Mortality, Temperature, and Public Health Provision:

Evidence from Mexico

François Cohen (Graduate Institute of International and Development Studies,

Geneva, Switzerland)

Antoine Dechezleprêtre (London School of Economics and Political Science,

London, UK)

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Abstract

We examine the impact of temperature on mortality across income groups in Mexico using

individual death records (1998-2010) and Census data. Random variations in temperature are

responsible for 8% of deaths in Mexico (45,000 deaths every year). However, extremely hot

days (>32°C) kill a comparatively low number of people (<400 annually): 88% of weather-

related deaths are induced by mildly cold days (10-20°C). Moreover, mildly cold

temperatures only kill in the bottom half of the income distribution. We show that the Seguro

Popular, a universal healthcare policy progressively rolled out since 2004, reduced cold-

related mortality among eligible people by about 30%.

Keywords: temperature; mortality; inequality; universal healthcare; distributed lag model

JEL codes: I14, Q54

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1. Introduction

Between 1980 and 2015, life expectancy increased from 61.7 to 71.8 years worldwide (Global Burden of Disease Study, 2016). This significant gain in longevity largely comes from developing countries. However, in the context of future climate change, the convergence in life expectancy that has been observed between high and low-income countries could slow down. This is because these countries are more exposed to the consequences of climate change and have lower adaptive capacity. Recent studies looking at the impacts of extreme temperatures in the US (Heutel, Miller and Molitor, 2017; Barreca, 2012; Deschenes and Greenstone, 2011; Deschenes and Moretti, 2009; Braga *et al.*, 2001) have found much smaller impacts than those found in rural India by Burgess *et al.* (2014), suggesting the importance that income may play in shaping the impacts of climate change on health both across and within countries. ^{3,4}

In this paper, we focus on the relationship between temperature, mortality and social inequalities in Mexico. Temperature shocks are the most direct way that climate change can affect mortality. We use daily mortality records from 1998 to 2010 for 2,100 Mexican municipalities, representing around 95% of the country's population. The dataset includes over 9 million observations of daily mortality rates, matched with weather data from the closest meteorological stations. Our objective is to measure the extent to which temperature stress may unequally affect different groups of the population, depending on their capacity to protect themselves from adverse weather conditions.⁵ To this effect, we match the characteristics of individuals as reported in death records to the Mexican census data. This

¹ The gap is still large: life expectancy at birth is over 75 years in high-income countries, against 60 in low-income countries (WHO, 2016).

² Fast-growing emerging economies (e.g. China, Brazil or Mexico) have experienced a drastic reduction in the mortality caused by transmissible diseases in a couple of decades. Despite these significant gains, developing countries are less prepared to face catastrophic events affecting health: the recent West Africa Ebola virus epidemic is one example; another is the HIV/AIDS pandemic, which is still causing more than a million deaths each year. Climate change in the twenty-first century will likely add to this burden by affecting many determinants of health: water, food supply, public infrastructure, housing, economic growth and conflict. WHO estimates that by 2030-2050 climate change could trigger 250,000 additional deaths per year globally (http://www.who.int/mediacentre/factsheets/fs266/en/).

³ Burgess et al. (2014) attribute the high mortality effect of heat in rural India to the impact of extreme weather on agricultural yields, which depresses farmers' wages and in turns pushes mortality up.

⁴⁴ Climate change may also widen the differences that exist within countries between the rich and the poor. These differences seem to have increased recently for non-transmissible diseases: in the United States, recent studies have documented a growing life expectancy gap between the affluent and less affluent, which has been associated to widening income inequality (Olshansky et al., 2012).

⁵ The relationship between income and health has been studied at least since Preston (1975). The health economics literature (see Deaton, 2003, for a review) has shown that it is difficult to disentangle the impact of income on improving health from the impact of good health on raising revenues. In the context of climate change, income determines the capacity that households have to invest in protective measures.

allows us to estimate the income level of each individual in our dataset at the time of their death and analyse the vulnerability to temperature shocks across income groups. To our knowledge, this paper is the first analysis of the heterogeneous relationship between temperature and mortality in a developing country that combines daily mortality data with individual income-level data. In the final section of the paper, we exploit the progressive implementation of a universal healthcare policy – the *Seguro Popular* – to analyse the impact of extending universal healthcare on reducing weather vulnerability. To our knowledge, this paper is also the first to assess the impact of public health policies on climate resilience.

The use of daily data has two major advantages. First, the inclusion of municipality-by-month-by-year fixed effects allows us to purge the estimates from a large number of confounding factors that might be correlated with both temperatures and mortality. Second, the use of distributed lag models \grave{a} la Deschenes and Moretti (2009) or Braga et al. (2001), allows us to report the effect of extreme temperatures on mortality up to a month after an unusually hot or cold day, which accounts for possible mortality displacement effects.⁶

We find that random variation in temperatures is responsible for the death of around 45,000 people every year in Mexico, representing 8% of annual deaths in the country. Consistent with previous epidemiological research, we find that the most vulnerable group to extreme temperatures is the elderly, in particular people over 75 years old, followed by young children. The predominant causes of death from excessive heat and cold are, in order of importance, circulatory, respiratory and metabolic diseases. We find no statistically significant difference between men and women. In terms of magnitudes, our results suggest at least a three-time stronger vulnerability to cold in Mexico than in the US when we compare our estimates to recent studies (Barreca, 2012; Deschenes and Greenstone, 2011; Deschenes and Moretti, 2009; Braga *et al.*, 2001). On the other hand, we find a modest, but statistically significant impact of heat on mortality whereas previous studies in the US did not find any impact (Deschenes and Moretti, 2009; Braga *et al.*, 2001). In contrast, the impact of temperatures on mortality in Mexico appears to be much smaller than the impacts found in India by Burgess *et al.* (2014).

A first interesting contribution of this study is to unveil the impact of mildly cold temperatures on mortality. Whereas the media usually pay attention to extreme heat and cold,

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⁶ While the impact of on-the-day temperatures on death has been widely reported in the medical literature, their impact on longevity is debated and depends on the magnitude of mortality displacement effects: extreme temperatures could simply accelerate the death of already weak people by only a few days (e.g. Deschenes and Moretti, 2009; Hajat et al., 2006; Hajat et al., 2005; Braga et al., 2001).

these events are infrequent and only account for a minority of weather-related deaths in our analysis. In a hot country like Mexico, even days with mean temperature below 20°C (68°F) are associated with statistically significant increases in the daily mortality rate. Therefore, while very cold days with mean temperature below 10°C are responsible for the death of around 4,700 people each year, we estimate that 88% of weather-induced deaths – around 40,000 people per year – occur because of temperatures between 10°C and 20°C. In contrast, extremely hot days over 32°C trigger a comparably small amount of additional deaths (around 400 annually). We present data on the very low rate of heating equipment across Mexico which may account for this impact of mildly cold temperatures on mortality.⁷

In terms of longevity, we find that the number of years of life lost due to cold days under 10°C is 50% larger for children under 5 than for people aged 75. This is not only because children under 5 have a longer life expectancy, but also because the Mexican population is very young: there are around four times more children under 5 than people over 75. These results are in sharp contrast with the ones found by Deschenes and Moretti (2009) for the US. These authors found a large effect of cold temperatures on the longevity of people over 75, and a negligible one on children under 5.

The second contribution of this research is to show that vulnerability to extreme weather is negatively correlated with personal income. Controlling for differences in the age structure across income groups, we show that vulnerability to unusual cold (defined as a day with mean temperature below 10°C) is 35% higher for people in the bottom half of the income distribution compared to people in the top half. Death following mildly cold days (10-20°C) appears to concern only people living below the national median personal income. Hence, the great majority of cold-related deaths only affect the poorest income groups. In contrast, we find no statistically significant differences in vulnerability to heat across income groups.

The final contribution of this study is to assess the impact that improved access to healthcare has on reducing weather-related vulnerability. Our epidemiological analysis shows that policies targeting the most vulnerable people (particularly young children and the elderly in low-income households) could significantly reduce weather-related mortality. However, such

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⁷ These figures are informative about the direct impact that climate change is likely to have on human health in Mexico. Using predictions from several climate change scenarios, we expect that climate change will considerably reduce direct exposure to the temperature levels that are highly detrimental to health in Mexico: by the end of the 21st century, our model predicts that the number of weather-related deaths in Mexico would decrease by 50 to 80% even in the absence of any adaptation. This finding stands in sharp contrast with most recent analyses of both developed and developing countries, which tend to predict that climate change will significantly increase temperature-induced mortality (e.g. Deschenes and Greenstone, 2011; Burgess et al. 2014), and they illustrate the vast heterogeneity in climate change impacts across countries and regions.

policies should not focus on extremely cold days – unlike, for example, early warning systems and other health communication – but provide protection all year round since mildly cold days are responsible for the clear majority of weather-related deaths. This suggests that expanding access to healthcare (particularly for vulnerable groups) may be able to significantly reduce weather vulnerability. During our study period, Mexico implemented two nationwide policies – the *Seguro Popular* and the *Fondo de Protección contra Gastos Catastróficos* – to increase access to healthcare for low-income households. These programmes have the specificity that they do not provide full coverage and only target a list of priority diseases. We exploit exogenous variation in the causes of death affecting the Mexican population to assess the effectiveness of public healthcare measures in reducing the weather vulnerability affecting low-income families. We focus on the short-term impacts of the programmes on the provision of medical services in life-threatening cases. While our analysis focuses on one specific aspect of the impacts of the schemes on mortality, it actually is the first assessment of the impact of these schemes on mortality.

We find that the schemes increased medical assistance in life-threatening cases by three percentage points. This is a relatively high figure considering that medical assistance in cases of emergency is already high in Mexico (at around 84% in our sample). We estimate that this increase in medical assistance saved around 3,000 lives in 2010 only. Interestingly, many of these lives were saved thanks to a reduction in weather-related mortality. We find that the schemes have reduced mortality during cold days (<10°C) by around 30% for the set of diseases that they cover, i.e. respiratory diseases and diabetes, which are particularly sensitive to cold.

The policy implications from this paper go beyond the frontiers of Mexico. Even in hot countries where the coldest temperatures almost never reach 0°C, cold remains a risk factor with potentially large health impacts. Low-income households, particularly in the developing world, are ill-equipped to protect themselves against it. This puts them at a higher risk at all ages, and particularly when they become older. Furthermore, these households are at risk over longer time periods in the year than richer households, since they appear to be vulnerable to

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⁸ Access to healthcare is a major issue in Mexico: according to our own calculations based on data from the 2000 Mexican Census, over 80% of people that belong to the first quartile of the income distribution do not have access to social security. More generally, Mexico is the OECD country with the lowest budget dedicated to health: \$1,052 per capita and in purchasing power parity, compared to \$3,814 on average in other OECD countries (OECD Health Statistics 2016).

⁹ Other papers looking at the Seguro Popular have focused on health spending (King et al., 2009), health expenditure and self-declared information on health issues (Barros, 2008), access to obstetrical services (Sosa-Rubi, Galarraga and Harris, 2009) and prenatal services (Harris and Sosa-Rubi, 2009).

even mildly cold temperatures. We show that access to universal healthcare can successfully reduce this high vulnerability.

The remaining of this paper is structured as follows. Section 2 discusses the previous empirical literature on the impact of weather on mortality. Section 3 describes the data. Our general results on the impact of temperatures on mortality are presented in Section 4. Results by quartiles of income are presented in section 5, and the impact assessment of universal healthcare on reducing weather-related mortality is presented in section 6. A conclusive section summarises our findings and briefly discusses the implications of our results.

2. Previous empirical literature on temperature and mortality

To quantify heat- and cold-related mortality, epidemiological studies usually rely on a Poisson regression framework and daily death counts for a city or a group of cities. Recent epidemiological studies (e.g. McMichael *et al.*, 2008; Hajat *et al.*, 2007; Hajat *et al.*, 2006; Curriero *et al.*, 2002) put in evidence a J-, U- or V-shape relationship between on-the-day temperature and daily mortality. Populations face lower mortality at a given threshold temperature, which differs from one location to another (e.g. due to acclimation) and may possibly change over time. Above and below this threshold, mortality increases and, the farther from the threshold, the greater is heat- or cold-related mortality. This is very much in line with the medical evidence that the human body starts being at risk outside a comfort zone which varies across individuals but is generally believed to lie in the range of 20°C to 25°C. To From a methodological perspective, such a nonlinear relationship between mortality and temperature calls for the use of temperature bins in panel data analyses (Deschenes and Greenstone, 2011): the impact between temperature and mortality is then separately evaluated at different levels of temperature stress.

Despite evidence from the medical literature that even mildly cold or hot days can negatively affect human health, the economic literature has primarily focused on the impact of extremely hot and cold days (see for example Deschenes and Greenstone, 2011; and Deschenes and Moretti, 2009), plausibly because these extreme weather events tend to concentrate media attention. However, while the impact of a mildly cold or hot day is definitely less dangerous than that of an extremely hot or cold day, days lying outside the typical human body comfort zone are much more frequent.

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¹⁰ See the Appendix A1 for a background discussion about the physiological impact of cold and heat on human health.

This misrepresentation of the relative burden of extreme temperatures in the media is particularly striking in the case of very hot days. Whereas unusually hot days receive media attention, the question of their actual impact on mortality remains controversial once account is taken of displacement effects, i.e. the impact of a day's temperature on the mortality levels of the following days. On-the-day extra mortality on hot days was often found to be offset by lower mortality levels in the following days, suggesting that mortality on hot days may largely correspond to a harvesting effect (Deschenes and Moretti, 2009; Hajat *et al.*, 2005; Braga *et al.*, 2001). ¹¹

However, uncertainty remains on the true mortality impact of hot days because extreme weather events may not only directly affect human physiology, but also reduce agricultural output, potable water availability or family income. These elements may in turn affect health or access to healthcare and lead to extra mortality. In order to account for these longer-term impacts, a few economic studies have used monthly or annual panel data rather than daily data (Burgess *et al.*, 2014; Barreca *et al.*, 2016; Barreca, 2012; Deschenes and Greenstone, 2011). These studies find a clear correlation between hot temperatures and monthly or annual mortality. Burgess *et al.* (2014) find a strong impact of extreme temperatures on annual mortality in India, plausibly because shocks on temperatures affect agricultural productivity, and therefore the food intake and income of populations located in rural areas.

The existence of such economic factors in addition to the standard epidemiologic ones suggest that people's vulnerability to cold and hot temperatures depends on their access to protection measures. Evidence is contradictory on the impact of income on vulnerability to cold and hot temperatures. Hajat *et al.* (2007) find no evidence that deprivation modifies the impact of heat or cold on mortality in England and Wales, except in rural populations where cold effects were slightly stronger in more deprived areas. Medina-Ramon and Schwartz (2007) look at 50 cities in the US and find no difference in the effect of cold across cities, whereas some of the heterogeneity observed for the impact of heat could be explained by several city characteristics. Overall, they observe the lowest increases in mortality due to extreme heat in those cities with a large proportion of central air conditioning. Similarly, Barreca *et al.* (2016) argue that there is a strong correlation between the reduced mortality due to heat that has been

¹¹ For example, Gouveia et al. (2003) study the relationship between temperature and mortality in Sao Paulo. They find a positive relationship between mortality and heat when they use on-the-day and the previous day's temperatures. However, this effect dissipates when using a 3-week lag structure. This suggests harvesting in the case of hot days, similarly to the study performed by El-Zein et al. (2004) in Lebanon. These authors focus on Greater Beirut and, while the effect of cold days on mortality is persistent, the statistically significant effect for hot days dissipates when 14 day lags are taken into account.

¹² See Bupa (2008) and Deschenes (2014) for thorough literature reviews.

observed in the US over time and the gradual deployment of air conditioning. Going one step further, Heutel, Miller and Molitor (2017) argue that the deployment of air conditioning explains regional differences in the health impact of heat on the elderly in the US. Deschenes and Greenstone (2011) find a positive but moderate correlation between annual mortality rates and temperature, and predict a 3% increase in age-adjusted mortality by the end of the 21st century due to climate change. At the same time, they find that electricity consumption would increase four times as much as age-corrected mortality under climate change in the US, implying that households could massively resort to air-conditioning so as to reduce the negative impact of climate change on human health. Other potential adaptations include migration to places with a more indulgent climate (Deschenes and Moretti, 2009) or a reduction in the time spent outdoors (Graff-Zivin and Neidell, 2010).

Differences in the ability of populations to adapt might exist within countries, but also between countries, in particular between industrialised and emerging countries. Focusing on emerging countries, McMichael *et al.* (2008) gather mortality, weather and pollution data from 12 cities in medium and low income countries. The paper shows different patterns of heat and cold mortality according to the city under study. For a few cities, McMichael *et al.* (2008) find a negative effect of high temperatures on mortality using 13-day lags. For some others, a steep effect of cold on mortality is found, and no effect of heat on mortality. The authors conclude that many populations are likely to have substantial vulnerability to climate change, but that additional research is needed to elucidate vulnerability within populations.

A few studies provide indications or quantify the impact of climate change on mortality. Medina-Ramon and Schwartz (2007) and McMichael et al. (2008) argue that the magnitude of the effects of hot days may not be compensated by the reduction in mortality associated with warmer winter periods. More formally, Deschenes and Greenstone (2011) use the output of the Hadley CM3 model (error corrected) for the A1F1 scenario (2070-2099) and find an increase in mortality in the US by 3% due to climate change by the end of the century. Taking into account humidity and using the same climate change hypotheses, Barreca (2012) finds a very small reduction in mortality for the US (-0.08%). However, this reduction would hide significant heterogeneity: in the colder States of the US, mortality would decrease whereas it would significantly increase (by 0.44-2.92% depending on the region) in the warmest and most humid States. For India, Burgess et al. (2014) find a significant increase in heat-related

¹³ The cities are Delhi, Monterrey, Mexico City, Chiang Mai, Bangkok, Salvador, São Paulo, Santiago, Cape Town, Ljubljana, Bucharest and Sofia.

mortality in rural areas. Climate change impacts translate into a large increase in mortality by the end of the century of 12-46% depending on the climate model employed.¹⁴

Due to the variability of results from one city to the other, either in developing countries or in industrialised ones, it is difficult to conclude from existing epidemiologic research to what extent emerging countries are more vulnerable to excessive heat or cold. Furthermore, by focusing on large cities, they do not provide any information on vulnerability in rural areas, where most heat-related excess mortality could occur.

In this study, we consider that weather vulnerability in emerging economies may substantially differ from that in developed countries. In particular, all developed countries have already experienced an epidemiological transition: cancers and other non-transmissible diseases have long become the major cause of death in these countries, whereas this phenomenon is recent in developing countries. Furthermore, elemental protection measures (e.g. proper clothing) are available to all in industrialised countries, and national programs such as Medicare and Medicaid provide universal healthcare coverage in life-threatening cases.

3. Data and summary statistics

To evaluate the relationship between temperature and mortality in Mexico, several data sources have been matched, in particular mortality data from the Mexican National Institute of Statistics and Geography (INEGI) and weather data from the National Climatological Database of Mexico. By combining data on mortality with weather data, we are able to construct daily municipal mortality rates for about 2,100 Mexican municipalities over 13 years (1998-2010). The final daily dataset includes over 9 million observations.

3.1 Mortality data

Our mortality data corresponds to the Mexican general mortality records (*defunciones generales*) from 1990 onwards as assembled by INEGI. The micro-data provides information about each case of death in Mexico, including cause, municipality and time of death along with socioeconomic information on the deceased. The records provide information on the exact day of death between 1998 and 2010, while only information on the month of death is provided in anterior records. A template of death certificate used in Mexico is provided in Appendix A2.

¹⁴ Either Hadley 3, A1F1 scenario or CCSM3, A2 scenario for 2070-2099

¹⁵ In 2008, there were 2,454 municipalities in Mexico (INEGI, 2008). Some municipalities are missing in our dataset because they are far from the weather stations, or close to weather stations that did not efficiently record both min. and max. temperatures.

Table 1 displays the average daily mortality rate by cause of death, gender and age, together with the average population within each group for 1998-2010. The average daily mortality rate across all municipalities is 1.3 deaths per 100,000 inhabitants. This figure is about twice as low as the current rate in the United States (see Deschênes and Moretti 2009), a feature that is explained by the larger proportion of young people in Mexico. The death rate is lowest for children aged 4-9 and rises non-linearly until it reaches 21.2 per 100,000 inhabitants for people aged 75 years and above.

Table 1: Summary of death statistics

		Average daily municipal mortality rate (deaths per 100,000 inhabitants)							
Group	Average population per municipality	All causes	Respiratory system diseases	Circulatory system diseases	Endocrine, nutritional and metabolic diseases	Infectious diseases	Neoplasms	Violent and accidental deaths	All other deaths
Total	44935	1.30	0.114 (8.8%)	0.295 (22.7%)	0.204 (15.7%)	0.049 (3.8%)	0.163 (12.5%)	0.142 (10.9%)	0.333 (25.6%)
Men	21886	1.47	0.127 (8.6%)	0.303 (20.6%)	0.195 (13.3%)	0.061 (4.1%)	0.162 (11%)	0.231 (15.7%)	0.391 (26.6%)
Women	23049	1.13	0.102	0.288 (25.5%)	0.212 (18.8%)	0.038 (3.3%)	0.164 (14.5%)	0.057	0.27 (23.9%)
Aged 0-4	4543	1.08	0.12 (11.1%)	0.014 (1.3%)	0.037	0.081 (7.5%)	0.015	0.081 (7.5%)	0.732 (67.7%)
Aged 4-9	4739	0.08	0.005	0.002	0.003	0.006 (7.4%)	0.014	0.027	0.024 (29.2%)
Aged 10-19	9227	0.15	0.005	0.007	0.004 (2.6%)	0.006	0.017 (11.1%)	0.079 (52.4%)	0.033 (22.3%)
Aged 20-34	11042	0.37	0.012 (3.3%)	0.023 (6.3%)	0.015	0.032 (8.6%)	0.03 (8.2%)	0.178 (48%)	0.08 (21.7%)
Aged 35-44	5674	0.66	0.025	0.075	0.06	0.051	0.089	0.174 (26.2%)	0.19 (28.6%)
Aged 45-54	3880	1.39	0.056	0.232 (16.7%)	0.248 (17.8%)	0.061 (4.4%)	0.235 (16.9%)	0.182	0.377 (27.1%)
Aged 55-64	2462	3.10	0.16 (5.2%)	0.658 (21.2%)	0.751 (24.2%)	0.091 (2.9%)	0.545 (17.6%)	0.207 (6.7%)	0.688 (22.2%)
Aged 65-74	1482	5.16	0.266 (5.2%)	1.09 (21.1%)	1.25 (24.2%)	0.151 (2.9%)	0.905 (17.5%)	0.343	1.155 (22.4%)
Aged 75+	963	21.25	2.92 (13.7%)	7.51 (35.3%)	3.41 (16%)	0.425	2.28 (10.7%)	0.591 (2.8%)	4.114 (19.4%)

Notes: The table shows cause-specific daily mortality rates in number of deaths per 100,000 inhabitants. The share of average group mortality is presented in brackets. The sample includes 2,289 municipalities over 11.65 years on average. All means are weighted by the relevant population group in municipalities.

We break down mortality rates by cause of death, based on the typology of the 10th version of the International Classification of Diseases (10-ICD) of the World Health Organisation (WHO). We consider seven types of cause of death: infectious and parasitic diseases; malign neoplasms; endocrine, nutritional and metabolic deaths (including diabetes which account for 80% of deaths in this category, followed by malnutrition with 11% of deaths); diseases of the

INEGI for the years of the national censuses (1990, 1995, 2000, 2005 and 2010). We perform a linear interpolation of the population for the years between two censuses to obtain estimates of the Mexican population of each municipality in each year between 1990 and 2010. This may introduce measurement error in the dependent variable, a problem known to reduce model efficiency but not the consistency of the estimates.

¹⁶ We calculate daily municipal mortality rates by dividing the amount of deaths in a municipality on a specific day with the population in this municipality. To do so, we use municipal population data available from the INECL for the years of the retional consumer (1000, 1005, 2000, 2005, and 2010). We parform a linear

circulatory system; diseases of the respiratory system; and violent and accidental deaths. ¹⁷ As it has been reported elsewhere, the primary cause of death is circulatory system diseases, which, as we have seen above, has been identified as affected by temperatures in the epidemiologic literature. The importance of each cause of death differs by age and gender. For example, the prevalence of violent and accidental death is four times greater among men than among women. It is also the main cause of death for people aged between 10 and 44. The importance of circulatory system diseases rises with age and peaks above 75, when it becomes the primary cause of death.

3.2 Weather and climate data

The National Climatological Database of Mexico provides daily temperature and precipitation records for around 5,500 operating and formerly operating land-based stations in Mexico. Information on the longitude and latitude of the stations is also provided. For this project, we have used daily records from 1989 onwards.

Figure 1 below presents the historical distribution of daily average temperature in Mexico from 1998 to 2010.¹⁸ The data has been aggregated at municipality level and weighted according to the population of each municipality to reflect the average exposure of Mexicans to low and high temperatures.¹⁹

We use 13 temperature bins: "below 10°C", "above 32°C" and eleven 2°C bins in between. In the empirical models presented hereafter, we use the same temperature bins to estimate the relationship between temperature and mortality. In Figure 1, each bar represents the average number of days in each temperature category for the average person in Mexico. The mode of

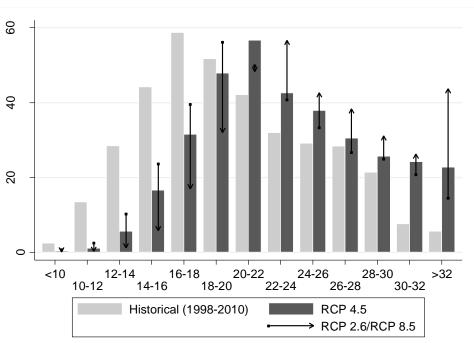
¹⁷ Their correspondence with 10-ICD codes (2010) is as follows: Infectious diseases (A-B); neoplasms (C); endocrine, nutritional and metabolic diseases (E); circulatory system diseases (I); respiratory system diseases (J); violent and accidental deaths (V, W, X and Y).

¹⁸ Daily average temperature is defined as the average between the maximum and the minimum temperature of that day, following recommendations by the World Meteorological Organization (2011). In some robustness checks we also directly look at the effect of minimum and maximum temperatures.

¹⁹ In order to compute mean temperatures and precipitations at municipal level, we match the municipalities in Mexico with the closest land-based stations. To do so, we use the information on the longitude and latitude of municipalities from the National Geostatistical Framework (marco geoestadístico nacional) of the INEGI. We calculate the longitude and latitude of the centroid of each municipality (averaging the coordinates of all the locations that are part of a municipality), and then the distance between this centroid and all the land-based stations of the climatological data. Based on their distance to the centroid of each municipality, land-based stations are matched with municipalities. We consider a land-based station to be within a municipality if it is less than 20km from its centroid. For municipalities that are in very isolated zones, we have less than 5 active stations in the 20km radius. In this case, we match each municipality with the five closest stations within a maximum radius of 50km. Once we have identified the land-based stations relevant to a municipality, we compute the daily mean temperature and precipitation levels in a municipality by averaging the records of all the stations considered to be relevant to a given municipality. This computation is performed for each day from 1989 onwards.

the distribution is between 16 and 18°C, and 50% of days lie in the range 14-22°C. At the extremes of the distribution, the average Mexican is exposed to 5.6 days per year below 10°C (50°F) and 2.5 days per year above 32°C (90°F). Mexico's climate is much warmer than that of the US, which fewer days below 16°C and many more days above 26°C. ²⁰ The distribution is also more spread out in the US.

Figure 1: Population-weighted number of days per year falling within each temperature bin (in °C) for historical data and 3 climate change scenarios based on GFDL CM3 model output (2075-2099)



Notes: The figure shows the distribution of daily mean temperatures across 13 temperature-day bins. Each light grey bar represents the average number of days in each temperature category over 1998-2010, weighted by total population in a municipality-year. The climate change results depend on the scenario chosen. The dark grey bar is for the RCP4.5 scenario whereas the arrows represent the impact of shifting from the RCP2.6 scenario (low emissions) to the RCP8.5 scenario (high emissions).

In addition, Figure 1 also provides estimates of the distribution of cold and hot days under climate change. These estimates are derived from the output of the third version of the Coupled Physical Model of the Geophysical Fluid Dynamics Laboratory (GFDL CM3) of the National Oceanic and Atmospheric Administration (NOAA).²¹

²⁰ Deschenes and Greenstone (2011) provide a distribution of daily mean temperatures in the U.S. On average, temperatures are much lower: there are around 120 days with a mean temperature below 10°C and 1.3 days with temperatures greater than 90°F (32.2°C).

²¹ We extract monthly average temperature forecasts for Mexico and 2075-2099 based on three IPCC emissions scenarios (RCP2.6, RCP4.5 and RCP8.5). We obtain the model output from the Atlas Climático Digital de México. This Atlas provides climate model output for Mexico online and is monitored by Centro de Ciencias de la Atmósfera of the Universidad Nacional Autónoma de México (UNAM). We extrapolate the number of days falling within each temperature bin for each climate scenario and municipality. To do so, we calculate the difference between the monthly average temperature as observed in the historical data (1998-2010) and the forecasts of GFDL CM3: this gives estimates of monthly increases in average temperature due to climate change. Assuming that the distribution of daily temperatures around the monthly average temperature in one location and

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3.3 Socioeconomic data

In addition, information from the Mexican 2000 census of population and housing is used in this paper to estimate the income of the deceased. In particular, we extract socioeconomic information on income, educational attainment, social insurance coverage, profession, age, etc. Furthermore, we also use survey data from the Mexican Survey of Household Income and Expenditure (ENIGH: Encuesta Nacional de Ingreso y Gasto de Hogares) between 1998 and 2010 to assess heating and cooling equipment ownership.²²

These data sources are described in detail in Appendix A3. Succinctly, the 2000 Census shows larges differences in the average personal income between the poorest and the richest households. People in the first income quartile have an average personal income which is 18 times lower than people in the top quartile. This large inequality is a feature of the Mexican economy that we will use in the next sections to investigate differences in the weathermortality relationship across income groups. In addition, these large inequalities strongly reflect in an insufficient healthcare coverage of the very poor: more than 80% of the people in the 1st quartile of income have no social security.

4. The effect of temperatures on mortality in Mexico

4.1 Method

We now present the results obtained with the distributed lag model of section 2. In Appendix A4, we also present the results obtained with a simpler model with no lags, therefore considering only the contemporaneous relationship between temperature and mortality.

Figure 2 displays the cumulative impact of temperature on 31-day mortality for the whole population and all causes of death as estimated with our distributed lag model. We find the classical U-shaped relationship between temperatures and mortality identified in previous studies. However, looking at the two extremes of the temperature distribution observed in Mexico, low temperatures appear to lead to much more extra mortality than high temperatures. A day with an average temperature below 10°C kills 6 to 7 times more than a day with an average temperature above 32°C. Interestingly, we find statistically significant impacts of days above 32°C on mortality, suggesting that extremely hot days displace death

the population distribution across municipalities would remain constant under climate change, we can evaluate the proportion of days falling within each temperature bin under each climate change scenario. The result of this exercise is synthetically provided in Figure 1 for the three climate scenarios. We observe that the distribution of daily temperature shifts sharply to the right in all scenarios with much fewer cold days and many more hot days by the end of the century.

²² The survey waves used in this research are for 1998, 2000, 2004, 2005, 2006, 2008 and 2010.

by more than one month and not only a few days, a finding in contrast with that of Deschênes and Moretti (2009) for the US. Furthermore, we find statistically significant and strong impacts on mortality of all temperatures bins below 20°C. In fact, the contrast between a day below 10°C and a day between 10-12°C is not sharp. A day between 10-12°C increases mortality by around 0.5 deaths per 100,000 inhabitants when a day below 10°C increases mortality by 0.7 deaths per inhabitants. Likewise, a day between 16-18°C increases mortality by 0.1 deaths per 100,000 inhabitants such that a week of mildly cold days at 16-18°C will have the same mortality impact as one unusually cold day below 10°C. The comparison is interesting when we consider that there are around 51 days at 16-18°C per year in Mexico and only 5.6 days per year below 10°C. In Mexico, the effects of temperatures below 20°C and above 32°C are long-lasting effects that can reduce people's longevity.

1.20
1.00
0.80
0.60
0.40
0.20
-0.20
-0.40

Temperature bins (Celsius degrees)

Figure 2: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants

Notes: The graph shows the cumulative effect of a day with a temperature within each bin based (relative to the 24-26°C category) obtained from a dynamic model with 30 lags. The diamonds show the sum of the coefficients on these thirty lags in each category. Dashed lines correspond to the 95% confidence interval. The dependent variable is daily mortality rate at the municipality level. 312,140 groups and 30.1 observations per group on average. The regression controls for daily precipitation level and includes a range of municipality-by-year-by-month fixed effects.

These results are consistent with the dynamic effects of heat and cold days on mortality as reported previously, for example by Deschenes and Moretti (2009). Like these authors, we find evidence of strong harvesting for hot days whereas the impact of cold days accumulates after the event (see all details in Appendix A5). A cold day below 10°C has a statistically significant effect on mortality every day during the first week, and we find statistically significant effects at 5, 14 and even 21 and 22 days after the cold day. By contrast, we find that a hot day above 32°C has a strong and immediate effect on mortality but this effect is

statistically significant only for the first two days, after which the coefficients become systematically negative although not statistically significantly so.

Table 2 combines the results presented in Figure 2 with the distribution of hot and cold days in Mexico shown in Figure 1. Days under 10°C are responsible for the death of around 4,700 people each year (95% confidence interval is 4,177–5,296). This represents 0.8% of the annual number of deaths in Mexico. However, because mild temperatures between 10 and 20 degrees are much more frequent, the total amount of extra mortality associated with moderately low temperatures below 20°C is around 43,700 per year²³ (95% confidence interval is 34,600-52,800), or 7.7% of the annual number of deaths in Mexico, suggesting that the impact of mild temperatures on mortality is much stronger than the impact of unusually cold days.²⁴ At the other extreme of the spectrum, extremely hot days over 32°C trigger a comparably small amount of additional deaths (around 370 annually, 95% CI = 92–663).

These estimates can be compared with the results of recent studies conducted with US panel data (Barreca, 2012; Deschenes and Greenstone, 2011; and Deschenes and Moretti, 2009) and Indian data (Burgess *et al.*, 2014). This comparison is presented in detail in Appendix A6, but in short, our results are higher in magnitude than the ones obtained in the US²⁵, and are far smaller in magnitude than the ones found by Burgess et al. (2014) for extremely hot days in India. Burgess et al. (2014) find strong effects on rural populations and not on urban populations. This is because unusually hot weather during the growing season sharply depresses agricultural yield and the wages of agricultural laborers in rural areas, which in turns pushes mortality up. For Mexico, we find no statistically significant difference between

²³ This includes the impact of days below 10°C. The estimate for the impact of mild temperatures alone (10°C-20°C) is therefore slightly below 40,000 deaths.

We are comparing days with an average temperature between 10°C and 20°C with days with an average temperature between 24°C and 26°C. Minimal temperatures at night can be cold (e.g. 0-10°C) for mildly cold days, whereas maximal temperatures can be high in the reference bin (depending on intra-day variations).

²⁵ We use the exact same model as Deschenes and Moretti (2009) in Appendix B10. For cold, we find a marginal increase of deaths per 100,000 inhabitants by 0.60 for a day below 10°C in Mexico with this model. Deschenes and Moretti (2009) find a marginal increase of 0.23 deaths per 100,000 inhabitants for a day below -1°C in the US. For heat, we find a marginal increase of deaths per 100,000 inhabitants by 0.13 for a day above 32°C in Mexico. Deschenes and Moretti (2009) find no statistically significant effect of hot days. Contrary to Deschenes and Moretti (2009), Barreca (2012) and Deschenes and Greenstone (2011) do find a mortality impact of hot days using monthly and annual data.

²⁶ For cold, the coefficient of the model by Burgess et al. (2014) is not statistically significant at the lower limit of 10°C or below possibly due to the small frequency of such cold days in their data. However, they find that the log annual mortality rate increases by 0.004 for each day between 10-12°C and by 0.007 for each day between 14°C. In other words, an additional day between 10-14°C increases the annual mortality rate by about 0.4-0.7% in India. For heat, they find that an additional day above 32°C increases the annual mortality rate by about 0.5-1%. On the opposite, we find that a day above 32°C leads to a marginal increase in the annual death rate by less than 0.03% in Mexico.

rural and urban areas (see Appendix B9). This suggests that the impact of heat is smaller in countries in which people in rural areas do not entirely rely on subsistence agriculture.

Table 2: Estimated number of deaths per year by temperature bin

	Deaths per year				
Average daily temperature	Average	95% confidence interval			
<10°C	4,736*	(4177; 5296)			
10-12°C	8,474*	(7457; 9491)			
12-14°C	11,259*	(9553; 12965)			
14-16°C	8,910*	(6470; 11349)			
16-18°C	7,399*	(4468; 10331)			
18-20°C	2,968*	(510; 5426)			
20-22°C	1,036	(-834; 2907)			
22-24°C	- 1,018	(-2561; 526)			
24-26°C	-	-			
26-28°C	-1,332	(-2665; 0)			
28-30°C	-900	(-1944; 144)			
30-32°C	2	(-539; 544)			
>32°C	378*	(92; 663)			

Notes: * denotes statistically significant at 5%. The 95% confidence interval in brackets only takes into account the uncertainty of the impact of temperature bins on mortality. It does not take into account the variability of hot and cold days in Mexico from one year to the other.

4.3 Implications for climate change

We can use our model to simulate the impact that climate change may have on mortality in Mexico. This is done in detail in Appendix A7. Because the frequency of cold and mildly cold days is expected to decrease, the number of deaths imputable to temperatures reduces with the forecasted temperatures of GFDL CM3 as compared with the historical ones. With the RCP2.6 scenario (low GHG emissions), temperature-related mortality would be twice as small. The RCP8.5 scenario (high GHG emissions) corresponds to an 80% reduction in the estimated relationship between mortality and temperature. We show later that weather-related mortality affects mostly people in the first two quartiles of the income distribution, suggesting that the reduction in the exposure to cold weather associated by climate change could lead to a reduction in mortality inequality. However, this analysis comes with serious warnings: climate change could also affect mortality through increased frequency of natural catastrophes and not only through temperatures; our analysis at the daily level does not allow for acclimatization; and we could be underestimating the impact of increased heat waves if the effect of heat grows non-linearly beyond 32°C days. In addition, our model includes municipality-by-month-by-year fixed effects which control for income, technologies, and for the general health of the population, three factors that climate change could influence. All in all, our results simply suggest that extra winter mortality will reduce and that mortality will be more equally spread across seasons.

4.4 Impacts by gender, age and cause of death

We now look at the impact of temperatures on mortality by gender, age and cause of death. This exercise is useful to identify the type of people at risk during cold waves. We focus on the two extremes of the distribution. Table 3 displays only the 31-day cumulative impact of a day with average temperature below 10°C (corresponding to the left-hand side of Figure 2) whereas Table 4 displays the 31-day cumulative impact of a day with average temperature above 32°C (corresponding to the right-hand side of Figure 2).

Table 3: Impact of a day under 10 Celsius degree on cumulative mortality

	Cause of death						
Group	All causes	Respiratory system diseases	Circulatory system diseases	Endocrine, nutritional and metabolic diseases	Infectious diseases	Neoplasms	Accidents and violent deaths
Total	0.703***	0.19***	0.21***	0.131***	0.02***	0.004	0.025
Total	(0.042)	(0.014)	(0.02)	(0.015)	(0.006)	(0.012)	(0.016)
Men	0.774***	0.209***	0.248***	0.129***	0.025**	0.016	0.008
IVICII	(0.063)	(0.02)	(0.029)	(0.02)	(0.01)	(0.018)	(0.028)
Women	0.635***	0.171***	0.174***	0.133***	0.015**	-0.007	0.044***
Wolliell	(0.052)	(0.018)	(0.028)	(0.021)	(0.007)	(0.017)	(0.014)
Agad 0.4	0.774***	0.37***	0.003	0.045**	0.116***	0.004	0.139***
Aged 0-4	(0.103)	(0.052)	(0.009)	(0.022)	(0.029)	(0.007)	(0.034)
A110	0.022	-0.009	-0.001	-0.004	-0.002	-0.004	0.029
Aged 4-9	(0.029)	(0.012)	(0.003)	(0.006)	(0.006)	(0.006)	(0.018)
A and 10 10	0.011	0.012**	-0.004	0.013**	0.001	-0.013**	-0.014
Aged 10-19	(0.027)	(0.006)	(0.004)	(0.005)	(0.004)	(0.007)	(0.021)
A and 20 24	0.148***	0.009	0.014	0.007	-0.002	0.027***	0.067*
Aged 20-34	(0.047)	(0.008)	(0.01)	(0.008)	(0.008)	(0.01)	(0.037)
A 1 25 44	0.236***	-0.006	0.044*	0.075***	-0.012	0.04	0.005
Aged 35-44	(0.086)	(0.013)	(0.024)	(0.023)	(0.016)	(0.025)	(0.052)
A 1 45 54	0.301**	0.063**	0.025	0.173***	0.033	-0.066	-0.051
Aged 45-54	(0.133)	(0.028)	(0.053)	(0.05)	(0.021)	(0.046)	(0.057)
A 155.64	0.96***	0.206***	0.404***	0.361***	0.067**	-0.046	-0.063
Aged 55-64	(0.234)	(0.055)	(0.104)	(0.103)	(0.033)	(0.089)	(0.069)
1 1 65 74	1.576***	0.336***	0.633***	0.591***	0.113**	-0.072	-0.09
Aged 65-74	(0.362)	(0.085)	(0.161)	(0.16)	(0.053)	(0.138)	(0.105)
A 175.	16.436***	4.574***	6.103***	2.229***	0.183	0.26	0.116
Aged 75+	(1.028)	(0.403)	(0.594)	(0.38)	(0.132)	(0.281)	(0.167)

Notes: All the coefficients come from a different regression and correspond to the 31-day long run cumulative effect of a day below 10°C on mortality, for specific age groups or causes of death. The dependent variable is always the daily mortality rate in deaths per 100,000 inhabitants and all regressions include the daily precipitation level as control. Standard errors in brackets. *, ***, ***: statistically significant at 10%, 5% and 1%. 312,140 groups, 30.1 observations per group. Reference day is 24-26 Celsius degrees.

The 31-day effects of cold are much stronger for people over 75: the coefficient for cold-related mortality is 16 times higher than for the whole population. In addition, the very young (<5 years old) and senior people (>55) are also vulnerable to cold. Cold appears to have a particularly strong impact on metabolic, circulatory and respiratory diseases.²⁷ These three

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²⁷ There is also an impact of cold on infectious diseases. This could look surprising since diseases like malaria transmit at higher temperatures. Recall that Mexico has already undergone its epidemiological transition. More

causes of death are estimated to concentrate 70% of deaths due to unusual cold. Interestingly, cold days induce more accidental and violent deaths, but only among women. As for extreme heat, because of the small number of days above 32°C differences between age groups are not statistically significant. However, the model seems to indicate that days above 32°C primarily kill people between 35 and 54 years old and then again above 75 years old. Most heat-related deaths seem to be due to circulatory system diseases (affecting men) and accidental and violent deaths (affecting women).

Table 4: Impact of a day over 32 Celsius degree on cumulative mortality

	Cause of death							
Group	All causes	Respiratory system diseases	Circulatory system diseases	Endocrine, nutritional and metabolic diseases	Infectious diseases	Neoplasms	Accidents and violent deaths	
Total	0.127***	0.003	0.061**	0.006	-0.002	0.014	0.041*	
Total	(0.049)	(0.013)	(0.025)	(0.016)	(0.008)	(0.016)	(0.025)	
Men	0.146*	-0.016	0.099***	-0.004	0.003	0.034	0.042	
IVICII	(0.076)	(0.021)	(0.037)	(0.022)	(0.012)	(0.023)	(0.045)	
Women	0.108*	0.022	0.022	0.015	-0.006	-0.005	0.04**	
Women	(0.057)	(0.014)	(0.032)	(0.022)	(0.009)	(0.022)	(0.018)	
Aged 0-4	-0.022	-0.001	0.004	-0.014	-0.011	-0.016	0.034	
Aged 0-4	(0.085)	(0.025)	(0.009)	(0.018)	(0.023)	(0.014)	(0.032)	
Aged 4-9	-0.007	0.007	-0.0003	-0.009	-0.006	0.004	-0.028	
Aged 4-9	(0.031)	(0.006)	(0.003)	(0.005)	(0.007)	(0.007)	(0.026)	
Aged 10-19	0.057	0.003	-0.0003	0.003	-0.002	0.003	0.047	
Aged 10-19	(0.041)	(0.004)	(0.005)	(0.004)	(0.004)	(0.008)	(0.038)	
A and 20 24	0.028	0.016*	-0.002	0.022**	-0.023	0.001	0.033	
Aged 20-34	(0.073)	(0.008)	(0.011)	(0.01)	(0.014)	(0.012)	(0.066)	
1 25 44	0.129	0.009	0.047	-0.007	0.009	0.006	0.092	
Aged 35-44	(0.1)	(0.016)	(0.032)	(0.022)	(0.02)	(0.029)	(0.072)	
A 1 45 54	0.094	0.006	0.006	0.067	0.02	0.033	0.072	
Aged 45-54	(0.157)	(0.022)	(0.074)	(0.053)	(0.03)	(0.058)	(0.082)	
A 1 55 CA	0.003	0.018	0.04	-0.159	-0.057	0.196*	-0.018	
Aged 55-64	(0.249)	(0.046)	(0.121)	(0.105)	(0.043)	(0.108)	(0.102)	
A 1 (5 74	0.033	0.028	0.078	-0.257	-0.093	0.3*	-0.01	
Aged 65-74	(0.389)	(0.071)	(0.189)	(0.166)	(0.068)	(0.171)	(0.152)	
A 1 75 .	1.423	-0.334	1.156*	0.75*	0.175	-0.712*	-0.079	
Aged 75+	(1.152)	(0.369)	(0.689)	(0.399)	(0.153)	(0.379)	(0.21)	

Notes: All the coefficients come from a different regression and correspond to the 31-day long run cumulative effect of a day over 32°C on mortality, for specific age groups or causes of death. The dependent variable is always the daily mortality rate in deaths per 100,000 inhabitants and all regressions include the daily precipitation level as control. Standard errors in brackets. *, ***, ***: statistically significant at 10%, 5% and 1%. 312,140 groups, 30.1 observations per group. Reference day is 24-26 Celsius degrees.

Deschenes and Moretti (2009) similarly find (for the US) that people over 75 are much more vulnerable than the rest of the population. The causes of cold-related deaths seem very different though: in the US, two-thirds of cold-related deaths have a cardiovascular origin and around 20% are caused by respiratory diseases. Diabetes and infectious diseases respectively accounts for only about 3% and 2% of cold related deaths. Looking at the relevant estimates for Mexico, we find that cardiovascular diseases account for around a third of cold-related deaths only, followed by respiratory diseases (27%) and metabolic ones (17%, including

than 85% of deaths caused by infectious diseases are triggered by gastroenteritis and colitis; hepatitis B; sepsis and HIV.

mostly diabetes). Infectious diseases also account for a small share (3%) of cold-related deaths. 28

The output of the regressions by age groups can be used to compute annual deaths by age groups. These are reported in Table 5 for cold (<10°C), mildly cold (10-20°C) and hot (>32°C) days. The great majority of deaths correspond to people aged 75 and over, mostly during mildly cold day. Children under 5 constitute the second age category in terms of number of deaths. Individuals over 75 are much more vulnerable than children under 5, explaining the large gap in deaths. However, there were only around 700,000 people over 75 in Mexico in 2010, whereas the country comprised around 10 million children under 5 this same year. Results by age group are not statistically significant for days above 32°C.

Table 5: Death estimates by age group and temperature level

Age group	<10°C	10-20°C	>32°C
0-4	458*	2,706*	-6
5-9	14	-345	-2
10-19	13	754*	31
20-34	225*	-355	19
35-44	203*	356	49
45-54	186*	2,579	26
55-64	378*	2,423	1
65-74	371*	2,297	3
75+	2,536*	24,756*	97

Notes: These are estimates of the annual number of deaths due to cold (<10°C), mildly cold (10-20°C) and hot (>32°C) as compared to a day with average temperature of 24-26°C. Estimates take into account the frequency of cold, mildly cold and got days.

The estimates by age group are informative about the impact of cold on longevity. We calculate the annual total of years of life lost associated with outdoor temperature exposure for the Mexican population by using the life expectancy estimates of the Mexican life table of 2010 available from the Global Health Observatory data repository. Results are synthesized in Table 6. The number of years of life lost due to cold days under 10°C is 50% larger for children under 5 than for people aged 75. For days between 10°C and 20°C, we find that the number of years of life lost is roughly equivalent between the two groups. Deschenes and Moretti (2009) provide similar calculations of years of life lost for the US. In total, they find that people over 75 suffer from 106,405 years of life lost annually. However, the cumulative

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²⁸ It is interesting to express these differences between the US and Mexico in absolute terms, even though Deschenes and Moretti (2009) look at days below 30°F (-1.1°C): the amount of cold-related deaths caused by cardiovascular diseases is somehow comparable between the two countries (around 0.16 deaths per 100,000 inhabitants in the US vs. 0.25 in Mexico). On the opposite, respiratory diseases kill four times more on an unusually cold day in Mexico (0.05 deaths per 100,000 inhabitants in the US vs. 0.21 in Mexico) and metabolic diseases incomparably more (0.008 deaths per 100,000 inhabitants in the US for diabetes vs. 0.13 for all metabolic diseases, incl. primarily diabetes, in Mexico). Better protection for specific sets of diseases in the US therefore might therefore explain a significant share of the difference in the magnitude of the effects measured in both countries.

number of years of life lost in a year for children under 5 was only 5,410. The impact of cold weather on infant mortality is therefore much higher in the case of Mexico.

Table 6: Years of life lost estimates by age group and temperature level

Age group	<10°C	10-20°C	>32°C
0-4	35,872*	212,115*	-456
5-9	1,040	-25,734	-144
10-19	898	50,675*	2,073
20-34	12,443*	-19,639	1,050
35-44	8,767*	20,167	2,117
45-54	6,282*	87,130	863
55-64	9,461*	60,652	13
65-74	6,413*	48,452	60
75+	23,766*	232,044*	908

Notes: These are estimates of the total number of years of life lost for each age category. They are obtained from multiplying the estimated number of deaths of table 5 with the remaining life expectancy of each age group, as provided by the life table of 2010 for Mexico which is accessible from the Global Health Observatory data repository. An asterisk (*) denotes statistically significant results at 5%.

This result implies that priorities for policy makers in both countries should be different. US policies to reduce weather-related mortality may need to focus on the elderly, whereas emerging countries like Mexico may need to tackle both infant mortality and the vulnerability of the elderly to unusual weather.²⁹

4.5 Robustness

We have conducted a series of robustness checks to confirm all the aforementioned findings. Those are described in detail in the Appendix but we summarize them in this section.

First, we have considered specifications in which the definition of the temperature bins is different. We separately estimate the effect of daily minimum and daily maximum temperatures instead of using the daily average temperature (Appendix B1). This allows considering whether intra-day temperature variations has a strong impact on mortality. We find that minimum temperatures below 0°C are associated with an increase in mortality of 0.6 deaths per 100,000 inhabitants. We record no statistically significant effect on mortality for unusually high minimum temperatures above 25°C. We find an extra mortality impact of around 0.36 deaths per 100,000 inhabitants when daily maximum temperatures are below 15°C, and a small effect when they are unusually high (+0.18 deaths per 100,000 inhabitants for maximum temperatures above 40°C). The magnitude of these effects is similar to the one found when using daily averages in our base model. We also study the impact of

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²⁹ The calculation of the years of life lost assumes the same life expectancy for those who died from cold and for those who did not. This is an approximation with no consequence for the international comparison: the US figures have been obtained with the same assumption (Deschenes and Moretti, 2009). However, we may overestimate the total years of life lost.

consecutively hot or cold days on mortality and find no evidence that consecutively hot or cold days induce more mortality than if spread throughout the month (Appendix B2).

We then consider the role of acclimatization (Appendix B3). We assume that the temperaturemortality relationship might depend on the usual temperature faced by households in a given location. Heutel, Miller and Molitor (2017) find radically different results on the health impact of climate change in the US when taking into account differences in regional sensitiveness. Instead of using absolute temperature bins, we calculate deviations from the average temperature in each location to construct relative temperature bins with a 2°C window. The average temperature in each municipality is obtained by averaging all daily temperatures over 1997-2013. Then we rerun our distributed lag model with the newly constructed temperature bins. These include deviations between -10°C and +10°C with respect to the average temperature in each municipality. There are some small differences in magnitudes with the results obtained using absolute temperature bins, but the main messages on the large impact of mild cold and the comparatively small effect of heat remain unaffected. When accounting for the frequency of unusually cold and hot days, we find that days with mean temperature of more than 10°C below the municipality average are responsible for the death of around 2,700 people annually (95% CI is 2,200-3,200). Mild cold (deviations of between -2°C to -10°C) induce the death of 26,700 people (95% CI: 23,600-29,700). On the other hand, unusually hot days - above the average by 10°C or more - would cause around 350 deaths (95% CI: 100-600). We also find statistically significant effects for days with temperatures between 6°C and 10°C above the municipal average: these would be responsible for the death of around 1,500 people (95% CI: 900-2,200). In Appendix B4, we also run the model separately for four different climatic regions in Mexico, and find no statistically different health responses across regions: confidence intervals might be too large for us to effectively assess differences in acclimation, and long run adaptation to historical temperatures, with this method.

We also consider that precipitation levels might have delayed impacts on mortality and correlate with the temperature-mortality relationship. We find no statistically significant impact of lagged precipitations on mortality (Appendix B5). We also look at the confounding effect of humidity (Appendix B6). Results are not substantially modified, but we find that mortality due to heat is higher under dry climates.

We have also tested the sensitivity of the results to different sub-samples and to various alternative specifications. More precisely, we check for coefficient stability by splitting the

sample into two periods (1998-2003 and 2004-2010) (see Appendix B7). We find a decrease in the temperature-mortality relationship between the 1998-2003 and the 2004-2010 periods. We also estimate different effects of temperature on mortality for week days and weekends (Appendix B8), and on rural vs. urban populations (Appendix B9). We find that cold-related mortality is higher during weekends, consistent with people spending more time outdoors. We find no statistically significant difference between rural and urban areas. Among other things, this last result pinpoints that pollution is unlikely to be the main contributing factor explaining the temperature-mortality relationship that we observe.

Furthermore, we ensure that our results are fully comparable with the study by Deschenes and Moretti (2009) to draw comparisons between the US and Mexico. We reduce the number of temperature bins in our model to match their base specification (Appendix B10).³⁰ Using the model by Deschenes and Moretti (2009) gives results that are very similar to the base model.³¹

Finally, we use different structures for the fixed effects. In the base specification, we have used fully interacted, municipality-by-year-by-month fixed effects. This restrains the comparison of mortality effects to days within the same month of the year within a given municipality and disregards the fact that changes in temperature may affect seasonal patterns, and in turn mortality. Above all, we could underestimate the mortality impacts of direct exposure to temperature in very cold or very hot months by comparing very cold days with already cold days, and very hot days with already hot days within a month. To the contrary, we find that relaxing the controls for within-municipality seasonal patterns attenuates estimated impacts (see Appendix B12). This attenuation is likely to be due to an estimation bias. When we allow the comparison of mortality impacts to take place within a municipality and a given month, but across different years, results are similar to the base specification,

³⁰ Instead of using temperature bins, Deschenes and Moretti (2009) compute two sets of regressions, using as the independent variable either: a) a dummy variable which take the value of 1 on unusually cold days (average temperature $<20^{\circ}F$ or $<30^{\circ}F$, depending on specification); or b) a dummy variable which take the value of 1 on extremely hot days (average temperature $>80^{\circ}F$ or $>90^{\circ}F$, depending on specification). They therefore calculate the impact of unusually cold or hot days on mortality as compared to the impact of any other day in the year.

³¹ We also use the specification by Deschenes and Moretti (2009) to check the correctness of the window period of 30 days of our base specification. This is something that cannot be done with a high amount of bins as in our base specification because the calculations are far too computationally intensive with our very large dataset. In Appendix B11, we run a distributed lag model with 60 lags instead of 30 using the specification by Deschenes and Moretti (2009). The output confirms the relevance of a model with only 30 lags since results do not vary much: the model provides identical results for cold, but fails to predict any effect of heat due to amplified statistical variability.

suggesting that the base specification does not underestimate the impact of hot and cold days on mortality (Appendix B12).

5. Impacts by income group

5.1 Method

In this section, we seek to understand if mortality effects are stronger among the poor. We suspect that differences in living conditions and access to healthcare play a central role in the vulnerability to temperature variations, because poorer households will not have the same access to protection measures such as heating or air-conditioning or access to healthcare.

Income is not reported on death certificates.³² We use data from the 2000 Mexican census to estimate income levels at the moment of death in our mortality dataset.³³ To do so, we run a simple regression with data from the Mexican census where we predict income y_h of each individual h with a series of independent variables also present on death certificates. The regression used to predict income is:

$$\log(y_h) = \psi W_h + \omega_{i,r} + \omega_h$$

Where y_h is personal income for individual h in 2000 Mexican pesos, calculated as total household income divided by the square root of the number of people in the household (to account for economies of scale within households). Because personal income has a skewed distribution, we take the natural log to improve the fitness of the model and the accuracy of

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³³ We therefore only exploit cross-sectional information to predict income quartiles. A complementary possibility would have been to use the data from the 2010 census as well. However, the 2010 census do not report total income, but only income from work. This is a limitation and we have therefore preferred to use the 2000 data only.

³² We started by running our baseline distributed lag model by separating people according to their profession. These specifications are reported in Appendix C1. Contrary to income, the profession of the deceased is available on death certificates. We would suspect that the type of profession is correlated with differences in vulnerability to extreme weather events since professions are correlated with a series of relevant socioeconomic factors. First, the type of profession is the primary determinant of wages. Second, sectorial agreements usually stem from differentiated access to healthcare. Some categories benefit from specific healthcare regimes (e.g. the military and civil servants) whereas informal workers are excluded from any regime. Furthermore, exposure to heat/cold during the day will be different from workers in offices and workers that spend most of their time outdoors. However, running separate regressions by occupation do not show clear differences in terms of vulnerability to temperatures, except for workers in agriculture, fisheries and hunting who appear to suffer from cold temperatures. In fact, professional categories are an imperfect depiction of the diversity of living conditions among Mexicans. Whereas the revenues of the 1st quartile are more than 16 times lower than the ones of the 4th quartile of income, the difference between professions is much less contrasted. Looking at different professions is therefore far from fully characterizing the spectrum of inequalities in Mexico, since the high heterogeneity within each of these groups is not accounted for. In addition, some groups with high income levels, e.g. public servants and directors, as very small in terms of population and we may lack statistical power to determine differentiated vulnerability to temperatures. This is why the method that consists in predicting income levels based on the 2000 Census characterises much better differences in living conditions among the Mexican population and is reported as our base specification in this section.

predictions. W_h is a vector of independent variables that include gender, age, civil status, occupation, education level and inscription to public or private healthcare. It also includes a quadratic term for age and interaction terms between age (and age squared) and occupation to account for experience at work. $\omega_{i,r}$ is a fixed effect that takes into account that income may vary by municipality. However, within a given municipality, we also distinguish between people living in urban areas (e.g. the city centre) from those who live in rural areas. This is because rural localities that are remote from the city centre may still depend on the same municipality administratively. Thus, $\omega_{i,r}$ is a municipality i by-urban/rural area r fixed effect $(r \in \{rural, urban\})$. Finally, ω_h is an idiosyncratic error term and ψ is a vector of coefficients estimated from the regression 35 . The output of this estimation is presented in Appendix C2. The model includes about 8,756,000 observations. The regression results are consistent with economic theory (higher experience or education is correlated with higher income) and the model captures a large share of the variation in revenues (R2=0.44).

We use these regression results to predict the income level of deceased people, for whom we have the socio-demographic information reported on the death certificates (see Appendix A3 for the list of demographic variables available and Appendix A2 for an example of a death certificate). We can make income predictions by restricting the independent variables used in the income regression to those that are also present on the death certificates.

We then use predicted income values to construct income quartiles. Based on the 2000 Mexican census, we first compute the proportion of people in each municipality i whose predicted income would have fallen within income quartile κ . We then calculate the proportion of deaths in each municipality with a predicted income in each quartile κ and compute daily mortality rates by income quartile for each municipality i at time t.

Summary statistics on the daily mortality rate obtained for each income quartile and the proportion of deaths belonging to each quartile by specific cause of death are reported in Appendix C3. As expected, mortality is higher for the first two quartiles. We furthermore find that endocrine, nutritional and metabolic diseases, along with neoplasms, play a smaller role in the mortality of the 1st quartile, while respiratory systems diseases take a higher toll.³⁶

³⁴ We expect revenues to be lower in rural areas that administratively depend on a municipality as compared to city centres.

³⁵ The regression coefficients are weighted by population size in each municipality so as to be representative of the Mexican population. The 2000 Census includes about 10% of the Mexican population.

³⁶ In Appendix C3, we furthermore provide details on the relative prevalence of the most common diseases for endocrine, nutritional and metabolic diseases; circulatory system diseases; and respiratory diseases. These are of

The daily mortality rates by income quartile can be used to run separate distributed lag models for each income quartile.³⁷ The advantage of this approach is its high flexibility since the mortality impact of each temperature bin is estimated separately for each income quartile. The results however rely on predicted income values due to the absence of such information on death certificates. The main drawback is a loss of precision in the estimates due to measurement error in the dependent variable.³⁸

However, when running separate regressions by income quartiles, demographics are likely to play a significant role in explaining the differences in vulnerability across income groups. We have shown previously that the elderly are by far the most vulnerable. On the other side, the lowest quartiles of income are older on average because access to pensions is insufficient. In addition, poor families tend to have more children. The very young and the very old are overrepresented in the lowest quartiles and these people are more vulnerable to the weather independently of their living conditions.

In the section below, we also provide results by quartile while correcting for the differences in the pyramid of ages across quartiles.³⁹ We can then interpret the residual difference in vulnerability across the quartiles of income as originating principally from differences in

particular interest to this research because they correspond to the main causes of weather-related deaths. Deaths related to malnutrition, heart failures, cerebrovascular diseases and chronic lower respiratory diseases are more

common among the first quartile.

37 Even though we are using predicted mortality rates, standard errors using clustering are valid and there is no need for bootstrapping: this is because these predicted rates are used as the dependent variable. Using predicted instead of actual values therefore increases measurement error in the dependent variable and this directly affects the statistical power of our regressions.

³⁸ The method could also be inconsistent if some households systematically underreport their income levels. We mitigate this risk by excluding observations with doubtful declarations from the regression. After the census, the Mexican administration crosschecks individual declarations on employment status: in the survey, some individuals declare that they do not work whereas this is the case. We suspect these individuals to have underreported their income levels and exclude them from the regression used to predict income levels. They represent 2.6% of the original sample. In addition, to make sure that our findings are robust to a different measure of living conditions, less subject to misreporting than income, we use a poverty index. The Mexican Council of Population (CONAPO) defines a marginality index based on a set of questions asked to Mexican households in the 2000 census. The answers to this set of questions are less easy to manipulate by dishonest declarants and are less sensitive than income. We define and predict a poverty index for each deceased person in lines that are very similar to the CONAPO and construct quartiles based on this alternative metric. The results and exact method are in Appendix C4 and corroborate the findings obtained with predicted income levels.

³⁹ For each quartile, we compute the total number of people with a given age in the Mexican population. We then create weights based on the relative age composition of the quartiles of income. We take the 1st quartile as a reference. For example, if there are twice as many people aged 63 in the 1st quartile as in the 4th quartile, we create a weight equal to 2 for the people aged 63 in the 4th quartile. We use these weights to produce agecorrected death counts for the 2nd, 3rd and 4th quartiles. In our example, one death of a person aged 63 in the 4th quartile would count as 2 deaths. This is because there are twice as fewer people aged 63 in the 4th quartile as in the 1st quartile of income. Therefore, if the age composition of the 4th quartiles was similar to the one of the 1st quartile, we could have expected the death of two people instead of one. We run regressions by quartile based on this correction. Note that, for people aged 100 or more, we have created one single age category since we had very few observations by quartile.

living conditions (and not demographics). However, there is an important caveat to this approach since income is not randomly allocated across households. It follows that we observe a correlation, and no causal impact, of income on mortality, after controlling for age. The most accurate interpretation is that our results reflect a situation in equilibrium in which both ill health determines low income, and low income determines ill health.

5.2 Results

We now run separate regressions of Equation 1 for each income quartile. We evaluate the impact of extreme temperatures after up to 31 days on each quartile, using distributed lag models. These results are reported in Figure 3 and do not control for the difference in age distribution across quartiles.

Results show a stronger vulnerability of the first two quartiles compared to the last two, and statistically insignificant impacts of temperatures at all temperature levels for the last quartile. In particular, we find a strong difference in vulnerability to cold between the first and the last quartiles. Vulnerability to unusually cold temperatures is more than 4 times higher for people in the first quartile as compared to people in the fourth quartile and the difference is statistically significant at 1% (see Table 7). In contrast, we do not find any statistically significant difference in the impact of unusually hot days on mortality across income quartiles. This is likely to be caused by some lack of statistical power since only a minority of weather-related death are associated with excessive heat.

To deepen our understanding of the correlation between income and weather-related deaths, we have run the quartile-specific econometric models for separate causes of death. Table 7 below summarises these results focusing on days below 10°C. Interestingly, we find impacts across all quartiles from endocrine, nutritional and metabolic diseases, circulatory system diseases and respiratory system diseases. The magnitude of the impact of cold days remains relatively constant for endocrine, nutritional and metabolic diseases, suggesting little margin for improvement. This can be explained since diabetes has become prevalent across all income groups in Mexico. In contrast, the magnitude of the effect diminishes sharply between the 1st and the 4th quartile for circulatory and respiratory system diseases.

⁴⁰ We have also tried to run the model for different age groups. Unfortunately, running the model by age group significantly reduces model efficiency and results are inconclusive. The reader may notice that efficiency is not always very high with the breakdowns by death causes and income groups. For example, we find a higher impact, of circulatory system diseases on the 2nd quartile as compared to the 1st quartile. This impact is likely to be driven by relatively low efficiency: the two point estimates are not statistically different from one another.

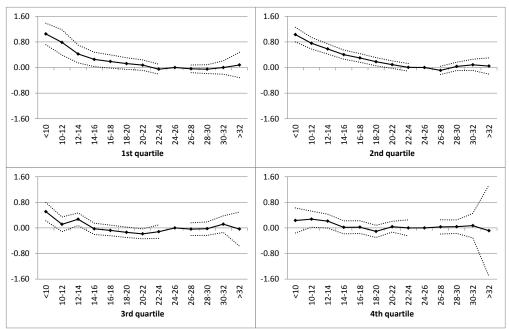


Figure 3: Impact of temperature bins on cumulative 31-day mortality, by income quartile

Note: The results for each quartile are taken from separate regressions. The dependent variable is the mortality per 100,000 inhabitants belonging to the quartile. The y-axis is mortality per 100,000 inhabitants and the x-axis corresponds to the cumulative impact after 31 days for each of the 2°C temperature bins in the regressions. The reference bin is 24-26°C. Onthe-day precipitations are used as controls, along with municipality-by-month-by-year fixed effects. The dashed lines represent the 95% confidence interval for each estimated set of coefficients.

In addition, Figure 3 clearly depicts an impact of cold temperatures at mild levels but only for the first and second quartiles. In Table 8, we report the magnitude of the impacts of both mildly and unusually cold weather (i.e. all temperature bins below 20°C) for each income quartile. For the 1st and 2nd quartiles of income, we find a statistically significant impact of cold below 20°C on mortality whereas no such impact is found for the third and the fourth quartiles. In other words, all of the death toll triggered by mildly cold days is borne by the population in the bottom half of the income distribution. Results by cause of death tend to corroborate that low-income households are more vulnerable to cardiovascular and respiratory diseases. 41,42

The policy implications of Table 8 are substantial. They suggest that the poor are not only much more vulnerable to unusually cold temperatures, but they are also vulnerable to temperatures to which richer households are not. This definitely puts poor households at risk since mildly cold days are relatively frequent.

⁴² Since we find no statistically different results between rural and urban areas in the core model, we can discard the eventuality that pollution, and not low temperatures, are strongly confounding the effect of cold weather on respiratory diseases.

⁴¹ We also find that richer households are statistically more vulnerable to neoplasms than poorer households. However, this information is derived from coefficients for neoplasms that are themselves not statistically different from 0 at the 5% level of significance.

<u>Table 7: Impact by income quartile and cause of death of a cold day below 10°C on</u> cumulative 31-day mortality

Cause of death	1 st quartile	2 nd quartile	3 rd quartile	4 th quartile	1 st vs. 4 th	First two versus last two
All causes	1.05***	1.03***	0.51***	0.23	+0.82***	+0.67***
	(0.17)	(0.12)	(0.15)	(0.2)	(0.26)	(0.16)
Respiratory system	0.31***	0.3***	0.08**	0.09***	+0.22***	+0.22***
diseases	(0.05)	(0.03)	(0.04)	(0.03)	(0.06)	(0.04)
Circulatory system	0.31***	0.39***	0.16***	0.08*	+0.23***	+0.23***
diseases	(0.08)	(0.06)	(0.06)	(0.05)	(0.1)	(0.06)
Endocrine, nutritional and	0.21***	0.2***	0.17***	0.19***	+0.03	+0.02
metabolic diseases	(0.07)	(0.05)	(0.05)	(0.06)	(0.09)	(0.05)
Infectious diseases	0.05	0.01	0.004	0.01	+0.03	+0.02
	(0.03)	(0.02)	(0.02)	(0.01)	(0.03)	(0.02)
Neoplasms	-0.01	-0.04	0.08*	0.06	-0.08	-0.10**
	(0.06)	(0.04)	(0.04)	(0.05)	(0.08)	(2.08)
Accidents and violent	0.04	0.02	-0.04	-0.2	+0.24	+0.15
deaths	(0.07)	(0.05)	(0.09)	(0.17)	(0.18)	(0.11)

Notes: All the coefficients come from a different regression and correspond to the 31-day long run cumulative effect of a day below 10°C on mortality, for specific quartiles and causes of death. The dependent variable is always the daily mortality rate in deaths per 100,000 inhabitants and all regressions include the daily precipitation level as control. Standard errors in brackets. *, ***, ***: statistically significant at 10%, 5% and 1%. 312,140 groups, 30.1 observations per group. Reference day is 24-26 Celsius degrees.

Table 8: Estimated deaths per year for temperatures below 20°C by income quartile

	Excess number of deaths per year						
Temperature level	<10°C	10-20°C	Total <20°C				
1 ot groutile	1,813***	11,909***	13,722***				
1st quartile	(1,232; 2,393)	(4,920; 18,899)	(6442; 21001)				
2 - 1	1,437***	21,055***	22,492***				
2nd quartile	(1118; 1757)	(14,024; 28,087)	(15,284; 29,701)				
2.1	731***	-1,306	-575				
3rd quartile	(321; 1142)	(-10,858; 8,246)	(-10,331; 9,182)				
441	404	1,936	2,340				
4th quartile	(-273; 1081)	(-8,753; 12,626)	(-8,676; 13,357)				
Entire manufation	4,385***	33,595***	37,980***				
Entire population	(3353; 5418)	(16,165; 51,025)	(20,050; 55,911)				

Notes: All estimated coefficients are in reference to a day with an average temperature of 24-26°C. Estimates are made with different distributions for cold days corresponding to population-weighted quartile-specific averages (they can be slightly different from the ones derived from Figure 2), for a total population of 114 million inhabitants equally spread across quartiles. Lower and upper bound of 95% confidence interval in brackets and do not account for the uncertainty in the variability of the weather. One, two and three stars respectively mean statistically significant at 10%, 5% and 1%.

Age-corrected results for cold days below 10°C are provided in Table 9. Point estimates show that the first two quartiles of income have comparable vulnerability levels. However, these income groups are 35% more vulnerable to unusual cold than the last two quartiles. This difference is statistically different at 5% (t-statistic of 2.04). Therefore, a sizeable difference in vulnerability levels correlates with differences in living conditions and social protection. Results by cause of death also corroborate that low-income households are more vulnerable to

cardiovascular and respiratory diseases. ⁴³ In addition, age-corrected regressions also corroborate that poor living conditions seem to make households vulnerable to mild cold. Full results for all temperature bins reported in Appendix C5 show statistically significant results for temperatures up to "below 20°C" for the 1st two quartiles, whereas results stop being statistically significant for temperatures above 14°C for the top two quartiles of income.

<u>Table 9: Impact by income quartile and cause of death of a cold day below 10°C on cumulative 31-day mortality correcting for differences in the pyramid of ages across quartiles are considered.</u>

Cause of death	1 st quartile	2 nd quartile	3 rd quartile	4 th quartile	1 st vs. 4 th	First two versus last two
All causes	1.05***	1.12***	0.75***	0.61***	+0.43	+0.36**
	(0.17)	(0.13)	(0.17)	(0.23)	(0.28)	(0.18)
Respiratory system	0.31***	0.34***	0.10**	0.21***	+0.10	+0.17***
diseases	(0.05)	(0.04)	(0.05)	(0.06)	(0.08)	(0.05)
Circulatory system	0.31***	0.43***	0.26***	0.18*	+0.13	+0.15*
diseases	(0.08)	(0.06)	(0.08)	(0.10)	(0.13)	(0.08)
Endocrine, nutritional and	0.21***	0.22***	0.22***	0.19**	+0.02	+0.01
metabolic diseases	(0.07)	(0.05)	(0.06)	(0.10)	(0.11)	(0.07)
Infectious diseases	0.05	0.01	0.01	-0.002	0.05	+0.02
	(0.03)	(0.02)	(0.02	(0.02)	(0.03)	(0.02)
Neoplasms	-0.01	-0.05	0.10*	0.12	-0.13	-0.14**
	(0.06)	(0.04)	(0.06)	(0.08)	(0.09)	(0.06)
Accidents and violent	0.04	0.02	-0.01	-0.11	+0.15	+0.09
deaths	(0.07)	(0.06)	(0.09)	(0.14)	(0.15)	(0.09)

Notes: All the coefficients come from a different regression and correspond to the 31-day long run cumulative effect of a day below 10°C on mortality, for specific quartiles and causes of death. The dependent variable is always the daily mortality rate in deaths per 100,000 inhabitants and all regressions include the daily precipitation level as control. Death counts and population levels have been weighted such that the pyramid of ages are comparable across age groups, taking the 1st quartile of income as a reference. Standard errors in brackets. *, **, ***: statistically significant at 10%, 5% and 1%. Reference day is 24-26 Celsius degrees.

These results by income groups are not surprising when we consider that low-income families have improper access to housing, drinkable water and health insurance (as reported on Census data – see Appendix A3). Specific protection against cold is also insufficient. Data from the Mexican survey of household income and expenditure shows that about 1% of households in the first income quartile own a heater, versus 7.9% for Mexicans in the fourth quartile (see Appendix C6 – the geographical distribution of heating and cooling appliance ownership is also provided). Similar contrasts are found when looking at air conditioning.

⁴³ We again also find that richer households are statistically more vulnerable to neoplasms than poorer households. However, this information is derived from coefficients for neoplasms that are not statistically significant at 5%.

6. Weather-related mortality and universal healthcare

During our study period, Mexico implemented a nationwide policy – the *Seguro Popular* – to increase access to healthcare for low-income households. Considering that developing countries may be financially constrained to protect their citizens from cold, targeted health programmes may offer the possibility to restrict the population of recipients to vulnerable groups. They can also restrict the range of covered diseases to those that are known to arise because of cold weather. Below, we provide evidence that the *Seguro Popular* has reduced mortality in general and weather-related mortality in particular. Our econometric setting enables us to correct for selection bias, which arises from the fact that the weakest people are also the ones most likely to contract a health plan.

The *Seguro Popular* was launched as a pilot exercise (2001-2003) to increase universal healthcare. Access to the *Seguro Popular* was open to all. In practice, it focused on people who were not eligible to employment-based health insurance, i.e. low-income households working in the informal sector. Enrolment was free in most cases even though a fee could be due if the family earned enough income. The fee then grew with income. By 2004, the Mexican government decided to progressively extend the programme to the entire population, municipality after municipality. In 2004, the Mexican government also promoted the *Fondo de Protección contra Gastos Catastróficos*, which provides financial support to families affected by a series of chronic, long-term diseases, in particular cancer and HIV. 46 Both

⁴⁴ Traditionally, low-income families working in the informal sector did not have access to healthcare insurance, and the country suffers from a chronic underfinancing of the public hospitals with free attendance. Mexico is the OECD country with the lowest budget dedicated to health: in 2015, current expenditure per capita in purchasing power parity was \$ 1,052, compared to \$ 3,814 on average in other OECD countries, and \$ 9,451 in the US (OECD Health Statistics 2016).

⁴⁵ This section contributes to the literature aiming at assessing the effectiveness of healthcare in reducing the mortality effect of unusual temperatures. Two recent studies have attempted to relate healthcare provision to weather-related mortality. Barreca et al. (2016) uses the number of doctors as a measure for healthcare provision to look at the impact of healthcare on mortality over the last century. They do not find any statistically significant impact. However, this could be because counting the number of doctors does not take into account the significant progress in medicine that occurred over the 20th century. Heutel, Miller and Molitor (2017) look at the impact of temperature on hospitalizations in the US. They find that temperatures are positively correlated with hospitalizations. This pattern differs from the U-shaped association that they find between temperature and mortality. However, they do not analyse the impact of healthcare provision (e.g. access to hospitals) on mortality.

⁴⁶ Furthermore, additional protection has been provided to children under 5 born after Dec. 1st 2006 with the implementation of a policy called the *Seguro Médico para una Nueva Generación*. We are not including this policy in the analysis since it has covered only a small minority of young children by 2007. For a small list of diseases covered under the *Seguro Popular* or the *Fondo de Protección contra Gastos Catastróficos*, conditions of age had to be filled to receive treatment. For some others, only diagnosis and prevention measures are covered (e.g. some cancers). We consider that a disease is covered by the scheme independently of these age conditions, and also as soon as its diagnosis (or some preventive action) is part of the scheme. This allows us to take into account spillover effects for transmissible diseases, or the positive effect of early detection on mortality in the case of severe chronic conditions.

programmes are still ongoing today. The extension of the *Seguro Popular* to the whole Mexican population depended on the enrolment of the existing medical infrastructure into the scheme or on the construction of new infrastructure. INEGI discloses the number of people that received medical attention under the *Seguro Popular* by municipality and year.⁴⁷ At its start in 2004, the *Seguro Popular* provided around 315,000 external consultations. This figure radically increased to 11 million in 2005, 20 million in 2006, 29 million in 2007, 38 million in 2008, 48 million in 2009 and 61 million in 2010.

A particularity of the *Seguro Popular* is that health coverage is restricted to a reduced list of priority diseases. It mostly includes preventive health actions (e.g. vaccines), ambulatory medicine (e.g. measles, tuberculosis), reproductive health, a selection of emergencies (in particular caused by hypertension and diabetes) and surgeries (e.g. appendectomy, treatment of fractures). The list of diseases covered by the *Seguro Popular* and the *Fondo de Protección contra Gastos Catastróficos* is updated every year. We have compiled this information for 2004-2010 using the catalogues published by the Mexican government, and recoded the information to clearly identify which diseases were covered by the scheme, using the ICD-10 nomenclature of diseases. According to our recompilation, in 2004, the *Seguro Popular* covered 734 ICD-10 codes, e.g. "A010 – Typhoid Fever". In 2010, it covered 1923 ICD-10 codes. For example, the 2010 nomenclature also included code "A02 – Salmonella infections". Appendix A8 displays the list of diseases covered by the *Seguro Popular* and the *Fondo de Protección contra Gastos Catastróficos* in 2010.

In the remaining of this section, we first show that the two schemes increased medical assistance before death, and that this subsequently impacted mortality. Then, we assess the extent to which a change in the coverage of the *Seguro Popular* correlates with a reduction in weather-related mortality in a reduced-form setting.

6.1 Method

We start by looking at the impact of universal health insurance on medical assistance and, subsequently, on mortality. To do so, we use two-stage least squares:⁴⁸

⁴⁷ The implementation of the *Fondo de Protección contra Gastos Catastróficos* was done through specialised institutions that needed to receive accreditation. The rollout of the programme was therefore very similar to the one of the *Seguro Popular*. We make the simplifying assumption that the municipalities who benefitted from the *Seguro Popular* were also the ones that benefitted from the *Fondo de Protección contra Gastos Catastróficos* since we unfortunately do not have this exact information.

⁴⁸ A reduced-form equation for this IV setting would consist in directly estimating the effect of $E_{i,d,m,t}$ on $Y_{i,d,m,t}$, i.e. the effect of eligibility to the schemes on mortality. It the present case, it makes sense to use 2SLS and properly identify the channel through which the *Seguro Popular* may affect mortality, i.e. health provision.

First stage: $M_{i,d,m,t} = m_1 E_{i,d,m,t} + m_2 C_{i,d,m,t} + F_{i,m,t} + v_{i,d,m,t}$

Second stage: $Y_{i,d,m,t} = m_3 M_{i,d,m,t} + m_4 C_{i,d,m,t} + \mu_{i,m,t} + \varepsilon_{i,d,m,t}$

For the first stage, we focus on the impact of universal health insurance on medical assistance. The dependent variable is medical assistance before death and is denoted $M_{i,d,m,t}$. It is derived from Mexican death certificates. These certificates report if an individual received medical assistance or not before dying. Therefore, $M_{i,d,m,t}$ is the share of people that received medical assistance before death among deceased people in municipality i and day d. To look at the effect of universal health insurance on medical assistance, we compute the share of people in municipality i and day d who were eligible to the Seguro Popular and the Fondo de Protección contra Gastos Catastróficos when they died. We refer to this share as $E_{i,d,m,t}$ in the set of equations above. Denoting G_t the municipalities covered by the schemes at time t, θ_t the set of diseases covered by the schemes and C_n the cause of death of individual n, then individual n in municipality i is entitled to the schemes if $(C_n \in \Theta_t)$ and $(i_n \in G_t)$. The share of deceased people eligible to the schemes is the sum of eligible deceased people divided by the total number of deaths in municipality i and day d. Importantly, we also control for the share of people that died from diseases covered by the schemes $(C_n \in \Theta_t)$, independently of the availability of the Seguro Popular in the municipality. We call this control variable the share of potentially covered cases and it is denoted $C_{i,d,m,t}$ in the equations above. It aims to serve as a benchmark against which the additional effect of the Seguro Popular on medical assistance is evaluated. Including this control variable is necessary because the share of potentially covered cases could be correlated with medical assistance, independently of the availability of the Seguro Popular in the municipality.

In this first stage regression, the variables for medical assistance $(M_{i,d,m,t})$, eligibility to the schemes $(E_{i,d,m,t})$ and the share of potentially covered cases $(C_{i,d,m,t})$ vary for each municipality and day of the year. This is important because we want to focus our analysis on short-run impacts. The regression therefore includes municipality by month by year fixed effects $(F_{i,m,t})$ and an error term that varies from one day to the other $(v_{i,d,m,t})$. m_1 and m_2 are coefficients to be estimated. Our identification strategy exploits the fact that we have disaggregated mortality data at daily level and we can thus follow the proportion of people who died from diseases covered by the schemes day after day, both in areas where the

⁴⁹ Because we estimate medical assistance, only days when at least one person dies are included in the regression. This creates measurement error on the degree of medical assistance and therefore an extra endogeneity issue that is in fact resolved thanks to the 2SLS setting.

schemes were active and in areas where they were not implemented yet. Our identification strategy consists in considering that shocks on death causes from one day to the other must have had differentiated impacts in covered vs. non-covered areas and across covered vs. non-covered diseases, since specific causes of death grant a right to medical assistance only for the people that live in the municipalities covered by the schemes.⁵⁰

The second stage of our 2SLS strategy consists in evaluating the impact of medical assistance on mortality (denoted $Y_{i,d,m,t}$). Our instrument for medical assistance is the share of people eligible to the schemes on day d, i.e. the share of people such that $(C_n \in \Theta_t)$ and $(i_n \in G_t)$, and we control for the share of potentially covered cases $(C_n \in \Theta_t)$. The coefficients to be estimated are m_3 and m_4 . $\mu_{i,m,t}$ is a municipality by month by year fixed effect and $\varepsilon_{i,d,m,t}$ is the error term.

For this IV regression to be valid, there must be no correlation between our instrument (the share of people who died in municipality i on day d and were eligible to the schemes) and the term of error, i.e. the daily shocks on mortality rates.

Recall that the econometric model includes municipality by-month by-year fixed effects: the only variations that remain to be explained are daily shocks on mortality levels and causes of death (in the municipalities that implement or do not implement the *Seguro Popular*). Changes in death causes on day d are likely to be reflective of the vulnerability of the people that are dying on this specific day. This may create correlation between our instrument and the term of error. We therefore need to control for the changes in vulnerability caused by changes in death causes. This is why we include $C_{i,d,m,t}$, the share of potentially covered cases in municipality i and day d, as a control variable. This variable accounts for the daily correlation between shocks on mortality rates and shocks on death causes. Controlling for this variable provides a very strong basis for the exclusion criterion to hold: it identifies the correlation between shocks on the share of potentially covered cases and mortality rates in all municipalities. In this context, shifts in the provision of medical assistance become the only

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⁵⁰ The rollout of the *Seguro Popular* was not conducted as a randomised controlled trial but depended on the availability of resources at municipality level to start providing public health assistance to uninsured Mexicans. Therefore, we cannot consider that the policy in itself had a strictly exogenous impact on medical assistance and mortality rates. The reason is that policy-makers could have decided to implement the *Seguro Popular* because they observed low medical assistance or high mortality rates. All the unobservables affecting the expected set of death causes for month m, year t and municipality i (including the endogeneity in the implementation of the *Seguro Popular*) are dealt with by the inclusion of the municipality by month by year fixed effects ($F_{i,m,t}$) in the regressions. For each panel of our dataset, with a given municipality-month, either the Seguro Popular is implemented (e.g. Tijuana, May 2010) or not yet available (e.g. Tijuana, January 2004).

way through which the instrument (day-to-day changes in the share of deceased people eligible to the *Seguro Popular*) could impact mortality.

Our process implicitly relies on some comparability between the panels that have and have not implemented the Seguro Popular since we consider that the correlation between shocks on vulnerability and death causes is the same between the two groups, after controlling for medical assistance.⁵¹ The acceptability of this assumption of comparability can be assessed. To do so, we investigate if there is any effect of adding additional controls allegedly correlated with vulnerability. We compute, for each municipality i and day d, a large set of sociodemographic variables about the identity of the deceased in municipality i and day d.⁵² We interact these variables with the presence of the Seguro Popular in municipality i on year t. We know for sure that all these variables (e.g. age) are strongly correlated with vulnerability. On the other side, if our assumption does not hold, then our instrument is also correlated with vulnerability. Therefore, the inclusion of such controls could significantly alter the estimated coefficients: this would be a textbook case of omitted variable bias where an additional control is correlated with both the dependent and independent variables. As shown in Table 10, including or not these sociodemographic controls have little impact on the results of our IV approach. Such stability in the estimates strongly supports the view that the exclusion restriction is respected.⁵³ In addition to this test, we also run placebo tests in

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⁵¹ The main problem with this assumption is that death causes could be manipulated by doctors so that more low-income families receive free treatment in the areas that implement the *Seguro Popular*. A shock on death causes may therefore not be fully comparable between the municipalities that implement and do not implement the *Seguro Popular*. In practice, several factors may prevent this from happening. First, the *Seguro Popular* only offers access to a certain type of interventions and drugs: manipulations may only occur in the cases for which the treatments of covered diseases can be applied to non-covered diseases. Second, this manipulation constitutes a fraud and doctors have to report on both diagnosis and treatment. Above all, diagnoses on death certificates are filled after death, i.e. when providing access to medical treatment is no longer useful. In our dataset, only 27.5% of death certificates are filled by the treating doctor. The remaining is either filled by another doctor (31%) or a forensic doctor (39%).

⁵² These demographic controls are described later on in the notes below table 10.

By doing so, we implicitly assume that this control is an exogenous variable. Else, 2SLS would be inconsistent. In particular, we exclude the possibility that there is reverse causality, i.e. that mortality rates influence our measure of the prevalence of potentially covered cases on day d or the following days. The absence of reverse causality is easily justifiable for non-transmissible diseases: breast cancer prevalence impacts mortality, but mortality caused by breast cancer does not increase the prevalence of breast cancer in the environment. For transmissible diseases, this assumption could be violated if acute outbursts of viruses start killing people. This phenomenon is unlikely to violate our assumption in the case of Mexico because the country has undergone its epidemiological transition: particularly virulent viruses only constitute a small fraction of deaths. Likewise, this assumption can somehow be tested gauged. We can add climatic variables in the IV regression to account for the prevalence of winter viruses. This constitutes the main channel through which today's mortality could impact death causes in a short time frame. Adding these controls has no effect on the coefficient estimated for the share of deaths caused by diseases covered by the scheme. This corroborates our identification strategy: if our strategy was not valid, then the temperature controls would capture some of the correlation between the share of people

Appendix A9 to make sure that our results are not driven by substantial differences in vulnerability between municipalities.

6.2 Results

The results of our estimation strategy are provided in Table 10. The first column reports OLS results and shows a slightly negative correlation between mortality and medical assistance: providing medical assistance to all would reduce daily mortality by 0.12 deaths per 100,000 inhabitants (9% of the average daily mortality rate). The second specification is the baseline output of our 2SLS strategy. The instrument is very strong, with a Wald F-statistic of 75.4. The third column includes sociodemographic controls likely to be correlated to individual vulnerability. The fourth column includes climatic controls, known to influence mortality levels and the prevalence of winter viruses. The coefficients for the variables of interest are stable in the specifications of the third and fourth columns, which corroborates the validity of our IV strategy. Finally, a fifth column corresponds to the reduced form approach of our IV strategy: we directly estimate the correlation between eligibility to the schemes and mortality. The reduced form approach estimates a similar marginal impact of the eligibility to the schemes on mortality.⁵⁴ In addition, we also ran another regression where we use two instruments, the share of people eligible to the schemes on day d and its first lag (on day d-1), in order to run an over-identification test. We find a p-value of 0.66 for this test, providing some reassurance that our instrument is valid.

Our first stage results (column 2, upper panel) show that a 1 percentage point increase in the share of cases eligible to the *Seguro Popular* and the *Fondo de Protección contra Gastos Catastróficos* yields a 0.03 percentage point increase in medical assistance before death. The second stage results show an impact of medical assistance on mortality which is 6 times greater than with the OLS regression: full medical coverage (in terms of medical assistance before death) would reduce daily mortality rate by 0.71 deaths per 100,000 inhabitants. ⁵⁵ In 2010, 592,000 deaths were registered, out of which 29% fell within the criteria of the *Seguro Popular* and the *Fondo de Protección contra Gastos Catastróficos*. It follows that the schemes increased medical assistance by 0.87 percentage points (29 x 0.03) and reduced the daily average mortality rate by 0.0062 deaths per 100,000 inhabitants (0.71/100 x 0.87).

who died from diseases covered by the schemes and the term of error, and we would get different estimates with and without these controls.

⁵⁴ i.e. while comparing column (2) and (5): 0.03 x -0.71 \approx -0.02.

⁵⁵ The OLS and IV settings are statistically different from one another at 1%.

Therefore, the schemes may have saved around 3,000 lives in 2010.⁵⁶ This figure is an underestimate of the total impact of the schemes on mortality since it only considers the impact that the schemes had on mortality caused by an improved level of medical assistance before death. All longer-term impacts (e.g. the impact of vaccinations in reducing the probability of contracting a disease) are not included in this figure.

Table 10: The impact of medical assistance before death on reducing mortality

	(1)	(2)	(3)	(4)	(5)
Independent variables	OLS	2SLS	2SLS	2SLS	OLS
First stage results: Medical assistance (%)					
Potentially covered cases (%)	-	0.08***	0.05***	0.08***	
		(0.004)	(0.004)	(0.004)	
Eligible to the schemes (%)	-	0.03***	0.03***	0.03***	
Temperature <10° C (cumulative effect)	_	(0.004)	(0.004)	(0.004) -0.001	
- Constitution of Constitution				(0.001)	
Stock-Yogo weak identification test	-	59.9	80.3	51.9	
(Kleibergen-Paap rk Wald F statistic: passed)					
Second stage results: Daily mortality (per 100,000)					
Medical assistance (%)	-0.12***	-0.71**	-0.63**	-0.69**	
	(0.005)	(0.33)	(0.29)	(0.33)	
Potentially covered cases (%)	-0.03***	0.03	0.02	0.03	-0.02**
T 100 G (1 d 65)	(0.004)	(0.03)	(0.02)	(0.03)	(0.009)
Temperature <10° C (cumulative effect)	-	-	-	0.65***	-
				(0.03)	0.02**
Eligible to the schemes (%)	_	-	_	_	-0.02** (0.01)
					` ′
Sociodemographic controls (†)	No	No	Yes	No	No

Notes: The table reports OLS and IV results obtained from fixed effect regressions with by municipality by month by year fixed effects. The dependent variable is daily mortality rate at the municipality level. The regression is weighted by the population in each municipality. One, two and three stars respectively mean statistically significant at 10%, 5% and 1%. Standard errors are clustered at municipality level. †: The sociodemographic controls included in the specification of column 3 are: the share of women, the share of deceased people (for municipality i and day d) falling into 9 age groups, 22 occupation categories (as provided in the Mexican death records), 7 educational attainment categories and 5 types of social security affiliation. We also control for the share of deceased that were residing in rural areas within municipality i. We furthermore allow each of these controls to have a different impact on municipalities that implement and do not implement the schemes with the corresponding interaction parameters.

Even then, 3,000 saved lives are far from being insignificant. Medical assistance before death is already high in Mexico (at around 84% in the entire sample) and the effects concentrate on the 50-60% of Mexicans with no social security. If all diseases and municipalities were covered, the total impact of the schemes would be a mortality reduction of -0.02 deaths per 100,000 inhabitants, i.e. an average mortality reduction by 1.7%. This reduction in the national mortality rate would exclusively come from the target population of the *Seguro Popular*: the half of the population that does not have social security.

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 $^{^{56}}$ For a population of around 120 million Mexicans, this is: $0.0062 / 100,000 \times 365 \times 120,000,000$.

6.3 Impact on weather-related vulnerability

We now turn our attention to assessing if the extension of medical coverage has reduced weather vulnerability. To start with, we looked at the distribution of the impacts of the *Seguro Popular* across seasons using the same IV setting. Results for our IV regression by season of the year are provided in Appendix A10. The coefficient for medical assistance is strong and statistically significant for the winter period (Dec. – Feb.) whereas it is smaller in magnitude and statistically insignificant for all the other seasons. Therefore, the *Seguro Popular* seems to have mostly reduced extra winter mortality. The strong impact of the schemes on winter mortality is consistent with the type of diseases covered, in particular diabetes and respiratory diseases such as pneumonias. In fact, people die more from the list of diseases covered by the schemes during the coldest months (see Appendix A10). Therefore, we could expect that an increase in health coverage for these diseases would reduce extra winter mortality.

A better way to look at the effect of the schemes on weather vulnerability would consist in including all the temperature bins in the model and interact them with medical assistance in a 2SLS setting. The problem that we encounter is that it creates as many endogenous variables as temperature lags in the model, e.g. an additional 30 endogenous variables if we consider a model with only one temperature bin (and 360 with all 12 bins). Instrumenting 30 variables produces serious problems for the strength of the instruments. Therefore, we conduct a similar exercise by adopting a reduced form approach where we directly use our instrument – the eligibility to the schemes – in the regression and interact it with temperature bins. The 2SLS results presented previously provide us with sufficient confidence that we will capture the impact of improved medical access on reducing weather vulnerability in such a reduced-form setting.

Results for this reduced form approach are provided in Table 11 for unusually cold temperatures (below 10°C) in the first column, and for mildly cold temperatures (below 20°C) in the second column. Columns 3 and 4 reproduce the regressions with sociodemographic control variables. The regressions include municipality-by-month-by-year fixed effects, and we only report the 31-day cumulative effects for the temperature bins and their interactions. The regressions also include two sorts of interactions with the temperature bins: one with the share of potentially covered cases ($C_n \in \Theta_t$), and another with the eligibility to the schemes ($C_n \in \Theta_t$ and $i_n \in G_t$). The first one is used as a control for the prevalence of specific causes of death during cold days. The second captures the effect of the *Seguro Popular* and the *Fondo de Protección contra Gastos Catastróficos*.

We find that the coverage of the schemes reduces vulnerability to cold temperatures below 10°C by around 30%.⁵⁷ This effect is statistically significant at the 5% level. In the second column, results suggest that the schemes reduce mortality during mildly cold days by around 17% for the set of diseases and municipalities that it covers.

Table 11: Reduced-form impact of eligibility to Seguro Popular on weather mortality

Independent variables	(1)	(2)	(3)	(4)
Potentially covered cases (%)	-0.02**	-0.05***	-0.01	-0.05***
	(0.01)	(0.01)	(0.01)	(0.01)
Eligible to the schemes (%)	-0.02	0.004	-0.01	0.004
	(0.01)	(0.02)	(0.01)	(0.01)
Temperature <10° C (cumulative effect):	0.64***	-	0.67***	
	(0.05)		(0.06)	
x Potentially covered cases (%)	0.14	-	0.12	
	(0.09)		(0.09)	
x Eligible to the schemes (%)	-0.23**	-	-0.22**	
	(0.10)		(0.10)	
Temperature <20° C (cumulative effect):	-	0.16***	-	-0.16***
		(0.02)		(0.02)
x Potentially covered cases (%)	-	0.06***	-	-0.05***
		(0.02)		(0.02)
x Eligible to the schemes (%)	-	-0.04*	-	-0.04**
		(0.02)		(0.02)
Sociodemographic controls (†)	No	No	Yes	Yes

Notes: The table reports OLS results obtained from fixed effect regressions with by municipality by month by year fixed effects. The dependent variable is daily mortality rate at the municipality level. The regression includes municipality-by-year-by-month fixed effects. The coefficients reported for the temperature bins and their interactions correspond to the dynamic cumulative effect over 31 days. The regression is weighted by the population in each municipality. *** indicates statistically significant at the 1% level, ** at the 5% level, and * at the 10% level. Standard errors are clustered at municipality level. †: these are the same as in Table 10.

7. Conclusion

Because investments in protective measures are determined by income, climate change is predicted to affect the poorest people in developing countries the most. This study analyses the heterogeneous impact of temperature shocks on mortality across income groups in Mexico using individual death records and Census data for the period 1998-2010. We find that random variation in temperatures is responsible for the death of around 45,000 people every year in Mexico, representing 8% of annual deaths in the country. However, extreme weather events only account for a small proportion of weather-related deaths: unusually cold days (<10°C) trigger around 4,700 deaths each year, extremely hot days (>32°C) kill less than 400 annually while 88% of weather-related deaths are induced by mildly cold days (10-20°C). The large effect of mildly cold days on mortality that we document has never been reported before, and we suspect this phenomenon to be specific to developing countries.

⁵⁷ 0.23/(0.64+0.14)

A consequence of our findings is that climate change should significantly reduce the number of weather-related deaths in Mexico by 50% to 80% by the end of the 21st century, even in the absence of any adaptation. This illustrates the vast heterogeneity in climate change impacts across countries and regions.

We find that vulnerability to weather shocks is strongly correlated with individual income, and that only people in the bottom half of the income distribution are vulnerable to mildly cold temperatures. The impact of unusually cold days (<10°C) is 35% greater for those living below the median average income. This suggests that not only are poorer households more vulnerable to cold, they also start being vulnerable at temperatures for which richer households are almost fully resilient. Differences in living conditions could explain these findings. For example, we find that only 1% of people in the bottom quarter of the income distribution are equipped with a heater.

Under these circumstances, there is a role for public policies to reduce the mortality inequalities caused by inclement weather. Healthcare systems can be used to reduce the mortality of vulnerable groups while targeting diseases that are known to respond to weather shocks. We exploit variation in universal healthcare coverage caused by the deployment of the Seguro Popular and the Fondo de Protección contra Gastos Catastróficos to assess their contribution to reducing weather vulnerability. We find that the schemes induced a 30% reduction in the mortality induced by days with mean temperature below 10°C and a 17% reduction in mortality during mildly cold days with mean temperature below 20°C for the diseases covered by the schemes. The overall welfare implications of weather vulnerability in developing countries are very large: in the sole case of Mexico, we estimate that forty thousand deaths each year are triggered by temperatures from which people from low-income households are inadequately protected. Furthermore, birth rates are higher in developing countries than in industrialised countries, implying that exposure to cold has a stronger impact on longevity because many young children are exposed. We show that access to universal healthcare can successfully reduce this high vulnerability, but more research is required to assess which protection measures are capable of reducing cold-related vulnerability in the most cost-effective manner.

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APPENDICES – FOR ONLINE PUBLICATION ONLY

Appendices are divided into 3 sections: the main appendices, from Appendix A1 to A10, robustness checks, from Appendix B1 to B12, and appendices on impacts by socioeconomic status, from Appendices C1 to C6. At the end, we have put the list of the references mentioned in the appendices.

MAIN APPENDICES

Appendix A1: Health risks of environmental exposure to heat and cold

Medical evidence

The good functioning of the human body requires core body temperature to be around 37°C. However, variations in ambient air temperatures, whether between seasons or throughout a day, induce heat transfers between the organism and the environment. Below or above a comfort zone within which ambient air temperatures are around 20-25°C, the body needs to activate heating or cooling responses.⁵⁸ The cooling and heating mechanisms of the human body put stress on the organism by themselves. Above all, they may not be sufficient to maintain core body temperature at 37°C, especially if the heat or the cold received is either intense or prolonged.

High ambient air temperatures can cause increases in core body temperature that are associated with dehydration and the development of pathologies. In a review, Basu and Samet (2005) pinpoint that hot temperatures are associated with excess mortality due to cardiovascular, respiratory, and cerebrovascular diseases. In fact, these pathologies develop much before the body enters severe hyperthermia: mild stress caused by ambient air temperatures above 25°C can be sufficient to trigger pathological responses. These pathologies arising because of heat are of the non-transmissible kind (e.g. heart attacks). In addition, mildly high temperatures can also open a window of opportunity for the development of transmissible pathologies. For example, the hosts of some viruses, such as

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⁵⁸ The human body relies on three sets of mechanisms to cope with changes in ambient air temperature: one triggering core body heating through voluntary or involuntary muscle contractions, shivering, tachycardia (the heart beats more quickly), vasoconstriction and rapid breathing to avoid hypothermia; another enabling core body cooling that principally consists of vasodilatation and sweating to avoid hyperthermia; and a neural function to monitor core body temperature (in the hypothalamus), activate either heating or cooling when required, and instigate a strong dislike for excessive heat and cold that encourages protective behaviours (Marriott and Carlson, 1996; Chenuel, 2012).

malaria or dengue, develop more easily in hot and humid environments, explaining higher incidence during hot and humid seasons (Colón-González et al., 2011). This constitutes another channel through which high ambient temperatures may provoke excess mortality.

Importantly, not everyone is vulnerable to heat the same way. Some people are at risk very promptly as soon as temperatures go above their comfort zone. Thermoregulation works inefficiently in some people, making them more vulnerable than others for a given temperature level. This is particularly the case for the elderly and younger children.⁵⁹

As much as high temperatures can overwhelm thermoregulation, cold days can also prevent core body temperature from being maintained at 37°C. Very serious cases of hypothermia (<32°C) impair cardiac, cerebrovascular and respiratory functions, which can lead to loss of consciousness and death (Colon *et al.*, 2011). However, strong hypothermia is uncommon whereas mild cold below the comfort zone is a very common situation which affects several functions of the organism, in particular the circulatory and respiratory functions. ⁶⁰ Like in the case of heat, people with inefficient thermoregulation systems or with preconditions will be more vulnerable to cold, and start being at risk for ambient air temperatures between 10°C and 20°C when others could sustain much lower temperatures. Older individuals respond poorly to cold stress (Young, 1991). This is because ageing is typically characterised by a loss in muscle mass and body fat. ⁶¹ Likewise, malnourished people are vulnerable to cold due to lack of body mass and because core body heating requires the consumption of calories beyond

⁵⁹ These groups tend to have low maximal aerobic power, high adiposity and small body stature and body mass compared with young adults. These characteristics imply relatively large surface area-to-mass ratio along with lower sweat rate and cardiac output. In addition, the elderly tend to have poor control of peripheral blood flow. Their hypothalamic system may also be less prompt in detecting hyperthermia and dehydration. All these factors reduce the efficiency of thermoregulation (Inbar et al., 2004). People with specific preconditions, such as diabetes, are more sensible to heat (Scott et al., 1987). Finally, risks depend on exposure. Occupation may play a major role (Thonneau, 1998): people spending much time outdoors and making physical efforts (which naturally produce heat in the body) are more exposed and therefore more at risk than people making less effort and staying indoors during hot days.

⁶⁰ This can be exemplified looking at the case of mild hypothermia (32-35°C) (Schubert, 1995). Circulatory effects include higher blood viscosity (by 4-6% for each °C) and higher risk of hypovolemia (decreased volume of circulating blood in the body). Mild hypothermia also affects the coagulation system through reversible platelet sequestration, decreases in enzymatic activity for clotting and increases in fibrinolitic activity. In addition, several organs are affected. The cardiac function suffers from higher stress (e.g. impairment of diastolic relaxation) such that mild hypothermia is correlated with higher risk of angina, myocardial and coronary ischemia. Likewise, lungs can be compromised: pulmonary oedemas have been found in patients after environmental exposure to cold (Morales and Strollo, 1993). More frequently, protective airway reflexes are reduced because of impairment of ciliary function. This predisposes to aspiration and pneumonia (Mallet, 2002). In addition, cerebral activity is reduced due to decreases in cerebral blood flow and cerebral metabolic rate of oxygen (by around 5% for each °C). Furthermore, low body temperature decreases the metabolic rate by 5-7% per °C and moderately affects both the hormonal and immunity systems: e.g. hypothermia reduces leukocyte mobility and the speed of phagocytosis (Schubert, 1995).

⁶¹ Muscle mass is the essential component of heat production in the body (Horvath, 1981) whereas body fat offers additional protection to cold.

the scope of what they may have in stock (Marriott and Carlson, 1996). In addition, some transmissible diseases develop more easily in cold environments. It is well-known that the transmission of air-borne viruses can be facilitated by low temperatures. Cold environments may also provide increased stability to enveloped viruses, such as influenza. This is why we observe waves of influenza throughout fall and winter. Colder temperatures may also encourage people to spend more time indoors, in closer proximity to one another and in poorly ventilated environments (Pica and Bouvier, 2014).

Consequently, ambient temperatures below or above a comfort zone of 20-25°C may be a contributing factor to the development of pathologies, and even trigger death, in particular among people with pre-existing health conditions. However, heat or cold will not be reported as the primary cause of hospitalisation or death except in the rare cases of severe hypothermia or hyperthermia. In milder cases, which likely constitute the majority of cold- or heat-related deaths, doctors are more likely to report the pathologies that might have arisen because of heat or cold exposure, such as heart attacks or influenza. For the statistician, this implies that looking directly at medical or death records for severe hypothermia and heat strokes underestimates the fraction of weather-related diseases or deaths.

Appendix A2: Template of death certificate used in Mexico

Mexican death certificates include information on many socio-demographic variables: date of birth, gender, civil status, nationality, profession, education level and affiliation to social security. This comes in addition to the information about usual place of residence and specific details about the death, in particular the place of death, date of death, cause of death and whether the deceased received medical assistance or not before dying.

A template of death certificate is provided hereafter (in Spanish).

SECRETARÍA DE SALUD CERTIFICADO DE DEFUNCIÓN 040000000 02 6. CURP DEL FALLECIDO(A) ESTADO CIVII DEL FALLECIDO PARA OBTENER EL ACTA DE DEFUNCIÓN ca O 10 O₆ No O 2 si O No Q 2 si O DE LA DEFUNCIÓN Y LAS DOS COPIAS AL MUERTES ACCIDENTALES si O 0, lad O₅

Figure B.1: 2004 Template of a death certificate (source: INEGI)

Appendix A3: Summary statistics from the 2000 Mexican Census

Table A3.1 presents some general socioeconomic information on the Mexican population based on the 2000 national Census. The information is split by income quartile, rural vs. urban populations and by type of profession (using the Mexican nomenclature of activities). Not surprisingly, rural populations are less educated, less likely to have access to social security and in general have an average income level about half that of people living in urban areas.

The difference in the average personal income between the poorer profession (agriculture) and the richer one (public servants and directors) is 1 to 6. Economic differences by quartiles are much sharper though since there is high heterogeneity within each profession. People in the first income quartile have an average personal income which is 18 times lower than people in the top quartile. This large inequality is a feature of the Mexican economy which we will use in the next sections to investigate differences in the weather-mortality relationship across income groups.

Table A3.1: Socioeconomic characteristics of the Mexican population based on 2000 Census

Population	Personal income*	No social security	Completed secondary school [†]	Age	Male	Share of population
Total	2,876	58.6%	37.1%	26.2	48.7%	100.0%
Rural	1,433	83.7%	17.3%	25.0	49.6%	25.4%
Urban	3,330	50.1%	43.8%	26.5	48.4%	74.6%
By quartile of income						
1st quartile	437	82.9%	18.6%	24.7	48.2%	25.0%
2nd quartile	1,155	60.8%	31.5%	24.5	48.7%	25.0%
3rd quartile	2,119	47.4%	42.3%	26.0	49.2%	25.0%
4th quartile	7,816	36.2%	59.7%	28.6	49.3%	25.0%
By type of profession						
Workers in agriculture, fisheries and hunting activities	1,552	87.1%	18.1%	38.2	92.7%	5.2%
Do not work (under 16)	2,371	62.5%	14.4%	7.7	50.0%	37.3%
Assistants in industrial and handmade production	2,397	62.1%	44.9%	28.5	85.3%	1.5%
Do not work (over 65)	2,647	49.4%	10.9%	74.4	36.5%	4.1%
Do not work (16-65)	2,648	62.4%	47.5%	34.3	21.2%	25.9%
Street vendors	2,679	81.4%	41.5%	38.6	68.8%	0.7%
Workers in industry of transformation	2,784	64.0%	46.9%	34.9	85.7%	5.5%
Workers in army and civil protection	3,059	21.4%	66.3%	36.5	94.3%	0.8%
Drivers of mobile machines and transports	3,061	54.6%	59.5%	35.8	99.3%	1.6%
Workers in personal services in institutions	3,116	47.0%	53.2%	34.2	60.4%	1.9%
Fixed machine operators	3,323	15.6%	61.3%	28.7	61.9%	1.9%
Domestic workers	3,753	78.2%	27.4%	34.0	12.2%	1.4%
Sellers, employees in trade and salesmen	3,817	57.9%	67.5%	35.0	60.6%	3.8%
Low-skilled workers in administrative tasks	4,124	24.1%	91.3%	31.0	38.4%	2.3%
Technicians	4,641	26.4%	91.4%	33.8	56.0%	1.0%
Overseers in industrial production	5,045	16.4%	84.0%	34.4	79.7%	0.6%
Workers in education	5,662	15.0%	98.9%	36.8	39.8%	1.4%
Medium-skilled workers in administrative tasks	5,973	18.3%	93.5%	35.8	67.6%	0.8%
Workers in art, sports and events	6,176	58.0%	81.3%	34.7	74.9%	0.3%
Certified professionals	7,758	32.0%	99.8%	36.5	63.2%	1.3%
Public servants and directors	10,453	29.0%	95.8%	39.7	74.0%	0.7%

Notes. The table shows average values of socioeconomic characteristics of the Mexican population based on the 2000 Census. Statistics are calculated using the sample weights provided by INEGI. *: Personal income (in 2000 Mexican pesos) is calculated as family income divided by the square root of the total number of people in the household. This calculation method allows accounting for economies of scale in larger households. †: includes people that were completing secondary school.

Appendix A4: Contemporaneous effect

Due to an omitted variable bias, correlating today's temperatures with today's mortality will lead to biased estimates of the impact of temperature on mortality if no account of the temperatures of the previous days is made. Figure A4.1 displays the impact of the day's temperature on mortality for all Mexicans and all causes of death when no lagged temperature bins are included in the model. This can help the reader assess the magnitude and the direction of the bias produced in this case.

The Mexican population appears to be very sensitive to high temperatures above 28°C. A statistically significant impact of temperatures below 14°C is also detected. However, an extremely hot day above 32°C is three times more lethal than an unusually cold day below 10°C. The temperature bin with the lowest mortality is 18-20°C.

0.50
0.40
0.30
0.20
0.10
0.00
-0.10
-0.20

Temperature bins (Celsius degrees)

Figure A4.1: Impact of the day's average temperature on daily mortality, in deaths per 100,000 inhabitants

Notes. Lines in dash correspond to the 95% confidence interval values obtained for each estimated coefficient. 312,140 groups and 30.1 observations per group. The regression results control for the day's precipitation level.

Therefore, the model with contemporaneous temperatures underestimates the effect of cold and over-estimates the impact of heat. Biases also appear when the contemporaneous model is run with a breakdown by gender, age and type of disease leading to death (see Table A4.1 and Table A4.2). In particular, men appear to be three to four times more strongly impacted by unusual cold – this result is not confirmed with a distributed lag model.

<u>Table A4.1: Impact of a day under 10 Celsius degree on mortality as compared to a reference</u> day of 18-20 degrees

	Cause of death						
Group	All causes	Infectious diseases	Neoplasms	Endocrine, nutritional and metabolic diseases	Circulatory system diseases	Respiratory system diseases	Violent and accidental
Total	0.0831***	0.0006	0.0014	0.0154***	0.0265***	0.0217***	-0.0107***
10111	(0.0099)	(0.0016)	(0.0029)	(0.0036)	(0.0047)	(0.0033)	(0.0037)
Men	0.13***	0.0016	0.0042	0.0244***	0.0489***	0.0268***	-0.025***
141011	(0.0149)	(0.0024)	(0.0044)	(0.0049)	(0.0069)	(0.0048)	(0.0063)
Women	0.0387***	-0.00037	-0.0013	0.0071	0.005	0.0167***	0.0033
Women	(0.0126)	(0.002)	(0.0039)	(0.0052)	(0.0064)	(0.0044)	(0.0033)
Aged 0-4	0.179***	0.0099	-0.0013	0.0081	0.0018	0.128***	0.0187**
Aged 0-4	(0.024)	(0.0068)	(0.0018)	(0.0052)	(0.0021)	(0.0123)	(0.0083)
Acad 4.0	0.0018	0.0015	-0.0019	-0.0001	-0.001	0.0019	-0.0013
Aged 4-9	(0.0065)	(0.0015)	(0.0017)	(0.0014)	(0.0007)	(0.002)	(0.0041)
A and 10 10	-0.0088	-0.0013	0.0009	0.0012	0.0008	0.00006	-0.0165***
Aged 10-19	(0.0066)	(0.0013)	(0.0017)	(0.0009)	(0.0011)	(0.0012)	(0.0052)
Aged 20-34	-0.0054	0.0001	0.0022	-0.0002	0.0022	0.0026	-0.0271***
Ageu 20-34	(0.0107)	(0.002)	(0.0025)	(0.002)	(0.0022)	(0.0019)	(0.008)
A 1 25 44	0.0285	0.0022	0.0038	0.0189***	-0.0019	-0.0012	-0.027**
Aged 35-44	(0.0194)	(0.0036)	(0.0061)	(0.0063)	(0.0062)	(0.0031)	(0.0108)
A 1 15 51	0.0678**	0.0053	-0.0073	0.03**	0.0067	-0.0025	-0.0144
Aged 45-54	(0.033)	(0.0072)	(0.011)	(0.0127)	(0.0123)	(0.0071)	(0.014)
A 1 55 C1	0.233***	0.0118	-0.0153	0.0621**	0.0844***	0.0369***	0.0053
Aged 55-64	(0.0559)	(0.0088)	(0.0206)	(0.0254)	(0.0249)	(0.0137)	(0.0171)
A 1 CE 74	0.372***	0.019	-0.0224	0.102***	0.131***	0.0591***	0.0065
Aged 65-74	(0.0861)	(0.0135)	(0.0318)	(0.0389)	(0.0383)	(0.0213)	(0.026)
A 1 75 ·	1.03***	-0.0601**	0.0049	0.0597	0.522***	0.168*	0.0359
Aged 75+	(0.238)	(0.0289)	(0.0683)	(0.0875)	(0.141)	(0.0946)	(0.0397)

Notes: Standard errors in brackets. *** indicates statistically significant at the 1% level, ** at the 5% level, and * at the 10% level. Unit is deaths in a day per 100,000 inhabitants (in subgroup). 312,140 groups and 30.1 observations per group.

<u>Table A4.2: Impact of a day over 32 Celsius degree on mortality as compared to a reference day of 18-20 degrees</u>

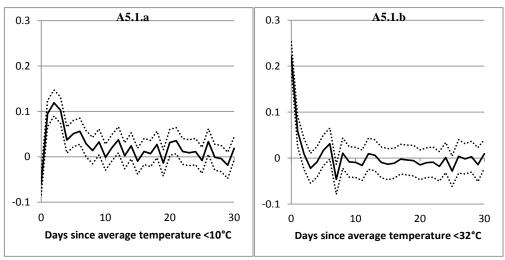
	Cause of death							
Group	All causes	Infectious diseases	Neoplasms	Endocrine, nutritional and metabolic diseases	Circulatory system diseases	Respiratory system diseases	Violent and accidental	
Total	0.269***	0.0072***	0.0213***	0.0374***	0.0837***	0.0227***	0.057***	
Total	(0.0152)	(0.0023)	(0.005)	(0.0051)	(0.0073)	(0.0039)	(0.0077)	
Men	0.29***	0.008**	0.0144**	0.0334***	0.0907***	0.018***	0.0952***	
Wien	(0.0233)	(0.0035)	(0.0071)	(0.007)	(0.011)	(0.0058)	(0.0137)	
Women	0.249***	0.0065**	0.0283***	0.041***	0.0766***	0.0274***	0.0191***	
Wolliell	(0.0184)	(0.003)	(0.007)	(0.0073)	(0.0095)	(0.0052)	(0.0058)	
Aged 0-4	0.138***	0.0337***	0.0029	0.0271***	0.0022	0.0129	-0.0042	
Aged 0-4	(0.0278)	(0.0102)	(0.0024)	(0.0078)	(0.0027)	(0.0086)	(0.0103)	
Aged 4-9	0.0145	0.0026	0.0009	-0.0025	0.0003	-0.0002	0.0117	
Aged 4-9	(0.0091)	(0.0021)	(0.0024)	(0.0017)	(0.0009)	(0.0018)	(0.0072)	
Aged 10-19	0.0415***	-0.0013	-0.00005	0.001	0.001	-0.0007	0.0359***	
Ageu 10-19	(0.0112)	(0.0014)	(0.0023)	(0.0016)	(0.0015)	(0.0014)	(0.0101)	
A and 20, 24	0.0843***	-0.004	0.0012	0.01***	0.0053	0.0044	0.0641***	
Aged 20-34	(0.0219)	(0.0037)	(0.0033)	(0.0033)	(0.0033)	(0.0036)	(0.0195)	
A and 25 44	0.155***	0.0004	0.0091	0.0069	0.0173*	0.0036	0.105***	
Aged 35-44	(0.0321)	(0.0057)	(0.0093)	(0.0063)	(0.0096)	(0.004)	(0.0244)	
A 1 45 54	0.157***	0.007	0.0335*	-0.019	0.0437**	0.0076	0.082***	
Aged 45-54	(0.0441)	(0.0086)	(0.0176)	(0.0152)	(0.0181)	(0.0072)	(0.0239)	
A 1 55 CA	0.206***	-0.0057	0.0171	0.0225	0.131***	-0.0002	0.0322	
Aged 55-64	(0.0774)	(0.0112)	(0.0314)	(0.0338)	(0.0379)	(0.0153)	(0.0279)	
A and 65 74	0.338***	-0.0078	0.0287	0.0347	0.212***	-0.0019	0.0539	
Aged 65-74	(0.122)	(0.018)	(0.0496)	(0.0531)	(0.0595)	(0.0233)	(0.0437)	
A 1 75 :	5.75***	0.147***	0.35***	1***	2.46***	0.705***	0.23***	
Aged 75+	(0.355)	(0.0474)	(0.116)	(0.129)	(0.214)	(0.118)	(0.0656)	

Notes: Standard errors in brackets. *** indicates statistically significant at the 1% level, ** at the 5% level, and * at the 10% level. Unit is deaths in a day per 100,000 inhabitants (in subgroup). 312,140 groups and 30.1 observations per group.

Appendix A5: Short-term dynamics of impacts in distributed lag model

Our results are consistent with the dynamic effects of heat and cold days on mortality as reported previously, for example by Deschenes and Moretti (2009). Like these authors, we find evidence of strong harvesting for hot days whereas the impact of cold days accumulates after the event. These short-term dynamics can be observed on Figure A5.1.a and A5.1.b, which present the impact on mortality of extremely hot/cold days on the day of the weather event and for each of the following 30 days. A cold day below 10°C has a statistically significant effect on mortality every day during the first week, and we find statistically significant effects at 5, 14 and even 21 and 22 days after the cold day. By contrast, we find that a hot day above 32°C has a strong and immediate effect on mortality but this effect is statistically significant only for the first two days, after which the coefficients become systematically negative although not statistically significantly so.

Figure A5.1: Impact within 31 days of a cold day (<10°C – panel a) or a hot day (>32°C – panel b) on daily mortality rate per 100,000 inhabitants



Note: These two graphs are obtained from the same regression, considering all Mexican people and all causes of death (1998-2010). Unit is deaths per 100,000 inhabitants. Each point corresponds to an estimated coefficient from the distributed lag model for days below 10°C (Panel a) or above 32°C (Panel b). Dashed lines correspond to the 95% confidence interval obtained for each estimated coefficient. 312,140 groups and 30.1 observations per group. All regressions control for the day's precipitation level.

Appendix A6: Comparison of main results with related studies

The methodology and data used in this paper are very close to Deschenes and Moretti (2009). These authors use a 30-day distributed lag model and estimate that one day below 30°F (-1.1°C) increases the mortality rate by 0.23 deaths per 100,000 inhabitants as compared to any other day in the year. On the other hand, we find that 31-day cumulated mortality is increased by 0.70 deaths per 100,000 inhabitants for an additional day below 10°C: their estimate is three times lower than ours while we look at hotter days since days below 50°F (10°C) correspond to the lower limit for unusually cold days in our data. If we use the exact same methodology as Deschenes and Moretti (2009) (as in Online Appendix B10), we then find an estimate of 0.60 deaths per 100,000 inhabitants and the lower bound of our 95% confidence interval is 0.53. This figure does not overlap with the upper bound of Deschenes and Moretti (2009)'s 95% confidence interval at around 0.28. Therefore, both estimates are statistically different from each other suggesting that Mexican residents are more vulnerable to cold than US residents.

Table A6.1: Comparison of the main results of similar panel data studies

Study	Country and period	Frequen-cy of data	Day below minus 1.1°C (10°F)	Day between 4.4-10°C (40-50°F)	Day below 10°C (50°F)	Days between 10-14°C (50°F-	Day above 32°C (or 90°F)
This study	Mexico (1998- 2010)	Daily			+0.70 deaths per 100,000 inhabitants		+0.13 deaths per 100,000 inhabitants
					Annual mortality rate increases by about 0.15%		Annual mortality rate increases by about 0.03%
Deschenes and Moretti (2009)	USA (1972- 1988)	Daily	+0.23 deaths per 100,000 inhabitants				Statistically insignificant
Barreca (2012)	USA (1973- 2002)	Monthly		+0.15 deaths per 100,000 inhabitants			+0.17 deaths per 100,000 inhabitants
Deschenes and Greenstone (2011)	USA (1968- 2002)	Annual	+0.69 deaths per 100,000 inhabitants				+0.92 deaths per 100,000 inhabitants
Burgess et al. (2014)	India (1957- 2000)	Annual				Annual mortality rate increases by about 0.4-0.7%	Annual mortality rate increases by about 0.5-1%

The estimates by Deschenes and Moretti (2009) are in line with those obtained in other studies for the US. Barreca (2012) finds that a day between 40°F and 50°F (4.4-10°C) increases the monthly mortality rate by 4.5 people per 100,000 inhabitants. This corresponds to a daily mortality rate of 0.15 people per 100,000 inhabitants (95% confidence interval = 0.09-0.22). Using annual data, Deschenes and Greenstone (2011) find that a day below 10°F

(-12°C) increases mortality by 0.69 people per 100,000 inhabitants and a day between 40°F and 50°F (4-10°C) by 0.27 people per 100,000 inhabitants as compared to a day between 50°F and 60°F (10-15.5°C). The upper bound of the 95% confidence interval for this last estimate is around 0.49 and therefore statistically below ours.

One reason why Mexicans could be more vulnerable to cold than Americans could be acclimation: since they live in a hot country, Mexicans may be less prepared to face low temperatures. However, our results suggest that Mexicans could also be more vulnerable to high temperatures. For a day above 90°F (32.2°C), Deschenes and Moretti (2009) find no evidence of an impact of heat on mortality after 30 days. They find a highly positive impact of temperatures on mortality on the first days of heat waves but the latter is compensated for in the short run due to a harvesting effect. For the same level of temperatures, we find a statistically significant and positive impact of hot days on 31-day cumulative mortality: with temperatures above 32°C, the mortality rate is, in average, higher by 0.13 deaths by 100,000 inhabitants in Mexico.

However, Barreca (2012) and Deschenes and Greenstone (2011) do find a mortality impact of hot days: respectively 0.17 and 0.92 deaths per 100,000 inhabitants for temperatures above 90°F (32°C). The impact found by Barreca (2012) using mortality data is therefore comparable to ours in magnitude. As for Deschenes and Greenstone (2011), they use annual data over a long time period (1968-2002) so as to capture indirect effects of temperatures on mortality through other channels (e.g. agricultural and industrial output, and therefore income, employment, access to healthcare, etc.). Their estimates would indicate stronger vulnerability in the US but are not as easily comparable to our results, not only because we use with daily data but also look at a different time period.

Let us now turn our eyes to the results obtained by Burgess et al. (2014) for India. These authors use a log-linear model to estimate the impact of temperatures on annual mortality. They find impacts of a much higher magnitude for India as compared to the US estimates of Deschenes and Moretti (2009). For cold, the coefficient of their model is not statistically significant at the lower limit of 10°C or below possibly due to the small frequency of such cold days in their data. However, they find that the log annual mortality rate increases by 0.004 for each day between 10-12°C and by 0.007 for each day between 14°C. In other words, an additional day between 10-14°C increases the annual mortality rate by about 0.4-0.7% in India. For heat, they find that an additional day above 32°C increases the annual mortality rate by about 0.5-1%.

We may compare these figures with ours, taking into considerations that our study uses daily data and therefore is not fully comparable. The average daily mortality rate is around 1.3 deaths per 100,000 inhabitants in Mexico. Converted to an annual rate, this corresponds to about 475 deaths per 100,000 inhabitants. In this context, our estimate of an extra 0.7 deaths per 100,000 inhabitants caused by a day below 10°C roughly represents a marginal increase of 0.15% in the annual death rate. Likewise, the estimate of 0.13 deaths per 100,000 due to a day above 32°C corresponds to a marginal increase in the annual death rate by less than 0.03%. The relative impact of cold on mortality in Mexico seems at least 2-3 times lower than in India whereas the estimated impact for heat is incomparably lower.

Appendix A7: Impacts of Climate Change

We calculate the number of weather-related deaths under climate change based on the output of the climate model GFDL CM3 for 2075-2099. Annual death estimates under climate change are provided in Table A7.1.62 Because the frequency of cold and mildly cold days is expected to decrease, the number of deaths imputable to temperatures reduces with the forecasted temperatures of GFDL CM3 as compared with the historical ones. With the RCP2.6 scenario (low GHG emissions), temperature-related mortality would be twice as small. The RCP8.5 scenario (high GHG emissions) corresponds to an 80% reduction in the estimated relationship between mortality and temperature. We show later that weather-related mortality affects mostly people in the first two quartiles of the income distribution, suggesting that the reduction in the exposure to cold weather associated by climate change could lead to a reduction in mortality inequality. Therefore, in Mexico, we predict that climate change will reduce the impact of short-term weather variability on mortality, with significant health benefits. However, this analysis comes with serious warnings: climate change could also affect mortality through increased frequency of natural catastrophes and not only through temperatures; our analysis at the daily level does not allow for acclimatization; and we could be underestimating the impact of increased heat waves if the effect of heat grows non-linearly beyond 32°C days. In addition, our model includes municipality-by-month-by-year fixed effects which control for income and for the general health of the population. Climate change may impact income, or the general health of the population, and these factors may in turn impact mortality. Our econometric specification cannot assess the magnitude of such indirect economic effects on mortality. Yet in section 5, we show that these factors can significantly modify the health response to temperature shocks, even when looking at short run impacts.

Table A7.1: Impact of temperatures on annual deaths in several climate scenarios

Number of deaths	Estimates	Compared to historical temperatures
Historical	41,335*	
	(27,299; 55,370)	
GFDL CM3:		
RCP2.6	18,152*	-23,183*
	(5,898; 30,405)	(-26,410; -19,956)
RCP4.5	12,842*	-28,493*
	(1,177; 24,506)	(-33,196; -23,790)
RCP8.5	7,513	-33,821*
	(-4,000; 19,026)	(-41,629; -26,014)

Note: * denotes statistically significant at 5%. The 95% confidence interval in brackets only take into account the uncertainty of the impact of temperature bins on mortality. It does not take into account the uncertainty of climate models in the distribution of daily temperatures.

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⁶² The distributions for hot and cold days obtained with this climate model are reported in Figure 1.

Appendix A8: List of diseases covered by the Seguro Popular and Fondo de

Protección contra Gastos Catastróficos in 2010

Diseases and treatments covered by the Seguro Popular in 2010

PUBLIC HEALTH

Newborn and children under 5 years of age

- 1 BCG Vaccine
- 2 Hepatitis B vaccine
- 3 Pentavalent vaccine with acellular pertussis component (DpaT + VIP + Hib)
- 4 SRP triple viral vaccine
- 5 Rotavirus vaccine
- 6 Influenza vaccine
- 7 DPT vaccine
- 8 Sabin trivalent oral polio vaccine
- 9 Preventive actions for newborn
- 10 Preventive actions for children under 5 years

Girls and boys from 5 to 9 years old

11 Preventive actions for children between the ages of 5 and 9

Teenagers 10 to 19 years old

- 12 Early detection of eating disorders
- 13 Preventive actions for adolescents aged 10 to 19
- 14 Hepatitis B vaccine

Adults 20 to 59 years old

- 15 SR double viral vaccine
- 16 Tetanus and diphtheria toxoid (Td)
- 17 Preventive actions for women aged 20-59
- 18 Preventive actions for men aged 20-59
- 19 Complete medical examination for women aged 40-59
- 20 Complete medical examination for men aged 40-59
- 21 Prevention and care of family and sexual violence in women
- Adults over 60 years and over
- 22 Pneumococcal vaccine for the elderly
- 23 Influenza vaccine for the elderly
- 24 Preventive actions for adults over 60 and older
- General / family consultation
- 25 Diagnosis and treatment of iron deficiency anemia and vitamin B12 deficiency
- 26 Diagnosis and treatment of vitamin A deficiency
- 27 Diagnosis and treatment of rubella
- 28 Diagnosis and treatment of measles
- 29 Diagnosis and treatment of chickenpox
- 30 Diagnosis and treatment of acute pharyngotonsillitis
- 31 Diagnosis and treatment of whooping cough
- 32 Diagnosis and treatment of non-suppurative otitis media
- 33 Diagnosis and treatment of acute rhinopharyngitis (common cold)
- 34 Diagnosis and treatment of conjunctivitis
- 35 Diagnosis and treatment of allergic rhinitis
- 36 Diagnosis and treatment of classical dengue fever
- 37 Diagnosis and outpatient treatment of acute diarrhea
- 38 Diagnosis and treatment of paratyphoid fever and other salmonellosis
- 39 Diagnosis and treatment of typhoid fever
- 40 Diagnosis and treatment of herpes zoster
- 41 Diagnosis and treatment of candidiasis
- 42 Diagnosis and treatment of gonorrhea
- 43 Diagnosis and treatment of chlamydia infections including trachoma-
- 44 Diagnosis and treatment of Trichomonas infections

- 45 Diagnosis and treatment of syphilis
- 46 Diagnosis and treatment of cystitis
- 47 Diagnosis and treatment of acute vaginitis
- 48 Diagnosis and treatment of acute vulvitis
- 49 Diagnosis and pharmacological treatment of intestinal amebiasis
- 50 Diagnosis and pharmacological treatment of hookworm and necatorisis
- 51 Diagnosis and pharmacological treatment of ascariasis
- 52 Diagnosis and pharmacological treatment of enterobiasis
- 53 Diagnosis and pharmacological treatment of echinococcosis
- 54 Diagnosis and pharmacological treatment of equistosomiasis (bilharziasis)
- 55 Diagnosis and pharmacological treatment of strongyloidiasis
- Pharmacological diagnosis and treatment of filariasis
- 57 Diagnosis and pharmacological treatment of giardiasis
- 58 Diagnosis and pharmacological treatment of tapeworms
- 59 Diagnosis and pharmacological treatment of trichuriasis
- 60 Diagnosis and pharmacological treatment of trichinosis
- 61 Diagnosis and treatment of scabies
- 62 Diagnosis and treatment of pediculosis and phthiriasis
- 63 Diagnosis and treatment of superficial mycoses
- 64 Diagnosis and treatment of onychomycosis
- 65 Diagnosis and treatment of infectious cellulitis
- 66 Diagnosis and treatment of allergic contact dermatitis
- 67 Diagnosis and treatment of atopic dermatitis
- 68 Diagnosis and treatment of irritant contact dermatitis
- 69 Diagnosis and treatment of diaper dermatitis
- 70 Diagnosis and treatment of exfoliative dermatitis
- 71 Diagnosis and treatment of seborrheic dermatitis
- 72 Diagnosis and treatment of common warts
- 73 Diagnosis and treatment of acne
- 74 Diagnosis and treatment of hepatitis A
- 75 Diagnosis and treatment of acute gastritis
- 76 Diagnosis and treatment of irritable bowel syndrome
- 77 Diagnosis and pharmacological treatment of diabetes mellitus $\boldsymbol{2}$
- 78 Diagnosis and pharmacological treatment of hypertension
- 79 Diagnosis and treatment of osteoarthritis
- 80 Diagnosis and treatment of low back pain
- 81 Other general medical care
- 82 Temporary Family Planning Methods: Hormonal Contraceptives (AH)
- 83 Temporary family planning methods: condoms
- 84 Temporary family planning methods: intrauterine device
- 85 Prenatal care in pregnancy

SPECIALTY CONSULTATION

- 86 Diagnosis and treatment of attention deficit hyperactivity disorder
- 87 Diagnosis and treatment of generalized developmental

disorders (Autism)

- 88 Diagnosis and treatment of dysmenorrhea
- 89 Menopause and climacteric care
- 90 Diagnosis and treatment of fibrocystic mastopathy
- 91 Diagnosis and treatment of endometrial hyperplasia
- 92 Diagnosis and treatment of subacute and chronic vaginitis
- 93 Diagnosis and treatment of endometriosis
- 94 Diagnosis and treatment of urethritis and urethral syndrome
- 95 Diagnosis and treatment of low grade intraepithelial squamous lesions
- 96 Diagnosis and treatment of high-grade intraepithelial squamous lesions
- 97 Diagnosis and treatment of malnutrition and obesity in children and adolescents
- 98 Diagnosis and treatment of Kwashiorkor
- 99 Diagnosis and treatment of nutritional marasmus
- 100 Diagnosis and treatment of malnutrition sequelae
- 101 Diagnosis and treatment of acute laryngotracheitis
- 102 Diagnosis and treatment of suppurative otitis media
- 103 Diagnosis and treatment of acute sinusitis
- 104 Diagnosis and treatment of asthma in adults
- 105 Diagnosis and treatment of asthma in children
- 106 Diagnosis and treatment of tuberculosis (TAES)
- 107 Diagnosis and treatment of drug-resistant tuberculosis
- 108 Prevention, diagnosis and treatment of psoriasis
- 109 Diagnosis and treatment of reflux esophagitis
- 110 Diagnosis and treatment of peptic ulcer
- 111 Diagnosis and treatment of dyslipidemia
- 112 Diagnosis and treatment of hyperthyroidism
- 113 Diagnosis and treatment of congenital and adult hypothyroidism
- 114 Diagnosis and pharmacological treatment of diabetes mellitus 1
- 115 Diagnosis and treatment of chronic heart failure
- 116 Diagnosis and treatment of osteoporosis
- 117 Diagnosis and treatment of gout
- 118 Diagnosis and treatment of rheumatoid arthritis
- 119 Diagnosis and treatment of affective disorders
- (Dysthymia, depression and bipolar affective disorder)
- 120 Diagnosis and treatment of anxiety disorders (generalized anxiety, distress and panic attacks and
- reactions to severe stress and adaptation disorders [posttraumatic stress disorder and adaptive disorder])
- 121 Diagnosis and treatment of psychotic disorders
- (Schizophrenia, delusions, psychotic and schizotypal)
- 122 Diagnosis and pharmacological treatment of epilepsy
- 123 Diagnosis and treatment of Parkinson's disease
- 124 Diagnosis and treatment of congenital dislocation of the hip
- 125 Rehabilitation of fractures
- 126 Rehabilitation of facial paralysis
- 127 Selective and indicated prevention of addictions (Counseling)
- 128 Diagnosis and treatment of addictions

ODONTOLOGY

- 129 Prevention of caries and periodontal disease
- 130 Sealing of dentures and fissures
- 131 Removal of caries and restoration of teeth with amalgam, resin or glass ionomer
- 132 Elimination of outbreaks of infection, abscesses (including drainage and pharmacotherapy)
- 133 Removal of teeth, including erupted and root rests (does not include third molar not erupted)

- 134 Diagnosis and treatment of pulpitis and pulp necrosis
- 135 Diagnosis and treatment of maxillary abscess
- 136 Third molar extraction

EMERGENCIES

- 137 Stabilization in emergencies due to hypertensive crisis
- 138 Emergency Stabilization of the Diabetic Patient
- 139 Urgent management of non-ketotic hyperglycemic syndrome
- 140 Stabilization in the emergency room for angina pectoris
- 141 Diagnosis and treatment of acute phenothiazine intoxication
- 142 Diagnosis and treatment of acute alkali intoxication
- 143 Diagnosis and treatment of acute food poisoning
- 144 Diagnosis and treatment of acute salicylate poisoning
- 145 Diagnosis and treatment of acute methyl alcohol intoxication
- 146 Diagnosis and treatment of acute organophosphate poisoning
- 147 Diagnosis and treatment of acute carbon monoxide poisoning
- 148 Diagnosis and treatment of snake bite
- 149 Diagnosis and treatment of alacranismo
- 150 Diagnosis and treatment of bee, spider and other arthropod stings
- 151 Management of biting and prevention of rabies in humans
- 152 Extraction of foreign bodies
- 153 Management of traumatic soft tissue injuries (healing and suturing)
- 154 Diagnosis and treatment of mild traumatic brain injury (Glasgow 14-15)
- 155 Emergency management of first-degree burns
- 156 Diagnosis and treatment of cervical sprain
- 157 Diagnosis and treatment of shoulder sprain
- 158 Diagnosis and Treatment of Elbow Sprain
- 159 Diagnosis and treatment of wrist and hand sprain
- 160 Diagnosis and treatment of sprained knee
- 161 Diagnosis and treatment of ankle and foot sprains

HOSPITALIZATION

- 162 Diagnosis and treatment of pyelonephritis
- 163 Diagnosis and treatment of bronchiolitis
- 164 Diagnosis and treatment of acute bronchitis
- 165 Diagnosis and treatment of meningitis
- 166 Diagnosis and treatment of mastoiditis
- 167 Diagnosis and treatment of inastoratis
- 168 Diagnosis and treatment of pneumonia in children
- 169 Diagnosis and treatment of pneumonia in adults and older adults
- 170 Diagnosis and treatment of amebic liver abscess
- 171 Diagnosis and treatment of pelvic inflammatory disease
- 172 Diagnosis and treatment of threatened abortion
- 173 Diagnosis and treatment of preterm delivery
- 174 Care of childbirth and physiological puerperium
- 175 Pelviperitonitis
- 176 Puerperal endometritis
- 177 Diagnosis and treatment of puerperal septic shock
- 178 Care of the newborn
- 179 Neonatal jaundice
- 180 Diagnosis and treatment of uncomplicated prematurity
- 181 Diagnosis and treatment of prematurity with hypothermia

- 182 Diagnosis and treatment of the newborn with low birth weight
- 183 Diagnosis and treatment of preeclampsia
- 184 Diagnosis and treatment of severe preeclampsia
- 185 Diagnosis and treatment of eclampsia
- 186 Puerperal obstetric haemorrhage
- 187 Bleeding from placenta previa or premature
- detachment of placenta normoinserta
- 188 Infection of episiotomy or obstetric surgical wound
- 189 Diagnosis and treatment of renal and ureteral lithiasis
- 190 Diagnosis and treatment of lower urinary lithiasis
- 191 Diagnosis and treatment of hemorrhagic dengue
- 192 Diagnosis and Treatment of Moderate Head Injury (Glasgow 9-13)
- 193 Diagnosis and conservative management of acute pancreatitis
- 194 Hospital management of seizures
- 195 Hospital management of hypertension
- 196 Diagnosis and treatment of acute heart failure (pulmonary edema)
- 197 Chronic Obstructive Pulmonary Disease
- 198 Diagnosis and treatment of peripheral neuropathy secondary to diabetes
- 199 Hospital management of second degree burns
- 200 Diagnosis and treatment of digestive haemorrhage
- 201 Diagnosis and treatment of HELLP syndrome
- 202 Diagnosis and treatment of chorioamniositis
- 203 Diagnosis and treatment of obstetric embolisms
- 204 Diagnosis and treatment of gestational diabetes
- 205 Diagnosis and treatment of functional heart disease in the pregnant woman
- 206 Diagnosis and treatment of deep venous thrombosis in the pregnant woman

GENERAL SURGERY

- 207 Exploratory laparotomy
- 208 Appendectomy
- 209 Splenectomy
- 210 Surgical treatment of diverticular disease
- 211 Surgical treatment of ischemia and intestinal infarction
- 212 Surgical treatment of intestinal obstruction
- 213 Surgical treatment of gastric and intestinal perforation
- 214 Surgical treatment of colonic volvulus
- 215 Surgical treatment of the rectal abscess
- 216 Surgical treatment of fistula and anal fissure
- 217 Hemorrhoidectomy
- 218 Surgical treatment of hiatal hernia
- 219 Surgical treatment of congenital pylorus hypertrophy
- 220 Crural Hernioplasty
- 221 Inguinal Hernioplasty
- 222 Umbilical Hernioplasty
- 223 Ventral Hernioplasty
- 224 Open cholecystectomy
- 225 Laparoscopic cholecystectomy
- 226 Surgical treatment of condylomata
- 227 Surgical Treatment of Breast Fibroadenoma
- 228 Surgical treatment of ovarian cysts
- 229 Surgical treatment of torsion of attachments
- 230 Salpingoclasia (Definitive method of family planning)
- 231 Surgical care of trophoblastic disease
- 232 Surgical treatment of ectopic pregnancy
- 233 Uterine therapeutic uterine by incomplete abortion
- 234 C-section and surgical puerperium care
- 235 Uterine repair

- 236 Endometrial ablation
- 237 Endometriosis Laparoscopy
- 238 Myomectomy
- 239 Abdominal hysterectomy
- 240 Vaginal hysterectomy
- 241 Colpoperineoplasty
- 242 Vasectomy (Definitive method of family planning)
- 243 Circumcision
- 244 Orchidopexy
- 245 Open Prostatectomy
- 246 Transurethral resection of prostate
- 247 Removal of cancerous skin lesion (melanoma not included)
- 248 Removal of benign tumor in soft tissues
- Tonsillectomy with or without adenoidectomy
- 250 Excision of juvenile pharyngeal papilloma
- 251 Palateplasty
- 252 Cleft Lip Repair
- 253 Muscular shortening surgery for strabismus
- 254 Muscular lengthening surgery for strabismus
- 255 Surgical treatment of glaucoma
- 256 Pterygium excision
- 257 Surgical treatment of hydrocephalus
- 258 Placement and removal of various catheters
- 259 Radical neck dissection
- 260 Thoracotomy, pleurotomy and chest drainage
- 261 Surgical treatment of congenital dislocation of the
- 262 Surgical treatment of the equine foot in children
- 263 Safenectomy
- 264 Surgical reduction by dislocations
- 265 Surgical reduction of clavicle fracture
- 266 Surgical reduction of hum fracture

List of diseases and treatments covered by the Fondo de Protección contra Gastos Catastróficos in 2010

Cervical Cancer Uterine Cancer Antiretroviral HIV / AIDS Treatment Cataract in Adults Congenital cataract Malignant Breast Tumor

Neonatal Intensive Care

Prematureness Respiratory insufficiency Sepsis

Cancer of children Astrocytoma Medulloblastoma Neuroblastoma Ependymoma

Other Tumors of the Central Nervous System

Wilms tumor

Other kidney tumors Acute lymphoblastic leukemia

Acute Myeloblastic Leukemia

Chronic Leukemias

Preleukemic Syndromes

Hepatoblastoma

Hepatocarcinoma

Osteosarcoma

Ewing's sarcoma

Non-Hodgkin's Lymphoma

Retinoblastoma

Soft Tissue Sarcoma

Sarcomas Gonadales

Estragonadales Germinal Tumors

Various germinal tumors

Carcinomas

Histiositosis

Extension of coverage of pediatric pathology

Cardiac Congenital Malformations

Esophageal atresia

Omphalocele

Gastrochisis

Atresia / Duodenal Stenosis

Atresia Intestinal

Atresia Anal

Hypoplasia / Renal Dysplasia

Ureter Retrocavo

Ectopic Meatoses

Ureteral Stenosis

Ureterocele

Vesical Extrophy

Hypospadias

Epispadias

Ureteral Stenosis

Ureteral Meat Stenosis

Spina Bifida

Appendix A9: Placebo tests on the impact of the Seguro Popular on medical assistance

To estimate the impact of the *Seguro Popular* on medical assistance, we have made the assumption that, without the *Seguro Popular*, a shock on potentially covered cases would have had the same effect in municipalities that implemented and did not implement the *Seguro Popular*. Hereafter, we have run placebo tests that corroborate this assumption.

More precisely, we define a variable of interest that corresponds to the share of people dying from diseases included in the 2010 lists of the schemes on day d, of month m and year t in municipality i. We call this share the share of *virtually covered cases* would the *Seguro Popular* be implemented in municipality i with the 2010 rules. We estimate the correlation between this variable, medical assistance and mortality using the IV approach presented in section 6. However, we reduce our sample to the 1998-2003 period, i.e. before the implementation of the schemes.

We estimate separate effects of virtually covered cases on medical assistance between a control group and a placebo group. The placebo group is made of the municipalities that first implemented the schemes by 2004. The control group is made of all the other municipalities. In this fashion, we can check if virtually covered cases had a different impact on medical assistance in the set of municipalities that implemented the *Seguro Popular* early on (i.e. in 2004) as compared to all the other municipalities. We find that, before the implementation of the schemes, a shock on the set of virtually covered cases had the same impact on medical assistance between the placebo and the control group. This finding is presented in Table A9.1. In column 1, we did not use any control variable, whereas a full list of demographic controls was added in column 2.

In complement, we redefine the placebo and control groups to include in the placebo group the municipalities that joined the *Seguro Popular* by 2005 and 2006 (instead of 2004). Results are similar with these changes of definition of the placebo group. More precisely, the regressions in Table A9.1, columns 3-6, show that there is no statistical difference of the impact of virtually covered cases on medical assistance between the placebo and control groups. Out of six regressions, we find a statistically significant (at 10%) effect of potentially covered cases in the placebo group (%) in only one specification (see Table A9.1, column 5). This coefficient is no longer statistically significant when sociodemographic controls are added (see Table A9.1, column 6). This is consistent with the principle that, out of 10

regressions, a variable of interest with zero impact could still display a statistically significant coefficient (at 10%) in one of them.

<u>Table 10: The impact of virtually covered cases in control and placebo groups on medical assistance (and mortality) before the implementation of the schemes (1998-2003)</u>

	(1)	(2)	(3)	(4)	(5)	(6)
Year of reference (to construct placebo group)	2004	2004	2005	2005	2006	2006
Independent variables	2SLS	2SLS	2SLS	2SLS	2SLS	2SLS
First stage results: Medical assistance (%)						
Virtually covered cases (%) x municipality in placebo group (%)	0.12*** (0.003) 0.006 (0.01)	0.09*** (0.003) 0.02 (0.01)	0.12*** (0.005) -0.01 (0.01)	0.10*** (0.004) -0.01 (0.01)	0.12*** (0.006) -0.01* (0.01)	0.09*** (0.005) -0.01 (0.01)
Stock-Yogo weak identification test (Kleibergen-Paap rk Wald F statistic: passed)	0.2	1.7	2.2	2.1	3.1	0.9
Second stage results: Daily mortality (per 100,000)						
Medical assistance (%)	-0.63 (3.82)	-0.69 (1.43)	1.03 (0.96)	1.01 (1.05)	-0.40 (0.65)	-0.48 (1.40)
Potentially covered cases (%)	0.03 (0.44)	0.03 (1.30)	-0.17 (0.11)	-0.13 (0.10)	-0.001 (0.01)	-0.01 (0.13)
Sociodemographic controls (†)	No	Yes	No	Yes	No	Yes

Notes: The table reports IV results obtained from fixed effect regressions with by municipality by month by year fixed effects. The dependent variable is daily mortality rate at the municipality level. The regression is weighted by the population in each municipality. One, two and three stars respectively mean statistically significant at 10%, 5% and 1%. Standard errors are clustered at municipality level. †: The sociodemographic controls included in the specification of columns 2, 4 and 6 are: the share of women, the share of deceased people (for municipality *i* and day *d*) falling into 9 age groups, 22 occupation categories (as provided in the Mexican death records), 7 educational attainment categories and 5 types of social security affiliation. We also control for the share of deceased that were residing in rural areas within municipality *i*. We furthermore allow each of these controls to have a different impact on municipalities that implement and do not implement the schemes with the corresponding interaction parameters.

Appendix A10: Impacts of the Seguro Popular by season of the year

We now turn our attention to assessing if the extension of medical coverage has reduced weather vulnerability. To start with, we have looked at the distribution of the impacts of the *Seguro Popular* across seasons using the same IV setting. Recall that our IV strategy used eligibility to the schemes to instrument for medical assistance. From a theoretical perspective, the second stage results that we obtain identify the effect of medical assistance on mortality for the people who are induced into receiving medical assistance thanks to their eligibility to the schemes. If eligibility to the schemes has a strong effect on cold-related diseases, then the estimated coefficient for medical assistance should be particularly high for the winter period, since eligibility to the schemes should primarily impact excess winter mortality.

Results for our IV regression by season of the year are provided in the table A10.1. The coefficient for medical assistance is strong and statistically significant for the winter period (Dec. – Feb.) whereas it is smaller in magnitude and statistically insignificant for all the other seasons. Therefore, the *Seguro Popular* seems to have mostly reduced extra winter mortality.

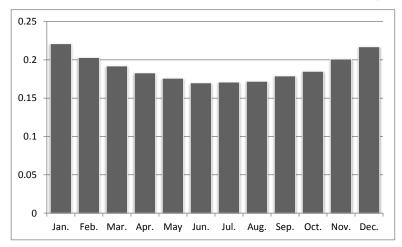
<u>Table A10.1: The impact of medical assistance before death on reducing mortality, according to the period of death</u>

Independent variables	Dec. – Feb.	Mar May	Jun. – Aug.	Sep. – Nov.
Medical assistance (%)	-3.32***	0.63	-0.60	-0.11
	(1.25)	(0.68)	(0.48)	(0.54)
Potentially covered cases (%)	0.29**	-0.12*	0.02	0.04
	(0.13)	(0.07)	(0.05)	(0.05)
Stock-Yogo weak identification test	12.4	15.7	28.9	19.3
(Kleibergen-Paap rk Wald F statistic: passed)				

Note: The table reports IV results obtained from fixed effect regressions with by municipality by month by year fixed effects. The dependent variable is daily mortality rate (per 100,000 inhabitants) at the municipality level. The regression includes municipality-by-year-by-month fixed effects. The instrument corresponds to the share of deceased people that were eligible to the schemes. One, two and three stars respectively mean statistically significant at 10%, 5% and 1%.

The strong impact of the schemes on winter mortality is consistent with the type of diseases covered, in particular diabetes and respiratory diseases such as pneumonias. In fact, people die more from the list of diseases covered by the schemes during the coldest months (see Figure A9.1 below). Therefore, we could have expected that an increase in health coverage for these diseases would reduce extra winter mortality.

Figure A9.1: Seasonality in the daily mortality rate from the diseases covered by the Seguro Popular and the Fondo de Protección contra Gastos Catastróficos (2004-2010)



Note: The y-axis represents the daily municipal mortality rate in deaths per 100,000 inhabitants. Only the causes of death covered by the schemes are included. Municipal rates have been weighted by the population of each municipality.

ROBUSTNESS CHECKS

Appendix B1: Separating the impact of minimum and maximum temperatures

We have correlated mortality with the average temperature in a day. We have therefore averaged minimum and maximum daily temperatures: the same effect at a given average temperature level is assumed and no consideration is made for within-day variation. To investigate this issue, we can run a specification of the distributed lag model where we calculate separate effects for minimum and maximum temperatures. No additional insight seems to be brought by such an exercise.

Figure B1.1: Impact of minimum daily temperature on 31-day cumulative mortality, in deaths per 100,000 inhabitants.

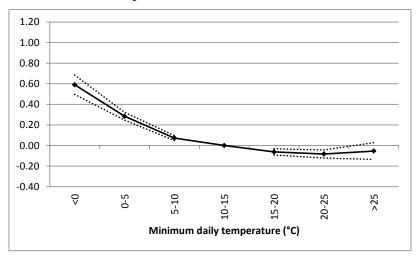
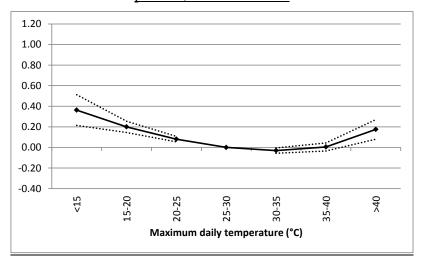


Figure B1.2: Impact of maximum daily temperature on 31-day cumulative mortality, in deaths per 100,000 inhabitants.

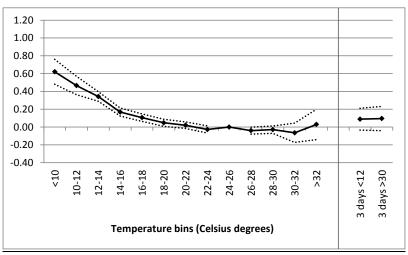


Notes. Lines in dash correspond to the 95% confidence interval values obtained for each sum of estimated coefficients. The regression results control for the day's precipitation level. The coefficients displayed Figure B1.1 and Figure B1.2 have been estimated jointly and come from the same fixed effect regression.

Appendix B2: Consecutive hot and cold days

One might fear that mortality might increase if extremely high or low temperatures remain for more than a day. We have tested this by adding two additional bins to the base specification: the first bin is equal to 1 if the last three days have undergone an average temperature below 12°C. The second bin is equal to 1 if the last three days have suffered from an average temperature above 30°C. The model then includes all the remaining temperature bins and 28 lags. The cumulative effect of the two new bins is positive but not statistically significant, suggesting that the base model with no such bins is a sufficient depiction of the relationship between mortality and temperatures.

Figure B2.1: Impact of consecutive hot and cold days on 28-day cumulative mortality, in deaths per 100,000 inhabitants.

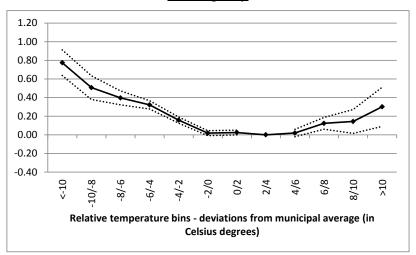


Appendix B3: Using relative temperature bins

Due to acclimatization, the mortality effect of the same cold or a hot day may differ from one location to another. As a robustness check for our mail model, we run a series of specifications where we assume that the impact of temperature on mortality depends on the difference between the temperatures faced during a given day and the ones that are usually experienced: instead of using absolute temperature bins, we calculate deviations from the average temperature in each location to construct relative temperature bins with a 2°C window. The average temperature in each municipality is obtained by averaging all daily temperatures over 1997-2013. Then we rerun our distributed lag model with the newly constructed temperature bins. These include deviations between -10°C and +10°C with respect to the average of each municipality.

The 31-day cumulative results for all the population and causes of deaths are displayed in Figure B3.1.

Figure B3.1: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants, using temperature bins relative to the average temperature in each municipality



Notes. Lines in dash correspond to the 95% confidence interval values obtained for each sum of estimated coefficients. 312,140 groups and 30.1 observations per group. The regression results control for the day's precipitation level.

They show little difference with the results obtained using absolute temperature bins, even though the effect of cold is smaller whereas the effect of heat is bigger. When accounting for the frequency of unusually cold and hot days, we find that days with temperatures below the average by 10°C or more would be responsible for the death of around 2,700 people annually (95% confidence interval is 2,200-3,200). Mild cold (deviations between -2°C to -10°C) are imputed the death of 26,700 people (95% confidence interval is 23,600-29,700). On the other hand, unusually hot days – above the average by 10°C or more – would cause around 350

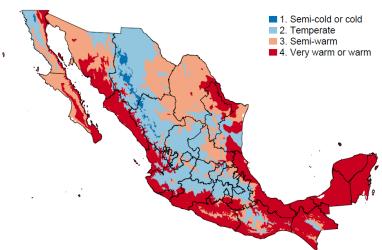
deaths (95% confidence interval is 100-600). We also find statistically significant effects for days with temperatures between 6°C and 10°C above the municipal average: they would be responsible for the death of around 1,500 people (95% confidence interval is 900-2,200).

In addition, we have run the distributive lag model for every subgroup and type of disease covered in this analysis and found results very similar to the ones uncovered with the distributive lag model.

Appendix B4: The temperature-mortality relationship by climate region

Mexico is a large country with very diverse climates. Due to adaptation and acclimation, it is likely that the temperature-mortality relationship is different in the hottest regions are compared to the coldest ones. The INEGI provides a detailed map of Mexico with a typology of 21 climates. We have simplified this typology and broken down Mexico into 4 climate categories (see Figure B4.1): very warm and warm (covering very dry, dry, semi-dry, humid and semi-humid regions that are also very warm and warm); semi-warm; temperate; and semi-cold or cold (covering respectively all semi-warm, temperate, semi-cold or cold regions independently of humidity).

Figure B4.1: Map of Mexico distinguishing between warm, semi-warm, temperate and cold <u>climates</u>



By overlapping the map of Mexican climates with the map of Mexican municipalities, we have matched the boundaries of the 2,456 Mexican municipalities with the boundaries of our four climatic categories. The map of INEGI defines municipalities as a set of data points that produce a polygon.

Our matching strategy assigns a climate to each point of the polygon that corresponds to the boundaries of a municipality. For each municipality, we calculate the number of delimiting data points that fall in a given climate. If this number exceeds half the total number of data points that constitute the boundary of the municipality, we consider that this municipality belongs to this climate.

This approximation allows us to classify municipalities into four main climate categories, for which we run the distributed lag model separately. The output of the separate regressions is provided below. In cold regions, we find very strong mortality impacts of hot days. However,

the sample is of limited scope since it includes only 11 municipalities. On the other hand, we find no striking difference between warm, semi-warm and temperate regions. However, the confidence interval for the impact of cold days on mortality is large in hot regions. This is likely to be due to the lack of cold events in hot regions.

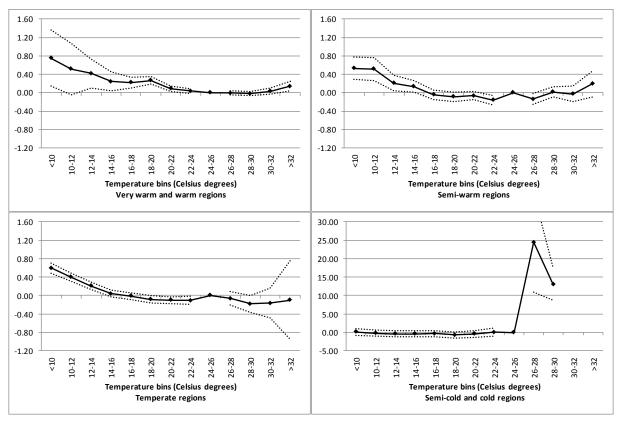


Figure B4.2: Mortality impacts by climate region in Mexico

Appendix B5: Polynomial model and interactions with precipitations

We have found a J-shaped temperature-mortality relationship in this paper, in which cold days have a stronger effect than hot days on mortality. Instead of using temperature bins, we can proxy this relationship using a polynomial form to describe the relationship between mortality and average temperatures:

$$Y_{i,d,m,t} = \sum_{k=0}^{K=30} \sum_{a=1}^{A} \theta_{-k}^{a} . T_{i,d-k,m,t}^{a} + \sigma . P_{i,d,m,t} + \mu_{i,m,t} + \varepsilon_{i,d,m,t}$$

In the equation above, we have replaced the temperature bins by the average temperature for municipality i in day d minus k of month m in year t, denoted $T_{i,d-k,m,t}$. We can then consider a nonlinear relationship by including $T_{i,d-k,m,t}^2$, $T_{i,d-k,m,t}^3$ and so on in the equation. We restrict ourselves to the case where A = 3.

Using a polynomial function instead of temperature bins reduces the amount of coefficients to be estimated by the model and therefore its computational intensity. We take advantage of this fact to better control for the confounding effect of precipitations, and interact precipitations with temperatures. To do so, we also include lagged precipitations over 30 days and assume a polynomial relationship between precipitations and temperatures. We also interact precipitations with temperatures.

Estimated coefficients are provided in the table below for three different specifications. They show very little difference in the estimated impacts and the cumulative effects of precipitations are not statistically significant. A polynomial form for the relationship between temperatures and mortality is preferred to a linear form.

<u>Table B5.1: Cumulative 31-day impact of temperatures and precipitations using polynomials</u> as functional forms in the fixed effect linear model

31-day effects	Dependent variable: daily mortality rate					
Tamparatura	-0.18***	-0.32***	-0.17***			
Temperature	(0.02)	(0.02)	(0.02)			
C guyana d	0.0048***		0.0041***			
Squared	(0.0012)		(0.0012)			
To the cube	-0.000035**		-0.000022			
To the cube	(0.000019)		(0.00002)			
Precipitations		0.02	0.0007			
Frecipitations		(0.03)	(0.0041)			
C guyana d			-0.000001			
Squared			(0.000001)			
To the cube			0.000000001			
To the cube			(0.000000001)			
Interaction:		-0.001	(0.0003)			
Temperature x precipitation		(-0.001)	0.0002			

Notes. Unit is deaths per 100,000 inhabitants. The first specification only controls for on-the-day precipitation levels.

Appendix B6: Confounding effect of humidity

Barreca (2012) shows that humidity interacts with temperatures in a way that can slightly alter mortality estimates, along with and their geographical distribution. In the regressions below, we use a specification similar to Deschenes and Moretti (2009) and introduce evaporation levels as an additional control variable with 30 lags. We also interact it with temperature bins. Table B6.1 displays the results obtained.

We find that mortality due to heat is higher under dry climates. Our results therefore do not match those of Barreca (2012) who find that higher humidity lead to higher mortality in hot and humid regions.

<u>Table B6.1: Impact of humidity on mortality using a specification similar to Deschenes and</u>

Moretti (2009)

Independent variables	(1)	(2)	(3)	(4)
Temperature <10° C	0.56*** (0.039)		0.97*** (0.106)	
Temperature >32° C		0.22*** (0.049)		-0.11 (0.216)
24h evaporation (in mm):	-0.02*** (0.003)	-0.03*** (0.003)	-0.02*** (0.003)	-0.03*** (0.003)
x Temperature <10° C			-0.13*** (0.031)	
x Temperature >32° C				0.04 (0.024)

Notes. Dependent variable is mortality per 100,000 inhabitants. All population and causes of death are considered. Reported effects are cumulated effects over 31 days. Standard errors in brackets. *** indicates statistically significant at the 1% level, ** at the 5% level, and * at the 10% level.

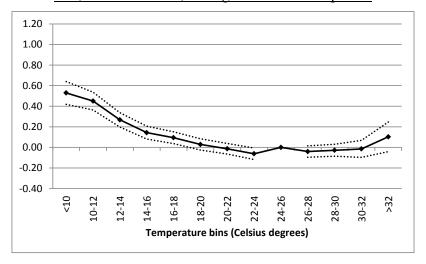
Appendix B7: Splitting the sample into two periods

We report below the results of the distributed lag model for all the population and all causes of death, splitting the sample into two periods: 1998-2003 and 2004-2010. As one could expect, the temperature-mortality relationship seems less strong in the later period, probably due to an improvement of living conditions. The two coefficients for temperatures below 12°C are statistically different between both periods.

1.20 1.00 0.80 0.60 0.40 0.20 0.00 -0.20 -0.40 10-12 12-14 14-16 20-22 16-22-Temperature bins (Celsius degrees)

Figure B7.1: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants, during the 1998-2003 period

Notes: Lines in dash correspond to the 95% confidence interval values obtained for each sum of estimated coefficients. The regression results control for the day's precipitation level.



<u>Figure B7.2: Impact of temperature bins on 31-day cumulative mortality, in deaths per</u> 100,000 inhabitants, during the 2004-2010 period

Notes: Lines in dash correspond to the 95% confidence interval values obtained for each sum of estimated coefficients. The regression results control for the day's precipitation level.

Appendix B8: Separate effects for week days and weekends

The figures below provide the 31-day cumulative mortality estimates for hot and cold days, depending on whether they fell during a week day or the weekend.

Figure B8.1: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants, for events occurring during week days

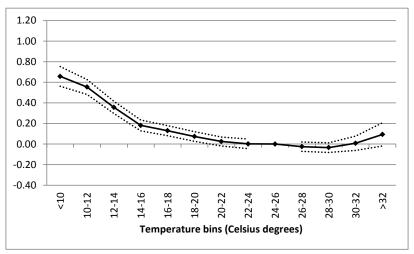
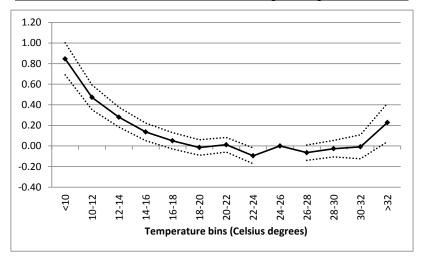


Figure B8.2: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants, for events occurring during the weekend



The estimates are roughly the same (no statistical difference for any temperature bin). This is consistent with the fact that most temperature-related deaths concentrate on the elderly.

Appendix B9: Effects in rural versus urban areas

We look here if short-run vulnerability to temperatures may differ between people living in urban areas vs. people living in rural areas. Results are displayed on Figure B9.1 and Figure B9.2. Even though the coefficient associated with day below 10°C in rural areas is above the one for urban areas, the curves are not statistically different from one another. This suggests that distance to city centres might have no strong implications for short-term weather-related mortality in the case of Mexico.

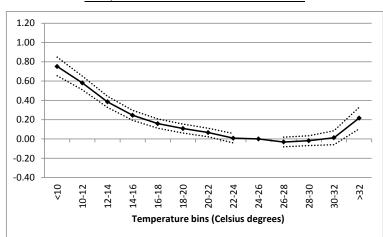


Figure B9.1: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants in urban areas

Notes. The graph shows the cumulative effect of a day with a temperature within each bin based (relative to the $24-26^{\circ}$ C category) obtained from a dynamic model with 30 lags run for populations living in urban areas only. The diamonds show the sum of the coefficients on these thirty lags in each category. Dashed lines correspond to the 95% confidence interval. The dependent variable is daily mortality rate at the municipality level. The regression controls for daily precipitation level and includes a range of municipality-by-year-by-month fixed effects.

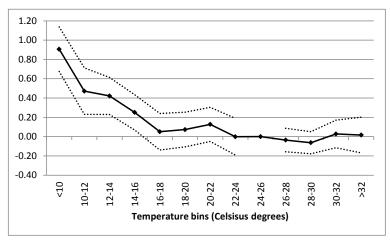


Figure B9.2: Impact of temperature bins on 31-day cumulative mortality, in deaths per 100,000 inhabitants in rural areas

Notes. The graph shows the cumulative effect of a day with a temperature within each bin based (relative to the 24-26°C category) obtained from a dynamic model with 30 lags run for populations living in rural areas only. The diamonds show the sum of the coefficients on these thirty lags in each category. Dashed lines correspond to the 95% confidence interval. The dependent variable is daily mortality rate at the municipality level. The regression controls for daily precipitation level and includes a range of municipality-by-year-by-month fixed effects.

Appendix B10: Results with the exact same method as in Deschenes and Moretti (2009)

There are two differences between the distributed lag model used in this paper and the one by Deschenes and Moretti (2009). First, these authors use total population as a weight and not the square root. This is because they put more emphasis on the results to be representative of the total population whereas we do not want them to be too much dependent on estimates for only a few big cities. In fact, either using total population or the square root has not significant impact on the results.

Second, instead of using temperature bins, Deschenes and Moretti (2009) use as independent variables: a) either a dummy variable which take the value of 1 on unusually cold days (average temperature <20°F or <30°F, depending on specification) and its lags; or b) a dummy variable which take the value of 1 on extremely hot days (average temperature >80°F or >90°F, depending on specification) and its lags. They therefore calculate the impact of unusually cold or hot days on mortality separately and as compared to the impact of any other day in the year. We chose not to do so because epidemiological studies (as cited previously) show that the temperature-mortality relationship very often is a U- or V-shaped function. At the bottom, there is a threshold (e.g. 20°C) with very low mortality. The more temperatures depart from this threshold, the more mortality increases. Therefore, evaluating the impact of extremely hot or cold temperatures as compared to the impact of any other day in the year is not ideal because not only extremely hot/cold days lead to extra mortality with respect to the bottom threshold. This method is likely to systematically underestimate the impact of temperatures on mortality, because it does not take the days with least temperature stress as a reference for calculating extra mortality.

We reproduce here the results obtained with the method of Deschenes and Moretti (2009), limiting ourselves to the ones for all causes of death. Using such a methodology leads to the same persistent effects of cold and heat on mortality using our dataset. The magnitude of the effects is also similar.

<u>Table B10.1: Cumulative 31-day impact of extraordinarily cold and hot days on daily mortality, deaths for 100,000 inhabitants by subgroup</u>

	Daily average temperature			
Group	<10° C	>32° C		
T. 4.1	0.6***	0.13***		
Total	(0.036)	(0.037)		
Man	0.6***	0.18***		
Men	(0.05)	(0.056)		
Waman	0.59***	0.08*		
Women	(0.04)	(0.045)		
A and O. 4	0.63***	-0.04		
Aged 0-4	(0.091)	(0.077)		
Aged 4-9	0.05**	0.01		
Ageu 4-9	(0.02)	(0.023)		
Aged 10-19	-0.01	0.04		
Aged 10-19	(0.02)	(0.028)		
Aged 20-34	0.09***	0.06		
Aged 20-34	(0.031)	(0.045)		
A god 25 44	0.22***	0.08		
Aged 35-44	(0.059)	(0.074)		
A J 45 54	0.42***	0.32***		
Aged 45-54	(0.098)	(0.117)		
A and 55 64	1.02***	0.32		
Aged 55-64	(0.208)	(0.218)		
A god 65 74	1.71***	0.53		
Aged 65-74	(0.35)	(0.362)		
A and 75 :	14.5***	0.88		
Aged 75+	(0.989)	(1.036)		

Notes. Standard errors in brackets. *** indicates statistically significant at the 1% level, ** at the 5% level, and * at the 10% level. Unit is deaths in a day per 100,000 inhabitants (in subgroup). All causes of death are considered.

Appendix B11: Distributed lag model with 60 lags

The results after 60 lags are consistent with the results at 30 days: the estimated coefficient for unusual cold with 60 lags is not statistically different from the estimated coefficients with 30 lags. The impact of hot days is however no longer statistically significant. If any, the impact of hot days is therefore expected to be rather small as already predicted with the model with 30 lags.

<u>Table B11.1: Cumulative 61-day impact of extraordinarily cold and hot days on daily</u> mortality, deaths for 100,000 inhabitants

	Daily average temperature				
Group	<10° C	>32° C			
Total population	0.56*** (0.055)	-0.02 (0.057)			

Notes. Standard errors in brackets. *** indicates statistically significant at the 1% level, ** at the 5% level, and * at the 10% level. Unit is deaths in a day per 100,000 inhabitants (in subgroup). All causes of death are considered.

Appendix B12: Changing the structure of the fixed effects

In this Appendix, we use different structures for the fixed effects. In the base specification, we have used fully interacted, municipality-by-year-by-month fixed effects. This restrains the comparison of mortality effects to days within the same month of the year within a given municipality and disregards the fact that changes in temperature may affect seasonal patterns, and in turn mortality. Above all, we could underestimate the mortality impacts of direct exposure to temperature in very cold or very hot months by comparing very cold days with already cold days, and very hot days with already hot days within a month. To the contrary, we find that relaxing the controls for within-municipality seasonal patterns attenuates estimated impacts (Table B12.1, columns 1-3). This attenuation is likely to be due to an estimation bias. When we allow the comparison of mortality impacts to take place within a municipality and a given month, but across different years, results are similar to the base specification, suggesting that the base specification does not underestimate the impact of hot and cold days on mortality (Table B12.1, column 4).

Table B12.1: Impact of days below 10°C and above 32°C using different sets of fixed effects

Fixed effects	Base specification	(1)	(2)	(3)	(4)
Municipality x Year x Month	X				
Municipality		X	X		
Year		X	X		X
Month		X			
Municipality x year				X	
Municipality x month					X
Climatic region x month			X	X	
D :11	0.70***	0.38***	0.32***	0.43***	0.71***
Day with temperature <10° C	(0.04)	(0.03)	(0.03)	(0.02)	(0.04)
Day with temperature >32° C	0.13**	0.10***	0.09***	0.11***	0.12**
Day with temperature >32 C	(0.05)	(0.03)	(0.02)	(0.02)	(0.05)

Note: Dependent variable is mortality per 100,000 inhabitants. Standard errors in brackets. *, ***, ***: statistically significant at 10%, 5% and 1%. Reference day is 24-26 Celsius degrees. The impact of hot and cold days are estimated using different regressions.

SOCIOECONOMIC STATUS

Appendix C1: Impact of temperatures by profession

The distributed lag model has been run separately according to the profession of the deceased to assess differences in vulnerability by profession. Results are summarised on the figures below.

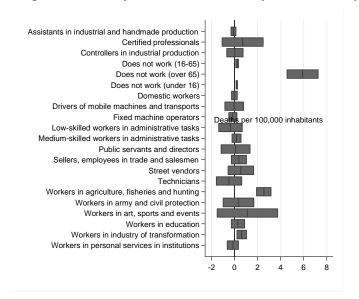
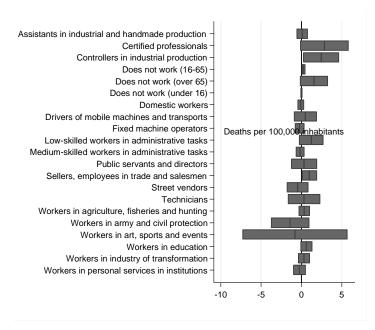


Figure C1.1: Impacts of 31-day cumulative mortality for a cold day below 10°C

 $\textbf{Notes.} \ \text{The grey areas represent the 95\% confidence interval for each estimated coefficient.}$





Notes. The grey areas represent the 95% confidence interval for each estimated coefficient.

Appendix C2: Regression output to predict income with 2000 Census

The regression run to predict personal income with a sample of 2000 Mexican Census is a panel data regression which includes a long list of fixed effects, in particular municipality fixed effects separately estimated for people living in the rural/urban part of a municipality. Table C2.1 provides a brief description of the regression used and just a couple of coefficients (as examples) for age and gender.

Table C2.1: Regression used to predict income levels

Dependent variable	Log(Personal income)		
Age	-0.0089***		
Age squared	(0.0002) 0.0001*** (0.00001)		
Female	0.0033** (0.0014)		
Fixed effects			
Civil status	Yes		
Occupation	Yes		
Social security affiliation	Yes		
Educational level	Yes		
Municipality and rural/urban area	Yes		
Interactions:			
Civil status x gender	Yes		
Occupation x age	Yes		
Occupation x age squared	Yes		
R2	0.44		
Number of observations	8,756,128		

Notes. Standard errors in brackets. *, **, ***: statistically significant at 10%, 5% and 1%.

Appendix C3: Income quartiles and death causes

Table C3.1: Mortality rates and causes of death by quartile

Cause of death	1 st quartile	2 nd quartile	3 rd quartile	4 th quartile	All quartiles			
Daily mortality rate	1.4	1.54	1.25	1.05	1.3			
Mortality of each income quartile by cause of death (as % of all deaths in quartile)								
Respiratory system diseases	10%	9%	8%	7%	9%			
Circulatory system diseases	24%	24%	24%	25%	24%			
Endocrine, nutritional and metabolic diseases	14%	18%	18%	16%	16%			
Infectious diseases	4%	3%	3%	4%	4%			
Neoplasms	11%	12%	14%	17%	13%			
Accidents and violent deaths	13%	10%	11%	12%	11%			
All other deaths	24%	23%	22%	19%	22%			

Note: The daily mortality rates are in deaths per 100,000 inhabitants.

<u>Table C3.2: Importance of specific diseases within the three categories of weather-sensitive causes of deaths</u>

Cause of death	1 st quartile	2 nd quartile	3 rd quartile	4 th quartile	All quartiles			
Endocrine, nutritional and metabolic diseases:								
Diabetes mellitus	73%	84%	88%	88%	83%			
Malnutrition	21%	10%	6%	5%	10%			
Circulatory system diseases:	Circulatory system diseases:							
Hypertensive diseases	12%	12%	12%	10%	11%			
Ischaemic heart diseases	41%	46%	49%	52%	47%			
Heart failure	11%	7%	5%	4%	7%			
Cerebrovascular diseases	27%	25%	24%	23%	25%			
Respiratory system diseases:	Respiratory system diseases:							
Pneumonia, (organism unspecified)	29%	28%	28%	28%	28%			
Other acute lower respiratory infections	4%	4%	4%	3%	4%			
Chronic lower respiratory diseases	52%	51%	49%	47%	50%			
Other respiratory diseases principally affecting the interstitium	4%	6%	7%	9%	6%			

Note: Percentages correspond to the share of deaths entailed by a specific disease within its category, e.g. 73% of deaths from endocrine, nutritional and metabolic diseases are due to diabetes mellitus for the 1st quartile of income.

Appendix C4: Effect of temperature on mortality by quartiles defined with a poverty indicator

Instead of using income levels to create quartiles of population, we can use alternative metrics of wellbeing and living conditions. Below, we use a composite indicator inspired from the marginality index of the Mexican Council of Population (CONAPO).

The index of the CONAPO classifies localities according to their degree of marginality (from very low to very high) and has been used by government to design social policies. The indicator of the CONAPO relies on eight variables available from the Mexican censuses. The Council calculates 1) the share of the population of aged 15 or more who is analphabetic; 2) the share of the population of aged 15 or more who did not complete primary education; 3) the average number of occupants per room; 4) the share of households without exclusive toilet; 5) the share of households without electricity; 6) the share of households without current water within their property; 7) the share of houses or flats with earthen floor; and 8) the share of houses or flats with no refrigerator.

We construct an individual-specific poverty indicator based on the features used by CONAPO to classify localities by level of marginality. Since we want an indicator which is equally reflective of poverty for children and adults, we only consider the last five characteristics listed above (4-8): children under a certain age are necessarily analphabetic and cannot have completed primary education. Likewise, a relatively high amount of occupants per room has not exactly the same relevance in terms of living conditions if these include small kids.

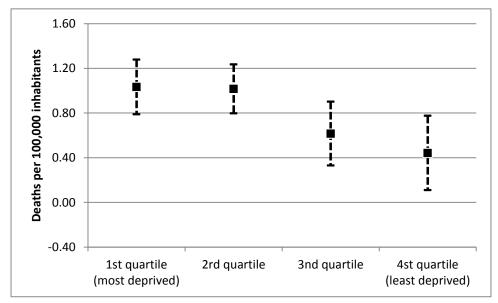
We compute an exclusion indicator that range from 0 (no exclusion) to 5 (strong exclusion) for each individual in the Census. If an individual belongs to a household that has exclusive toilets, electricity, current water, a proper floor (not an earthen one) and a refrigerator, then the poverty indicator equals 0. If one of these elements is missing, the indicator is equal to one; if two of these elements are missing, the indicator is equal to two; and so on. The maximum value of 5 is given to households that have no exclusive toilets, no electricity, no current water, an earthen floor in the house and no refrigerator. These are obviously consistent with very precarious living conditions.

Once the indicator has been computed for each person in the 2000 Census, the exact same methodology is applied as for income to create quartiles. In short, we run a linear regression to predict the value taken by the poverty indicator based on a series of observables that are both present in the Census and in the mortality data. We then make out-of-sample predictions

of the indicator on the deceased to proxy living conditions at the moment of death. Then we separate the population of the deceased and the living in four groups (from low to high living conditions) and run the econometric model separately for the four groups of people.

The results of such process are presented below and confirm higher vulnerability for poorer households.

Figure C4.1: Impact by quartiles (based on poverty indicator) of a cold day below 10°C on cumulative 31-day mortality



 $\textbf{Notes.} \ \text{The grey areas represent the 95\% confidence interval for each estimated set of coefficients}.$

Appendix C5: Age-corrected regressions by income quartiles

Figure C5.1 below presents the full results of the age-corrected regressions by income quartiles for all causes of death. Statistically significant impacts are found for all temperature bins below 16°C for the 1st income quartile, and below 20°C for the 2nd income quartile. Results stop being statistically significant after 14°C in the case of the 3rd and 4th income quartiles.

1.60 1.60 0.80 0.80 0.00 0.00 -0.80 -0.80 -1.60 -1.60 20-22 22-24 18-20 22-24 20-22 1st quartile 2nd quartile 1.60 1.60 0.80 0.80 0.00 0.00 -0.80 -0.80 -1.60 -1.60 12-14 16-18 18-20 20-22 28-30 16-18 18-20 24-26 10-12 14-16 22-24 24-26 26-28 30-32 10-12 14-16 22-24 26-28 28-30 30-32 20-. 3rd quartile 4th quartile

Figure C5.1: Impact of temperature bins on cumulative 31-day mortality, by income quartile after correcting for differences in age structure across quartiles

Notes. The results for each quartile are taken from separate regressions. Deaths have been weighted to account for differences in the pyramid of ages across quartiles, taking the first quartile of income as a reference. The dependent variable is the mortality per 100,000 inhabitants belonging to the quartile. The y-axis is mortality per 100,000 inhabitants and the x-axis corresponds to the cumulative impact after 31 days for each of the 2°C temperature bins in the regressions. The reference bin is 24-26°C. On-the-day precipitations are used as controls, along with by-month-by year-by municipality fixed effects. The dashed lines represent the 95% confidence interval for each estimated set of coefficients.

Appendix C6: Heating and cooling appliance ownership in the Mexican income and family expenditure surveys

We have gathered ownership data from the Mexican income and family expenditure surveys (*Encuestas Nacionales de Ingresos y Gastos de los Hogares*) for the following waves of the survey: 1998, 2000, 2002, 2004, 2005, 2006, 2008 and 2010. Questions are slightly different from one year to the other. In 1998 and 2000, the questions were:

- Do you have heaters in your home? If so, how many heaters do you have?
- Do you have air conditioning in your home? If so, how many air conditioners do you have?

In 2002, the questions were changed to distinguish central systems from individual appliances:

- Do you have central heating in your home?
- Do you have room heaters in your home? If so, how many?
- Do you have central air conditioning in your home?
- Do you have room air conditioners in your home? If so, how many?

For central heating and central air conditioning, respondents could answer "yes, of exclusive use", "yes, shared" or "no".

In 2004, 2005 and 2006, the same questions were asked but respondents could no longer precise if central heating or air conditioning was of exclusive use or shared. Answers were restricted to "yes" or "no".

In 2008 and 2010, only two questions were asked:

- Does this house/flat have heating?
- Does this house/flat have air conditioning?

Respondents could respond either "yes" or "no".

We provide summary statistics on appliance ownership at national level in Table C6.1. A further breakdown by Mexican State is provided hereafter.

Table C6.1: Ownership rates of heating and cooling appliances by income quartile

	Heaters and/or heating systems				Heaters and/or heating systems			ing systems
Quartile	1st	2nd	3rd	4th	1 st	2nd	3rd	4th
1998	0.6%	1.2%	2.0%	6.4%	2.3%	4.2%	8.5%	18.0%
2000	0.9%	1.5%	2.6%	8.6%	3.9%	7.8%	9.4%	22.0%
2002	4.0%	3.2%	3.9%	8.6%	4.8%	6.4%	10.4%	21.3%
2004	0.7%	1.2%	2.7%	9.7%	4.1%	7.9%	14.1%	27.1%
2005	0.6%	1.3%	3.0%	9.7%	3.8%	8.5%	14.3%	28.3%
2006	0.9%	1.7%	3.2%	10.7%	4.2%	8.9%	14.4%	29.2%
2008	0.1%	0.6%	1.1%	4.5%	4.9%	8.7%	13.7%	25.7%
2010	0.2%	0.5%	1.5%	5.1%	5.8%	9.6%	14.9%	28.1%
All years	1.0%	1.4%	2.5%	7.9%	4.2%	7.8%	12.5%	25.0%

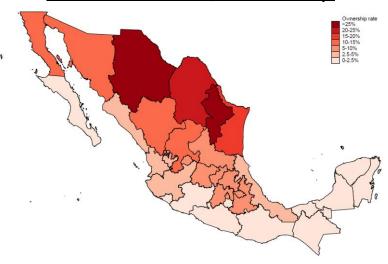
Note: the left hand-side panel reports the ownership rate of heaters and/or heating systems by income quartile as reported in the Mexican national income and expenditure survey; the right hand-side panel reports the ownership rate of air conditioners and/or air conditioning systems. The questions vary from year to year, explaining differences in average level between years. For example 2002-2006 surveys distinguish between central and room appliances, others do not. See Appendix V for the exact questions. Consequently, it is unfortunately not possible to interpret the evolution across time by income quartile. Source: Encuesta Nacional de Ingreso y Gasto de Hogares, various years.

Figure C6.1: Share of households from 1st income quartile declaring that they have at least one heater in their house (2004-2006 survey)



Source: Encuesta Nacional de Ingreso y Gasto de Hogares (2004, 2005, 2006).

Figure C6.2: Share of households from 4th income quartile declaring that they have at least one heater in their house (2004-2006 survey)



Source: Encuesta Nacional de Ingreso y Gasto de Hogares (2004, 2005, 2006).

Figure C6.3: Share of households from 1st income quartile declaring that they have at least one air conditioner in their house (2004-2006 survey)



Source: Encuesta Nacional de Ingreso y Gasto de Hogares (2004, 2005, 2006).

Figure C6.4: Share of households from 4th income quartile declaring that they have at least one air conditioner in their house (2004-2006 survey)



Source: Encuesta Nacional de Ingreso y Gasto de Hogares (2004, 2005, 2006).

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