

The Effects of Survey Nonresponse on Inferences

About Volunteer Work in the United States

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Volunteer work is an important component of economic activity in the United States (Abraham and Mackie 2005) and it also plays a key role in noneconomic spheres of American society (Wilson 2000). Who volunteers and why? What sorts of work do volunteers perform? How does volunteering affect volunteers? What cultural and political factors influence the amount and nature of volunteering in a society? How does volunteering shape culture and the polity? These are among the issues that make volunteering of interest across the social sciences.

As is true for many aspects of social life, information about volunteering comes largely from sample surveys. These surveys yield widely varying estimates of volunteering. For instance, the 2002 Supplement to the Current Population Survey (CPS), conducted by the Census Bureau, reported that 28 percent of American adults were volunteers (Boraas, 2003), whereas the 1996 National Household Education Survey (NHES), conducted by Westat, reported a figure of 39 percent (National Center for Education Statistics, 1997), and the 1996 Independent Sector (IS) survey, conducted by the Gallup organization, estimated 49 percent (Hodgkinson and Weitzman, 1996). These surveys differed in many ways (including the definition of volunteering), but one of the most pronounced was the degree to which they were successful in collecting information from all sample members. The response rate to the CPS supplement was 82 percent (Census Technical Documentation); the rate for NHES was 59 percent (National Center for Education Statistics, 1997); and, although we have been unable to obtain the rate for the IS survey, Gallup response rates rarely exceed 40 percent. The variation in these rates corresponds to the variation in the volunteering estimates: the higher the response rate the lower the volunteering estimate.

Although this is, at best, suggestive evidence for the proposition that survey estimates of volunteering are inversely related to survey response rates, such a proposition makes sense of a second puzzling survey finding about volunteering – change over time in volunteering appears

quite different from change over time in other forms of social capital. In *Bowling Alone*, Putnam (2000: 127) noted that “Trends in volunteering over the last several decades are more complicated and in some respects more intriguing than the uniform decline that characterizes most dimensions of social capital in America in this period.” As shown in Figure 1, nationwide Gallup surveys recorded a sharp *increase* between 1977 and 1991 in affirmative responses to the question “Do you, yourself, happen to be involved in any charity or social service activities, such as helping the poor, the sick or the elderly?” If these data are taken at face value, they indicate that volunteering in America almost doubled in little more than a decade! But over this same period, survey response rates declined markedly due to increasing difficulties in both contacting people and persuading them to be interviewed (see Curtin, Presser, and Singer 2005, for the experiences of one major survey). Response rates are unavailable for the Figure 1 surveys, but the Gallup Organization experienced a general decline during this period similar to that of most survey organizations. Further, the first five surveys shown in Figure 1 (from 1977 through 1986) were conducted face-to-face, whereas the other four (between 1987 and 1991) were conducted by telephone. Thus apart from the general decline in response rates, the later studies would have had lower response rates than the earlier ones because response rates are almost always lower on the phone than in-person (Groves and Kahn, 1979). Moreover, the 1988 survey, which produced the highest estimate of volunteering, almost certainly had the lowest response rate, as it was conducted over only two days compared to seven days for the 1987 survey and four days for all the other surveys. (Shorter field periods produce lower response rates because they allow less time to contact and persuade respondents.)

FIGURE 1 ABOUT HERE

Are there grounds for expecting that a survey’s response rate would affect its estimate of volunteering? Although the *potential* for nonresponse bias grows as the level of nonresponse

increases, several recent studies, on a wide range of topics, have found that *actual* nonresponse bias is largely unaffected by increases in the nonresponse rate. Keeter et al. (2000), Curtin, Presser, and Singer (2000), and Merkle and Edelman (2002) report little, if any, link between nonresponse rates and bias, and a meta-analysis by Groves (2006) shows only a weak relationship between nonresponse level and bias in studies that had validation measures. These results suggest that many of the variables measured in surveys are only slightly correlated, if at all, with the causes of nonresponse (cf., Groves, Presser, and Dipko, 2004 and Abraham, Maitland and Bianchi, 2006).

Volunteering is likely to be an exception to this pattern. The two major determinants of response rate -- contactability (the ease of being able to make contact with an individual) and amenability (an individual's willingness to be interviewed) -- both seem likely to be influenced by the same factors that influence volunteering: social integration, altruism, and a sense of responsibility. Indeed survey participation is similar to volunteering in that survey respondents are asked to help an organization by giving their time (Knack, 1992).¹

Nonresponse bias is not the only form of error that may affect survey estimates of volunteering. The estimates are also subject to recall error and social desirability bias. These two problems are minimized in the American Time Use Survey (ATUS), which, beginning in 2003, obtained detailed reports from a sample of Americans about how they spent their time on the day preceding the interview (Bureau of Labor Statistics and U.S. Census Bureau, 2005). Recall error is curtailed because of the very short reference period, and social desirability bias is unlikely because respondents are simply asked to report, in chronological order, everything they did yesterday. Only after respondents report all their activities does the interviewer ask whether

¹ In fact, Keeter et al.'s (2000: 140) methodological experiment found significantly more volunteering in a survey with a relatively low response rate than in a comparable survey with a much higher response rate. But in the context of the large number of comparisons made in that study (the great majority of which showed no difference), sampling error seemed a plausible interpretation for the finding.

any of them involved volunteering. Because the ATUS response rate has been only about 50 percent, however, the ATUS estimate of volunteering may be subject to considerable nonresponse bias.

In this paper, we assess the nonresponse error in the ATUS estimate of volunteering. This is possible because a random subsample of the ATUS was previously part of the Current Population Survey (CPS) Volunteering Supplement, which attains a much higher response rate than the ATUS (over 80 percent). We compare the CPS volunteering estimates for two groups of Supplement respondents: those who became ATUS respondents and those who became ATUS nonrespondents. We find a large difference consistent with the hypothesis that ATUS respondents are more likely than ATUS nonrespondents to be volunteers. In the CPS, ATUS respondents reported volunteering at a rate 75 percent higher than did ATUS nonrespondents.

We then examine the contribution of the various sources of nonresponse to the overall result, and find that each source (e.g., refusal, noncontact, and language barrier) is associated with less volunteering, though there is some variation in the magnitude of the effects. Further, the nonresponse bias is evident within demographic and other subgroups, which means that conventional statistical adjustments for nonresponse (e.g., weighting class adjustments and response propensity models) would not correct the bias.

Our next step is to investigate the impact of the nonresponse error on estimates of the correlates of volunteering activity. We do this by comparing models of volunteering in the CPS for ATUS respondents with those for ATUS nonrespondents. The results for the two groups are quite similar, suggesting that although nonresponse bias has a powerful effect on the univariate distribution of volunteering, it generally does not affect inferences about the causes of volunteering.

Finally, we show that predictions based on these CPS results are confirmed by the volunteering activity reported in the ATUS itself. The ATUS univariate estimate of volunteer hours is much higher than the CPS estimate but indistinguishable from the estimate in the CPS using only cases that later became ATUS respondents. In addition, inferences from the ATUS about the relation between volunteer hours and other respondent characteristics are similar to those from the CPS.

DATA

The data used in our analysis come from two sources: the September 2003 Current Population Survey Volunteering Supplement, and the 2003 and 2004 American Time Use Survey.

Current Population Survey (CPS) Volunteering Supplement

Every month, the U.S. Census Bureau conducts the Current Population Survey (CPS) with an area probability sample of almost 60,000 households in the 50 states and the District of Columbia. In order to facilitate state-level estimates, smaller states are oversampled. The majority of interviews are conducted by telephone but many are done in-person. The response rate to the main interview averages more than 90 percent.

A volunteering supplement has been administered to CPS sample members each September since 2002. The supplement is administered after the main interview, either in English or in Spanish, and asks about all household members age 15 and older. In the September 2003 supplement, information was collected for 95,337 individuals. The overall response rate (main interview response rate multiplied by supplement response rate) was 81 percent.

The first two questions in the 2003 supplement were:

Since September 1st of last year, have [you/NAME] done any volunteer activities through or for an organization? IF NO: Sometimes people don't

think of activities they do infrequently or activities they do for children's schools, or youth organizations as volunteer activities. Since September 1st of last year, have [you/NAME] done any of these types of volunteer activities?

Respondents who replied affirmatively were then asked for the number of organizations for which volunteer work was done, and the numbers of weeks and hours per week worked for each organization (or, if the respondent said the hours per week varied, the number of hours for the year).

American Time Use Survey (ATUS)

The American Time Use Survey (ATUS), an annual survey begun in 2003, is also conducted by the U.S. Census Bureau. The ATUS sample is chosen randomly from households participating in the CPS. CPS households are contacted for interviews in four consecutive months (Months in Sample, or MIS, 1-4); leave the sample for the next eight months; and then return for four additional months (MIS 5-8, which occur exactly one year after MIS 1-4).² In any given month, approximately one-eighth of the CPS sample is in its final month (MIS-8). The ATUS sample is selected from households that successfully complete the MIS-8 main interview (about 93 percent of the MIS-8 sample do so). One randomly chosen individual age 15 years or older from each selected household is designated for ATUS participation. The ATUS sample is selected proportional to a state's population (thereby offsetting the CPS oversampling of residents of small states). In addition, households with an Hispanic or non-Hispanic black householder, as well as households with children, are oversampled. We correct for the oversampling by using selection weights in our analyses.

² The CPS tracks dwelling units, not individuals or households. Thus, as individuals move into or out of the dwelling unit, they will move into and out of the CPS sample.

ATUS interviews are distributed across the weeks of the year, with random allocations to reporting days: one-quarter Saturdays, one-quarter Sundays, and one-half spread equally across the five weekdays. The interviews are administered by telephone, either in English or in Spanish. The roughly 5 percent of the sample for whom no telephone number is available are sent a letter asking them to call a toll free number on a specified day and offered a \$40 incentive for doing so. Respondents are asked to describe their primary activities, and how long each lasted, from 4:00 a.m. the previous day until 4:00 a.m. the interview day. After the 24 hours are accounted for, respondents are asked whether they did any “volunteer activities for or through an organization” during the day and those who say “No,” are prompted with “Sometimes people don’t think of activities they do for schools, or youth, or religious organizations as volunteer activities.” Anyone answering affirmatively is then asked “Which of the activities you told me about were volunteer activities?”

Attempts to contact individuals for the ATUS begin about two months after the household’s final CPS interview and continue for up to eight successive weeks. To minimize the lag between the September 2003 Volunteering Supplement responses and the ATUS response outcome, we use only the random subsample of cases that were MIS 5-8 in September 2003 (i.e., those whose final CPS interview took place in September, October, November or December 2003). Some of these 2003 Volunteering Supplement respondents were selected for the 2003 ATUS and some for the 2004 ATUS.

Of the 95,337 individuals 15 or older for whom information was collected in the September 2003 CPS Volunteering Supplement, there were 9,004 subsequently selected for the ATUS who were MIS-5 to MIS-8 in September 2003. The ATUS weighted response rate for

these cases was 57 percent.³ Accounting for the non-response to the CPS MIS-8 basic interview (as noted above, about 93 percent of the sample completed that interview), the effective response rate for the ATUS cases in our sample was about 53 percent. This is just slightly above the response rates for the entire ATUS.⁴

The Bureau of Labor Statistics provides joint ATUS-CPS files that contain identifying information from the ATUS and a battery of demographic and labor force participation items from the CPS final interview (MIS-8).⁵ We merged these variables with the appropriate year's ATUS interview files (2003 or 2004) and then linked the resulting file to the September 2003 CPS supplement file.

Although the results presented in this article are based on MIS 5-8 cases from the 2003 CPS Volunteering Supplement that became part of the 2003 or 2004 American Time Use Survey, we also conducted an analysis of cases from the 2002 CPS Volunteering Supplement that became part of the 2003 ATUS. That analysis yielded conclusions that are the same as those we present here.

METHODS

Our analysis begins with September 2003 CPS Supplement univariate and bivariate distributions of volunteer activity for individuals later selected for the ATUS. These analyses are augmented by multivariate models of the decision to volunteer and of the choice regarding how

³ Ignoring the unequal probabilities of selection (e.g., households with children were oversampled), yields an unweighted response rate for these cases of 56 percent.

⁴ In 2003, a total of 20,720 ATUS interviews were completed, for a weighted response rate of 54 percent. Budget cuts reduced the total number of interviews to 13,973 in 2004, for a weighted response rate of 53 percent. Accounting for nonresponse to the MIS-8 CPS interviews, these figures imply effective response rates of about 50 percent. The slightly higher response rate in our subsample dovetails with those cases having completed the September Volunteering Supplement, whereas other cases selected for the ATUS were either nonrespondents to the Supplement or not in the September sample. On the calculation of response rates, see Abraham, Maitland and Bianchi (2006). For other survey details see Bureau of Labor Statistics and U.S. Census Bureau (2005).

⁵ We examined the correlations between these variables as measured in the final month in sample and in the month of the Volunteering Supplement, and generally found little change. The one exception is labor force participation, which shows movement mainly between adjacent categories of hours worked.

many hours to volunteer. If the same factors determine the two decisions, a Tobit specification might be appropriate for explaining both outcomes (Tobin 1958). Formally, the Tobit model can be written:

$$y_i^* = x_i\beta + u_i, i = 1, 2, \dots, n$$

$$(1) \quad y_i = y_i^* \text{ if } y_i^* > \alpha$$

$$y_i = 0 \text{ if } y_i^* \leq \alpha$$

where y_i^* is a latent variable that is observable only when its value lies above a censoring threshold, y_i is the observed variable, x_i is a vector of explanatory factors, β is a vector of coefficients, α is the censoring threshold, and u_i is a residual that is assumed to be independent and identically distributed (i.i.d.) from a normal distribution with mean zero and variance σ^2 . In our case, y_i^* would be the latent propensity to volunteer, with actual volunteer hours y_i equal to y_i^* when $y_i^* > 0$ and equal to 0 otherwise. Although the Tobit specification has some appeal, a formal test strongly rejected its assumption that a single set of parameters can explain both the decision to volunteer and the number of hours volunteered.⁶

A more flexible alternative specification is a two-part model of the sort proposed by Cragg (1971). In the two-part model, a logit or probit equation is used to model whether or not a behavior of interest occurs, and a separate regression equation is used to model the intensity of the behavior for those who engage in it. Two-part models have been used, among other applications, to study spending on consumer durables (Cragg, 1971), vacation spending (Melenberg and Van Soest, 1996), and the demand for cigarettes (Raptou et al., 2005), and frequently are employed by health economists interested in explaining health care outcomes and medical care spending (see, for example, Duan et al 1983). All of these applications have in common (a) that different factors are believed to determine an outcome's incidence (positive

⁶ See Greene (2003), p. 770, for details concerning the test of the Tobit model restrictions.

versus zero) as opposed to its intensity (the positive values' magnitude) and (b) that, among those who engage in the behavior, values close to zero are rare. In our analysis, we use a probit equation to model whether or not a person reports any volunteer activity, and an ordinary least squares regression to model hours of volunteer activity for those who report volunteering.

Noting that:

$$(2) \quad E(y|X) = P(y > 0|X) \times E(y|y > 0, X)$$

the effect of the change in any explanatory variable on average volunteer hours can be written as:

$$(3) \quad \frac{\delta E(y)}{\delta X_c} = \frac{\delta P(y > 0)}{\delta X_c} \times E(y|y > 0, X) + P(y > 0|X) \times \frac{\delta E(y|y > 0)}{\delta X_c}$$

In this equation, the first term on the right-hand side represents the effect on overall average volunteer hours due to the change in the probability of volunteering and the second term represents the effect due to the change in average volunteer hours among those who volunteer.

Unless otherwise stated, all the estimates we report in this paper were calculated using ATUS selection (i.e., base) weights.⁷ Standard errors were calculated using a replicate variance method proposed by Fay (1989) that accounts for the increase in variance associated with the clustering and weighting in the ATUS sample relative to the variance that would be expected in a simple random sample of the same size. The replicate weights we used to implement this procedure were provided by BLS. The DESCRIP procedure in SAS-callable SUDAAN was used to compute the standard errors reported in Tables 1, 2 and 4, and also to carry out tests of statistical significance for differences in estimates across subgroups. Stata survey commands (SVY) were used to estimate the multivariate models and associated standard errors reported in Table 3.

⁷ Since our aim is to assess nonresponse bias, we chose not to use the ATUS final weights that incorporate post-stratification, which is intended to adjust for nonresponse. In fact, as we note below, our findings imply that post-stratification will not correct for the nonresponse bias in the volunteering estimates.

RESULTS

CPS Univariate Estimates of Volunteering

We begin by comparing the volunteering reports from the entire CPS supplement sample to those from our overlap subsample (which, two to eight months later, became ATUS respondents or nonrespondents). Comparing the first two rows of Table 1, we see that the estimates from the overlap sample are very close to those from the full sample. About 29 percent of each group reported having volunteered in the past year, and the average volunteer time reported by the two groups was very similar (37.9 hours versus 37.4 hours). The estimate of mean hours among those who did any volunteering also is very similar for the two groups (131.4 in the overlap subsample versus 129.8 hours in the full sample). In each case, the estimated value for our overlap sample is statistically indistinguishable from that for the CPS Supplement respondents who were not selected for the ATUS.

TABLE 1 ABOUT HERE

The third and fourth rows of Table 1 show that individuals who became ATUS respondents were over 75 percent more likely than those who became ATUS nonrespondents to report in the CPS that they had done volunteer work: 36 percent versus 20 percent. Likewise ATUS respondents reported an average of over 75 percent more hours worked as a volunteer than did ATUS nonrespondents: 46.6 hours versus 26.4 hours. The second finding is entirely a function of the first – among those who said they had volunteered, there is no difference between respondents and nonrespondents in the average number of volunteer hours.⁸

The remaining rows of Table 1 disaggregate the nonrespondents into five subgroups: two for which telephone contact was not undertaken (because the respondent moved, did not have a

⁸ Interviewers were instructed to try to conduct CPS Supplement interviews with each household member, but proxy reports were accepted where that was not feasible. As a check on whether proxy reports affected our results, we redid the analyses including only single person households, in which proxy reports are not possible. None of our conclusions about the differences between respondents and nonrespondents was altered. **VERIFY**

phone, or information about the phone number was lacking); one that was not successfully contacted (despite good contact information); one that refused; and a residual group, which consisted largely of individuals who spoke neither English nor Spanish. All five groups had much lower rates of volunteering than did respondents, but there were differences among the groups.

ATUS contacts are not attempted with individuals who no longer live at the address at which the CPS interview was conducted. Likewise, contacts are not attempted with individuals for whom phone numbers are missing (aside from sending out a letter requesting the individual phone a toll-free number, to which relatively few people respond despite the promised incentive payment). These two groups show lower rates of volunteering than do unsuccessful contacts and refusals, and the volunteering rates of the “other” (mainly language barrier) group are even lower. We interpret these results in terms of social integration: Geographical mobility, lacking phone service, and speaking neither English nor Spanish are all indicators of weak social integration. The differences between the two remaining groups, refusals and unsuccessful contacts, are not statistically significant. Though both show significantly higher rates of volunteering than other nonrespondents, their volunteering rates are still much lower than those of respondents.⁹

CPS Bivariate Estimates of Volunteering

The conventional approach to dealing with nonresponse bias is some form of statistical adjustment based on weighting classes or propensity models. As noted by Groves (2006: 653), however, “All of these adjustment techniques require assumptions that groups of respondents and nonrespondents share response propensities and distributional properties on survey

⁹ The differences among the five nonrespondent groups in hours volunteered parallel the differences in volunteering rates, though because the standard errors of these estimates are relatively larger the differences are generally not statistically significant. Among those who claimed to have volunteered, the differences in average hours between respondents and each of the nonrespondent groups is not statistically significant (with the exception of the “other” group).

measures.” To the extent that respondents and nonrespondents within a weighting class do not share the same expected values on the survey variable, the adjustment will fail. Unlike the typical situation in which this assumption is untestable, we can examine whether the CPS volunteering reports of ATUS respondents and nonrespondents are the same within subclasses of background variables that have been shown to be related to volunteering.

Wilson and Musick (1997; see also Musick, Wilson, and Bynum, 2000) argue that volunteering is a function of three kinds of resources: human capital, social capital and cultural capital. The CPS has no measures of cultural resources, but it does contain indicators of both human capital and social capital, as well as other background characteristics. Our principal measure of human capital is educational attainment, which we supplement with household income as a rough proxy for unobserved skills that may affect earnings power. The CPS contains six measures that we use as indicators of social capital: labor force participation, marital status, children in the household, presence of other adults in the household, home ownership, and presence of a telephone in the household. Finally, we account for five background characteristics: sex, age, race/ethnicity, urbanicity and region of residence. Taken together, these constitute a potentially rich set of variables for nonresponse adjustment.

TABLE 2a ABOUT HERE

The second column in Table 2a shows that human capital, social capital, and other background variables are related to rates of volunteering in the expected ways. For example, individuals with more education are more likely to report volunteering, as are those with higher incomes, married people, those with school-age children, homeowners, women, and non-Hispanic non-blacks.

The more important result in Table 2a is that, within every subgroup, respondents report higher volunteering rates than nonrespondents. There is some variation in the size of the

difference (e.g., the difference is larger for older respondents than for younger ones), but the most striking aspect of the table is the extent to which nonrespondents are distinctive. In 51 of the 52 subgroups, the volunteering rate among respondents is at least 25 percent greater than among nonrespondents (and in 49 cases it is at least 40 percent higher). This is compelling evidence against the assumption that respondents and nonrespondents within subclasses share the same expected values on volunteering, and thus indicates that statistical adjustment for the nonresponse bias is apt to fail.

Table 2b presents the same analysis, substituting overall average volunteer hours for volunteer rates. The second column generally shows the expected associations between hours and the independent variables. Individuals who are better educated and those who have higher incomes report more volunteer hours, as do married people, homeowners, women, and non-Hispanic non-blacks. More importantly, in 50 of the 52 subgroups, respondents report higher levels of mean volunteer hours than do nonrespondents, and the difference in reported hours between respondents and nonrespondents is not statistically significant in the two cases where the reverse is true.

TABLE 2b ABOUT HERE

The fact that the volunteering difference between respondents and nonrespondents is similarly large within subgroups is bad news for univariate analyses, because it means that statistical adjustment for nonresponse will not correct the bias in the estimate of the overall amount of volunteering in the population. But it is good news for bivariate (and possibly, multivariate) analysis. To the extent that the nonresponse bias is constant across subgroups, inferences about the relationships between volunteering and the variables used to define the subgroups will be unbiased. That is, although the level of volunteering will be exaggerated, it

will be exaggerated to the same extent across subgroups, and thus differences between subgroups will be unaffected.

Although comparing the second and fourth columns in Tables 2a and 2b indicates that the bivariate associations are generally immune to nonresponse bias, given the interrelationships among the variables we would like to know whether the same can be said about multivariate analyses.

CPS Multivariate Estimates of Volunteering

In this section, we estimate multivariate models of volunteering using the human capital, social capital, and other demographic indicators, first for the full ATUS overlap sample and then for the subsample of ATUS respondents. Adding a dummy variable distinguishing ATUS respondents and nonrespondents to the full sample model allows us to test whether nonresponse has a significant impact on estimates after simultaneously controlling for all the other characteristics. Comparison of the full sample and ATUS respondent models allows us to test whether inferences about the factors that affect volunteering are robust to nonresponse. To the extent that coefficients in the model including only respondents differ from those in the model based on the full sample, nonresponse bias affects multivariate inference. To the extent that coefficients in the model are not altered as we move from the full sample to the respondents-only sample, multivariate inferences are not compromised by nonresponse bias.

The first three columns in Table 3 summarize the results of using the full sample to estimate a two-part multivariate volunteering model. The first column of the table, from a probit analysis of the propensity to volunteer, shows the percentage point change in the volunteering rate associated with having the particular indicated characteristic rather than the full base set of

characteristics.¹⁰ The coefficients in the second column, from an ordinary least squares analysis, show the effects of the same characteristics on annual volunteer hours among those who volunteer. Drawing on equation (3), the two sets of coefficients can be used in conjunction with the estimates of volunteer rate $[P(y>0 | X)]$ and of volunteer hours among those who volunteer $[E(y | y>0, X)]$ to estimate the effects of each variable on overall average volunteer hours (for a person with the base group characteristics). These estimates are shown in the third column of Table 3.

TABLE 3 ABOUT HERE

As can be seen in the first column of Table 3, most of the explanatory variables are significantly related to the propensity to volunteer. The differences among groups generally parallel those observed in the bivariate results (Table 2a), but tend to be smaller (as would be expected given the intercorrelation among the independent variables). The main exception involves the association between marital status and the propensity to volunteer: The higher volunteering rate among married respondents shown in Table 2a reflects these individuals' other characteristics rather than their marital status.

In contrast, the multivariate model reveals that only a few characteristics are significantly related to volunteer hours among those who volunteered (second column of Table 3). Among volunteers, those who had not graduated from high school and those without telephone service reported fewer hours, and those older than 65 reported more hours. In addition, compared to volunteers not in the labor force, those who worked less than 45 hours per week reported fewer hours. (The remaining statistically significant effect, for cases whose urbanicity was missing, is based on only 6 respondents.)

¹⁰The reference person is defined by the modal category of each characteristic. She is a married female non-Hispanic non-black high school graduate age 31-45 who lives in the south and owns a home in the suburbs that has telephone service. She is not in the labor force and has an annual household income in the range \$20,000-\$39,999. She has no children or non-immediate family members resident in the household.

These results are not comparable to the bivariate results in Table 2b since they are based on only respondents who volunteered, whereas Table 2b includes all respondents. However, the estimates in the third column of Table 3, which show the effects on overall volunteer hours, can be compared to the bivariate estimates reported in Table 2b. The two sets of results are generally similar, with education, income, presence of school-age children, and not having a telephone in the household associated with sizeable differences in volunteer hours. Except for the difference associated with household telephone status, these large differences in average volunteer hours are due primarily to differences in volunteer rates rather than hours worked among those who volunteer (the difference by household telephone status reflects both of these effects).

By adding a dummy variable for whether a person is an ATUS respondent to the volunteer rate and volunteer hours equations reported in the first two columns of Table 3, we can measure the degree to which ATUS respondents differ from ATUS nonrespondents in their reported volunteer behavior once all the background characteristics are held constant. The coefficient on the respondent dummy variable in the volunteer propensity equation is highly significant and indicates that, among those with the reference set of personal characteristics, respondents are 10 percentage points more likely to volunteer than nonrespondents (results not shown). As was true absent controls for other characteristics, conditional on volunteering, the hours reported by respondents do not differ significantly from those reported by nonrespondents. Taken together, the two sets of coefficient estimates imply that, all else the same, respondents volunteer annually about 13 hours more than nonrespondents.¹¹

By comparing the results in the last three columns of Table 3 with those in the first three columns, we can assess the extent to which using data for respondents alone affects the multivariate associations between volunteering and the background characteristics. Consistent

¹¹ Adding the ATUS response dummy variable to these equations has little effect on the coefficients estimated for the other characteristic variables. Full results are available upon request.

with the message of Tables 2a and 2b, the exclusion of the nonrespondents has little impact on multivariate inferences. The coefficients in the “respondent” models generally are very similar to those in the “full sample” models. As a more formal test, we compared models estimated for the full sample that contained all of the background characteristics plus a dummy variable for whether the person was an ATUS respondent to a series of models that added interactions between the ATUS respondent dummy and selected background variables. The first of these alternate models added interactions between the respondent dummy and the four education dummies, the second added interactions with the four household income dummies, and so on. For eight of the sets of background variables (labor force status, marital status, children in the household, others in the household, telephone service, sex, urbanicity and region) the interaction terms were not statistically significant, and in three of the five cases in which they were significant (education, income and housing tenure), the inferences that would be drawn about the relationship between volunteering and the background characteristic are very similar in the full versus respondent samples.¹² Only for race and age would a substantive inference be altered. In the case of age, we have no explanation for the effect (which involves respondents older than 65). The finding for race may best be treated as a chance finding, since it is due to Hispanic respondents being less apt to volunteer than Hispanic nonrespondents, a pattern that was not replicated in our analysis of the 2002 CPS volunteer supplement.

ATUS Estimates of Volunteering

Among those in the overlap sample who said in the CPS that they had volunteered, 70 percent became ATUS respondents compared to just 52 percent who said they had not volunteered. We expect that this difference means that nonresponse in the ATUS will (a) cause

¹² Education and income are positively related to volunteering in both samples, but the effect of education is slightly stronger in the full sample whereas the effect of income is slightly stronger in the respondent sample. Renters are a little less likely to volunteer than are homeowners in both the full and respondent samples; the housing tenure interactions are significant as a group only because of the small number of cases that are missing data on housing tenure.

that survey's volunteering estimates to be too high, but (b), based on our CPS findings, will not affect the ATUS estimates of the correlates of volunteering.

Unlike the CPS (which has a twelve-month reporting period), the single day reporting period of the ATUS means that the ATUS cannot produce a meaningful estimate of the proportion of volunteers. The ATUS can, however, be used to produce an estimate of mean annual volunteer hours that we compare to the CPS hours estimate. The first step in constructing the ATUS estimate of mean annual hours is to calculate the weighted mean of the hours reported by survey respondents on their diary day, using the ATUS sample (base) weights adjusted to account for the differing probabilities of assignment to weekday versus weekend days. Multiplying this estimate of mean daily volunteer hours by 365 yields an estimate of mean annual volunteer hours.

As noted earlier, the CPS estimate of average volunteer hours for the period September 2002 to September 2003 is 37.4 hours for the full ATUS overlap sample and 46.6 hours for the ATUS respondents among that group (Table 1). If there were no other differences between the surveys, we might therefore expect the ATUS estimate to be within sampling error of 46.6 hours. In fact, that is what we observe: the ATUS estimate of mean volunteer hours for calendar year 2003 is 49.8 hours with a standard error of 2.4 hours (Table 4). Although the ATUS estimate of volunteer hours suffers from substantial nonresponse bias that makes it much too high, the CPS estimate might suffer from recall error due to the difficulty of the twelve-month recall task (leading to a downward bias) and from social desirability bias (leading to an upward bias). Since the ATUS is not apt to suffer significantly from either of these two error sources, the similarity

of the ATUS estimate and the CPS estimate based on ATUS respondents suggests that any bias in the CPS estimates from these sources roughly balances out on average.¹³

TABLE 4 ABOUT HERE

Table 4 also shows that, although the ATUS estimate of average volunteer hours is much larger than that from the full CPS sample, the ATUS estimates of the associations between respondent background characteristics and volunteer hours are similar to those from the CPS (Table 2b). Thus at least in this case conclusions about the determinants of volunteering are generally unaffected by nonresponse.¹⁴

DISCUSSION

It has often been assumed that the large variation in survey estimates of volunteering (including those from the CPS volunteering supplement) is due to both the difficulty respondents have recalling an entire year's activities and the risk that respondents might exaggerate the extent of their volunteering. Scholars thus eagerly awaited the American Time Use Survey, which provides volunteering estimates that are unlikely to suffer from these problems. With this in mind, Abraham and Mackie (2005: 147) wrote that "When they become available, data on volunteer activity from the new American Time Use Survey (ATUS) should be more reliable than any that currently exist." Our results, however, indicate that although the ATUS estimates may be less susceptible to measurement error than the CPS volunteering supplement estimates, this is outweighed by the much greater nonresponse error of the ATUS.

¹³ We can also compare the 2003 CPS estimate for ATUS respondents (46.6 hours with a standard error of 2.6 hours) to the ATUS estimate for the same respondents (58.2 hours with a standard error of 4.5 hours), but this comparison is more severely compromised by different reference periods.

¹⁴ Our inferences about the impact of ATUS nonresponse are based on the characteristics of those ATUS nonrespondents (somewhat more than half) for whom we have CPS Volunteering Supplement observations. The remaining ATUS nonresponse overlaps with nonresponse to the CPS Supplement. It seems very likely that the latter nonrespondents (for whom we have no data) are also disproportionately nonvolunteers, but to the extent this is incorrect it could compromise our conclusions.

We believe our finding of nonresponse bias in the ATUS sample has implications for many other surveys of volunteering. In the introduction, for instance, we noted the discrepancy in volunteering rates between the 1996 National Household Education Survey (NHES; NCES, 1997) -- 39 percent – and the 2002 CPS Supplement (Boraas, 2003) – 28 percent. The higher volunteering rate in the 1996 NHES as compared to the 2002 CPS supplement is entirely consistent with the two surveys' relative response rates (59 percent versus 81 percent).

Similarly, the volunteering rate of 55 percent from Grimm et al.'s (2005) survey of teenagers is likely a function of that survey's relatively low response rate (44 percent). Thus Grimm et al.'s conclusion that teenagers volunteer at a rate “more than one a half times the adult rate of 29 percent as established by the...2004 Current Population Survey figures, which used the same questions and definitions as the Youth Volunteering Survey” is almost surely wrong. The difference is most likely a function of nonresponse bias – that is, of teenagers who were not volunteers being much less likely to participate in the Grimm et al., survey. (The CPS estimates show that teenagers are *less* apt to be volunteers than are middle-aged adults.)

If we are correct that our results are due to the strong connection between the causes of volunteering and the causes of survey participation, then surveys will also tend to overestimate other activities that have an altruistic aspect. At least two studies have reported results consistent with this. Kennickell (2005), for instance, found that individuals who declared a charitable contributions deduction on their federal income tax return were more likely to be respondents on the Survey of Consumer Finances than those who did not declare that deduction. Likewise, Couper, Singer, and Kulka (1998) found that households in which respondents said they

participated in community activities were more likely to have returned their decennial census form than households in which respondents said they did not participate in such activities.¹⁵

Moreover, to the extent that the size of the nonresponse bias in volunteering (and related activities) is affected by the level of nonresponse, the secular decline in response rates means that inferences about changes over time in volunteering (and related activities) may be distorted by the bias. Volunteering estimates may increase over time (see Figure 1) only because the composition of the sample changes as response rates decline. This could account for why some trends in volunteering over the last several decades are so different from trends in other dimensions of social capital (Putnam 2000: 127).

Our results suggest that bivariate and multivariate inferences about the relationship of volunteering (and possibly related activities) to other respondent characteristics are relatively immune to nonresponse bias. In this sense, the results are similar to the pattern that Schuman and Presser (1981) called “form-resistant correlations.” They used the term to describe the finding that changes in survey question wording that affect univariate distributions very often do not affect bivariate or multivariate distributions. The findings presented here suggest the same may be true of nonresponse.

¹⁵ Although the evidence is more indirect, Robinson (1989) suggests that surveys are also apt to overestimate participation in the arts because those who engage in such activities are more likely to participate in surveys than those who don't engage in the arts.

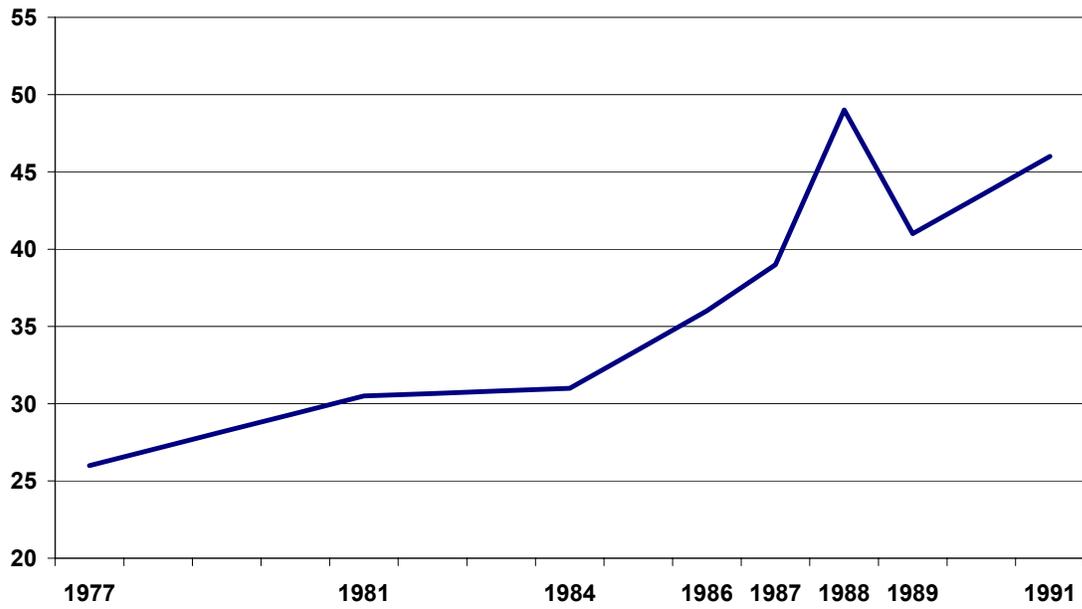
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Fig. 1. Percent Volunteers in Gallup Surveys



SOURCE: Roper Center Accession Nos.
0047361, 0031892, 0030064, 0032803, 0026314, 0026426, 0047980, 0028183, and 0236260

TABLE 1: Volunteer rates and mean hours spent volunteering, September 2003 CPS Volunteer supplement, various samples

	N	Percent Who Volunteer	Volunteer Hours	
			Volunteer Mean	Overall Mean
All volunteer supplement respondents	95337	28.8 (0.1)	129.8 (1.6)	37.4 (0.5)
Volunteer supplement respondents in ATUS sample	9004	28.9 (0.6)	131.4 (5.2)	37.9 (1.7)
ATUS respondents	5079	35.5 (0.8)	131.3 (6.5)	46.6 (2.6)
ATUS nonrespondents	3925	20.1 (0.8)	131.6 (10.1)	26.4 (2.2)
Contact not attempted	946	17.4 (1.6)	137.3 (22.7)	23.8 (4.4)
Inadequate/missing contact information	920	17.0 (1.3)	124.6 (22.6)	21.2 (4.1)
Unsuccessful contact	518	22.9 (1.9)	110.8 (18.2)	25.3 (4.7)
Refusals	1439	23.8 (1.2)	139.4 (4.0)	33.2 (4.0)
Other	102	6.7 (3.2)	70.5 (28.4)	4.7 (2.3)

Note: The column labeled "Volunteer Mean" reports average annual volunteer hours for those who volunteer; the column labeled "Overall Mean" averages volunteer hours across the full sample, including those with zero volunteer hours. Volunteer hours are imputed for 143 individuals who reported that they volunteered but did not report amount of time they spent. The first row is weighted using the CPS supplement non-response weight and the remaining rows are weighted using the ATUS sample weight. Except in the first row, standard errors reported in parentheses have been adjusted to account for the clustering and weighting of the survey sample; the information necessary to make this adjustment is not available on the CPS supplement public use data files.

TABLE 2A: Volunteer rates calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics

	Full ATUS Overlap Sample			Respondents			Non- respondents		
	N	Rate		N	Rate		N	Rate	
		(s.e.)			(s.e.)			(s.e.)	
Full sample	9004	28.9 (0.6)		5079	35.5 (0.8)		3925	20.1 (0.8)	
Human Capital Indicators									
Education									
Less than high school	1718	17.4 (1.1)	**	830	24.0 (1.7)		888	11.0 (1.2)	**
High school graduate	2715	21.8 (0.9)	--	1432	26.2 (1.2)	--	1283	16.8 (1.3)	--
Some college	2363	31.9 (1.1)	**	1375	36.6 (1.5)	**	988	25.1 (1.4)	**
Bachelor's degree	1469	43.4 (1.5)	**	929	50.8 (1.9)	**	540	30.0 (2.0)	**
Graduate degree	739	44.2 (2.0)	**	513	50.7 (2.5)	**	226	29.9 (3.0)	**
Household income									
Missing	1360	21.3 (1.2)		612	27.9 (1.9)		748	15.4 (1.4)	
Under \$20,000	1734	16.9 (1.1)	**	833	20.6 (1.6)	**	901	13.6 (1.4)	*
\$20,000 to \$39,999	2123	22.1 (1.1)	--	1215	24.6 (1.4)	--	908	18.9 (1.6)	--
\$40,000 to \$74,999	2095	33.2 (1.0)	**	1287	40.5 (1.3)	**	808	21.8 (1.7)	
\$75,000 or more	1692	45.0 (1.4)	**	1132	51.3 (1.8)	**	560	32.2 (2.2)	**
Social Capital Indicators									
Labor force status									
Not in labor force	3576	24.5 (0.9)	--	1945	32.2 (1.2)	--	1631	15.3 (0.9)	--
Work <35 hrs/wk	827	38.8 (2.1)	**	521	41.9 (2.8)	**	306	33.5 (3.0)	**
Work 35-44 hrs/wk	3074	27.3 (0.9)	*	1708	33.5 (1.3)	*	1366	19.1 (1.1)	**
Work 45 plus hrs/wk	1047	37.3 (1.6)	**	633	42.9 (2.1)	**	414	28.5 (2.4)	**
Work hours vary	480	34.9 (2.6)	**	272	41.0 (3.5)	**	208	26.7 (3.7)	**

TABLE 2A: Volunteer rates calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics (continued)

	Full ATUS Overlap Sample			Respondents			Non- respondents		
	N	Rate		N	Rate		N	Rate	
		(s.e.)			(s.e.)			(s.e.)	
Marital status									
Married	4346	34.4 (0.8)	--	2714	40.4 (1.1)	--	1632	24.2 (1.2)	--
Widowed	756	22.3 (1.6)	**	415	31.2 (2.6)	**	341	11.9 (1.8)	**
Divorced	1141	23.6 (1.5)	**	628	29.0 (2.2)	**	513	17.1 (1.7)	**
Spouse absent	432	20.9 (2.5)	**	214	21.6 (3.2)	**	218	20.1 (3.4)	
Never married	2329	22.9 (1.1)	**	1108	29.0 (1.7)	**	1221	17.1 (1.2)	**
Children in household									
No children under age 6	7293	28.6 (0.6)	--	4135	35.1 (0.8)	--	3158	19.7 (0.8)	--
Children under age 6	1711	30.4 (1.3)		944	37.3 (1.9)		767	21.8 (1.7)	
No children age 6-17	5928	25.1 (0.7)	--	3330	31.5 (1.0)	--	2598	16.7 (0.8)	--
Children age 6-17	3076	36.4 (1.1)	**	1749	43.6 (1.5)	**	1327	26.9 (1.5)	**
Others in household									
No relatives	7272	30.5 (0.7)	--	4197	36.9 (0.9)	--	3075	21.1 (0.8)	--
One or more relatives	1732	25.0 (1.2)	**	882	31.5 (1.8)	**	850	17.8 (1.4)	*
No non-relatives	8286	30.3 (0.6)	--	4763	36.8 (0.9)	--	3523	21.2 (0.8)	--
One or more non-relatives	718	15.2 (1.4)	**	316	19.0 (2.4)	**	402	12.2 (1.6)	**
Housing tenure									
Missing	258	35.2 (3.3)		142	42.7 (4.6)		116	24.7 (4.8)	
Owner	6136	31.5 (0.7)	--	3694	38.2 (1.0)	--	2442	21.4 (0.9)	--
Renter	2610	20.6 (1.0)	**	1243	24.9 (1.4)	**	1367	16.5 (1.2)	**

TABLE 2A: Volunteer rates calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics (continued)

	Full ATUS Overlap Sample			Respondents			Non- respondents		
	N	Rate		N	Rate		N	Rate	
		(s.e.)			(s.e.)			(s.e.)	
Telephone status									
Telephone household	8594	29.5 (0.6)	--	4933	36.0 (0.8)	--	3661	20.6 (0.8)	--
Non-telephone household	410	12.6 (1.9)	**	146	15.5 (3.1)	**	264	11.0 (2.1)	**
Other Characteristics									
Sex									
Male	4023	25.5 (0.9)	**	2202	31.9 (1.3)	**	1821	17.5 (0.9)	**
Female	4981	31.9 (0.7)	--	2877	38.5 (0.9)	--	2104	22.5 (1.1)	--
Age									
Age 15-30	2096	25.3 (1.1)	**	1025	30.6 (1.7)	**	1071	20.0 (1.4)	*
Age 31-45	2830	34.0 (1.0)	--	1591	41.0 (1.4)	--	1239	24.8 (1.4)	--
Age 46-55	1486	33.9 (1.3)		918	39.0 (1.9)		568	25.1 (2.2)	
Age 56-65	1080	27.7 (1.3)	**	708	34.1 (1.8)	**	372	14.9 (2.1)	**
Over age 65	1512	21.1 (1.2)	**	837	29.4 (1.7)	**	675	10.9 (1.4)	**
Race/ethnicity									
Hispanic	1115	14.8 (1.2)	**	548	17.6 (1.9)	**	567	12.1 (1.5)	**
Non-Hispanic black	1267	20.5 (1.3)	**	561	25.2 (2.4)	**	706	16.9 (1.6)	**
Non-Hispanic non-black	6622	32.1 (0.6)	--	3970	38.7 (0.9)	--	2652	22.2 (0.9)	--
Urbanicity of residence									
Central city	2224	24.3 (1.3)	**	1131	31.0 (1.8)	**	1093	17.2 (1.5)	*
Balance of MSA	3774	31.0 (0.8)	--	2154	38.3 (1.2)	--	1620	21.1 (1.1)	--
Other metropolitan	1278	28.1 (1.5)		739	34.3 (2.0)		539	18.8 (2.1)	
Non-metropolitan	1705	30.1 (1.3)		1039	34.7 (1.7)		666	22.9 (1.8)	
Not identified	23	26.5 (13.8)		16	32.2 (15.7)		7	13.5 (14.4)	

TABLE 2A: Volunteer rates calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics (continued)

	Full ATUS Overlap Sample		Respondents		Non- respondents	
	Rate		Rate		Rate	
	N	(s.e.)	N	(s.e.)	N	(s.e.)
Region of residence						
Northeast	1739	26.2 (1.1)	994	32.7 (1.6)	745	17.6 (1.5)
South	3316	27.3 (0.9)	-- 1754	32.9 (1.3)	-- 1562	20.6 (1.2)
West	1858	29.2 (1.5)	1064	36.4 (1.9)	794	19.3 (1.6)
Midwest	2091	33.2 (1.1)	** 1267	40.1 (1.4)	** 824	22.0 (1.7)

Note: All estimates are weighted using the ATUS sample weights. Standard errors reported in parentheses have been adjusted to account for the clustering and weighting of the survey sample.

** Significantly different from reference group mean at 0.01 level

* Significantly different from reference group mean at 0.05 level

-- Reference group

TABLE 2B: Volunteer hours calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics

	Full ATUS Overlap Sample			Respondents		Non- respondents			
	Overlap Sample								
	N	Hours (s.e.)		N	Hours (s.e.)	N	Hours (s.e.)		
Full sample	9004	37.9 (1.7)		5079	46.6 (2.6)		3925	26.4 (2.2)	
Human Capital Indicators									
Education									
Less than high school	1718	15.9 (2.2)	**	830	21.0 (2.7)	**	888	11.0 (3.3)	*
High school graduate	2715	28.9 (3.2)	--	1432	36.2 (5.2)	--	1283	20.4 (3.1)	--
Some college	2363	44.6 (3.5)	**	1375	50.3 (4.7)		988	36.3 (5.5)	*
Bachelor's degree	1469	60.5 (5.0)	**	929	70.3 (6.7)	**	540	43.0 (8.1)	*
Graduate degree	739	58.7 (5.9)	**	513	67.0 (7.6)	**	226	40.4 (8.2)	*
Household income									
Missing	1360	32.6 (4.2)		612	37.2 (4.4)		748	28.4 (7.0)	
Under \$20,000	1734	27.6 (4.2)		833	33.3 (7.5)		901	22.5 (4.5)	
\$20,000 to \$39,999	2123	31.9 (3.5)	--	1215	38.8 (5.7)	--	908	23.0 (3.0)	--
\$40,000 to \$74,999	2095	38.2 (3.3)		1287	48.6 (4.9)		808	21.7 (3.5)	
\$75,000 or more	1692	55.3 (4.0)	**	1132	62.9 (4.8)	**	560	39.8 (6.3)	*
Social Capital Indicators									
Labor force status									
Not in labor force	3576	40.1 (2.8)	--	1945	50.8 (3.7)	--	1631	27.4 (4.1)	--
Work <35 hrs/wk	827	45.6 (4.6)		521	46.7 (5.8)		306	43.6 (7.8)	
Work 35-44 hrs/wk	3074	28.7 (2.3)	**	1708	34.7 (3.5)	**	1366	20.7 (2.5)	
Work 45 plus hrs/wk	1047	45.6 (6.6)		633	55.8 (9.9)		414	29.5 (5.5)	
Work hours vary	480	48.5 (6.9)		272	68.5 (10.9)		208	21.3 (5.0)	

TABLE 2B: Volunteer hours calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics (continued)

	Full ATUS Overlap Sample			Respondents		Non- respondents			
	Overlap Sample								
	N	Hours (s.e.)		N	Hours (s.e.)	N	Hours (s.e.)		
Marital status									
Married	4346	44.8 (2.7)	--	2714	53.5 (3.9)	--	1632	30.0 (3.2)	--
Widowed	756	47.7 (5.8)		415	69.9 (9.8)		341	21.7 (7.1)	
Divorced	1141	28.0 (3.4)	**	628	33.5 (4.8)	**	513	21.2 (4.7)	
Spouse absent	432	26.6 (5.4)	**	214	26.4 (8.6)	**	218	26.8 (6.8)	
Never married	2329	27.4 (3.1)	**	1108	30.9 (4.5)	**	1221	24.1 (4.3)	
Children in household									
No children under age 6	7293	39.1 (2.0)	--	4135	47.9 (2.9)	--	3158	27.3 (2.5)	--
Children under age 6	1711	31.8 (2.9)	*	944	39.6 (4.3)		767	22.0 (3.7)	
No children age 6-17	5928	35.7 (2.0)	--	3330	44.2 (2.9)	--	2598	24.3 (2.7)	--
Children age 6-17	3076	42.4 (2.8)		1749	51.3 (4.5)		1327	30.6 (3.7)	
Others in household									
No relatives	7272	39.9 (2.0)	--	4197	48.0 (2.9)	--	3075	28.1 (2.6)	--
One or more relatives	1732	33.1 (3.3)		882	42.6 (4.8)		850	22.8 (4.2)	
No non-relatives	8286	39.4 (1.7)	--	4763	47.6 (2.6)	--	3523	28.0 (2.4)	--
One or more non-relatives	718	23.5 (5.8)	**	316	34.0 (11.7)		402	15.0 (4.4)	*
Housing tenure									
Missing	258	49.2 (12.0)		142	49.4 (13.3)		116	48.9 (23.5)	
Owner	6136	40.5 (2.2)	--	3694	49.9 (3.3)	--	2442	26.2 (2.5)	--
Renter	2610	29.2 (2.9)	**	1243	34.6 (4.7)	**	1367	24.1 (3.7)	

TABLE 2B: Volunteer hours calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics (continued)

	Full ATUS Overlap Sample			Respondents		Non- respondents			
	Overlap Sample								
	N	Hours (s.e.)		N	Hours (s.e.)	N	Hours (s.e.)		
Telephone status									
Telephone household	8594	39.0 (1.8)	--	4933	47.4 (2.7)	--	3661	27.4 (2.3)	--
Non-telephone household	410	10.3 (2.6)	**	146	12.8 (4.2)	**	264	8.8 (3.7)	**
Other Characteristics									
Sex									
Male	4023	33.0 (2.5)	**	2202	40.6 (3.5)	*	1821	23.6 (3.3)	--
Female	4981	42.3 (2.5)	--	2877	51.6 (3.7)	--	2104	29.0 (2.9)	
Age									
Age 15-30	2096	27.9 (3.3)	*	1025	31.3 (4.9)		1071	24.6 (4.6)	
Age 31-45	2830	37.5 (3.1)	--	1591	44.4 (4.4)	--	1239	28.4 (3.5)	--
Age 46-55	1486	45.8 (4.3)		918	54.2 (6.3)		568	31.5 (5.3)	
Age 56-65	1080	40.7 (4.6)		708	47.2 (6.3)		372	27.7 (6.7)	
Over age 65	1512	43.3 (4.4)		837	61.3 (6.1)	*	675	20.9 (5.5)	
Race/ethnicity									
Hispanic	1115	18.0 (3.2)	**	548	14.7 (2.6)	**	567	21.3 (5.7)	
Non-Hispanic black	1267	27.6 (4.2)	*	561	31.6 (5.4)	**	706	24.5 (6.4)	
Non-Hispanic non-black	6622	42.3 (2.1)	--	3970	52.0 (3.1)	--	2652	27.7 (2.4)	--
Urbanicity of residence									
Central city	2224	34.1 (3.6)		1131	46.4 (6.3)		1093	21.3 (3.0)	
Balance of MSA	3774	38.3 (2.4)	--	2154	46.2 (3.0)	--	1620	27.5 (3.7)	--
Other metropolitan	1278	33.1 (3.3)		739	39.9 (4.8)		539	22.9 (4.0)	
Non-metropolitan	1705	45.7 (4.8)		1039	52.8 (7.4)		666	34.5 (6.2)	
Not identified	23	15.9 (8.5)	*	16	21.2 (11.2)	*	7	4.1 (4.3)	**

TABLE 2B: Volunteer hours calculated from the 2003 CPS volunteer supplement, ATUS sample members, respondents and non-respondents, by selected demographic characteristics (continued)

	Full ATUS Overlap Sample		Respondents		Non- respondents	
	Overlap Sample		Hours		Hours	
	N	(s.e.)	N	(s.e.)	N	(s.e.)
	Hours		Hours		Hours	
Region of residence						
Northeast	1739	36.0 (3.8)	994	45.6 (5.0)	745	23.3 (5.9)
South	3316	35.1 (2.5)	-- 1754	41.1 (3.8)	-- 1562	28.1 (3.5)
West	1858	43.2 (3.9)	1064	55.9 (6.1)	* 794	25.9 (3.9)
Midwest	2091	38.6 (3.0)	1267	46.0 (4.7)	824	26.6 (4.0)

Note: All estimates are weighted using the ATUS sample weights. Standard errors reported in parentheses have been adjusted to account for the clustering and weighting of the survey sample.

** Significantly different from reference group mean at 0.01 level

* Significantly different from reference group mean at 0.05 level

-- Reference group

Table 3: Two-part models of volunteer activity, September 2003 CPS Volunteer supplement, full ATUS overlap sample and ATUS respondent sample

	Full ATUS Overlap Sample		ATUS Respondent Sample	
	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)
		Calculated Effect on Overall Hours		Calculated Effect on Overall Hours
Human Capital Indicators				
Education				
Less than high school	-0.046** (0.016)	-55.6* (22.9)	-0.024 (0.023)	-51.9 (30.5)
Some college	0.081** (0.016)	17.8 (17.4)	0.076** (0.020)	9.7 (23.0)
Bachelor's degree	0.182** (0.020)	25.3 (18.2)	0.197** (0.027)	15.3 (24.0)
Graduate degree	0.175** (0.026)	19.2 (18.6)	0.181** (0.032)	11.6 (23.8)
Household income				
Missing	-0.022 (0.019)	13.0 (24.5)	0.021 (0.025)	-16.5 (28.3)
Under \$20,000	-0.026 (0.019)	14.2 (27.9)	-0.008 (0.024)	-0.2 (41.9)
\$40,000-\$74,999	0.072**	-18.7 (17.7)	0.126** (0.020)	-25.9 (26.1)
\$75,000 or more	0.131** (0.021)	-11.5 (20.6)	0.176** (0.026)	-24.8 (29.5)
Social Capital Indicators				
Labor force status				
Work < 35 hrs/wk	0.063** (0.023)	-40.4* (16.1)	0.018 (0.029)	-35.7* (17.6)
Work 35-44 hrs/wk	-0.046** (0.015)	-53.7** (13.2)	-0.058** (0.019)	-43.5** (14.5)
Work 45 plus hours/wk	-0.009 (0.018)	-42.2 (23.4)	-0.030 (0.024)	-21.5 (29.1)
Work hours vary	0.031 (0.027)	-23.7 (21.9)	0.000 (0.034)	21.8 (28.2)

Table 3: Two-part models of volunteer activity, September 2003 CPS Volunteer supplement, full ATUS overlap sample and ATUS respondent sample (continued)

	Full ATUS Overlap Sample			ATUS Respondent Sample		
	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Calculated Effect on Overall Hours	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Calculated Effect on Overall Hours
Marital status						
Widowed	0.026 (0.027)	34.7 (27.8)	12.2	0.040 (0.035)	39.7 (34.5)	15.2
Divorced	-0.022 (0.018)	-20.2 (20.2)	-7.9	-0.011 (0.023)	-35.4 (26.4)	-10.5
Spouse absent	-0.014 (0.030)	-8.0 (24.0)	-3.8	-0.035 (0.037)	-18.0 (36.2)	-8.9
Never married	-0.030 (0.018)	7.2 (14.9)	-1.9	-0.015 (0.023)	-12.6 (18.5)	-5.1
Children in household						
Children under age 6	-0.019 (0.014)	-14.2 (11.3)	-6.1	-0.015 (0.019)	-11.7 (13.1)	-4.9
Children age 6-17	0.125 (0.015)	11.0 (11.5)	18.4	0.133 (0.021)	20.7 (15.1)	21.6
Others in household						
One or more relatives	-0.055 (0.013)	13.1 (14.0)	-3.6	-0.049 (0.017)	20.6 (17.4)	-0.6
One or more nonrelatives	-0.104 (0.019)	28.4 (36.8)	-5.7	-0.109 (0.026)	66.6 (56.8)	3.9
Housing tenure						
Missing	0.039 (0.034)	14.5 (31.2)	8.6	0.044 (0.048)	-15.2 (27.2)	1.4
Renter	-0.034 (0.013)	13.4 (13.4)	-0.8	-0.038 (0.018)	8.8 (15.9)	-2.3
Telephone status						
Non-telephone household	-0.078 (0.026)	-53.0 (26.1)	-23.3	-0.089 (0.039)	-37.7 (31.1)	-20.7

Table 3: Two-part models of volunteer activity, September 2003 CPS Volunteer supplement, full ATUS overlap sample and ATUS respondent sample (continued)

	Full ATUS Overlap Sample		ATUS Respondent Sample			
	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Calculated Effect on Overall Hours	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Calculated Effect on Overall Hours
<i>Other Characteristics</i>						
Sex						
Male	-0.060 (0.011)	4.4 (11.7)	-6.4	-0.061 (0.013)	-3.9 (14.1)	-8.5
Age						
Under age 30	0.014 (0.017)	-0.5 (16.7)	1.7	-0.007 (0.020)	10.2 (21.0)	1.8
Age 46-55	0.018 (0.018)	26.4 (15.8)	9.0	0.002 (0.020)	32.5 (19.0)	8.6
Age 56-65	-0.019 (0.020)	31.0 (20.1)	5.5	-0.011 (0.024)	30.4 (21.1)	6.5
Over age 65	-0.056 (0.024)	57.2 (24.7)	7.6	-0.034 (0.031)	72.1 (25.4)	14.5
Race/ethnicity						
Hispanic	-0.097 (0.018)	-6.6 (22.8)	-13.8	-0.113 (0.023)	-47.7 (19.8)	-26.2
Non-Hispanic black	-0.048 (0.016)	2.8 (21.9)	-5.3	-0.054 (0.029)	-4.0 (25.8)	-7.6
Urbanicity of residence						
Central city	-0.013 (0.015)	12.2 (15.2)	1.4	-0.006 (0.020)	28.7 (19.2)	6.8
Other metropolitan	0.002 (0.018)	-2.1 (12.6)	-0.3	0.000 (0.022)	4.3 (15.8)	1.1
Non-metropolitan	0.038 (0.016)	30.0 (17.9)	12.4	0.029 (0.018)	27.4 (22.8)	10.6
Not identified	-0.006 (0.146)	-66.5 (27.9)	-17.7	0.005 (0.155)	-52.8 (41.5)	-13.1

Table 3: Two-part models of volunteer activity, September 2003 CPS Volunteer supplement, full ATUS overlap sample and ATUS respondent sample (continued)

	Full ATUS Overlap Sample			ATUS Respondent Sample		
	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Calculated Effect on Overall Hours	Vol. (0/1) dF/dx (std. err.)	Hrs. for Volunteers Coef. (std. err.)	Calculated Effect on Overall Hours
Region of residence						
Northeast	-0.027 (0.014)	4.0 (16.1)	-2.4	-0.017 (0.019)	16.4 (18.1)	2.2
Midwest	0.026 (0.015)	-17.2 (11.4)	-1.2	0.040 * (0.019)	-16.3 (15.1)	0.6
West	-0.001 (0.016)	16.5 (15.0)	4.1	0.023 (0.020)	26.8 (20.1)	9.8
Constant		125.0 ** (25.4)			122.6 ** (31.7)	
N	9004	2609		5079	1804	
R2		0.0356			0.0430	

Note: The volunteer propensity equation is a probit model. The numbers reported in the first and fourth columns are estimates of change in the volunteering rate implied by the probit coefficients. The volunteer hours equations in the second and fifth columns are least squares regressions estimated for those with positive volunteer hours. Estimated effects on overall hours appear in the third and sixth columns. All estimates are weighted using the ATUS sample weights. Standard errors reported in parentheses have been adjusted to account for the clustering and weighting of the survey sample.

** Statistically significant at the 0.01 level and

* Statistically significant at the 0.05 level.

TABLE 4: Volunteer hours calculated from the 2003 American Time Use Survey, by selected demographic characteristics

	N	Hours (s.e.)	
Full sample	20720	49.8 (2.4)	
<i>Human Capital Indicators</i>			
Education			
Less than high school	3634	36.4 (6.9)	
High school graduate	5790	38.0 (3.7)	--
Some college	5524	53.6 (4.1)	**
Bachelor's degree	3668	68.7 (6.2)	**
Graduate degree	2104	78.4 (7.3)	**
Household income			
Missing	2451	36.5 (4.6)	
Under \$20,000	3506	42.2 (6.2)	
\$20,000 to \$39,999	4726	44.3 (3.9)	--
\$40,000 to \$74,999	5330	56.9 (6.3)	
\$75,000 or more	4707	59.2 (4.8)	*
<i>Social Capital Indicators</i>			
Labor force status			
Not in labor force	7657	63.4 (4.7)	--
Work <35 hrs/wk	2125	64.9 (7.1)	
Work 35-44 hrs/wk	7098	34.2 (2.6)	**
Work 45 plus hrs/wk	2754	43.1 (5.8)	**
Work hours vary	1086	41.6 (10.5)	
Marital status			
Married	10825	58.0 (2.9)	--
Widowed	1655	49.0 (7.6)	
Divorced	2510	37.4 (5.7)	**

TABLE 4: Volunteer hours calculated from the 2003 American Time Use Survey, by selected demographic characteristics (continued)

	N	Hours (s.e.)	
Marital status (continued)			
Spouse absent	914	40.9 (14.0)	
Never married	4816	40.3 (5.4)	**
Children in household			
No children under age 6	16778	51.6 (2.7)	--
Children under age 6	3942	40.5 (4.9)	*
No children age 6-17	13217	44.0 (2.9)	--
Children age 6-17	7503	61.4 (4.0)	**
Others in household			
No relatives	16945	54.5 (2.6)	--
One or more relatives	3775	38.1 (5.3)	**
No non-relatives	19366	51.0 (2.4)	--
One or more non-relatives	1354	37.9 (8.3)	
Housing tenure			
Owner	15584	55.7 (2.8)	--
Renter	5136	31.6	**
Telephone status			
Telephone household	20168	50.6 (2.5)	--
Non-telephone household	552	27.2 (10.8)	*
Other Characteristics			
Sex			
Male	9049	47.7 (3.8)	
Female	11671	51.8 (2.8)	--
Age			
Age 15-30	4327	35.3 (5.5)	*
Age 31-45	6775	49.7 (4.2)	--
Age 46-55	3779	59.0 (5.6)	

TABLE 4: Volunteer hours calculated from the 2003 American Time Use Survey, by selected demographic characteristics (continued)

	N	Hours (s.e.)	
Age (continued)			
Age 56-65	2652	49.6 (5.1)	
Over age 65	3187	68.1 (5.9)	**
Race/ethnicity			
Hispanic	2300	28.7 (5.0)	**
Non-Hispanic black	2396	57.0 (13.1)	
Non-Hispanic non-black	16024	52.2 (2.2)	--
Urbanicity of residence			
Central city	4603	39.7 (6.4)	
Balance of MSA	8894	49.1 (3.4)	--
Other metropolitan	3039	59.2 (5.5)	
Non-metropolitan	4131	54.9 (5.0)	
Not identified	53	128.4 (64.1)	
Region of residence			
Northeast	4119	49.3 (6.7)	
South	7135	45.0 (2.9)	--
West	4259	53.9 (5.3)	
Midwest	5207	53.2 (5.0)	

Note: All estimates are weighted using the ATUS final weights. Standard errors reported in parentheses have been adjusted to account for the clustering and weighting of the survey sample.

** Significantly different from reference group mean at 0.01 level

* Significantly different from reference group mean at 0.05 level

-- Reference group