeks	1-12	1-11	1-11	1-11	1-11	1-10	1-10	1-10
N	540	305	410	428	424	134	306	314

Each column represents two separate regressions. All regressions include individual fixed effects and period fixed effects (separately for each subgroup). Column (1) concerns listed vacancies, columns (2)-(5) concern applications and columns (6)-(8) concern interviews. Columns (1)-(2) are linear regressions, columns (3)-(5) are negative binomial regressions, and columns (6)-(8) are Poisson regression models. In columns (3)-(8) we report [exp(coefficient) - 1], which is the percentage effect. Standard errors clustered by individual in parentheses (except for the neg. bin. model). \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Table 25: Robustness: intervention effect using weekly data

	Listed (1) Breadth	(2) Breadth	App (3) In lab	Applications (4) b Outside lab	(5) Both	(6) In lab	Interviews (7) Outside lab	(8) Both
Treatment	0.11*	0.0038 (0.15)	0.077	-0.056	-0.028	0.56	0.24 (0.23)	0.29 (0.25)
Treatment X occupationally broad	-0.055	-0.22 (0.17)	-0.065 (0.12)	-0.12 (0.085)	-0.12 (0.075)	-0.44 (0.32)	-0.11 (0.26)	-0.18 (0.23)
X occupationally narrow	$0.27^{***}$ $(0.086)$	0.26 $(0.22)$	0.21 $(0.16)$	-0.0020 $(0.089)$	0.051 $(0.083)$	1.32* (1.18)	$0.61^*$ $(0.42)$	0.73* $(0.48)$
Model Observation weeks N	Linear 1-12 2392	Linear 1-12 934	Neg. Bin. 1-11 2251	Neg. Bin. 1-11 2016	Neg. Bin. 1-11 1984	Poisson 1-10 2098	Poisson 1-10 1776	Poisson 1-10 1744

(2)-(5) concern applications and columns (6)-(8) concern interviews. Columns (1)-(2) are linear regressions, columns (3)-(5) are negative binomial regressions, and columns (6)-(8) are Poisson regression models. In columns (3)-(8) we report [exp(coefficient) - 1], which is the percentage effect. Standard errors clustered by individual in parentheses (except for the neg. bin. model). \* p < 0.10, \*\*\* p < 0.05, \*\*\* p < 0.01for each subgroup), individual random effects and individual characteristics. Column (1) concerns listed vacancies, columns Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately

Table 26: Robustness: intervention effect using linear models

	(1) In lab	Applications (2) Outside lab	(3) Both	(4) In lab	Interviews (5) Outside lab	(6) Both
Treatment	0.17	-0.21 (0.87)	0.024	0.039	0.12 (0.084)	0.17
Treatment X occupationally broad	0.093	-0.028 (1.33)	0.045	0.00070 (0.041)	0.057	0.023 (0.13)
X occupationally narrow	0.25 $(0.63)$	-0.38 (1.03)	-0.020 (1.44)	0.074 (0.068)	0.18* $(0.10)$	0.30** $(0.15)$
Model Observation weeks N	Linear 1-11 541	Linear 1-11 490	Linear 1-11 487	Linear 1-10 540	Linear 1-10 466	Linear 1-10 464

Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately for each subgroup), individual random effects and individual characteristics. Columns (1)-(3) concern applications and columns (4)-(6) concern interviews. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Table 27: Robustness: intervention effect excluding each individual's last one or two weeks

		App	Applications			Interviews	
	$\begin{array}{c} (1) \\ \text{Breadth} \end{array}$	$ \begin{array}{c} (2) \\ \text{In lab} \end{array} $	(3) Outside lab	(4) Both	$\begin{array}{c} (5) \\ \text{In lab} \end{array}$	(6) Outside lab	(7) Both
Treatment	0.0037		-0.021		0.38	0.37*	0.42*
	(0.20)	(0.15)	(0.091)	(0.091)	(0.66)	(0.26)	(0.28)
Treatment	1	1 0	6	0	1	1 0 0	
X occupationally broad	-0.47**	-0.053	-0.028	-0.023	-0.57	-0.035	-0.098
	(0.22)	(0.18)	(0.13)	(0.12)	(0.29)	(0.28)	(0.24)
X occupationally narrow	0.47	0.26	-0.014	0.082	1.18	0.80**	0.96**
	(0.29)	(0.25)	(0.13)	(0.13)	(1.23)	(0.44)	(0.52)
Model	Linear	Neg. bin.	Neg. bin.	Neg. bin.	Poisson	Poisson	Poisson
Observation weeks	Varying	Varying	Varying	Varying	Varying	Varying	Varying
Z	302	499	487	484	473	464	462

Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed  $[\exp(\text{coefficient})-1]$ , which is the percentage effect. Standard errors clustered by individual in parentheses (except for the neg. bin. model). \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01effects (separately for each subgroup), individual random effects and individual characteristics. Columns (1)-(4) concern applications and columns (5)-(7) concern interviews. Column (1) is a linear regression, columns (2)-(4) are negative binomial regressions, and columns (5)-(7) are Poisson regression models. In columns (2)-(7) we report

Table 28: Robustness: intervention effect using a different breadth measure based on transitions observed in the BHPS

	Listed (1)	(2)	(3)	Applications (4)	(5)	(9)	Interviews (7)	<u> </u>
	$\frac{1}{2}$	Breadth	$\ln \log$	Outside lab	$\operatorname{Both}$	$\ln \log $	Outside lab	$\operatorname{Both}$
Treatment	0.0053** $(0.0027)$	0.0028 (0.0097)						
Treatment X occupationally broad	-0.0012 (0.0024)	-0.0098 (0.0094)	-0.042 (0.21)	-0.049 (0.13)	-0.070 (0.13)	-0.84 (0.74)	0.16 (0.28)	0.031
X occupationally narrow	0.0076 $(0.0053)$	0.015 $(0.015)$	0.20 $(0.20)$	-0.018 (0.13)	0.091 $(0.12)$	1.07** $(0.50)$	$0.50** \\ (0.25)$	$0.64^{**}$ $(0.25)$
St.Dev. dep. var. Model Observation weeks N	.029 Linear 1-12 540	.047 Linear 1-11 305	Neg. bin. 1-11 541	Neg. bin. 1-11 490	Neg. bin. 1-11 487	Poisson 1-10 540	Poisson 1-10 466	Poisson 1-10 464

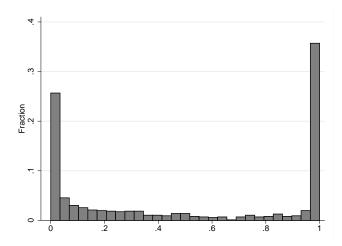
 $\exp(\operatorname{coefficient}) - 1]$ , which is the percentage effect. Standard errors clustered by individual in parentheses (except for the neg. bin. model). The outcome measure in column (1) is breadth of listed vacancies based on the BHPS transitions. The outcome measure in column (2) is breadth of applications based on the BHPS transitions. The groups used in this table for each subgroup), individual random effects and individual characteristics. Column (1) concerns listed vacancies, columns Note that the blank spaces in this table are specifications that are unaffected by the breadth measure and would thus produce Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately (5) are negative binomial regressions, and columns (6)-(8) are Poisson regression models. In columns (3)-(8) we report "occupationally broad" and "occupationally narrow") are defined using the breadth measure based on BHPS transitions. (2)-(5) concern applications and columns (6)-(8) concern interviews. Columns (1)-(2) are linear regressions, columns (3)the same result as our baseline specification. \* p < 0.10, \*\* p < 0.05, \*\*\* p < 0.01

Table 29: Robustness: intervention effect using alternative interface usage instrumented with treatment assignment

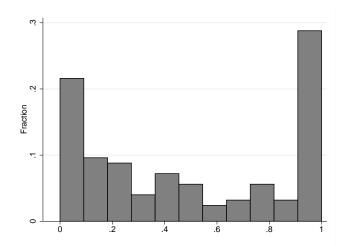
	ᆼ		Appli	Applications			Interviews	
	$\begin{array}{c} (1) \\ \text{Breadth} \end{array}$	(2) Breadth	$ \begin{array}{c} (3) \\ \text{In lab} \end{array} $	(4) Outside lab	$\begin{array}{c} (5) \\ \text{Both} \end{array}$	$\begin{array}{c} (6) \\ \text{In lab} \end{array}$	(7) Outside lab	$\begin{array}{c} (8) \\ \text{Both} \end{array}$
Alt. interface use		0.036		-0.42		0.072	0.24	0.32
	(0.12)	(0.38)	(0.85)	(1.64)	(2.29)	(0.080)	(0.16)	(0.20)
Alt. interface use								
X occupationally broad	-0.18**	-1.10**	0.18	-0.21	-0.025	0.0014	0.13	0.053
	(0.090)	(0.55)	(1.38)	(3.12)	(4.36)	(0.099)	(0.29)	(0.32)
X occupationally narrow	0.54***	$0.84^{*}$	0.41	-0.58	-0.016	0.12	0.30*	0.49**
	(0.16)	(0.46)	(0.99)	(1.66)	(2.31)	(0.11)	(0.16)	(0.24)
Model	Linear IV	Linear IV	Linear IV	Linear IV	Linear IV	Linear IV	Linear IV	Linear IV
Observation weeks	1-12	1-11	1-11	1-11	1-11	1-10	1-10	1-10
Z	540	305	541	490	487	540	466	464

for each subgroup), individual random effects and individual characteristics. Column (1) concerns listed vacancies, columns (2)-(5) concern applications and columns (6)-(8) concern interviews. All columns are linear IV regressions in which the use of the alternative interface is instrumented for by treatment assignment. Standard errors clustered by individual in parentheses. \* p < 0.10, \*\* p < 0.05, Each column represents two separate regressions. All regressions include time-slot fixed effects, period fixed effects (separately \*\*\* p < 0.01

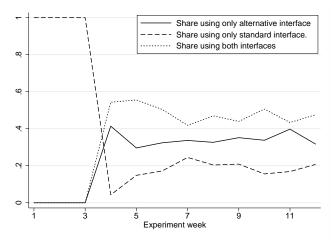
Figure 16: Usage of the alternative interface (contains only the treatment group participants in weeks 4-12)



(a) Share of listed vacancies from alternative interface (user by week observations)



(b) Share of listed vacancies from alternative interface (one observation per user)  $\,$ 



(c) Alternative interface usage over time

Table 30: Correlation between different broadness measures for listed vacancies

	M listed	G4 listed	G3 listed	G2 listed	G1 listed
M listed	1				
G4 listed	.97	1			
G3 listed	.99	.97	1		
G2 listed	.98	.94	.97	1	
G1 listed	.96	.91	.94	.98	1

M is the broadness measure used in the empirical analysis, Gx is the Gini-Simpson measure applied to the x-digit SOC code. Correlation are computed based on individual observations, collapsed into two periods as is done in the empirical analysis.

Table 31: Correlation between different broadness measures for applications

	M applied	G4 applied	G3 applied	G2 applied	G1 applied
M applied	1				
G4 applied	.73	1			
G3 applied	.80	.93	1		
G2 applied	.83	.87	.95	1	
G1 applied	.79	.79	.87	.91	1

M is the broadness measure used in the empirical analysis, Gx is the Gini-Simpson measure applied to the x-digit SOC code. Correlation are computed based on individual observations, collapsed into two periods as is done in the empirical analysis.

Table 32: Characteristics of the treatment and control group (based on the first week initial survey), for different groups of survivors

	Control group		Treatment group			T-test (p-value)			
	survivors in:		in:	survivors in:		for equality in:			
	week 1	week $4$	week $12$	week 1	week $4$	week $12$	week 1	week $4$	week $12$
Demographics:									
female (%)	42	43	34	43	41	41	0.83	0.87	0.43
age	36	36	37	36	37	40	0.85	0.32	0.15
high educ $a$ (%)	44	43	48	41	41	43	0.63	0.77	0.55
survey qualification level	4.2	4.3	4.4	4.4	4.4	4.6	0.36	0.44	0.48
white $(\%)$	80	81	81	80	81	77	0.97	0.97	0.59
number of children	0.66	0.68	0.77	0.38	0.39	0.46	0.01	0.03	0.08
couple (%)	25	24	26	21	20	19	0.41	0.35	0.3
any children (%)	31	30	33	24	24	29	0.17	0.33	0.62
Job search history:									
vacancies applied for	75	69	92	53	43	34	0.18	0.09	0.01
interviews attended	0.43	0.39	0.33	0.54	0.49	0.51	0.28	0.25	0.15
jobs offered	0.37	0.42	0.44	0.48	0.43	0.47	0.43	0.96	0.9
at least one offer $(\%)$	20	21	23	20	18	19	0.91	0.50	0.52
days unempl. (mean)	290	291	305	228	182	190	0.39	0.13	0.14
days unempl. (median)	81	84	87	77	78	80			
less than 183 days	0.75	0.76	0.73	0.78	0.79	0.76	0.60	0.54	0.64
less than 366 days	0.84	0.85	0.82	0.87	0.89	0.86	0.54	0.41	0.51
jobseekers allow. $(\pounds)$	49	49	46	56	59	65	0.46	0.38	0.27
housing benefits $(\pounds)$	65	67	81	62	62	74	0.90	0.82	0.81
other benefits (£)	9.7	11	1.6	18	19	26	0.41	0.5	0.21
Weekly search weeks 1-3:									
listed	493	513	477	493	464	415	1	0.32	0.33
viewed	25	25	26	26	25	24	0.57	0.81	0.36
saved	10	10	12	11	10	9.7	0.54	0.86	0.32
applied	3.3	3.8	4.6	2.5	2.7	2.6	0.14	0.13	0.035
interview	0.098	0.11	0.11	0.083	0.096	0.08	0.66	0.65	0.55
applications other	9.3	9.2	11	7.4	7.5	6.7	0.13	0.17	0.027
interviews other	0.54	0.51	0.32	0.47	0.47	0.52	0.48	0.69	0.11
broadness listed <sup><math>b</math></sup>	3.2	3.2	3.2	3.3	3.2	3.1	0.50	0.57	0.39
broadness applied <sup><math>b</math></sup>	3	3	3	3.2	3.2	3.1	0.34	0.40	0.45
hours spend $^c$	11	11	11	12	12	12	0.15	0.34	0.61
concern health $(1-10)$	1.5	1.3	1.8	1.7	1.8	2.1	0.48	0.12	0.47
conc. financial (1-10)	7.2	7.3	7.1	7	6.9	7.1	0.47	0.29	0.93
conc. competition $(1-10)$	7.4	7.5	7.3	7.2	7.2	7.3	0.43	0.37	0.97
met caseworker (%)	32	32	30	28	28	27	0.48	0.45	0.58
Observations	152	127	73	143	123	79			

- 8.2 Experimental instructions and supplemental documents
- 8.2.1 Consent form

# Consent Form for Participants: "How Do Unemployed Search for Jobs?"

Thank you for your willingness to consider taking part in this study. Please read the information below carefully. By signing the consent form below, you indicate that you have understood the purpose of the study, you have been made aware of your rights and you have agreed with the terms and conditions of the study.

### **Purpose of the study**

The study is undertaken to understand better how people search for jobs. The study aims to observe how people search for real jobs. The goal is to document parts of the job search process.

#### How will this work?

The study will be conducted over a period of 12 weeks and you are asked to take part to one weekly session of 2 hours taking place at a pre-agreed time slot. You will be asked to come to our computer facilities, located at the School of Economics, 31 Buccleuch Place, EH8 9JT Edinburgh. There will be a maximum of 30 participants present at the same time in the facilities. The research team aims to provide an environment that is conducive to the job search of participants and hopes that participants will attend for the duration of the study or up to the point you find a job.

You will be able to spend most time each week to search for job vacancies. These job vacancies are obtained from two sources:

- Our main data source is the vacancy database of Universal Jobmatch and coincides with those used at Jobcentre Plus.
- Additionally, our database includes a small number of vacancies (no more than 2 per 100 vacancies) that is added for research purposes. These "research vacancies" are included to understand better which types of vacancies people are interested in even if these are not currently offered. If you express interest in such a vacancy, you will be immediately informed that this is a research vacancy before you start any application.

We will track the pages you consult, what vacancies you are looking at and consider applying to. This information will never be linked to any of your personal information such as your name and address, which will be stored separately. Your personal information will never be given out to anyone and will be accessible only to selected members of the research team.

You will also be asked some survey questions about your job search in the past week and your wellbeing. In the initial week, we will also ask a number of questions about your background and unemployment history. Six month after the end of your participation we will send you a survey about your labour market experience and your well-being.

Note that we ask all participants to stay for the full 2 hours in the laboratory. But if you do not want to search for jobs anymore, we provide some alternative ways in which you can use the computer and internet facilities.

If you are unable to participate to a session, please inform us as soon as possible (under <a href="mailto:jobsearch@ed.ac.uk">jobsearch@ed.ac.uk</a> or 0131 6508324). The research team will attempt to provide additional slots in case a participant misses his time slots for justified reasons (e.g., job interviews, illness).

## **Important notes**

- Participation to this study is entirely voluntary. You should by no means feel complied to participate. You can also withdraw from the study at any time if you wish to do so.
- Since the study is to gain understanding in how people search for jobs, the research team holds no particular view on how individuals should search for jobs. Thus, you should search for jobs in the same way as you would normally do.
- The study is conducted by the research team, and no personalized information is shared with any other organization. Therefore, no information will be shared with Job Centre Plus or the Department of Work and Pensions. If you would like to obtain a record of your search activities, e.g. to use for discussion with your case worker, you can obtain a printed record to take along at the end of each session.
- You should be aware that **participation in this study does not provide any additional benefits**, and in particular it does not provide particular help in job search. In particular, you **should follow your usual job search strategy**, such as for example looking at other job vacancies beyond those provided in our database, searching from home via the internet, and contacting friends and acquaintances. You should not take the time within the study as an indication of the appropriate time to spend on searching for a job.
- All the data collected during your time in our computer facility is anonymous. Your search activities will not be matched to your identity in any way. You will be attributed a randomly generated number at the first session and all data records will be matched to that number.
- We will ask you for a telephone number that we can use to contact you. We will only contact you to remind you of the time slot you have been allocated to and to inform you of any changes in schedule. Of course the telephone number will not be matched to the data we collect in the laboratory.
- You have the right to withdraw entirely from the study (i.e. ask us to delete all the data records associated with you) at any point during the study.
- The impersonal data collected will be used for research purposes (and ONLY for research purposes). Personal data will never be given out, and will be eliminated after the study is completed. The results of the study will be published in peer-reviewed scientific journals.

#### **Compensation**

You will be compensated for your efforts of coming to and participating in each session in our computer facility with a compensation of £12.50 per visit (2 hours) to the laboratory. Additionally, if you participated in all four sessions in the first four weeks you are entitled to a £50 clothing voucher for job market attire as compensation for arranging the visit every week. The same holds for weeks 5 to 8 and for weeks 9 to 12.

## **Eligibility**

Participants have to be at least 18 years of age, permanent residents of the UK and living in Edinburgh (or within a distance of 5 miles from Edinburgh). You should be seeking for a job for a period of 4 weeks or less at the start date of the study.

#### Signature

If any of the material above is unclear to you, or if you have any doubts and would like clarification, please consult a member of the research team before proceeding.

If you are willing to take part in this study, please sign the consent form below:

I certify that I voluntarily participate in this research study. I certify that I read and understood the information above, and am eligible for taking part in this study.

(please print your name)
(please sign)
(place and time of signature)

# 8.2.2 Lab instructions

# **UNIVERSITY JOB SEARCH STUDY: INSTRUCTIONS**

Please do not start using the computer before we indicate you to do so.

We will read these instructions aloud at the start of the first session.

#### INTRODUCTION

Welcome and thank you for coming here today. Before we explain how each session will work, we would like to raise your attention to the following:

- **Health and Safety**: There will always be one person from the research team in the computer room. There is one toilet on this floor that you are free to use. In case of fire, please do follow the signs for fire exit. The main exit is through the staircase you have used to come up here.
- No smoking: Smoking is not allowed in this building.
- **Silence**: Since there are many of you in the room, we would appreciate if you would keep silent, so that everyone can concentrate on their computer activity.
- Mobile phones: Mobile phones must either be switched off or be on "silent" during each
  session. We would appreciate if you leave it on only if you are expecting an important phone
  call. And if you do receive a phone call, please leave the room and take the call outside (in
  the staircase).
- Food and drinks are not allowed in this room.
- Questions: Please do not hesitate to call us if you have a question.

## WHAT IS THE STUDY ABOUT?

The goal of the study is to understand how people search for jobs. Importantly, we hold no preconceptions regarding how people should search for jobs. We designed this study to find out what people usually do and what strategies are most successful. At the moment, we do not know what these are. We are interested in finding out common patterns in search strategies, and kindly ask you to search exactly in the same way as you normally would.

# WHAT WILL HAPPEN IN EACH SESSION

When you come in, you will be assigned to a computer station. We may provide specific instructions at the beginning of the session, so please do wait for us to indicate the start of the session. We will now describe how each session will proceed.

#### 1. LOGIN

You have received a unique login number and password that you can use to login on the website here and also from home. You will be able to access your records using this login information.

#### 2. SURVEY

**Each weekly session** will start with a **short survey**, asking questions about your past week and job search. After filling the survey, you will be re-directed towards the job search engine's main page.

**For the first session**, we will ask you to fill in a longer survey asking you questions about your background, qualifications and job search experience so far. You will only need to answer this initial survey once, in this session. It should take 20 minutes to fill in this initial survey.

#### 3. THE JOB SEARCH ENGINE

We have designed our own job search engine. It allows you to search through all UK vacancies that are also recorded in Universal Jobmatch.

We ask you to search for jobs using this search engine only for a minimum of 30 minutes.

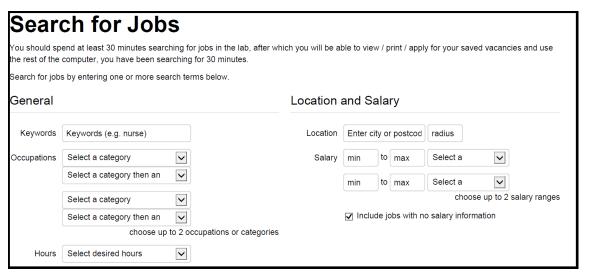
You can search using various criteria (keywords, occupations, location, salary, preferred hours). Importantly, you do not have to specify all of these. You just need to fill at least one of them.

If you specify more than one criterion, it is important to note that the computer will search for vacancies that satisfy <u>all</u> the criteria at the same time. For example, if you enter a keyword <u>and</u> you also select an occupation, it will search for vacancies that match <u>both at the same time</u>. Vacancies that match the keyword but not the occupation will not be shown.

Within some categories you can fill in more than one field. For example, within "occupations" you can specify up to two of them. If you do fill in two occupations, the computer that match either the first OR the second occupation. Vacancies that match one occupation but not the other will still be shown. You can also specify more than one pay range. This allows you to specify, for example, the hourly wages and the yearly wages that you are willing to accept. If you only specify hourly wages, it will not show vacancies that only specify yearly wages.

If you fill in your preferred hours, for example full time work, it will only list vacancies where the employer ticked a box that it is full-time work. Vacancies where the employer did not explicitly state that it is full-time work will not be shown.

If you leave a field empty, the computer will not use that criterion to restrict your search.



Once you have defined your search criteria, you can press the search button at the bottom of the screen and a list of vacancies fitting your criteria will appear. You can click on each individual vacancy to get more information about it. You can then either

- Save the job (if you are interested in applying)
- Do not save the job (if you are not interested)

**If you save the job,** the computer will keep a record of the vacancy. You will be able to see all records of all saved vacancies at the end of the session.

If you do not want to save the job and want to go back to the search results, we will first ask you a few questions about why you are not interested in the job. Your answers are very important to us.

You can modify your search criteria at any point and launch a new search.

Note that we have also created a small number of vacancies ourselves (about 2% of the database), which are there for research purposes only. This is to learn whether you would find these vacancies attractive and would consider applying to them if they were available. We kept them to a minimum not to disturb your search. These vacancies will appear as all the other vacancies and may appear in your search results. But we will inform you at the end of the 30 minutes of any vacancy that may not be real. You will be able to see the list of your saved vacancies immediately after the 30 minutes are over, and we will indicate if any of them was an artificial one.

We may try alternative interfaces for the job search engine in the coming weeks. We will inform you if we do so and will explain the changes at that point in time.

#### 4. FREE USE OF THE FACILITIES (after 30 minutes)

We will let you know when the first 30 minutes are over. You will then be free to use the computer for other purposes. You can of course keep searching using our job search engine, or you can do other things, such as write your CV, write a letter, or even send e-mails. You can use the facilities for up to 2 hours.

If you do not wish to continue searching or use the computer for other purposes, you are free to leave.

#### **END OF THE SESSION**

We can print a record of your job search for the day (just call us once you have finished), but only if that is your wish. You are free to show these records to your adviser at the Job Centre. They informed us that this would count as a proof of search activity.

Compensation: In general, you will receive a total of £11 as a compensation for your travel and meal expenses. This time, as you will soon discover in the initial survey, we do offer you the possibility of investing part of this compensation in this initial session. This is not compulsory. But if you do choose an investment option, your earnings will then be a function of what investment you have chosen.

Please collect your compensation from the registration room. You will get an envelope and be asked to sign a receipt. Note that the Job Centre has agreed that these £11 are a compensation for expenses and are not an income.

#### **IMPORTANT NOTES**

#### LOG IN FROM HOME OR FROM ANOTHER COMPUTER

You will be able to use our search engine from home or from another computer as well. You just need to log in on the website and use your login information. You will be able to see all the vacancies you saved and will be able to retrieve all the relevant information about them.

Note that as indicated in the consent form, all records saved are anonymous. These will not be matched to your names at any point.

#### YOUR COMMITMENT

Note that it is very important for us that you come back every week and search in our facilities, unless of course you have found a job. If for one reason or the other you do have to cancel your session in a given week, please let us know as soon as possible. We will either try to reallocate you to another slot or ask you to search from home in that particular week. If you have found a job, please do let us know. This is of course of key importance for our study.

Also, importantly, you will receive a £50 clothing voucher for <u>each four consecutive weeks</u> you come. The first voucher will be distributed in the fourth week, that is, three weeks from now. The second voucher will be distributed in the eighth week and the third voucher in the twelfth week.

Thank you very much for your attention. If you have any questions, please raise your hand and we will come to you.

8.2.3 Lab instructions alternative interface

# **PLEASE READ**

## **NEW JOB SEARCH INTERFACE**

#### **IMPORTANT CHANGES**

We have designed a new search interface that should give you a better idea of jobs that might be relevant to you. This new interface suggests additional types of jobs (occupations) that are related to your preferred occupation.

You will be asked to specify your preferred occupation and the interface will return suggestions of other occupations that may be of interest to you. They may not all be relevant, but hopefully some will be relevant and will allow you to broaden your search horizon.

We use two methodologies to do this:

The first is using information from national labour market statistics, which follows workers over time and record in what occupation they are employed. The data records transitions between occupations and we can identify the most common occupations people switch to from a given occupation. We will ask you to indicate your preferred occupation using a keyword search and selecting the relevant title in a drop-down menu. The second is using information on transferable skills across occupations from an American website (called O\*net). For each occupation, we will suggest up to 10 related occupations that require similar skills.

Since the databases are different for each of the two routes, we will ask you to specify your preferred occupation twice and select it in the menu of possible occupations. So we will ask you again to indicate your preferred occupation using a keyword search and selecting the relevant title in a drop-down menu.

Once you have specified your preferred occupation for each of the two methodologies, you can then click "Save and Start Searching"

and you will be taken to a new screen that will suggest these new occupations to you.

The occupations will be listed in two columns:

The left column suggests occupation based on the first methodology (based on the UK labour market transitions). The right column suggests occupations based on the second methodology (O\*net related occupations).

You can select or unselect the occupations you find relevant and would like to include in your search.

We also have information about how competitive the labour market is for a given set of occupations. We have constructed "heat maps" that use recent labour market statistics for Scotland and show you where jobs may be easier to get (because there are many jobs relative to the number of interested job seekers). These maps are based on broad categories of jobs, not on each very specific occupation. You can click on the button "heat map" to see the relevant map. We would like you to try this new interface from now on.

It is nevertheless possible to switch back to the old interface that you have used in the previous weeks. You will see a button on the screen indicating "use old interface". If you click it, you will be taken to the old search engine interface. From there you can also return the new interface.

Thank you very much for your attention.

8.2.4 Baseline survey questionnaire

# **INITIAL SURVEY**

We will start by asking a few questions about your background and personality. Please fill in the answers as appropriate.

Gende	r: [drop down m	enu]
	Male Female	
Countr		down menu with all countries in alphabetical order]
Ethnici	<b>ty:</b> [drop down r	menu]
	Caucasian whit	e
	East Asian	
	Black African	
	Black Caribbea	n
	Indian	
	Pakistani	
	Bangladeshi	
	Other	
Age: _	[number]	
What a	are the first 3 let	ters of the postcode of your residence? [EH1 until EH17 as dropdown menu
Qualifi	cations (tick the	appropriate box): [drop down menu]
	0	Ph.D.
	0	Postgraduate Masters degree
	0	Undergraduate Degree
	0	Other higher education
	0	A level / Higher or equivalent (secondary education)
	0	GCSE Other qualification
	0	No qualification
	O	ivo qualification
Data w	ou hocamo unor	nnloved: / / [numbers]
Date yo	ou became uner	mployed: / [numbers]
Date of	f registration wi	th Job Seeker Allowance: / / [numbers]

# Job experience

From (date) to (date)	Employer	Job title	Reason for departure
[numeric fields]	[open field]	[open field]	[drop down menu]
(month) (year)			Temporary contract
			Redundancy
			Voluntary quit

(mone	(year)					Redundancy Voluntary quit
Les. Les. Les. Les. Les. Les.	o you think you think you stan 4 week sthan 8 week sthan 12 week sthan 6 mont sthan a year will take me mo	s s ks hs	•	drop dow	n menu]	
[drop dowr	cupation woul n menu with to ocation (and r	he detailed			able in unive	rsal job match]
City:	P	ostcode: _		_ Radi	us: (m	niles)
In what ran	ge of salaries	are you lo	oking for a jol	)?		
£ month]	[number] to i	<u> </u>	[number]	[drop	down menu	: per hour, per week, per
What type	of contract are	e you lookir	ng for? (you ca	n select mo	ore than one a	answer if appropriate)
. 🗆	Full Time					
. 🗆	Contract					
. 🗆	Part Time					
. 🗆	Placement S	Student				
. 🗆	Temp					
. 🗆	Other					
How many	vacancies did	you apply s	since you have	become ui	nemployed? _	[Number]

How many job interviews did you get so far? \_\_\_\_ [Number]

How many job offers did you get so far? [Number]	
What are your most important concerns at the moment (rate 10 (very strong concern)).	on scale from 0 (not a concern at all) to
My financial situation is deteriorating Personal difficulties prevent me from focusing on job search Health-related problems hinder my job search activities	[number] [number] [number]

#### Risk preferences question

We now offer you the possibility to do a gamble with some of the compensation you will receive for today's session. You do not have to participate. If you participate, we will reduce your compensation by £2.80, but you will earn an amount of money depending on the gamble you choose and the outcome of the gamble.

We propose you 5 gambles. You can only choose one of them. Indicate your choice at the bottom of the page.

Each gamble corresponds to a flip of a coin and has two possible outcomes (Heads or Tail). We indicate below what you would win in each case. We will flip a coin at the end of the session, when you leave the room. Note that you do not have to play and you can simply choose to keep £2.80.

#### Gamble 1

TAIL: £2.40 HEADS: £3.60

Gamble 2

TAIL: £2.00 HEADS: £4.40

Gamble 3

TAIL: £1.60 HEADS: £5.20

Gamble 4

TAIL: £1.20 HEADS: £6.00

Gamble 5

TAIL: £0.20 HEADS: £7.00

Your choice [drop down menu]

Ш	1 keep £2.80
	I play Gamble 1

- ☐ I play Gamble 2
- ☐ I play Gamble 3
- □ I play Gamble 4
- ☐ I play Gamble 5

#### Time preferences questions

At the end of the session, one participant in the room will be selected at random and will receive lottery tickets (in addition to the compensation promised). Each ticket gives the chance to win up to £250,000. Note that the lottery tickets will be sent at the date indicated to the person's home address, so you will not need to collect them here.

Could you please indicate for each of the 15 choices below which option you would prefer. If you are selected, we will select one of the 15 choices at random and send you the relevant number of tickets at the date chosen.

Choice 1:	☐ 5 lottery tickets today	☐ 6 lottery tickets in a week
Choice 2:	$\square$ 5 lottery tickets today	☐ 7 lottery tickets in a week
Choice 3:	$\square$ 5 lottery tickets today	$\square$ 8 lottery tickets in a week
Choice 4:	$\square$ 5 lottery tickets today	$\square$ 9 lottery tickets in a week
Choice 5:	$\square$ 5 lottery tickets today	$\hfill\Box$ 10 lottery tickets in a week
Choice 6:	$\square$ 5 lottery tickets today	☐ 6 lottery tickets in 4 weeks
Choice 7:	$\square$ 5 lottery tickets today	☐ 7 lottery tickets in 4 weeks
Choice 8:	$\square$ 5 lottery tickets today	☐ 8 lottery tickets in 4 weeks
Choice 9:	$\square$ 5 lottery tickets today	☐ 9 lottery tickets in 4 weeks
Choice 10:	$\square$ 5 lottery tickets today	$\square$ 10 lottery tickets in 4 weeks
Choice 11:	$\square$ 5 lottery tickets in 8 weeks	☐ 6 lottery tickets in 12 weeks
Choice 12:	$\square$ 5 lottery tickets in 8 weeks	☐ 7 lottery tickets in 12 weeks
Choice 13:	$\square$ 5 lottery tickets in 8 weeks	☐ 8 lottery tickets in 12 weeks
Choice 14:	☐ 5 lottery tickets in 8 weeks	☐ 9 lottery tickets in 12 weeks
Choice 15:	☐ 5 lottery tickets in 8 weeks	☐ 10 lottery tickets in 12 weeks

8.2.5 Weekly survey questionnaire

# Weekly job survey

How many hours did you spend searching for jobs? \* For the following questions please exclude any searching done during the previous session here at the university or applications made as a result. Did you search for jobs using any of the following (you can select more than one answer if appropriate) ☐ DirectGov / Universal Jobmatch ☐ Other internet websites □ Newspapers ☐ Through friends / family / acquaintances ☐ Through the jobcentre Through a private employment agency Approached employers directly (handing in CVs etc.) Please specify any other ways you looked for a job 4 How many other vacancies did you apply to? \* Please tell us the title, employer and salary information for any jobs you applied for (if known) How many interviews did you go to? \* How many job offers did you get? \* Did you accept a job offer? \* O Yes No If you have worked in a temporary or part-time job in the past week please tell us about it (title, employer, hours, part/full-time, salary information)

We will now ask a few questions about your other search activities over the past week.

If you took part in any training since last weeks session please tell us what this was

Did you meet a case worker at the jobcenter? *
○ Yes ○ No
Are jobs that you encounter in your other search activities broadly similar to those that you encounter when searching here at the university? *
<sup>○</sup> Very similar <sup>○</sup> Similar <sup>○</sup> Different <sup>○</sup> Very different
Finally we will ask a few general questions.
What are your most important concerns at the moment (rate on scale from $0$ (not a concern at all) to $10$ (very strong concern))
My financial situation is deteriorating *
Personal difficulties prevent me from focusing on job search *
There is strong competition for jobs *
Health-related problems hinder my job search activities*
Do you have any feedback for us on our search engine and computer interface?
<b>●</b>

# 8.2.6 Heat maps

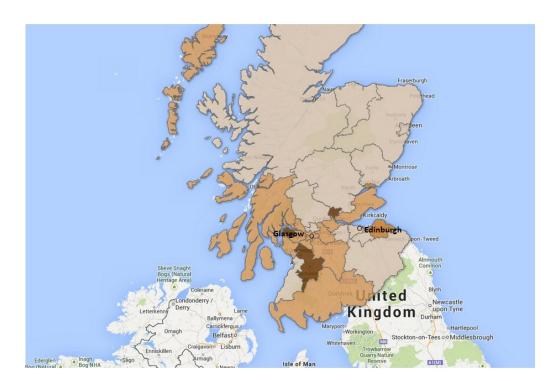


Figure 17: Example of a heatmap

The darker the color, the higher the number of job seekers per vacancy in the particular occupation.